Meaningful Intervals

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Summary

- · Uncertainty and variability are distinct concepts
- fundamental difference between intervals for the mean vs process (CI/PI)
- Intervals only meaningful in a probabilistic framework
- Model assumptions must be explicit, interpretable, testable, and related to past volatility
- sophisticated methods (e.g. bootstrap) don't avoid need to check suitability of model. Bootstrap does not make Mack right.
- Regression formulation of standard link ratio methods and extensions-ELRF modeling framework includes Mack, Murphy & much more
- Link ratio methods Mack & quasi-Poisson GLM are structure-less, give uninformative indications, no descriptors of the features in the data. Give incorrect calendar period liability stream
- Even if diagnostics are perfect, mean reserve may still be wide of the mark
- CVs by accident year fail on basic principle of insurance (statistics)
- On updating, estimates of mean ultimates may be grossly inconsistent
- · Comments also apply to Munich Chain Ladder

More information will be available at

CABINET ROOM

6:30pm-11:30pm

Thursday September 18th

Summary c'td

- PTF (and MPTF) modeling framework for building single-/multi-triangle models that can capture trend structure and volatility in real data
- identified model in PTF framework describes the trend structure and volatility succinctly (four pictures). All assumptions tested and validated.
- Model satisfies axiomatic trend properties
- Real loss triangle can be regarded as sample path from fitted probabilistic model. Can't tell the difference between real and simulated
- Distributions, prediction intervals are conditional on an explicit set of assumptions that are tested and validated by the data
- Readily obtain percentile and V@R tables for total reserve and aggregates, by calendar year and accident year.
- Obtain consistent estimates of prior year ultimates on updating
- Calendar year liability stream is critical for capital allocation and cost of capital calculations. (What does it depend on?)
- · Pricing future underwriting years
- · No two companies are the same in respect of volatility and correlations
- All the above illustrated with many real data sets including <u>data from</u> <u>Murphy et al "Manually Adjustable Link Ratio Model for Reserving"</u>. A surprising finding!

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Variability and Uncertainty

- different concepts; not interchangeable

"Variability is a phenomenon in the physical world to be measured, analyzed and where appropriate explained. By contrast uncertainty is an aspect of knowledge."

Sir David Cox

A basic forecasting problem

Consider the following simple example –

n observations $Y_1...Y_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$

$$Y_i = \mu + \varepsilon_i \quad \varepsilon_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$$

Now want to forecast another observation...

(Actually, don't really need normality for most of the exposition, but it's a handy starting point.)

Example: Coin vs Roulette Wheel

Coin

100 tosses fair coin (#H?) Mean = 50 Std Dev = 5

CI [50,50]

In 95% of experiments with the coin the number of heads will be in interval [40,60].

"Roulette Wheel"

No. 0,1, ..., 100 Mean = 50 Std Dev = 29 CI [50,50]



In 95% of experiments with the wheel, observed number will be in interval [2, 97].

Where do you need more risk capital?

Introduce uncertainty into our knowledge - if coin or roulette wheel are mutilated then conclusions could be made only on the basis of observed data.

A basic forecasting problem

$$Y_{n+1} = \mu + \varepsilon_{n+1}$$

$$\hat{Y}_{n+1} = \hat{\mu} + \hat{\epsilon}_{n+1}$$
 forecast of the error term

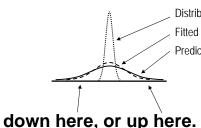
Variance of the forecast is

=
$$\sigma^2$$
 /n + σ^2 = "parameter uncertainty+
process variability"

So again, imagine that the distributions are normal.

Prediction distribution wider than fitted

The next observation might lie -



Distribution of $\overset{\bullet}{\mu}$ used for CI $\,$ - relates to "range" for mean Fitted distribution

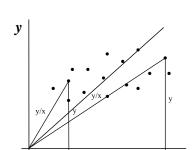
Prediction interval for Y unknown – relates to "range" for future *observed*

(implications for risk capital)

ELRF (Extended Link Ratio Family) x is cumu. at dev. j-1 and y is cum. at dev. j

 Link Ratios are a comparison of columns

X = Cum. (a) j-1



We can graph the ratios

of Y:X - line through O?

Using ratios \Rightarrow E(Y|x) = β x

Y = Cum. @ i

Mack (1993)

$$y = bx + \varepsilon$$
 : $V(\varepsilon) = \sigma^2 x^{\delta}$

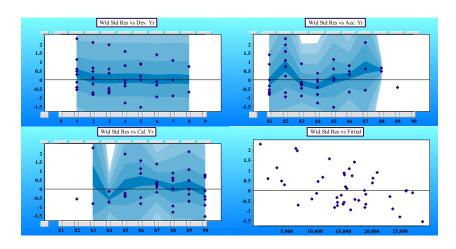
Minimize
$$\sum_{w} w (y - bx)^{2}$$
where
$$w = \frac{1}{x} \delta$$

1.
$$\delta = 1$$
, $\hat{b} = \frac{\sum x \frac{y}{x}}{\sum x} = \frac{\sum y}{\sum x}$

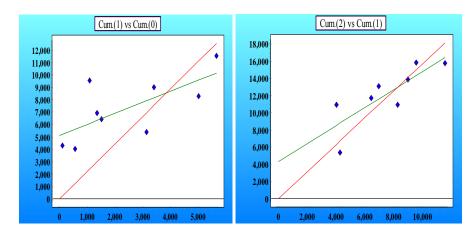
Chain Ladder Ratio (<u>Volume Weighted Average</u>)

2.
$$\delta = 2$$
, $\hat{b} = \frac{1}{n} \sum \frac{y}{x}$
Arithmetic Average

Mack and Murphy data IL(C) Mack (=volume weighted average) weighted standardized residuals. Note trend versus fitted values



Mack and Murphy data IL(C) Need intercepts



Intercept (Murphy (1994))

$$y = a + bx + \varepsilon$$
 : $V(\varepsilon) = \sigma^2 x^{\delta}$

Since y already includes x: y = x + p, ie p = y - x

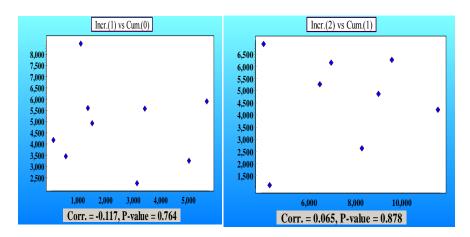
$$p = a + (b-1)x + \varepsilon : V(\varepsilon) = \sigma^2 x^{\delta}$$

Incremental Cumulative

at j at j -1

Is b -1 significant? Venter (1996)

Mack and Murphy data IL(C) Link Ratios=1. Zilch Predictive power



$$p=a+(b-1)x+\varepsilon : V(\varepsilon)=\sigma^2x^\delta$$
 Case (i) $b>1$ $a=0$

Case (ii)
$$b = 1$$
 $a \neq 0$
 $\hat{a} = \text{Ave}(\text{Incrementals})$
Abandon Ratios - No predictive power

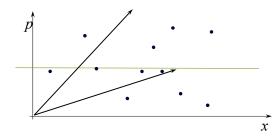
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Is assumption E(p|x) = a + (b-1)x tenable?

Note: If corr(x, p) = 0, then corr((b-1)x, p) = 0

If x, p uncorrelated, no ratio has predictive power

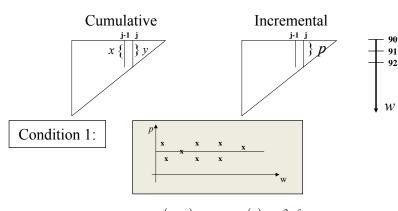
Ratio selection by actuarial judgement can't overcome zero correlation.



Corr. often close to 0

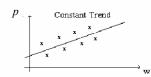
- -Sometimes not.

 Does this imply ratios are a good model?
- ranges?

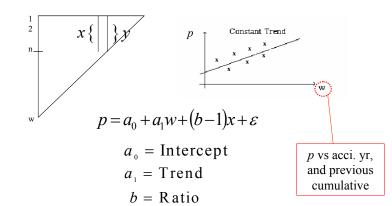


 $p = a + (b-1)x + \varepsilon : V(\varepsilon) = \sigma^2 x^{\delta}$

Condition 2:

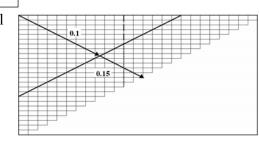


Now Introduce Trend Parameter For Incrementals



Condition 3:

Incremental



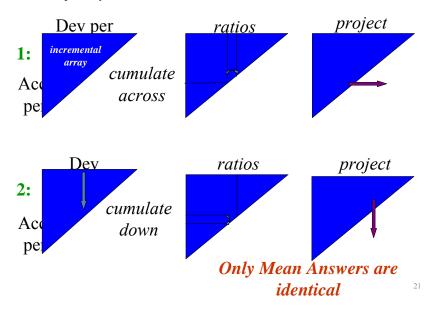
Review 3 conditions:

Condition 1: Zero trend

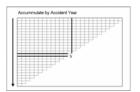
Condition 2: Constant trend, positive or negative

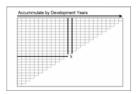
Condition 3: Non-constant trend

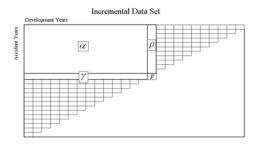
Mack=Chain Ladder (volume weighted average) treats accident years like development years. Can cumulate across or down. Does not matter!



Mack does not distinguish between accident years and development years







$$p = \beta \left(\frac{\alpha + \gamma}{\alpha} - 1 \right) = \frac{\beta \gamma}{\alpha}$$

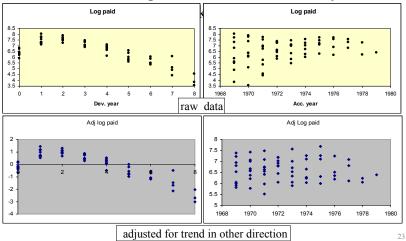
$$p = \gamma \left(\frac{\alpha + \beta}{\alpha} - 1 \right) = \frac{\gamma \beta}{\alpha}$$

The standard deviations are different because of different conditioning

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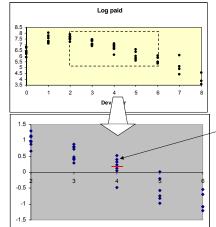
The Chain ladder (Volume weighted aveage) - Transpose Invariance property

Chain ladder does not distinguish between accident and development directions. But they are



Mack= The Chain Ladder (Volume Weighted Average)

Additionally, chain ladder (and ratio methods in general) ignore abundant information in nearby data.



- * If you left out a point, how would you guess what it was?
- observations at same dev. period *very* informative.

The quasi-Poisson GLM

- Model incrementals as quasi-Poisson (also called overdispersed Poisson), twoway cross-classification model with loglink.
- "Fit" doesn't reproduce what you'd think of as chain ladder (expected value is not ratio times previous cumulative), but forecasts do.

One step ahead prediction errors

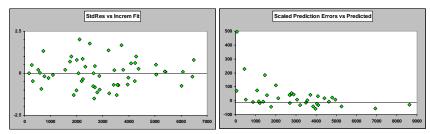
 - predictive behaviour of quasi-Poisson revealed by one-step ahead prediction errors.



Leave data out and predict it → validation

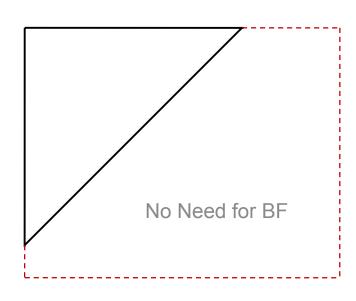
- P.E. = (predicted observed) (vs resid = fit obs)
- can standardize by $\sqrt{predictive variance}$
- Two crucial plots:
- prediction errors vs predicted (~ resids vs fit, better)
- P.E.s vs CY (res vs CY can show some issues)

One step ahead prediction errors



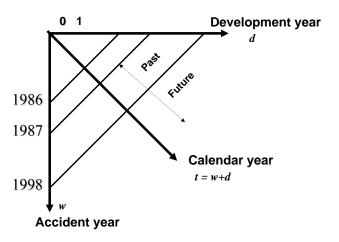
Res vs fit looks fine - BUT - PEs vs pred shows problem!

- -Predictive behaviour revealed:
 - underpredicts small, overpredicts large

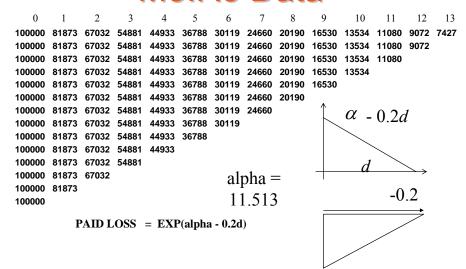


Probabilistic Modelling

Trends occur in three directions:

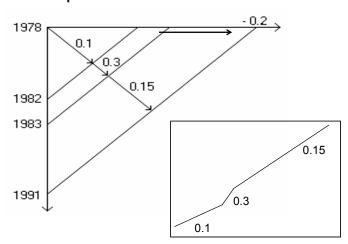


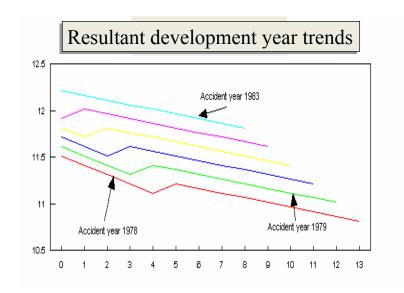
M3IR5 Data

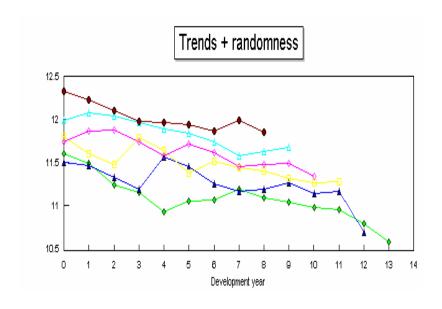


Probabilistic Modelling

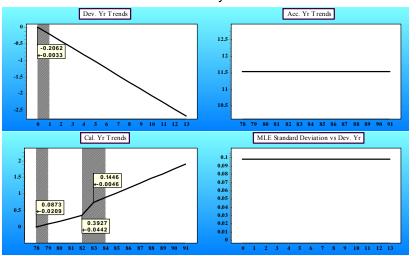
Axiomatic Properties of Trends





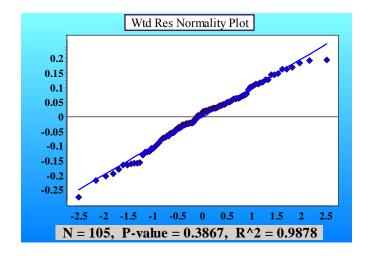


MODEL DISPLAYS. Graph bottom represents process variability



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Normal distribution about trend structure- integral part of model



Forecast lognormals for each cell
All assumptions are explicit. Process variability and
parameter uncertainty included

1984					
1504	205,644	220,996	169,549	166,858	15,289
1985	224,587	211,182	198,582	186,737	175,603
1900	221,660	247,187	207,918	18,780	17,816
4000	259,547	244,060	229,502	215,816	202,951
1986	220,334	234,427	23,094	21,896	20,799
4007	299,956	282,062	265,241	249,428	234,563
1987	271,278	28,430	26,939	25,576	24,325
4000	346,664	325,989	306,553	288,281	271,105
1988	35,037	33,181	31,483	29,927	28,496
4000	400,654	376,764	354,306	333,193	313,345
1989	40,913	38,797	36,858	35,076	33,433
4000	463,061	435,456	409,506	385,110	362,175
1990	47,859	45,440	43,218	41,171	39,280
4004	535,200	503,303	473,316	445,126	418,623
1991	56,078	53,304	50,750	48,391	46,206
	1995	1996	1997	1998	1999
al. Per.	2,506,808	2,278,761	2,052,087	1,824,784	1,594,672
Total	122,636	119,405	115,402	110,321	103,865

Simulate from forecast correlated lognormals Percentiles (Quantiles) and V@R statistics All assumptions are explicit. Process variability and parameter uncertainty included

Quantile Statistics and Value									
01		Sample			Kernel				
%	Quantile	# S.D.'s	V-a-R		Quantile	# S.D.'s	V-a-R		
99.995	26.970	3.820	3.544		27.145	4.008	3.718		
99.99	26.937	3.783	3.510		27.065	3.922	3.639		
99.98	26.866	3.707	3,439		26.970	3.820	3.544		
99.97	26.803	3.640	3.377		26.904	3.748	3.477		
99.96	26.773	3.607	3.347		26.850	3.690	3.423		
99.95	26.755	3.587	3.328		26.802	3.639	3.376		
99.94	26.749	3.581	3.323		26.759	3.592	3.333		
99.93	26.703	3.532	3.277		26.719	3.549	3.293		
99.92	26.691	3.519	3.265		26.682	3.508	3.255		
99.91	26.587	3.406	3.160		26.646	3.469	3.219		
99.9	26.567	3.385	3.141		26.611	3.432	3.185		
99.8	26.299	3.096	2.872		26.353	3.154	2.927		
99.7	26.152	2.937	2.725		26.201	2.991	2.775		
99.6	26.049	2.827	2.623		26.096	2.877	2.670		
F	05.004		0.500		20.045	A 704	0 500		

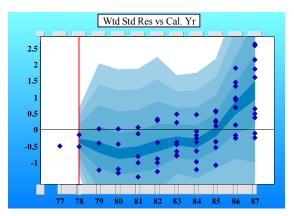
PROBABILISTIC MODEL Trends+ variation about trends **Based on Ratios** S2

Simulated triangles cannot be distinguished from real data – similar trends, trend changes in same periods, same amount of random variation about trends

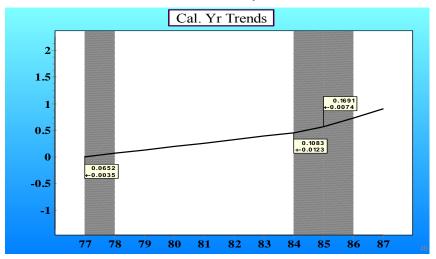
Models project past volatility into the future

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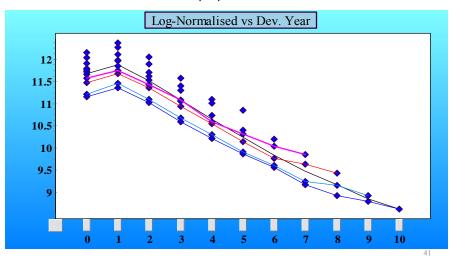
TG ABC PTF-Calendar period residuals adjusted for zero trend only



TG ABC PTF-Calendar Year Trends. Have control on future assumptions

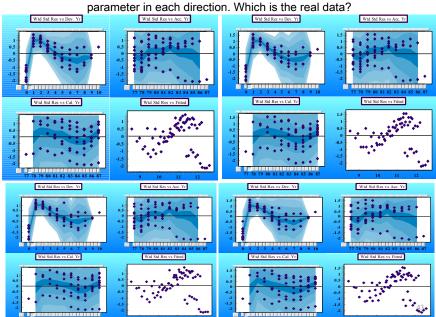


TG ABC
As you move down the accident years the "kick-up" is one development period earlier. Real data satisfies axiomatic trend properties



TG ABC: Three simulated from data model. Which is real data? 0.0473 +-0.0105 0.0439 +-0.0095 -0.3362 +-0.0100 -0.3303 +-0.0090 0.1610 +-0.0131 +-0.0149 -0.3994 +-0.0131 -0.3976 +-0.0119 2 3 4 5 6 Cal. Yr Trends 0.0989 +-0.0111 0.1083 +-0.0123 0.0684 +-0.0031 0.0652 +-0.0035 Acc. Yr Trends Dev. Yr Trends 0.1552 -0.0096 -0.3365 +-0.0072 -0.3209 0.1591 Cal. Yr Trends MLE Variance vs Dev. Yr MLE Variance vs Dev. Yr 0.1019 +-0.0088 0.0645 +-0.0025 0.0601 +-0.0037

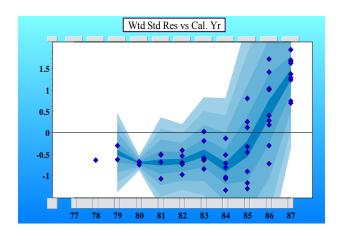
TG ABC: Three simulated, one real. Residuals of fitting only one parameter in each direction. Which is the real data?



TG ABC

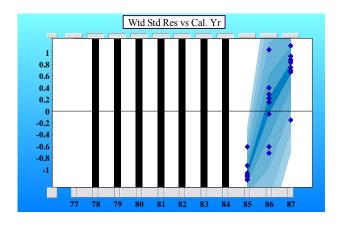
ELRF- Mack (volume weighted average link ratios) Residuals versus calendar year. Cannot capture calendar year trend structure.

No control on assumptions going forward



TG ABC

ELRF- Mack applied to last three years. No control on future assumptions



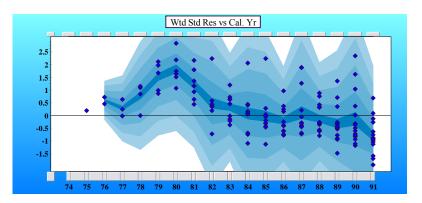
TG LR HIGH

ELRF- Mack

Trend estimated by method >> Trend in data

Mack= 896,133T +_104,117T

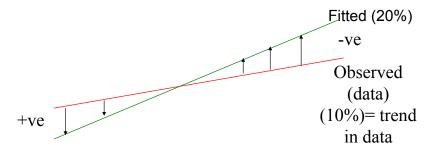
Arithmetic averages= 1,167,464T +- 234,466T



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TG LR HIGH

Residuals= Trend in data- Trend estimated by method

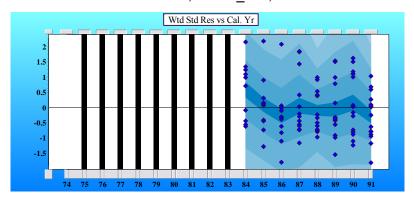


Residual = Data trend –Fitted Trend

TG LR HIGH

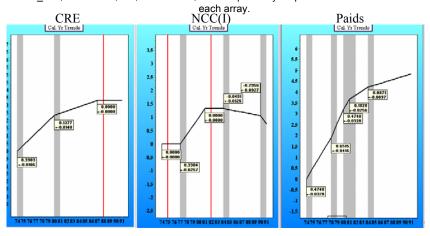
Trend in data close to trend in method. Best ELRF model, link ratios =1, and still not very good. Trend assumption of method not known 468,439T+- 47,346T Very different answer!

Mack= 896,133T +_104,117T



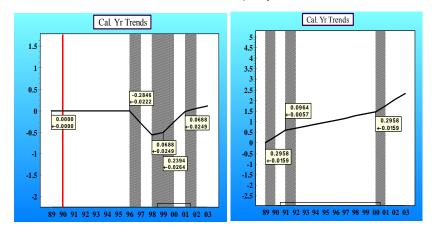
TG LR High

PTF calendar year structure for CREs, Number of claims closed and Paids Based on information from CRE and NCC(I) if anything paid trend in future likely to come down further. If continue with 8.71%+_0.97% obtain 593,506T+_42,191T. If revert to (extremely unlikely) 18.28%+_2.56% obtain 751,912T+_79,509T. If assume 47.48%+_3.2% for one year, thereafter 18.28%+_2.56% (completely unreasonable) obtain 1,007,497T+_112,234T. Note Mack and arithmetic average give 896,133T+_104,117T and 1,167,464T+-234,466T respectively. Important to monitor trends in



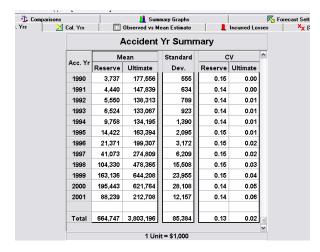
COMPANY XYZ

CREs versus Paids When was the company sold?

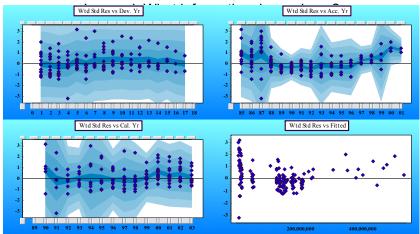


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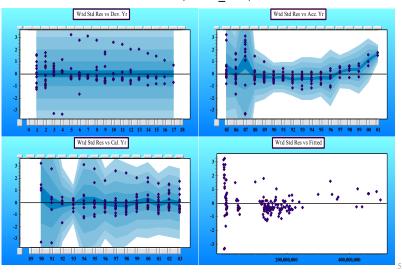
PTF: Forecast means and SDs based on easily interpretable assumptions. Trend assumption can be monitored 664,747+_85,384. Can also revert to longer term trend based on explanation for recent high trend



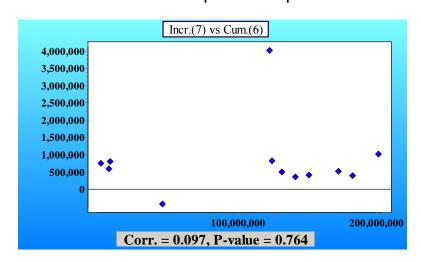
COMPANY XYZ Mack (volume weighted average link ratios) applied to



COMPANY XYZ Arithmetic Averages applied to IL(C). Reserves=2,100,714T +_1,071,680T Mack =697,371T+_234,088T



COMPANY XYZ Incurred data: No predictive power 6-7



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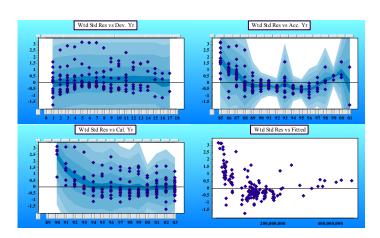
COMPANY XYZ

Mack applied to PL(C) is 427,024T+_53,715T (Residuals below)
Arithmetic Average PL(C) is 711,144T+_233,066T

Mack IL(C) 697,371T+_234,088T

Arithmetic average to IL(C) 2,100,714T +_1,071,680T Munich Chain Ladder?

PTF 664,747+_ 85,384 (Explicit assumptions)

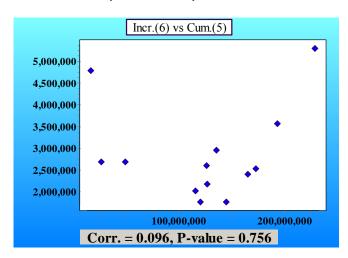


Mack PL(C) 427,024+ 53,715. Note inconsistent CVs

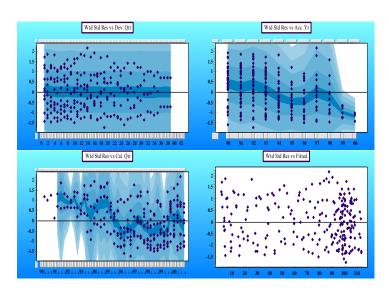
PTF 664,747+_85,384 explicit easily interpretable assumptions. Can revert to longer term trend and can

Acc. Yrs Mal. Yrs X	(%) Differer	nces 🚹 Co	mparisons 👢	Summary Grap	hs 🖽 BF &	ELR				
		Accident Yr Summary								
		М	ean	Standard	0	:v				
	Acc. Yr	Reserve	Ultimate	Dev.	Reserve	Ultimate				
	1991	8,022	151,421	1,261	0.16	0.01				
	1992	8,264	141,026	1,436	0.17	0.01				
	1993	8,737	135,280	1,650	0.19	0.01				
	1994	9,584	134,021	1,901	0.20	0.01				
	1995	13,142	162,114	2,806	0.21	0.02				
	1996	17,794	195,730	3,865	0.22	0.02				
	1997	27,401	261,137	6,535	0.24	0.03				
	1998	53,455	427,490	12,151	0.23	0.03				
	1999	89,976	571,049	23,699	0.26	0.04				
	2000	113,002	539,322	28,573	0.25	0.05				
	2001	52,474	176,943	18,355	0.35	0.10				
	Total	427,024	3,565,474	53,715	0.13	0.02				
				it = \$1.000						

COMPANY XYZ PL(C) No predictive power 5-6

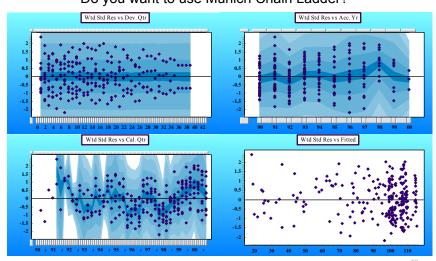


Seg 1 Mack PL(C)= 400,996T+_ 22,009T

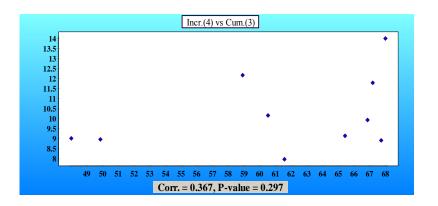


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Seg 1
Mack IL(C) 339,031T+_10,600T
Mack PL(C) 400,996T+_22,009T
Do you want to use Munich Chain Ladder?



IL(C) 3-4 no predictive power



Reserve Increases from year to year- Debunking a Myth

Here is a simple example that illustrates the main ideas that reserve increases do not represent under reserving. Indeed, they are necessary in order to manintain consisent estimates of prior year ultimates as the company writes new underwriting (accident) years).

On a logarithmic scale the data were generated as follows

 $Y(w,d) = 10 - 0.3^{*}d + 0.05$ (w+d-1) where w is the accident year 1,...,7 and d is the devlopment year 0,..., 5.

The numbers down each column increase by 0.05 on a log scale (approximately 5% annual). The numbers along each row decrease by 0.25 (=-0.3+0.05) on a log scale We have assumed that the paid losses run-off after five years. Even if this is the case for 1999, this may not be the case for subsequent accident years especially if inflation is 'high'

Reserves and ultimates as at year end 2004

		Ac	cident	Period v	s Develo	opment l	Period		
Cal. Per. T	otal	0	1	2	3	4	5	Reserve	Ultimate
1999	22,026	22,026	17,154	13,360	10,405	8,103	6,311	0	77,369
2000	40,310	23,156	18,034	14,045	10,938	8,519	6,634	6,634	81,325
2001	66,736	24,343	18,958	14,766	11,499	8,966	6,974	16,930	85,494
2002	68,999	25,591	19,930	15,522	12,088	9,414	7,332	28,835	89,878
2003	80,639	26,903	20,952	16,318	12,708	9,897	7,708	46,631	94,486
2004	91,085	28,283	22,026	17,154	13,360	10,405	8,103	71,048	99,331
Total Fitted	/Paid		2005	2006	2007	2008	2009	Total Reserve	Total Ultimate
Cal. Per. Total	358,796		66,022	46,251	30,589	18,112	8,103	169,078	527,873
Service and the service of the servi				1	Unit = S1				

Reserves and ultimates as at year end 2005

		Ac	cident	Period v	s Develo	opment l	Period		
Cal. Per.	Total	0	1	2	3	4	6	Reserve	Ultimate
1999	22,026	22,026	17,154	13,360	10,405	8,103	6,311	0	77,35
2000	40,310	23,156	18,034	14,045	10,938	8,519	6,634	0	81,32
2001	55,736	24,343	18,958	14,765	11,499	8,955	6,974	6,974	85,49
2002	68,999	25,591	19,930	15,522	12,088	9,414	7,332	16,746	89,87
2003	80,639	26,903	20,952	16,318	12,708	9,897	7,708	30,313	94,48
2004	91,085	28,283	22,026	17,154	13,360	10,405	8,103	49,022	99,33
2005	95,766	29,733	23,156	18,034	14,046	10,938	8,519	74,691	104,42
Total Fitted	I/Paid		2006	2007	2008	2009	2010	Total Reserve	Total Ultimate
Cal. Per. Total	464,660		69,407	48,623	32,167	19,041	8,519	177,746	632,29

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Reserve and ultimate as at year end 2004

Accident	Mean	Ultimate
Year	Reserve	
1999	0	77,359
2000	6,634	81,326
2001	15,930	85,494
2002	28,835	89,878
2003	46,631	94,480
2004	71,048	99,331
Total	169,078	527,873

Reserve and ultimate as at year end 2005

Accident	Mean	Ultimate		
Year	Reserve	3		
1999	0	77,359		
2000	0	81,326		
2001	6,974	85,495		
2002	16,746	89,878		
2003	30,313	94,486		
2004	49,022	99,331		
2005	74,691	104,424		
Total	177,746	632.298		

	tatio of t ultimate				
to year t-1					
	1.051267467				
	1.051275745				
	1.051266156				
	1.051269499				
	1.051277438				
	1.051273016				
	1.051266279				
atio	of Reserves				

N.B.

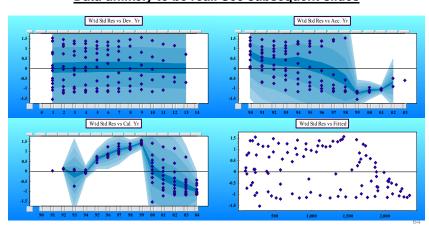
- 1. Estimtes of ultimate losses by accident year (1999- 2004) remain the same on update at end of 2005
- 2. The ratio of ultimate for year t to year t-1 is 1.05
- 3. Increase in total reserves from 2004 to 2005 is 1.05

Reserving"

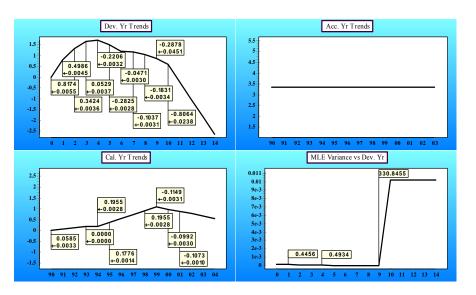
Mack (volume weighted averages) standardized residuals. Murphy only shows residuals versus accident year. Reserve too large by a factor of

Reserve is 10,437 +_ 936

Data unlikely to be real. See subsequent slides



Data from Murphy et al "Manually Adjustable Link Ratio Model for reserving" PTF designed model. Process Variability almost zero.



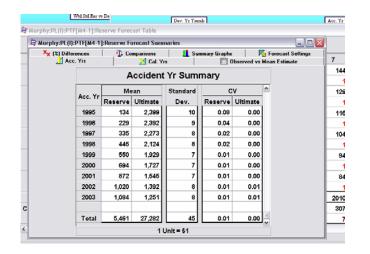
Data from Murphy et al "Manually Adjustable Link Ratio Model for Reserving" PTF Forecast table. Note mean fitted (black) very close to observed (blue). An absolute perfect fit, one of the reasons it is not likely to be real data

		Accid	lent Pe	riod vs	Develo	pment	Period		
	Cal. Per. Total	0	1	2	3	4	5	6	7
4000	2,229	70	194	285	363	344	248	168	144
1998	2,233	70	192	288	367	345	248	169	
4000	2,915	86	173	258	326	309	222	151	129
1999	2,911	88	173	257	328	308	225	1	
2000	2,737	76	157	232	293	277	200	135	116
	2,733	76	159	231	289	278	1	1	
	2,534	69	141	208	263	249	180	122	104
2001	2,536	68	139	204	262	2	1	1	
	2,299	62	126	187	236	224	161	109	9
2002	2,294	58	127	187	2	2	1	1	
	2,074	56	113	168	212	201	145	98	84
2003	2,061	53	114	1	2	2	1	1	
	Total Fitted/Paid		2004	2005	2006	2007	2008	2009	2010
Cal. Per.	21,855		1,816	1,529	1,223	908	635	440	307
Total	21,821		1,816	12	11	10	9	8	7

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Data from Murphy et al "Manually Adjustable Link Ratio Model for Reserving" PTF Accident year Summary Reserve 5,461+_45.

Why is Mack 10,437+_936 more than double??



Data from Murphy et al "Manually Adjustable Link Ratio Model for Reserving"

PTF model validation

Prediction errors two calendar years removed. That is at end 2002

Reserve forecasts beyond 2004 as at end 2002, 2003 and 2004.

Extremely stable!

