## GLM I

An Introduction to Generalized Linear Models

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## Outline

B Overview of Statistical Modeling
B Linear Models

- ANOVA
- Simple Linear Regression
- Multiple Linear Regression
- Categorical Variables
- Transformations

B Generalized Linear Models

- Why GLM?
- From Linear to GLM
- Basic Components of GLM's
- Common GLM structures

B References

## Modeling Schematic



## General Steps in Modeling

Goal: Explain how a variable of interest depends on some other variable(s).
Once the relationship (i.e., a model) between the dependent and independent variables is established, one can make predictions about the dependent variable from the independent variables.

1. Collect/build potential models and data with which to test models
2. Parameterize models from observed data
3. Evaluate if observed data follow or violate model assumptions
4. Evaluate model fit using appropriate statistical tests

- Explanatory or predictive power
- Significance of parameters associated with independent variables

5. Modify model
6. Repeat

## Basic Linear Model Structures

B ANOVA: $Y_{i j}=\mu+\Psi_{i}+e_{i j}$

- Assumptions: errors are independent and follow $\mathrm{N}\left(0, \sigma_{\mathrm{e}}{ }^{2}\right)$ - Normal distribution with mean of zero and constant variance $\sigma_{e}{ }^{2}$
$\sum \psi_{i}=0 \quad i=1, \ldots \ldots, \mathrm{k}$ (fixed effects model)
$\Psi_{i \sim} \mathrm{~N}\left(0, \sigma_{\psi}{ }^{2}\right)$ (random effects model)
B Simple Linear Regression: $y_{i}=b_{0}+b_{1} x_{i}+e_{i}$
- Assumptions:
- linear relationship
- errors are independent and follow $\mathrm{N}\left(0, \sigma_{\mathrm{e}}{ }^{2}\right)$
ß Multiple Regression: $y_{i}=b_{0}+b_{1} x_{1 i}+\ldots+b_{n} x_{n i}+e_{i}$
- Assumptions: same as simple regression, but with $n$ independent random variables (RV's)

B Transformed Regression : transform x, y, or both; maintain assumption that errors are $\mathrm{N}\left(0, \sigma_{\mathrm{e}}{ }^{2}\right)$

$$
\begin{aligned}
& y_{i}=\exp \left(x_{i}\right) \\
& \log \left(y_{i}\right)=x_{i}
\end{aligned}
$$

## One-way ANOVA

$Y_{i j}=\mu+\Psi_{i}+e_{i j}$
$Y_{i j}$ is the $j^{\text {th }}$ observation on the $i^{\text {th }}$ treatment

$$
\begin{aligned}
& j=1, \ldots . n_{i} \\
& i=1, \ldots . k \text { treatments or levels }
\end{aligned}
$$

$\mu$ is the common effect for the whole experiment
$\Psi_{i}$ is the $i^{\text {th }}$ treatment effect
$\mathrm{e}_{i j}$ is random error associated with observation $\mathrm{Y}_{i j}, \mathrm{e}_{i j} N\left(0, \sigma_{\mathrm{e}}{ }^{2}\right)$
B ANOVAs can be used to test whether observations come from different populations or from the same population
"Is there a statistically significant difference between two groups of claims?"
Is the frequency of default on subprime loans different than that for prime loans? Personal Auto: Is claim severity different for Urban vs Rural locations?

## One-way ANOVA

$Y_{i j}=\mu+\Psi_{i}+e_{i j}$
B Assumptions of ANOVA model

- independent observations
- equal population variances
- Normally distributed errors with mean of 0
- Balanced sample sizes (equal \# of observations in each group)

B Prediction:

- (observation) $\mathrm{Y}=$ common mean + treatment effect.
- Null hypothesis is no treatment effect. Can use contrasts or categorical regression to investigate treatment effects.


## One-way ANOVA

B Potential Assumption violations:

- Implicit factors: lack of independence within sample (e.g., serial correlation)
- Lack of independence between samples (e.g., samples over time on same subject)
- Outliers: apparent non-normality by a few data points
- Unequal population variances
- Unbalanced sample sizes
ß How to assess:
- Evaluate "experimental design" -- how was data generated? (independence)
- Graphical plots (outliers, normality)
- Equality of variances test (Levene's test)


## Simple Regression

B Model: $\mathrm{Y}_{i}=\mathrm{b}_{0}+\mathrm{b}_{1} \mathrm{X}_{i}+\mathrm{e}_{i}$

- $\quad Y$ is the dependent variable explained by $X$, the independent variable
- Y: mortgage claim frequency depends on X : Seriousness of delinquency
- Y: claim severity depends on X: Accident year
- Want to estimate how Y depends on X using observed data
- Prediction: $Y=b_{0}+b_{1} x^{*}$ for some new $x^{*}$ (usually with some confidence interval)


## Simple Regression

B Model: $\mathrm{Y}_{i}=\mathrm{b}_{0}+\mathrm{b}_{1} \mathrm{X}_{i}+\mathrm{e}_{i}$

- Assumptions:

1) model is correct (there exists a linear relationship)
2) errors are independent
3) variance of $e_{i}$ constant
4) $e_{i} \sim N\left(0, \sigma_{e}{ }^{2}\right)$

In terms of robustness, 1) is most important, 4) is least important Parameterize:
Fit $\mathrm{b}_{0}$ and $\mathrm{b}_{1}$ using Least Squares:
minimize: $\sum\left[y_{i}-\left(b_{0}+b_{1} x_{i}\right)\right]^{2}$

## Simple Regression

- The method of least squares is a formalization of best fitting a line through data with a ruler and a pencil
- Based on a correlative relationship between the independent and dependent variables


Note: All data in this presentation are for illustrative purposes only

$$
\beta=\frac{\sum_{i=1}^{N}\left(Y_{i}-\bar{Y}\right)\left(X_{i}-\bar{X}\right)}{\sum_{i=1}^{N}\left(X_{i}-\bar{X}\right)^{2}}, \quad a=\bar{Y}-\beta \bar{X}
$$

Slope Intercept

## Simple Regression

| ANOVA |  | SS |
| :--- | ---: | ---: |
| Regression |  | SS |
| Residual | 1 | $2,139,093,999$ |
| Total | 18 | $191,480,781$ |

B How much of the sum of squares is explained by the regression?
SS = Sum Squared Errors
SSTotal = SSRegression + SSResidual (Residual also called Error)
SSTotal $=\sum\left(\mathrm{y}_{i}-\mathrm{y}\right)^{2}$
SSRegression $=\mathrm{b}_{1 \text { est }}{ }^{*}\left[\sum \mathrm{X}_{i} \mathrm{y}_{i}-1 / \ln \left(\sum \mathrm{X}_{i}\right)\left(\sum \mathrm{y}_{\mathrm{i}}\right)\right]$
SSResidual $=\sum\left(y_{i}-y_{i e s t}\right)^{2}$
= SSTotal - SSRegression

## Simple Regression

| ANOVA | $d f$ | $S S$ | $M S$ | $F$ | Significance $F$ |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Regression | 1 | $2,139,093,999$ | $2,139,093,999$ | 201.0838 | 0.0000 |
| Residual | 18 | $191,480,781$ | $10,637,821$ |  |  |
| Total | 19 | $2,330,574,780$ |  |  |  |

MS = SS divided by degrees of freedom
R2: (SS Regression/SS Total)

- percentage of variance explained by linear relationship

F statistic: (MS Regression/MS Residual)

- significance of regression:
- tests $\mathrm{H}_{0}: \mathrm{b}_{1}=0$ v. $\mathrm{H}_{\mathrm{A}}: \mathrm{b}_{1} \neq 0$


## Simple Regression

|  | Coefficients | Standard Error | t Stat |  | P-value | Lower 95\% |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Intercept | $-3,552,486.3$ | $252,767.6$ | -14.054 | 0.0000 | $-4,083,531$ | $-3,021,441$ |
| Accident Year | $1,793.5$ | 126.5 | 14.180 | 0.0000 | 1,528 | 2,059 |

T statistics: $\left(b_{i e s t}-H_{o}\left(b_{i}\right)\right) /$ s.e. $\left(b_{i e s t}\right)$

- significance of coefficients
- $T^{2}=F$ for $b_{1}$ in simple regression


## Simple Regression

B $p$-values test the null hypotheses that the parameters $b_{1}=0$ or $b_{0}=0$. If $b_{1}$ is 0 , then there is no linear relationship between the independent variable $Y$ (severity) and the dependent variable X (accident year). If $\mathrm{b}_{0}$ is 0 , then the intercept is 0 .

SUMMARY OUTPUT


|  | Coefficients | Standard Error | t Stat |  | P-value | Lower 95\% |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Intercept | $-3,552,486.3$ | $252,767.6$ |  | -14.054 | 0.0000 | $-4,083,531$ |
| Accident Year | $1,793.5$ | 126.5 |  | 14.180 | 0.0000 | 1,528 |

## Residuals Plot

B Looks at ( $\mathrm{Y}_{\text {obs }}$ - $\mathrm{y}_{\text {pred }}$ ) vs. $\mathrm{y}_{\text {pred }}$
B Can assess linearity assumption, constant variance of errors, and look for outliers
B Residuals should be random scatter around 0, standard residuals should lie between -2 and 2
B With small data sets, it can be difficult to asess


## Normal Probability Plot

ß Can evaluate assumption $\mathrm{e}_{i} \sim \mathrm{~N}\left(0, \sigma_{\mathrm{e}}{ }^{2}\right)$

- Plot should be a straight line with intercept $\mu$ and slope $\sigma_{e}{ }^{2}$
- Can be difficult to assess with small sample sizes



## Residuals

B If absolute size of residuals increases as predicted value increases, may indicate nonconstant variance

May indicate need to transform dependent variable
B May need to use weighted regression
B May indicate a nonlinear relationship


## Distribution of Observations

ß Average claim amounts for Rural drivers is normally distributed as are average claim amounts for Urban drivers
ß Mean for Urban drivers is twice that of Rural drivers
$ß$ The variance of the observations is equal for Rural and Urban
B The total distribution of average claim amounts is not Normally distributed

- here it is bimodal



## Distribution of Observations

B The basic form of the regression model is $Y=b_{0}+b_{1} X+e$
ß $\mu_{i}=\mathrm{E}\left[\mathrm{Y}_{i}\right]=\mathrm{E}\left[\mathrm{b}_{0}+\mathrm{b}_{1} \mathrm{X}_{i}+\mathrm{e}_{i}\right]=\mathrm{b}_{0}+\mathrm{b}_{1} \mathrm{X}_{i}+\mathrm{E}\left[\mathrm{e}_{1}\right]=\mathrm{b}_{0}+\mathrm{b}_{1} \mathrm{X}_{i}$
$B$ The mean value of $Y$, rather than $Y$ itself, is a linear function of $X$
$\mathcal{B}$ The observations $\mathrm{Y}_{i}$ are normally distributed about their mean $\mu_{i} \mathrm{Y}_{i} \sim \mathrm{~N}\left(\mu_{i}, \sigma_{\mathrm{e}}{ }^{2}\right)$
B Each $Y_{i}$ can have a different mean $\mu_{i}$ but the variance $\sigma_{e}{ }^{2}$ is the same for each observation


## Predicting with regression

B MSResiduals (also called $\mathrm{se}_{\mathrm{e}}{ }^{2}$ ) is an estimate of $\sigma_{\mathrm{e}}{ }^{2}$, the variability of the errors
B Estimated Y has a lower standard error than Predicted Y , but both have the same point estimate $-\mu_{i}$

- $Y_{\text {pred }}=b_{0}+b_{1} x^{*}$ for some new $x^{*}$
- $Y_{\text {est }}=b_{0}+b_{1} x^{*}$ for some new $x^{*}$

B standard error for both use $\mathrm{S}_{\mathrm{e}}$ in formula
B $Y_{\text {pred }}$ standard error accounts for random variability around the line in addition to the uncertainty of the line
B Typically give a confidence interval around the point estimate (e.g. 95\%)
B ( $\left.\mathrm{Y}_{\text {pred }} \pm \mathrm{se}\left(\mathrm{Y}_{\text {pred }}\right)^{\star} \mathrm{T}_{0.025, ~ D F}\right)$
B Predictions should only be made within the range or slightly outside of observed data. Extrapolation can lead to erroneous predictions

## Multiple Regression

B $Y=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\ldots+\beta_{n} X_{n}+\varepsilon$
B $E[\underline{Y}]=\underline{\beta} X$

B Same assumptions as simple regression

1) model is correct (there exists a linear relationship)
2) errors are independent
3) variance of $e_{i}$ constant
4) $e_{i} \sim N\left(0, \sigma_{e}{ }^{2}\right)$

B Added assumption the n variables are independent

## Multiple Regression

ß Uses more than one variable in regression model

- R-sq always goes up as add variables
- Adjusted R-Square puts models on more equal footing
- Many variables may be insignificant

B Approaches to model building

- Forward Selection - Add in variables, keep if "significant"
- Backward Elimination - Start with all variables, remove if not "significant"
- Fully Stepwise Procedures - Combination of Forward and Backward


## Multiple Regression

B Goal : Find a simple model that explains things well with assumptions met

- Model assumes all predictor variables independent of one anotheras add more, they may not be (multicollinearity-strong linear relationships among the X's)
- As you increase the number of parameters (one for each variable in regression) you lose degrees of freedom
- want to keep df as high as possible for general predictive power
- problem of over-fitting


## Multiple Regression

B Multicollinearity arises when there are strong linear relationships among the x's
B May see:

- High pairwise correlations amongst the x's
- Large changes in coefficients when another variable added or deleted
- Large change in coefficients when data point added or deleted
- Large standard deviations of the coefficients

B Some solutions to combat overfitting and multicollinearity

- Stepwise Regression (Forwards, Backwards, Exhaustive) -- Order matters
- Drop one or more highly correlated variables
- Use Factor Analysis or Principle Components Analysis to combine correlated variables into a smaller number of new uncorrelated variables


## Multiple Regression

## B F significant and Adj R-sq high

ß Degrees of freedom ~ \# observations - \# parameters
B Any parameter with a t-stat with absolute value less than 2 is not significant
SUMMARY OUTPUT

| Regression Statistics |  |
| :--- | ---: |
| Multiple R | 0.97 |
| R Square | 0.94 |
| Adjusted R Square | 0.94 |
| Standard Error | 0.05 |
| Observations | 586 |



## Multiple Regression

B Residuals Plot


Milliman

## Multiple Regression

B Normal Probability Plot


## Categorical Variables

B Explanatory variables can be discrete or continuous
B Discrete variables generally referred to as "factors"
B Values each factor takes on referred to as "levels"
B Discrete variables also called Categorical variables
B In the multiple regression example given, all variables were discrete except HPA (embedded home price appreciation)

## Categorical Variables

B Assign each level a "Dummy" variable

- A binary valued variable
- $X=1$ means member of category and 0 otherwise
- Always a reference category
- defined by being 0 for all other levels
- If only one factor in model, then reference level will be intercept of regression
- If a category is not omitted, there will be linear dependency
- "Intrinsic Aliasing"


## Categorical Variables

B Example: Loan - To - Value (LTV)

- Grouped for premium - 5 Levels
- <=85\%,

LTV85

- 85.01\%-90\%,

LTV90

- 90.01\%-95\%,

LTV95

- 95.01\%-97\%,

LTV97

- >97\%

Reference

- Generally postively correlated with claim frequency
- Allowing each level it's own dummy variable allows for the possiblity of non-monotonic relationship
- Each modeled coefficient will be relative to reference level

| Loan \# | LTV | X1 <br> LTV85 | X2 <br> LTV90 | X3 <br> LTV95 | X4 <br> LTV97 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 1 | 97 | 0 | 0 | 0 | 1 |
| 2 | 93 | 0 | 0 | 1 | 0 |
| 3 | 95 | 0 | 0 | 1 | 0 |
| 4 | 85 | 1 | 0 | 0 | 0 |
| 5 | 100 | 0 | 0 | 0 | 0 |

## Transformations

B A possible solution to nonlinear relationship or unequal variance of errors
ß Transform predictor variables, response variable, or both
B Examples:

- $\mathrm{Y}^{\prime}=\log (\mathrm{Y})$
- $X^{\prime}=\log (X)$
- $X^{\prime}=1 / X$
$-Y^{\prime}=\sqrt{ } Y$
B Substitute transformed variable into regression equation
B Maintain assumption that errors are $\mathrm{N}\left(0, \sigma_{\mathrm{e}}{ }^{2}\right)$


## Why GLM?

B What if the variance of the errors increases with predicted values?

- More variability associated with larger claim sizes
$ß$ What if the values for the response variable are strictly positive?
- assumption of normality violates this restriction
ß If the response variable is strictly non-negative, intuitively the variance of $Y$ tends to zero as the mean of $X$ tends to zero
- Variance is a function of the mean
ß What if predictor variables do not enter additively?
- Many insurance risks tend to vary multiplicatively with rating factors


## Classic Linear Model to Generalized Linear Model

B LM:

- $\mathbf{X}$ is a matrix of the independent variables
- Each column is a variable
- Each row is an observation
- $\underline{\beta}$ is a vector of parameter coefficients
- $\underline{\varepsilon}$ is a vector of residuals


## B GLM:

- $\mathbf{X}, \underline{\beta}$ mean same as in LM
- $\underline{\varepsilon}$ is still vector of residuals
- $g$ is called the "link function"


## LM

$\underline{Y}=\underline{\beta} \mathbf{X}+\underline{\varepsilon}$
$\mathrm{E}[\underline{\mathrm{Y}}]=\underline{\beta} \mathbf{X}$
$\mathrm{E}[\underline{Y}]=\underline{\mu}=\underline{n}$
$\varepsilon \sim N\left(0, \sigma_{\mathrm{e}}{ }^{2}\right)$

## GLM

$g(\underline{\mu})=\underline{\eta}=\underline{\beta} \mathbf{X}$
$\mathrm{E}[\underline{Y}]=\underline{\mu}=g^{-1}(\underline{n})$
$\underline{Y}=g^{-1}(\underline{\eta})+\underline{\varepsilon}$
$\varepsilon \sim$ exponential family

## Classic Linear Model to Generalized Linear Model

## B LM:

1) Random Component: Each component of $\underline{Y}$ is independent and normally distributed. The mean $\mu_{i}$ allowed to differ, but all $Y_{i}$ have common variance $\sigma_{e}{ }^{2}$
2) Systematic Component : The n covariates combine to give the "linear predictor"

$$
\underline{\eta}=\underline{\beta} \mathbf{X}
$$

3) Link Function : The relationship between the random and systematic components is specified via a link function. In linear model, link function is identity fnc.

$$
E[Y]=\underline{\mu}=\eta
$$

B GLM:

1) Random Component : Each component of $\underline{Y}$ is independent and from one of the exponential family of distributions
2) Systematic Component: The n covariates are combined to give the "linear predictor"

$$
\underline{n}=\underline{\beta} \mathbf{X}
$$

3) Link Function : The relationship between the random and systematic components is specified via a link function g , that is differentiable and monotonic

$$
\mathrm{E}[\underline{Y}]=\underline{\mu}=g^{-1}(\underline{n})
$$

## Linear Transformation versus a GLM

B Linear transformation uses transformed variables

- GLM transforms the mean
- GLM not trying to transform Y in a way that approximates uniform variability

ß The error structure
- Linear transformation retains assumption $\mathrm{Y}_{i} \sim \mathrm{~N}\left(\mu_{i}, \sigma_{e}{ }^{2}\right)$
- GLM relaxes normality
- GLM allows for non-uniform variance
- Variance of each observation $Y_{i}$ is a function of the mean $E\left[Y_{i}\right]=\mu_{i}$


## The Link Function

B Example: the log link function $g(x)=\ln (x) ; g^{-1}(x)=e^{x}$
B Suppose Premium (Y) is an multiplicative function of Policyholder Age $\left(X_{1}\right)$ and Rating Area $\left(X_{2}\right)$ with estimated parameters $\beta_{1}, \beta_{2}$

- $\eta_{i}=\beta_{1} X_{1}+\beta_{2} X_{2}$
$-\mathrm{g}\left(\mu_{i}\right)=\eta_{i}$
- $\mathrm{E}\left[\mathrm{Y}_{\mathrm{i}}\right]=\mu_{i}=g^{-1}\left(\eta_{i}\right)$
$-E\left[Y_{i}\right]=\exp \left(\beta_{1} X_{1}+\beta_{2} X_{2}\right)$
$-E[\underline{Y}]=g^{-1}(\underline{\beta} \mathbf{X})$
$-E\left[Y_{i}\right]=\exp \left(\beta_{1} X_{1}\right) \cdot \exp \left(\beta_{2} X_{2}\right)=\mu_{i}$
$-g\left(\mu_{i}\right)=\ln \left[\exp \left(\beta_{1} X_{1}\right) \cdot \exp \left(\beta_{2} X_{2}\right)\right]=\eta_{i}=\beta_{1} X_{1}+\beta_{2} X_{2}$
- The GLM here estimates logs of multiplicative effects


## Examples of Link Functions

## B Identity

$$
-g(x)=x
$$

$$
g^{-1}(x)=x
$$

B Reciprocal

$$
\text { - } g(x)=1 / x \quad g^{-1}(x)=1 / x
$$

## B Log

$$
\text { - } g(x)=\ln (x) \quad g^{-1}(x)=e^{x}
$$

B Logistic

$$
\text { - } g(x)=\ln (x /(1-x)) \quad g^{-1}(x)=e^{x /\left(1+e^{x}\right)}
$$

## Error Structure

B Exponential Family

- Distribution completely specified in terms of its mean and variance
- The variance of $Y_{i}$ is a function of its mean

B Members of the Exponential Family

- Normal (Gaussian) -- used in classic regression
- Poisson (common for frequency)
- Binomial
- Negative Binomial
- Gamma (common for severity)
- Inverse Gaussian
- Tweedie (common for pure premium)


## General Examples of Error/Link Combinations

B Traditional Linear Model

- response variable: a continuous variable
- error distribution: normal
- link function: identity

B Logistic Regression

- response variable: a proportion
- error distribution: binomial
- link function: logit
ß Poisson Regression in Log Linear Model
- response variable: a count
- error distribution: Poisson
- link function: log

B Gamma Model with Log Link

- response variable: a positive, continuous variable
- error distribution: gamma
- link function: log


## Specific Examples of Error/Link Combinations

| Observed <br> Response | Link Fnc | Error Structure | Variance Fnc |
| :--- | :--- | :--- | :--- |
| Claim Frequency | Log | Poisson | $\mu$ |
| Claim Severity | Log | Gamma | $\mu^{2}$ |
| Pure Premium | Log | Tweedie | $\mu^{\mathrm{p}(1<\mathrm{p}<2)}$ |
| Retention Rate | Logit | Binomial | $\mu(1-\mu)$ |

## References

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