The Recent Review and Changes to the National Council on Compensation Insurance's Individual Risk Experience Rating Plan

by Jon Evans

ABSTRACT

NCCI recently completed an extended review of its Experience Rating (ER) Plan. Although no major changes had been made for many years, testing indicated that ER Plan performance was still generally good. The primary cause of deteriorating performance was the use of a fixed split point between primary and excess losses while average claim severity increased dramatically. The review process uncovered many interesting facets of actuarial methodology related to experience rating, but the changes coming out of the review did not fundamentally change the structure of the plan. NCCI has implemented an increase in the split point from \$5,000 to \$15,000+ inflation (over three years), and subsequent procedures to periodically increase the split point in the future corresponding to an index of claim severity. Along with the split point increase, the maximum cap on modification factors was changed. As part of the review, NCCI also made changes to several components of the calculation of primary and excess experience period expected losses to conform to changes in NCCI's class ratemaking procedures. A well-constructed experience rating plan can perform very well for a very long time with appropriate indexation applied to components. Simplicity, consistency, transparency, and an automatic indexation are particularly important for industrywide bureau plans such as the NCCI Experience Rating Plan.

KEYWORDS

NCCI, individual risk rating, experience rating, workers compensation, credibility, split credibility

1. Introduction

Experience rating for individual workers compensation risks dates back to the beginning of the Casualty Actuarial Society (CAS) and workers compensation insurance in the United States. Early volumes of the Proceedings of the Casualty Actuarial Society (PCAS) contain numerous papers on experience rating individual workers compensation risks and the credibility of individual risk experience (Dorweiler 1934; Edwards and Woodward 1915; Fisher 1916; Greene 1915; Johnson 1941; Kormes 1934, 1935; Michelbacher 1918; Perryman 1937; Senior 1925; Smick 1939; Whitney 1918). This area of ratemaking is also somewhat unique in CAS history in that from the beginning, fitting credibility values and performance testing of consequent modification factors has involved a "predictive" framework, simultaneously comparing data on previous loss experience and subsequent loss experience (particularly see Dorweiler 1934).

This paper describes a recent multi-year review of the NCCI's Experience Rating Plan. NCCI Staff performed many different analyses. These analyses were presented for review and discussion at periodic meetings of NCCI's Individual Risk Rating Work-

ing Group. The IRRWG consists of actuaries representing workers compensation insurers that are affiliates of NCCI.

The review process confirmed that the ER Plan was performing well into the first years of the 21st century. In more recent years, quintile testing (Figure 1) began to show slight upward slope in modified relative pure loss ratios. (See the first part of Section 2 and Section 2.1.4 for a detailed description of quintile testing.) This was ultimately diagnosed to be a consequence of the split point between primary losses and excess losses remaining fixed while severity had increased significantly. The severity index associated with the split point had remained relatively flat for some years after NCCI's 1998 adjustments to the ER Plan. However, by Policy Year 2006 the severity index had more than doubled (Figure 2) and anticipated trends placed it on course to triple by around 2011–2013.

In response, NCCI has implemented a major increase in the split point from \$5,000 to \$15,000+ inflation (to be phased in over three years), and subsequent procedures to periodically increase the split point in the future corresponding to an index of claim severity. Along with the split point increase, the maximum cap on modification factors was changed.

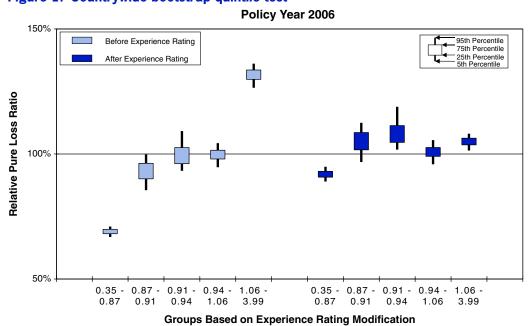
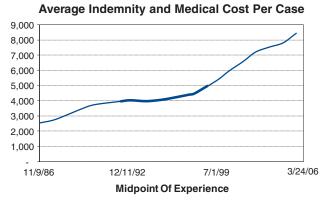


Figure 1. Countrywide bootstrap quintile test

Figure 2. Countrywide third report



While the ER review was being conducted, NCCI also implemented several changes to its class ratemaking methodology (Daley 2009). To accommodate class ratemaking changes in the ER Plan, changes were made to the calculation of D-ratios, which determine the fraction of expected experience period loss which is primary, and Expected Loss Rates (ELRs), which are rates of experience period expected ratable losses.

1.1. Research context

The content of this paper is primarily related to individual risk rating and credibility. Although the application is to workers compensation specifically, the methods shown are generally applicable to other casualty insurance. Background and specific details on the NCCI Experience Rating Plan will not be repeated here. Before reading this paper all readers, particularly those who do not routinely deal with the details of the plan, would be well advised to review one or more of several readily available references (Evans 2014; Gillam 1992a, 1992b 1992c; Mahler 1998; Meyers 1985; NCCI 1998; Venter 1987;). The most complete documentation of the specifics of the plan can be found in [NCCI, updated annually]. A number of papers from the early decades of the Proceedings of the CAS deal with individual risk rating in workers compensation (Dorweiler 1934; Edwards and Woodward 1915; Fisher 1916; Greene 1915; Johnson 1941; Kormes 1934, 1935; Michelbacher 1918; Perryman 1937; Senior 1925; Smick 1939; Whitney 1918). There are also a few other directly relevant PCAS papers from the middle of the 20th century (Bailey 1961; Stevens 1959; Uhthoff 1959).

1.2. Outline

Section 2 will describe some highlights of the review. The major changes to the ER Plan resulting from the review will be discussed in Section 3.

2. Review of the experience rating plan

The purpose of experience rating is to improve the estimate of future expected losses for an individual risk using previous actual loss experience for that risk. The basic formula for the experience rating modification factor, or *mod*, is (2.1).

$$\frac{A_{p} + WA_{e} + (1 - W)E_{e} + B}{E + B}$$
 (2.1)

Ap = actual primary ratable loss from the experience period

Ae = actual excess ratable loss from experience period

Ep = expected primary ratable loss from the experience period

Ee = expected excess ratable loss from experience period

E = Ep + Ee = expected ratable loss from experience period

B = ballast value

W = weight value

An alternative form that more directly illustrates the credibility values used in the mod formula is (2.2).

$$1 + Z_{p} \frac{A_{p} - E_{p}}{E} + Z_{e} \frac{A_{e} - E_{e}}{E}$$
 (2.2)

Zp = primary credibility

Ze = excess credibility

Ratable loss in the experience period includes a subset of total loss, determined through various exclusions such as individual loss limit, 70% exclusion of medical only losses, etc. Various specifics of the basic formula components and other aspects of the mod calculation have been changed in recent decades through special NCCI item filings. These filings include the revised Experience Rating Plan (RERP) in 1991, the Graduated Experience Rating Tables (GERT) in 1995, the Experience Rating Adjustment (ERA) in 1998, and the recent split point and maximum mod changes in 2012 based on the review described in this paper (and contained in Item Filing E-1402). As of this writing (2013), among states where NCCI files loss costs or rates, almost all have adopted ERA and Item E-1402. See Appendix A in Evans (2014) for more details on the history, such as the changing definition of primary and excess losses.

Experience Rating Plan performance is measured by the extent to which manual basis pure loss ratios vary by mod, increasing as mod values increase, and the uniformity of modified basis loss ratios across different mod values. It can be shown empirically through a quintile test of effective period relative pure loss ratios that the mod performs fairly well according to these two criteria (Figure 1). For the quintile test, risks are sorted by mod value and then grouped into five quintiles, each having an equal number of risks. Relative pure loss ratios on a manual and modified basis, respectively, for each quintile are shown in a chart. The review focused on how various aspects of the calculation affected performance of the plan, extensively using quintile testing to measure performance.

2.1. Stages of the review

The review was intended to be fairly comprehensive, and aspects of the plan to be reviewed were organized into four categories to be reviewed in roughly sequential stages, or tiers:

- 1. Severity Index
- 2. Data and Caps

- 3. Expected Losses
- 4. Credibility

2.1.1. Severity index

Severity indexation is used to update the maximum mod cap, the state accident limit, and the weight and ballast credibility values as part of NCCI class rate filings, which generally occur at the state level on an annual basis. The NCCI Experience Rating Plan Adjustment (ERA) of 1998 (NCCI 1998) also provided for indexation changes of the split point separating primary and excess losses on a national basis, but with no regular schedule or connection to periodic rate filings specified. Severity indexation is used to ensure that future performance will not be eroded by inflation in average claim amounts.

NCCI produces several different indices for experience rating that are substantially the same, but in practice are specified differently (Table 1):

- 1. State Average Cost Per Case (SACC), for all claims including medical only
- 2. State Reference Point (SRP) $\approx 250 \times \text{SACC}$
- 3. State Accident Limit (SAL) $\approx 25 \times SACC$
- 4. "G value" ≈ SACC / 1000

The ERA filing referenced a countrywide Average Cost Per Case (ACC), very similar to the State Average Cost Per Case, in connection with the split point (Figure 2).

An alternative index merits consideration if it varies between states and/or over time distinctly from

Table 1. Severity index example calculation*

•	•			
Policy Period	Total Total Undeveloped Undeveloped Incurred Period Cases Losses		Average Cost Per Case	
7/2003–6/2004	22,255	\$144,819,582	\$6,507	
7/2002–6/2003	22,939	\$150,662,284	\$6,568	
7/2001–6/2002	23,259	\$150,531,909	\$6,472	
Total	68,453	\$446,013,775	\$6,515	
State Average Cost Pe State Reference Point State Accident Limit (S "G" Value = SRP / 250 *Taken from an NCCI I	\$6,515 \$1,630,000 \$163,000 6.50			

the current index. Note, as will be described later, the severity index used for split point adjustments is on a countrywide basis and does not vary by state, unlike other severity indices used in the ER Plan. Alternative indexes that are very highly correlated to each other can be pared down to a single representative alternative. Secondly, the percentage claims and fraction of losses exceeding the state accident limit implied by an index should be relatively constant between states and over time. Several alternatives were investigated:

- State Average Claim Cost with 70% Exclusion of Med Only Losses
- State Average Cost Per Lost Time Claim
- State Average Cost Per Serious Claim
- State Average Weekly Wage (SAWW) capped at 150k on an annual basis
- Medical CPI

Only the State Average Cost per Lost Time Claim showed any promise of being both distinct from the current State Average Claim Cost basis and potentially more constant in terms of the excess percentage criteria (Figures 3, 4, and 5). However, ultimately this potential advantage was judged not great enough to warrant further investigation or the potential expense of such a fundamental change to the indexation bases.

2.1.2. Data and caps

This tier included topics such as the per claim limit (State Accident Limit), multiple claim limit, exclusions of catastrophic losses, maximum mod cap, 70% medical only exclusion, experience period, and special state and class exceptions to standard calculations. These features of the plan are less geared toward statistical or performance optimization but are selected with regard to practical considerations and buffering the impact of the mod in special cases for individual risks (as shown in Figure 6). For example, 70% of medical only losses are excluded from the mod calculation to remove the potential incentive for employers to not report small medical only claims.

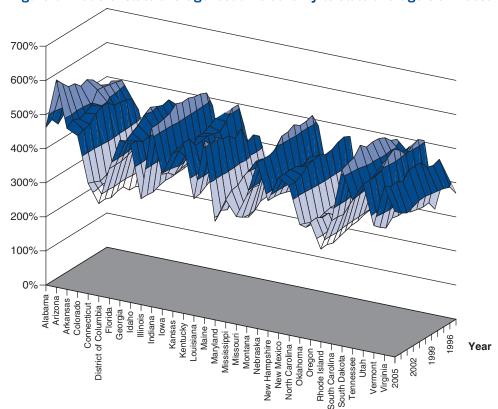


Figure 3. Ratio of state average lost time severity to state average claim cost

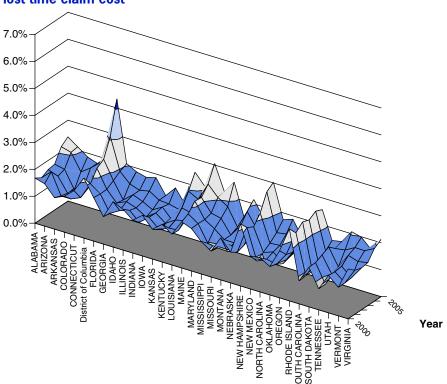


Figure 4. Percentage of all claims exceeding 6 times state average lost time claim cost



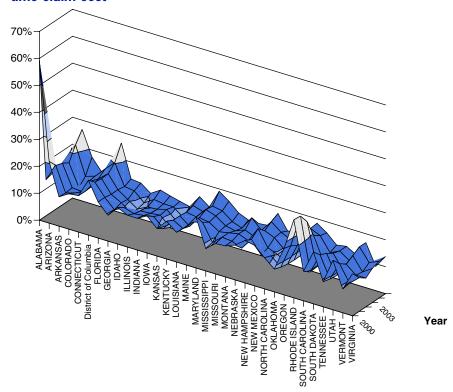


Figure 6. Individual risk impact of state accident limit and mod cap

Florida

Intrastate Rated Risks

Effective Period 1/1/2005-12/31/2005

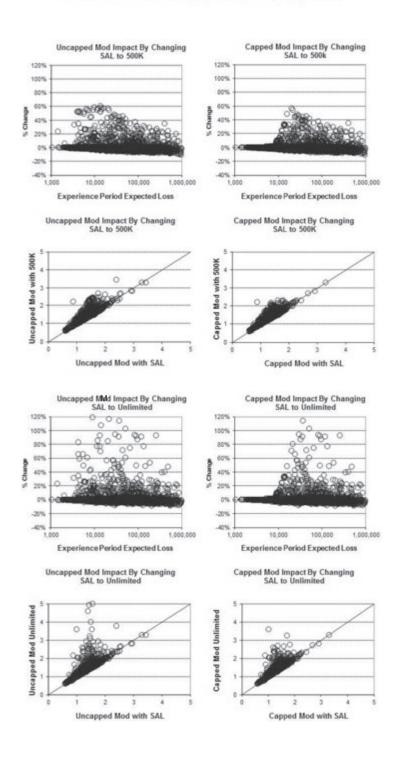


Figure 6 shows some examples of the impact of the maximum mod cap and the State Accident Limit on the mod at the individual risk level.

Analysis and discussion with members of the IRRWG uncovered no areas of great concern, aside from a longstanding issue with the practice of net reporting, which occurs in a small number of states where losses net of deductibles are used in experience rating. In most states requiring net reporting, actual experience period losses utilized in the mod calculation are net of policy deductibles, but experience period expected losses must be calculated gross of deductibles. This introduces a subsidy between policyholders without deductibles and those that choose deductibles. Previous NCCI attempts to adjust expected loss calculations (another tier of the review) for deductibles had been rejected by regulators in some net reporting states. In those jurisdictions, the subsidy is acknowledged by regulators and considered appropriate public policy as a means to encourage use of deductibles (the assumption being deductibles are a safety incentive). The 70% med only deduction overlaps some of the loss that would be under deductibles and partially mitigates the actuarial problem caused by the mismatch between actual and expected experience period losses in net reporting states.

2.1.3. Expected losses

Expected Loss Rates (ELRs) are multiplied by payroll in the experience period to produce total expected ratable losses for calculating the mod. The D-ratio, an estimate of the fraction of ratable losses which are primary, is used to separate this total into primary and excess components. Several statistical measurements of actual to expected losses for the experience period were explored, but none were found particularly insightful or resulted in any recommendations for changes to ELR or D-ratio calculations.

The calculation methodology and details underlying the ELRs and D-ratios were reviewed and found to be basically sound.

There was some concern that performance might be unequal between risks in different hazard groups. Haz-

ard groups are a partition of employment classifications to reflect claim severity. For quintile testing purposes, risks were "assigned" to hazard groups according to the classification which generated the largest manual premium. Quintile testing demonstrated that performance was effectively uniform across hazard groups (Figures 7–11). Although performance was uniform, there were some differences in the average mods between hazard groups, such as a general decrease in average mod with increasing hazard group. These differences were somewhat equalized by class ratemaking changes, not directly part of the ER review, in how ELRs and D-ratios are calculated.

A more immediate concern arose due to changes in class ratemaking that affected the ELR and D-ratio calculations and were being implemented simultaneous to the ER review (Daley 2009). These changes primarily consisted of the consolidation of non-serious and serious indemnity partial pure premiums into a single indemnity partial pure premium, calculation of loss development factors using losses limited at 500k, the use of excess loss factor (ELF) derived excess loss provisions instead of provisions based on spreading actual excess losses, and the shift from four to seven hazard groups. Significant changes were made to the ELR and D-ratio calculations to accommodate changes to class ratemaking, as described in Sections 3.3 and 3.4.

2.1.4. Credibility

The performance of credibility underlying the mod is demonstrated by the patterns in the modified basis relative pure loss ratios in the quintile test (Figure 12).

If the modified loss ratios show upward or positive "slope", credibility is too low. Downward slope indicates credibility is too high. The quintile test was the analytical workhorse of the review and similar empirical predictive tests have always been central to experience ratemaking (Dorweiler 1934; Stevens 1959). Over the course of the review, quintile tests were performed on many specific categories, such as by state, by hazard group, by risk size, and combinations thereof, and with mods as currently calculated as well as possible

Figure 7. Lowest quintile

2003 Modified Pure Loss Ratios By HG

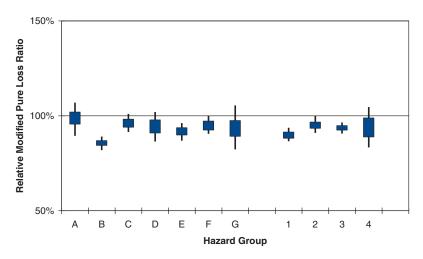


Figure 8. Low quintile

2003 Modified Pure Loss Ratios By HG

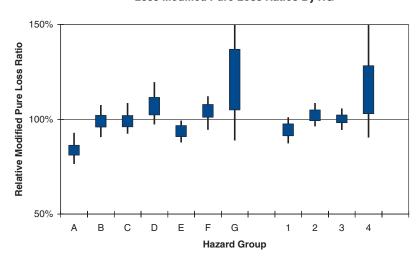


Figure 9. Middle quintile

2003 Modified Pure Loss Ratios By HG

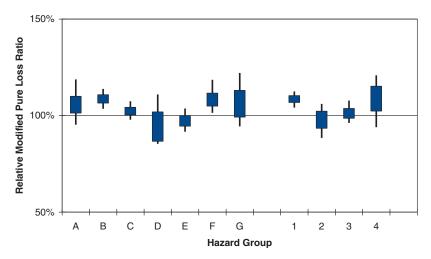


Figure 10. High quintile
2003 Modified Pure Loss Ratios By HG

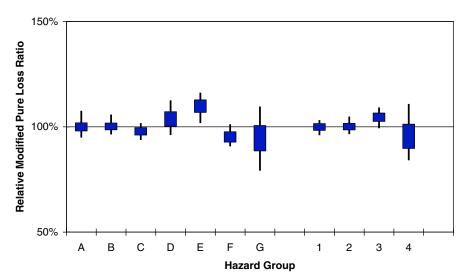
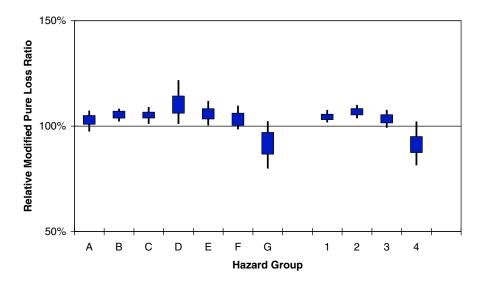


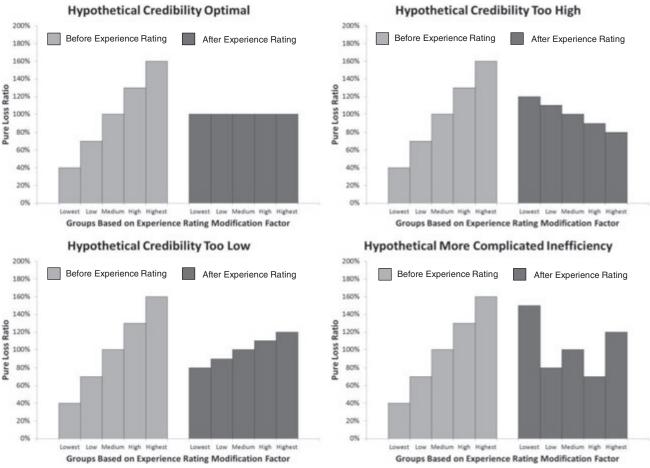
Figure 11. Highest quintile
2003 Modified Pure Loss Ratios By HG



alternative mod calculations. To illustrate the credibility of a quintile test, a particular concern in categories with sparse data, bootstrap-derived confidence intervals around the relative pure loss ratios (Figure 1) were usually presented instead of point estimates.

In recent years quintile testing began to show slight upward slope (Figure 1). As previously noted, this was ultimately diagnosed to be a consequence of the split point remaining fixed while severity had increased dramatically (Figure 2). The review did explore several different alternative models (rank correlation criteria, Generalized Linear Models, etc.) to the underlying parameterization of the weight and ballast credibility values. These alternatives were challenging to implement, particularly with regard to handling credibility differences by size of risk, a consideration that is well handled through parameter and process assumptions underlying the current parameterization model (Gillam 1992a; NCCI 1998). An additional con-

Figure 12. Interpreting quintile tests



cern was that changes in ELRs and D-ratios resulting from class ratemaking changes would not be available in data for credibility fitting purposes for some years in the future. Ultimately, NCCI determined that current parameterization worked well when the indexation updates were fully implemented. Since the indexation of weights, ballasts, and the State Accident Limit were automatically tied to routine rate filings they had kept up, but the split point had fallen behind.

Quintile testing that maintained the underlying credibility parameterization for alternative split points was performed. Two numerical statistics were calculated. The "Old Quintile Statistic" measures the degree to which the modified pure loss ratios are uniform; this statistic should be as small as possible. The "New Quintile Statistic" measures the amount of manual loss ratio variation eliminated by applying the mod; this

value should be as big as possible. Review of these statistics suggested that by Policy Year 2006 the optimal split point, where the old statistic tended to be lowest and the new statistic tended to be highest, was likely slightly higher than \$10,000 (Figure 13, Table 2 and 3). Severity trends subsequent to the Policy Year 2006 pointed to an optimal split point that would reach upwards of \$15,000 sometime in the years following Policy Year 2011.

Note, the relative pure loss ratios after experience rating are flatter with the \$10,000 split point (Figure 13) than with a \$5,000 split point (Figure 1), which shows a slight upward slope. Similar quintile test charts were produced corresponding to the split point and risk size combinations (Table 2 and 3) and were reviewed. The results were generally consistent with the hypothetical charts (Figure 13). There was a slope pattern after experience rating that increased

Figure 13. Countrywide bootstrap quintile test

Policy Year 2006 (using split point 10,000)

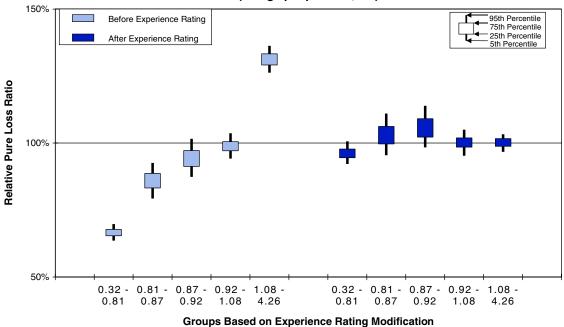


Table 2. Old quintile test statistic = B^*/A^*

A* = variance of un-modified loss ratios without bootstrapping

B* = variance of modified loss ratios without bootstrapping

B* = varianc	e of modified loss	s ratios without boot	tstrapping				
PY 2002 Risk Size							
Split Point Countrywide		1000–10,000	10,000-100,000	100,000-1M	1M-10M		
2500	0.064	0.258	0.097	0.057	0.034		
3750	0.037	0.163	0.071	0.054	0.026		
5000	0.021	0.107	0.032	0.043	0.015		
7500	0.007	0.034	0.006	0.032	0.019		
10000	0.007	0.027	0.011	0.021	0.013		
15000	0.043	0.109	0.095	0.020	0.009		
20000	0.093	0.218	0.271	0.022	0.014		
25000	0.191	0.415	0.524	0.035	0.015		
50000	1.027	2.211	3.614	0.137	0.066		
PY 2006		Risk Size					
Split Point	Countrywide	1000–10,000	10,000-100,000	100,000-1M	1M-10M		
2500	0.167	0.296	0.134	0.080	0.231		
3750	0.098	0.211	0.105	0.078	0.224		
5000	0.081	0.109	0.071	0.064	0.205		
7500	0.082	0.235	0.027	0.048	0.176		
10000	0.026	0.236	0.014	0.037	0.179		
15000	0.023	0.352	0.075	0.020	0.179		
20000	0.060	0.628	0.149	0.013	0.176		
25000	0.120	0.885	0.229 0.008		0.150		
50000	50000 0.870		1.527	0.068	0.180		

Table 3. New quintile test statistic = sign(A-B)* |A-B|0.5

A = variance of un-modified loss ratios with bootstrapping

B = variance	of modified loss	ratios with bootstra	apping					
PY 2002		Risk Size						
Split Point	Countrywide	1000–10,000	10,000-100,000	100,000-1M	1M-10M			
2500	0.218	0.239	0.241	0.275	0.273			
3750	0.225	0.246	0.252	0.281	0.271			
5000	0.234	0.252	0.255	0.279	0.264			
7500	0.241	0.253	0.260	0.279	0.271			
10000	0.247	0.241	0.263	0.282	0.272			
15000	0.240	0.228	0.249	0.282	0.269			
20000	0.235	0.205	0.223	0.282	0.264			
25000	0.221	0.166	0.173	0.278	0.266			
50000	-0.052	-0.212	-0.395	0.259	0.252			
PY 2006		Risk Size						
Split Point Countrywide		1000-10,000	10,000-100,000	100,000-1M	1M-10M			
2500	0.181	0.157	0.231	0.299	0.225			
3750	0.187	0.170	0.237	0.304	0.229			
5000	0.192	0.177	0.238	0.307	0.231			
7500	0.197	0.167	0.248	0.306	0.230			
10000	0.206	0.159	0.251	0.308	0.226			
15000	0.203	0.126	0.253	0.308	0.230			
20000	0.198	0.078	0.235	0.309	0.227			
25000	0.190	-0.034	0.218	0.306	0.224			
50000	0.063	-0.239	-0.173	0.290	0.227			

upward as the split point was lowered and equivalently decreased downward as the split point was decreased.

Another aspect of credibility is the minimum threshold for experience rating by size of risk. Credibility and the cap on the maximum mod both decrease with decreasing risk size. For very small risks, the variation in modified premium versus manual premium is small enough that it makes experience rating impractical for these risks.

The minimum premium threshold table was set in the early 1980s, with only a few minor changes since, to correspond to roughly 10 employees based on state average weekly wages and the state average rate level. To be eligible for experience rating, subject premium for the experience period either exceeds Column A for the most recent 24 months or average annual subject premium for the entire experience period exceeds the Column B level in the threshold table (Table 4). The current thresholds are much lower in real dollars than when established in the 1980s.

Severity indexation would be more desirable from a credibility standpoint since it would imply a standard of a minimum expected number of claims. Based on review of the pervasive use of severity indexation in other parts of the plan, it would be desirable for consistency and simplicity to use severity indexation for the eligibility threshold also. In the review, NCCI explored several different possible severity index based standards for eligibility. However, as of this writing (early 2014), NCCI has no specific plans regarding new eligibility thresholds.

Table 4. Current eligibility thresholds

	State Table of Subject Premium Eligibility Amounts							
State	Col A	Col B	State	Col A	Col B	State	Col A	Col B
AL	10,000	5,000	IA	7,500	3,750	NH	11,000	5,500
AK	5,000	2,500	KS	4,500	2,250	NM	9,000	4,500
AZ	6,000	3,000	KY	10,000	5,000	NC	5,000	2,500
AR	8,000	4,000	LA	10,000	5,000	OK	10,000	5,000
CO	8,000	4,000	ME	9,000	4,500	OR	5,000	2,500
CT	11,000	5,500	MD	10,000	5,000	RI	10,000	5,000
DC	7,000	3,500	MN	6,000	3,000	SC	9,000	4,500
FL	10,000	5,000	MS	9,000	4,500	SD	7,500	3,750
GA	10,000	5,000	MO	7,000	3,500	TN	9,000	4,500
HI	5,000	2,500	MT	5,000	2,500	UT	7,000	3,500
ID	6,000	3,000	NE	6,000	3,000	VT	8,000	4,000
IL	10,000	5,000	NV	6,000	3,000	VA	7,000	3,500
IN	5,000	2,500						

2.2. Other findings

2.2.1. Rank versus expected loss

In the course of the review, the use of a criterion of 0% rank correlation to fit credibility values, as introduced by Glenn Meyers (1985), was explored. It became apparent that the criterion Meyers used of 0% rank correlation between modification factors and subsequent actual modified pure loss ratios was generally not consistent with the key criterion of equal expected values for subsequent modified pure loss ratios, as evidenced empirically by flat modified relative pure loss ratios in a quintile test. The high degree of departure between these two criteria was surprising and the author is preparing a separate paper for the CAS on this topic.

2.2.2. Optimal number of quantiles

Quintile testing is central to performance testing and credibility fitting for the ER Plan (Gillam 1992a; Venter 1987). A natural question arises as to why NCCI uses five quantiles (quintiles) for this purpose. The author presented an explanation of the optimal number of quantiles in terms of a "noise to signal ratio." The key result was that this noise to signal ratio is proportional to k^{1.5} where k is the number of quantiles and inversely proportional to n^{0.5} where n is

the number of risks of approximately the same size. Thus there is a tremendous penalty, in terms of data volume required, for adding more quantiles to the test. Doubling the number of quantiles requires eight times as much data to maintain the same resolution, in terms of the noise to signal ratio. The author is preparing a separate paper for the CAS on this topic.

3. Changes to the experience rating plan

Several changes emerged from the review, some having already been implemented and others pending in the near future as of this writing.

3.1. Split point

Split points of about \$10,000 for Policy Year 2006 and about \$15,000 for Policy Year 2013 were determined to be appropriate updates for the split point. These selections were made with consideration to both historical changes in the severity index and quintile testing. The split point was initially raised to \$10,000 with further gradual changes over a three year period to reach the equivalent of \$15,000 plus index changes over the intervening time. Subsequent indexation of the split point based on the country-

wide Average Claim Cost will be an automatic part of NCCI's annual rate filings.

At any given time the split may vary by state due to differences in the schedule of rate filings by state, but will generally coincide when all of the seasonal filings have been made and before the next filing season.

3.2. Mod cap

The previous mod cap formula was (3.1)

$$1 + 0.00005(E + 2E/G)$$
 (3.1)

where E is experience period expected losses and G is the G value form of the severity index (see Section 2.1.1). This was updated to (3.2).

$$1.10 + 0.0004(E/G)$$
 (3.2)

This new formula will allow slightly more room for debit mods on the smallest risks. The new formula is also entirely a function of implied expected number of claims (E/G \sim expected losses/average severity \approx expected claim counts) between states and over time, whereas the previous formula was partially dependent on the absolute value of expected losses.

3.3. D-ratio calculation

Part of NCCI's recent revision of class ratemaking methods involved a shift from the three categories of partial pure premiums (serious indemnity, non-serious indemnity, and medical) to only the two categories of indemnity and medical (Daley 2009). Partial D-ratios had been calculated for each premium component statewide and then weighted by the partial premiums in each class to produce the class D-ratio. To

compensate for this loss of refinement, the new partial D-ratios for medical and indemnity losses were estimated separately for each hazard group. Since D-ratio factors should tend to decrease for higher hazard groups, a monotonicity smoothing algorithm was applied to indicated D-ratio factors. This smoother iterates several times, averaging a D-ratio factor with factors in the adjacent hazard group(s) in cases where a factor is lower than the factor in the adjacent higher hazard group and/or higher than the factor in the adjacent lower hazard group.

3.4. ELR calculation

The Expected Loss Rates (ELRs) are applied to experience period payrolls to calculate expected ratable losses. One key part of the calculation of ELRs is the Excess Loss Adjustment Factor (ELAF), which removes expected losses above the State Accident Limit. The ELAF calculation had been based on some simple excess loss ratio curves for undeveloped losses by entry ratio from the 1970s (Harwayne 1976), represented in terms of quadratic and cubic polynomials for Fatal, Permanent Total, and Major Permanent Partial injury type categories. Recent NCCI Excess Loss Factors (ELF) were not well suited for this purpose as they reflect losses that are developed and stochastically dispersed to ultimate values instead of the immature values for ratable experience period losses (Corro 2006). A set of replacement excess ratio curves for the ELAFs, still represented in terms of polynomials (3.3) but fit separately to Fatal, Permanent Total, Permanent Partial, Temporary Total, and Medical Only, was fit to more recent NCCI claims data (Table 5). The new excess ratio curves, still on

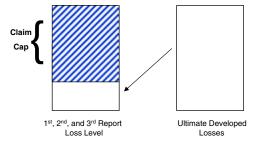
Table 5. Previous and revised ELR excess ratio curve coefficients

Previous Coefficients					Revised Coefficients		
Injury Type Category	a	b	С	Injury Type Category	a	b	С
Fatal	0.000000	2.310000	0.185000	Fatal	0.000000	1.017249	0.696561
Permanent total	0.000000	2.310000	0.185000	Permanent total	0.003148	0.054149	1.299625
Major permanent partial	0.167000	2.044000	0.805000	Permanent partial	0.000000	0.384102	1.086560
				Temporary total	0.001267	0.000000	0.876747
				Medical only	0.003165	0.000000	1.199899

Figure 14. Previous and revised ELR calculation

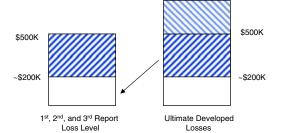
Previous

- De-develop ultimate unlimited losses
- Use 1970s era excess ratio curves to remove undeveloped losses excess of the SAL



Revised

- Remove the ratemaking excess ratio
- De-develop ultimate limited losses
- Apply updated undeveloped excess ratio curves to remove the layer between \$500,000 and the SAL



an undeveloped basis, were significantly higher for higher entry ratios.

The new class ratemaking procedures use loss development of claims limited to \$500k, with an excess provision derived from ELFs added to estimate ultimate unlimited losses. The previous ELAF procedure had "de-developed" ultimate expected losses and then subtracted an excess provision derived from the older undeveloped excess ratio curves. The new ELAF procedure first removes the ELF based excess provision at ultimate, de-develops losses, then subtracts an excess provision, derived from the updated undeveloped excess ratios curves, for the layer from the SAL to the 500k class ratemaking limit (Figure 14). It is worth noting that if at some point an SAL exceeds 500k, the calculation will add rather than subtract an excess layer provision, which is still actuarially sound.

$$E(r)$$
 = Undeveloped Excess Ratio
= $\frac{1}{ar^3 + br^2 + cr + 1}$ r = Entry Ratio (3.3)

4. Conclusions

The parameterization and severity indexation scheme for credibility underlying the NCCI ER plan implemented in the 1990s has well stood the test of

time. Recent performance testing indicates that experience rating continues to dramatically improve estimates of individual risk future expected losses versus the expected losses underlying the class loss costs. However, the indexation of parameters has lagged with regard to the split point, leading to a gradual deterioration of performance over time due to the resulting low credibility.

There is value to simple representation and presentation of the components of the ER formula even when such simplicity does not materially affect the actual calculation. Such simplicity facilitates understanding on the part of the many people over many years in many different circumstances who must deal with the formula.

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Abbreviations and notations

ACC, Average Cost Per Case

Ae, Actual Excess Ratable Loss From Experience Period

Ap, Actual Primary Ratable Loss From The Experience Period B, Ballast Value

C, Multiple of SACC Used for Eligibility Threshold

CAS, Casualty Actuarial Society

CPI, Consumer Price Index D, D-ratio

E, Experience Period Expected Ratable Loss

E(r), Undeveloped Excess Loss Ratio

Ee, expected excess ratable loss from experience Period

Ep, Expected Primary Ratable Loss From The Experience Period

ELAF, Excess Loss Adjustment Factor

ELF, Excess Loss Factor

ELR, Expected Loss Rate

ER, Experience Rating

ERA, Experience Rating Adjustment

G, G Value

GERT, Graduated Experience Rating Tables

HG, Hazard Group

IRRWG, Individual Risk Rating Working Group

k, Number of Quantiles

Mod, Experience Rating Modification Factor

n, Number of Risks in Data

NCCI, The National Council on Compensation Insurance

PCAS, Proceedings of the Casualty Actuarial Society

PY, Policy Year

r, Entry Ratio

RERP, Revised Experience Rating Plan

SACC, State Average Cost Per Case

SAL, State Accident Limit

SAWW, State Average Weekly Wage

SRP, State Reference Point

TX, Texas

W, Weight Value

WCSP, Workers Compensation Statistical Plan

y, Implied Expected Number of Claims

Ze, Excess Credibility

Zp, Primary Credibility