GLM I An Introduction to Generalized Linear Models

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Outline

- § Overview of Statistical Modeling
- § Linear Models
 - ANOVA
 - Simple Linear Regression
 - Multiple Linear Regression
 - Categorical Variables
 - Transformations
- § Generalized Linear Models
 - Why GLM?
 - From Linear to GLM
 - Basic Components of GLM's
 - Common GLM structures
- § References



Modeling Schematic

Independent vars/ Predictors Weights Dependent var/ Response Claims Loan Age Losses Region **Exposures** Claims Loan-to-Value (LTV) Premium Persistency Credit Score Statistical Model **Model Results Parameters Validation Statistics**



General Steps in Modeling

Goal: Explain how a variable of interest depends on some other variable(s).

Once the relationship (i.e., a model) between the dependent and independent variables is established, one can make predictions about the dependent variable from the independent variables.

- 1. Collect/build potential models and data with which to test models
- 2. Parameterize models from observed data
- 3. Evaluate if observed data follow or violate model assumptions
- 4. Evaluate model fit using appropriate statistical tests
 - Explanatory or predictive power
 - Significance of parameters associated with independent variables
- 5. Modify model
- 6. Repeat



Basic Linear Model Structures

- § ANOVA: $Y_{ij} = \mu + \psi_i + e_{ij}$
 - Assumptions: errors are independent and follow $N(0,\sigma_e^2)$ Normal distribution with mean of zero and constant variance σ_e^2

$$\sum \psi_i = 0 \quad i = 1,....,k \text{ (fixed effects model)}$$

$$\psi_{i} \sim N(0,\sigma_{\psi}^2) \text{ (random effects model)}$$

- § Simple Linear Regression : $y_i = b_o + b_1x_i + e_i$
 - Assumptions:
 - linear relationship
 - errors are independent and follow $N(0,\sigma_e^2)$
- § Multiple Regression : $y_i = b_o + b_1 x_{1i} + \dots + b_n x_{ni} + e_i$
 - Assumptions: same as simple regression, but with n independent random variables (RV's)
- § Transformed Regression : transform x, y, or both; maintain assumption that errors are $N(0,\sigma_e^2)$

$$y_i = \exp(x_i)$$

 $\log(y_i) = x_i$



One-way ANOVA

```
Y_{ij} = \mu + \psi_i + e_{ij}
Y_{ij} is the j<sup>th</sup> observation on the i<sup>th</sup> treatment
j = 1, \dots, n_i
i = 1, \dots, k treatments or levels
\mu is the common effect for the whole experiment
\psi_i is the i<sup>th</sup> treatment effect
e_{ij} is random error associated with observation Y_{ij}, e_{ij} \sim N(0, \sigma_e^2)
S = ANOVAs can be used to test whether observations come from different populations or from the same population
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"Is there a statistically significant difference between two groups of claims?"

Is the frequency of default on subprime loans different than that for prime loans?

Personal Auto: Is claim severity different for Urban vs Rural locations?



One-way ANOVA

$$Y_{ij} = \mu + \psi_i + e_{ij}$$

- § Assumptions of ANOVA model
 - independent observations
 - equal population variances
 - Normally distributed errors with mean of 0
 - Balanced sample sizes (equal # of observations in each group)

§ Prediction:

- (observation) Y = common mean + treatment effect.
- Null hypothesis is no treatment effect. Can use contrasts or categorical regression to investigate treatment effects.



One-way ANOVA

- § Potential Assumption violations:
 - Implicit factors: lack of independence within sample (e.g., serial correlation)
 - Lack of independence between samples (e.g., samples over time on same subject)
 - Outliers: apparent non-normality by a few data points
 - Unequal population variances
 - Unbalanced sample sizes
- § How to assess:
 - Evaluate "experimental design" -- how was data generated? (independence)
 - Graphical plots (outliers, normality)
 - Equality of variances test (Levene's test)



- § Model: $Y_i = b_o + b_1 X_i + e_i$
 - Y is the dependent variable explained by X, the independent variable
 - Y: mortgage claim frequency depends on X: Seriousness of delinquency
 - Y: claim severity depends on X: Accident year
 - Want to estimate how Y depends on X using observed data
 - Prediction: $Y = b_o + b_1 x^*$ for some new x^* (usually with some confidence interval)



- § Model: $Y_i = b_o + b_1 X_i + e_i$
 - Assumptions:
 - 1) model is correct (there exists a linear relationship)
 - 2) errors are independent
 - 3) variance of e_i constant
 - 4) $e_i \sim N(0, \sigma_e^2)$

In terms of robustness, 1) is most important, 4) is least important

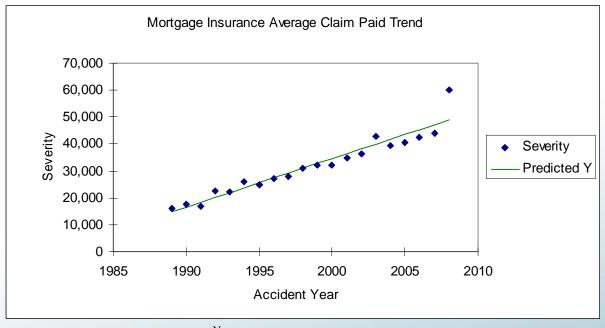
Parameterize:

Fit bo and b1 using Least Squares:

minimize:
$$\sum [y_i - (b_o + b_1 x_i)]^2$$



- The method of least squares is a formalization of best fitting a line through data with a ruler and a pencil
- Based on a correlative relationship between the independent and dependent variables



Note: All data in this presentation are for illustrative purposes only

$$b = \frac{\sum_{i=1}^{N} (Y_i - \overline{Y})(X_i - \overline{X})}{\sum_{i=1}^{N} (X_i - \overline{X})^2}, \quad a = \overline{Y} - b\overline{X}$$
Slope Intercept



ANOVA Regression

 Regression
 1 2,139,093,999

 Residual
 18 191,480,781

 Total
 19 2,330,574,780

df

SS

§ How much of the sum of squares is explained by the regression?

SSTotal = SSRegression + SSResidual (Residual also called Error)

$$SSTotal = \sum (y_i - y)^2$$

SSRegression =
$$b_1 est^* [\sum X_i y_i - 1/n(\sum X_i)(\sum y_i)]$$

$$SSResidual = \sum (y_i - y_{i est})^2$$



ANOVA

	df	SS	MS	F	Significance F
Regression	1	2,139,093,999	2,139,093,999	201.0838	0.0000
Residual	18	191,480,781	10,637,821		
Total	19	2,330,574,780			

MS = SS divided by degrees of freedom

R²: (SS Regression/SS Total)

percentage of variance explained by linear relationship

F statistic: (MS Regression/MS Residual)

- significance of regression:
 - tests H₀: b₁=0 v. H_A: b₁≠0



	Coefficients	Standard Error	t Stat	P-value	Lower 95%	Upper 95%
Intercept	-3,552,486.3	252,767.6	-14.054	0.0000	-4,083,531	-3,021,441
Accident Year	1,793.5	126.5	14.180	0.0000	1,528	2,059

<u>T statistics</u>: $(b_{i est} - H_o(b_i)) / s.e.(b_{i est})$

- significance of coefficients
- $T^2 = F$ for b_1 in simple regression



§ p-values test the null hypotheses that the parameters $b_1 = 0$ or $b_0 = 0$. If b_1 is 0, then there is no linear relationship between the independent variable Y (severity) and the dependent variable X (accident year). If b_0 is 0, then the intercept is 0.

SUMMARY OUTPUT

Regression State	istics
Multiple R	0.958
R Square	0.918
Adjusted R Square	0.913
Standard Error	3261.57
Observations	20

- 92% of the variance is explained by the regression
- The probability of observing this data given that $b_1 = 0$ is <0.00001, the significance of F
- Both parameters are significant
- $F = T^2$ for X Variable 1 201.08 = $(14.1804)^2$

ANOVA

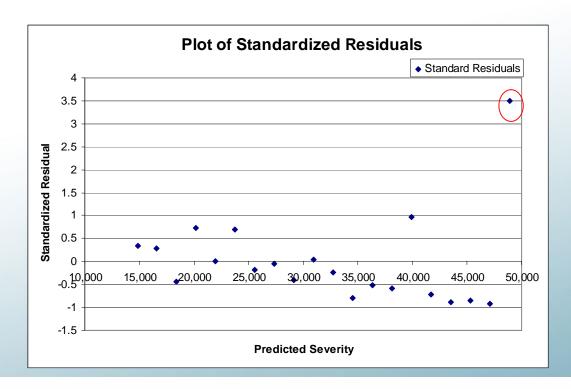
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Residuals Plot

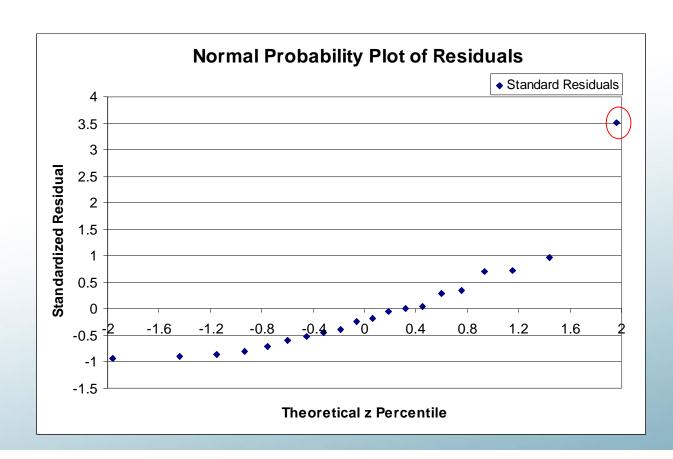
- § Looks at $(y_{obs} y_{pred})$ vs. y_{pred}
- § Can assess linearity assumption, constant variance of errors, and look for outliers
- § Residuals should be random scatter around 0, standard residuals should lie between -2 and 2
- § With small data sets, it can be difficult to asess





Normal Probability Plot

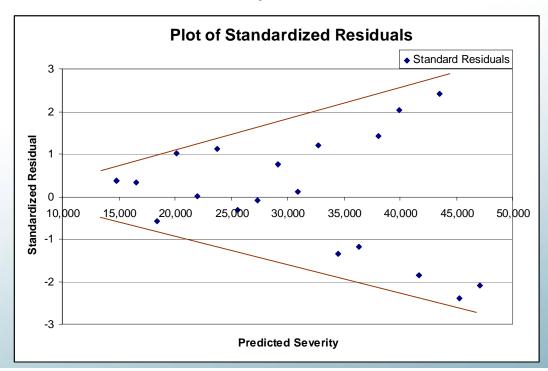
- **§** Can evaluate assumption $e_i \sim N(0, \sigma_e^2)$
 - Plot should be a straight line with intercept μ and slope σ_e^2
 - Can be difficult to assess with small sample sizes





Residuals

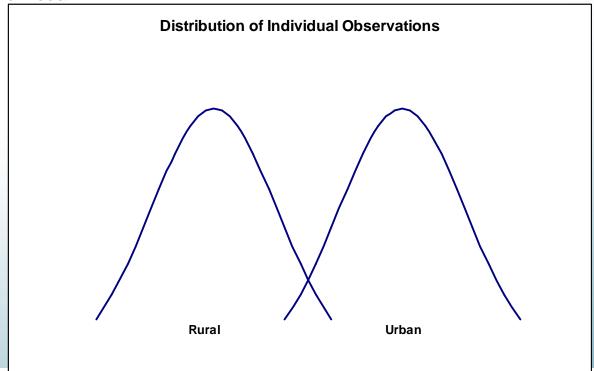
- § If absolute size of residuals increases as predicted value increases, may indicate nonconstant variance
- § May indicate need to transform dependent variable
- § May need to use weighted regression
- § May indicate a nonlinear relationship





Distribution of Observations

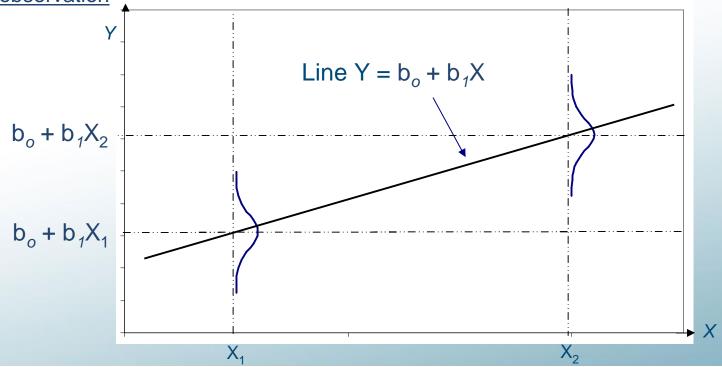
- § Average claim amounts for Rural drivers is normally distributed as are average claim amounts for Urban drivers
- § Mean for Urban drivers is twice that of Rural drivers
- § The variance of the observations is equal for Rural and Urban
- § The total distribution of average claim amounts is not Normally distributed
 - here it is bimodal





Distribution of Observations

- § The basic form of the regression model is $Y = b_o + b_1X + e$
- $\mu_i = E[Y_i] = E[b_o + b_1X_i + e_i] = b_o + b_1X_i + E[e_i] = b_o + b_1X_i$
- § The mean value of Y, rather than Y itself, is a linear function of X
- § The observations Y_i are normally distributed about their mean μ_i $Y_i \sim N(\mu_i$, σ_e^2)
- § Each Y_i can have a different mean μ_i but the variance σ_e^2 is the same for each observation





Predicting with regression

- § MSResiduals (also called s_e^2) is an estimate of σ_e^2 , the variability of the errors
- § Estimated Y has a lower standard error than Predicted Y, but both have the same point estimate μ_i
 - $Y_{pred} = b_o + b_1 x^*$ for some new x^*
 - $Y_{est} = b_o + b_1 x^*$ for some new x^*
- § standard error for both use se in formula
- § Y_{pred} standard error accounts for random variability around the line in addition to the uncertainty of the line
- § Typically give a confidence interval around the point estimate (e.g. 95%)
- $Y_{pred} \pm se(Y_{pred})*T_{0.025, DF}$
- § Predictions should only be made within the range or slightly outside of observed data. Extrapolation can lead to erroneous predictions



- § $Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + ... + \beta_n X_n + \varepsilon$
- $\S \quad E[\underline{Y}] = \underline{\beta} X$
- § Same assumptions as simple regression
 - 1) model is correct (there exists a linear relationship)
 - 2) errors are independent
 - 3) variance of e_i constant
 - 4) $e_i \sim N(0, \sigma_e^2)$
- § Added assumption the n variables are independent



- § Uses more than one variable in regression model
 - R-sq always goes up as add variables
 - Adjusted R-Square puts models on more equal footing
 - Many variables may be insignificant
- § Approaches to model building
 - Forward Selection Add in variables, keep if "significant"
 - Backward Elimination Start with all variables, remove if not "significant"
 - Fully Stepwise Procedures Combination of Forward and Backward



- § Goal : Find a simple model that explains things well with assumptions met
 - Model assumes all predictor variables independent of one another as add more, they may not be (multicollinearity—strong linear relationships among the X's)
 - As you increase the number of parameters (one for each variable in regression) you lose degrees of freedom
 - want to keep df as high as possible for general predictive power
 - problem of over-fitting



- § Multicollinearity arises when there are strong linear relationships among the x's
- § May see:
 - High pairwise correlations amongst the x's
 - Large changes in coefficients when another variable added or deleted
 - Large change in coefficients when data point added or deleted
 - Large standard deviations of the coefficients
- § Some solutions to combat overfitting and multicollinearity
 - Stepwise Regression (Forwards, Backwards, Exhaustive) -- Order matters
 - Drop one or more highly correlated variables
 - Use Factor Analysis or Principle Components Analysis to combine correlated variables into a smaller number of new uncorrelated variables



- § F significant and Adj R-sq high
- § Degrees of freedom ~ # observations # parameters
- § Any parameter with a t-stat with absolute value less than 2 is not significant

SUMMARY OUTPUT

Regression Statistics					
Multiple R		0.97			
R Square	(0.94			
Adjusted R Square		0.94	/		
Standard Error		0.05			
Observations		586			

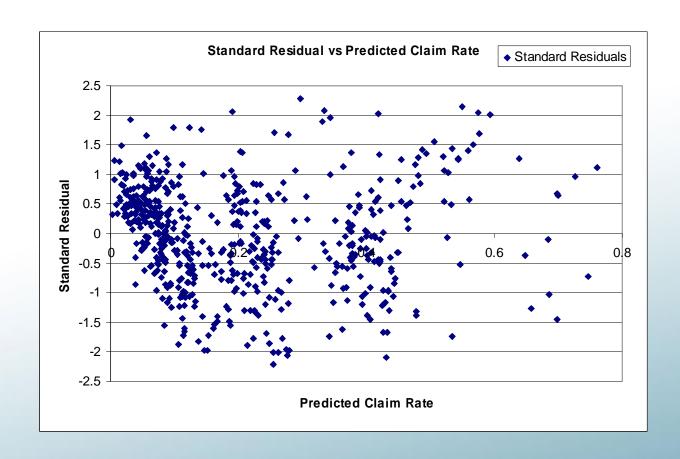
ANOVA

	df		SS	MS	F	Significance F
Regression		10	17.716	1.772	849.031	< 0.00001
Residual	(575	1.200	0.002		
Total		585	18.916			

	Coefficients	Standard Error	t Stat	P-value	Lower 95%	Upper 95%
Intercept	1.30	0.03	41.4	0.00	1.24	1.36
ltv85	-0.10	0.01	/ -12.9	/ 0.00	-0.11	-0.09
ltv90	-0.07	0.01	/ -9.1	0.00	-0.08	-0.06
ltv95	-0.04	0.01	-9.1	0.00	-0.05	-0.03
ltv97	-0.02	0.01	-6.0	0.00	-0.03	-0.01
ss30	-0.75	0.01	-55.3	0.00	-0.77	-0.73
ss60	-0.61	0.01	-56.0	0.00	-0.63	-0.59
ss90	-0.45	0.01	-53.5	0.00	-0.47	-0.43
ss120	-0.35	0.01	-40.1	/ 0.00	-0.37	-0.33
ssFCL	-0.24	0.01	-22.8	0.00	-0.26	-0.22
HPA	-0.48	0.03	-18.0	0.00	-0.53	-0.43

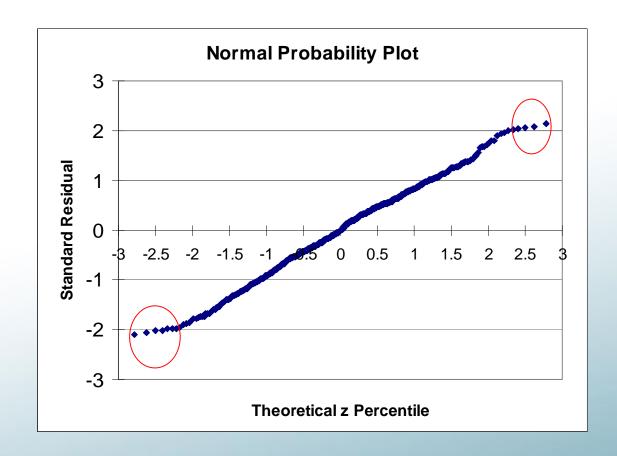


§ Residuals Plot





§ Normal Probability Plot





Categorical Variables

- § Explanatory variables can be discrete or continuous
- § Discrete variables generally referred to as "factors"
- § Values each factor takes on referred to as "levels"
- § Discrete variables also called Categorical variables
- § In the multiple regression example given, all variables were discrete except HPA (embedded home price appreciation)



Categorical Variables

- § Assign each level a "Dummy" variable
 - A binary valued variable
 - X=1 means member of category and 0 otherwise
 - Always a reference category
 - defined by being 0 for all other levels
 - If only one factor in model, then reference level will be intercept of regression
 - If a category is not omitted, there will be linear dependency
 - "Intrinsic Aliasing"



Categorical Variables

- § Example: Loan To Value (LTV)
 - Grouped for premium 5 Levels
 - <=85%, LTV85
 - 85.01% 90%, LTV90
 - 90.01% 95%, LTV95
 - 95.01% 97%, LTV97
 - >97% Reference
 - Generally postively correlated with claim frequency
 - Allowing each level it's own dummy variable allows for the possiblity of non-monotonic relationship
 - Each modeled coefficient will be relative to reference level

	X1	X2	Х3	X4
LTV	LTV85	LTV90	LTV95	LTV97
97	0	0	0	1
93	0	0	1	0
95	0	0	1	0
85	1	0	0	0
100	0	0	0	0
	97 93 95 85	LTV LTV85 97 0 93 0 95 0 85 1	LTV LTV85 LTV90 97 0 0 93 0 0 95 0 0 85 1 0	LTV LTV85 LTV90 LTV95 97 0 0 0 93 0 0 1 95 0 0 1 85 1 0 0



Transformations

- § A possible solution to nonlinear relationship or unequal variance of errors
- § Transform predictor variables, response variable, or both
- § Examples:
 - Y' = log(Y)
 - -X' = log(X)
 - X' = 1/X
 - $Y' = \sqrt{Y}$
- § Substitute transformed variable into regression equation
- § Maintain assumption that errors are $N(0,\sigma_e^2)$



Why GLM?

- **§** What if the variance of the errors increases with predicted values?
 - More variability associated with larger claim sizes
- **§** What if the values for the response variable are strictly positive?
 - assumption of normality violates this restriction
- § If the response variable is strictly non-negative, intuitively the variance of Y tends to zero as the mean of X tends to zero
 - Variance is a function of the mean
- § What if predictor variables do not enter additively?
 - Many insurance risks tend to vary multiplicatively with rating factors



Classic Linear Model to Generalized Linear Model

§ <u>LM</u>:

- X is a matrix of the independent variables
 - Each column is a variable
 - Each row is an observation
- $\underline{\beta}$ is a vector of parameter coefficients
- ε is a vector of residuals

§ <u>GLM</u>:

- X, <u>β</u> mean same as in LM
- $\underline{\varepsilon}$ is still vector of residuals
- g is called the "link function"

<u>LM</u>

$$\underline{Y} = \underline{\beta} X + \underline{\varepsilon}$$

$$E[\underline{Y}] = \underline{\beta} X$$

$$E[\underline{Y}] = \underline{\mu} = \underline{\eta}$$

$$\varepsilon \sim N(0, \sigma_e^2)$$

GLM

$$g(\underline{\mu}) = \underline{\eta} = \underline{\beta} X$$

$$\mathsf{E}[\underline{\mathsf{Y}}] = \underline{\mu} = g^{-1}(\underline{\eta})$$

$$\underline{Y} = g^{-1}(\underline{\eta}) + \underline{\varepsilon}$$

 $\varepsilon \sim \text{exponential family}$



Classic Linear Model to Generalized Linear Model

§ LM:

- 1) Random Component: Each component of \underline{Y} is independent and normally distributed. The mean μ_i allowed to differ, but all Y_i have common variance σ_e^2
- 2) Systematic Component: The n covariates combine to give the "linear predictor"

$$\underline{\eta} = \underline{\beta} X$$

3) Link Function: The relationship between the random and systematic components is specified via a link function. In linear model, link function is identity fnc.

$$E[\underline{Y}] = \underline{\mu} = \underline{\eta}$$

§ GLM:

- Random Component: Each component of Y is independent and from one of the exponential family of distributions
- 2) Systematic Component: The n covariates are combined to give the "linear predictor"

$$\underline{\eta} = \underline{\beta} X$$

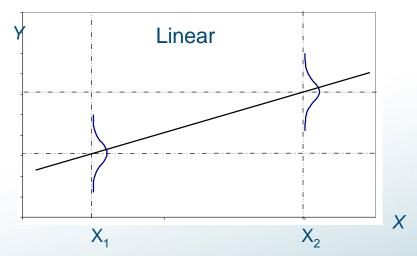
3) Link Function: The relationship between the random and systematic components is specified via a link function g, that is differentiable and monotonic

$$\mathsf{E}[\underline{\mathsf{Y}}] = \underline{\mu} = g^{-1}(\underline{\eta})$$



Linear Transformation versus a GLM

- § Linear transformation uses transformed variables
 - GLM transforms the mean
 - GLM not trying to transform Y in a way that approximates uniform variability



- § The error structure
 - Linear transformation retains assumption $Y_i \sim N(\mu_i, \sigma_e^2)$
 - GLM relaxes normality
 - GLM allows for non-uniform variance
 - Variance of each observation Y_i is a function of the mean $E[Y_i] = \mu_i$



The Link Function

- § Example: the log link function $g(x) = \ln(x)$; $g^{-1}(x) = e^x$
- § Suppose Premium (Y) is an multiplicative function of Policyholder Age (X_1) and Rating Area (X_2) with estimated parameters β_1 , β_2

$$- \eta_i = \beta_1 X_1 + \beta_2 X_2$$

$$- g(\mu_i) = \eta_i$$

-
$$E[Y_i] = \mu_i = g^{-1}(\eta_i)$$

-
$$E[Y_i] = \exp(\beta_1 X_1 + \beta_2 X_2)$$

$$- E[\underline{Y}] = g^{-1}(\underline{\beta} X)$$

$$- E[Y_i] = \exp(\beta_1 X_1) \cdot \exp(\beta_2 X_2) = \mu_i$$

$$- g(\mu_i) = \ln \left[\exp (\beta_1 X_1) \cdot \exp(\beta_2 X_2) \right] = \eta_i = \beta_1 X_1 + \beta_2 X_2$$

The GLM here estimates logs of multiplicative effects



Examples of Link Functions

Identity

$$-g(x)=x$$

$$g^{-1}(x) = x$$

§ Reciprocal

$$- g(x) = 1/x$$

$$g^{-1}(x) = 1/x$$

§ Log

$$g(x) = In(x)$$

$$g^{-1}(x) = e^x$$

§ Logistic

-
$$g(x) = \ln(x/(1-x))$$
 $g^{-1}(x) = e^{x}/(1+e^{x})$

$$g^{-1}(x) = e^{x}/(1+e^{x})$$



Error Structure

- § Exponential Family
 - Distribution completely specified in terms of its mean and variance
 - The variance of Y_i is a function of its mean
- § Members of the Exponential Family
 - Normal (Gaussian) -- used in classic regression
 - Poisson (common for frequency)
 - Binomial
 - Negative Binomial
 - Gamma (common for severity)
 - Inverse Gaussian
 - Tweedie (common for pure premium)



General Examples of Error/Link Combinations

- § Traditional Linear Model
 - response variable: a continuous variable
 - error distribution: normal
 - link function: identity
- § Logistic Regression
 - response variable: a proportion
 - error distribution: binomial
 - link function: logit
- § Poisson Regression in Log Linear Model
 - response variable: a count
 - error distribution: Poisson
 - link function: log
- § Gamma Model with Log Link
 - response variable: a positive, continuous variable
 - error distribution: gamma
 - link function: log



Specific Examples of Error/Link Combinations

Observed Response	Link Fnc	Error Structure	Variance Fnc
Claim Frequency	Log	Poisson	μ
Claim Severity	Log	Gamma	µ ²
Pure Premium	Log	Tweedie	µ ^p (1 <p<2)< td=""></p<2)<>
Retention Rate	Logit	Binomial	μ(1-μ)



References

- § Anderson, D.; Feldblum, S; Modlin, C; Schirmacher, D.; Schirmacher, E.; and Thandi, N., "A Practitioner's Guide to Generalized Linear Models" (Second Edition), CAS Study Note, May 2005.
- § Devore, Jay L. *Probability and Statistics for Engineering and the Sciences 3rd ed.*, Duxbury Press.
- § McCullagh, P. and J.A. Nelder. *Generalized Linear Models*, 2nd Ed., Chapman & Hall/CRC
- § SAS Institute, Inc. SAS Help and Documentation v 9.1.3

