Loss Simulation Model Testing and Enhancement

Casualty Loss Reserve Seminar

By

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Agenda

- Research Overview
- Model Testing
- Real Data
- Model Enhancement
- Further Development

I. Research Overview

Background – Why use the LSM

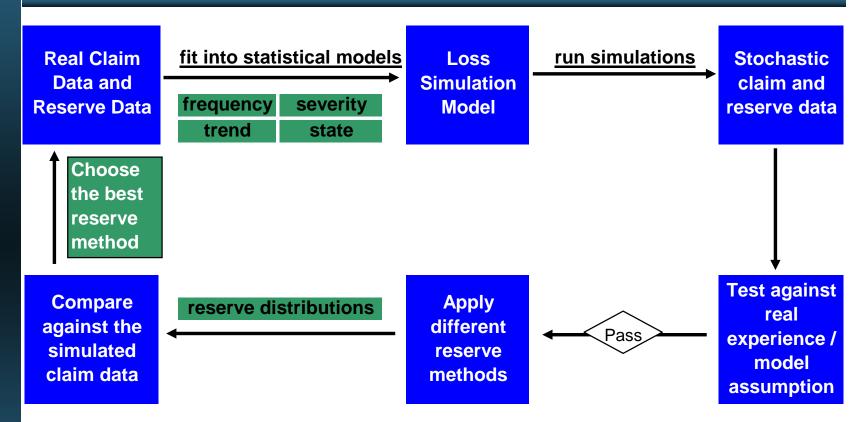
Reserving is a challenging task which requires a lot of judgements on assumption setting

The loss simulation model (LSM) is a tool created by the CAS Loss Simulation Model Working Party (LSMWP) to generate claims that can be used to test loss reserving methods and models

It helps us understand the impact of assumptions on reserving from a different perspective – distribution based on simulations that resemble the real experience

In addition, stochastic reserving is also a popular trend.

Background – How to use the LSM

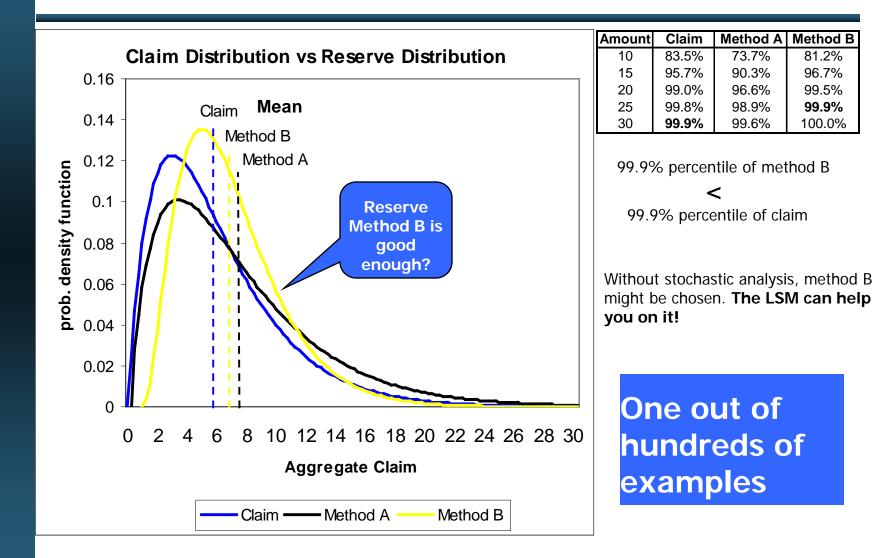


We do not expect an accurate estimation of the claim amount.

We are more concerned about the adequacy of our reserve.

At what probability that the reserve is expected to be below the final payment?

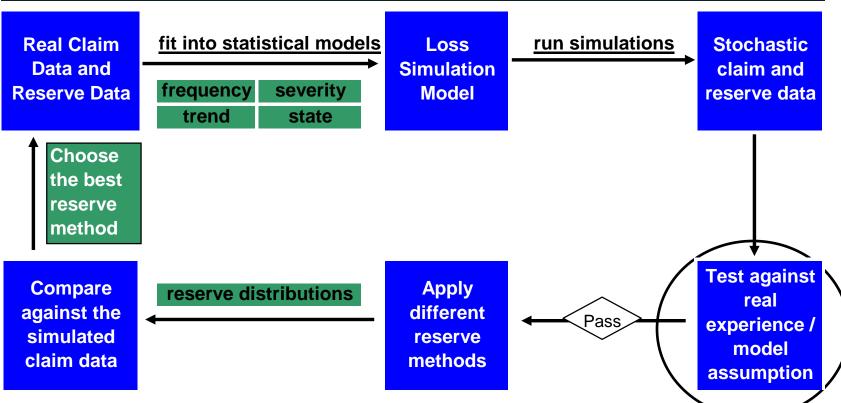
Background – How to use the LSM



Overview

- Test some items suggested but not fully addressed in the CAS LSMWP summary report "Modeling Loss Emergence and Settlement Processes"
- > Fit real claim data to models.
- Build two-state regime-switching feature in the LSM to add an extra layer of flexibility to describe claim data.
- Software: LSM and R. The source code of model testing and model fitting using R is provided.

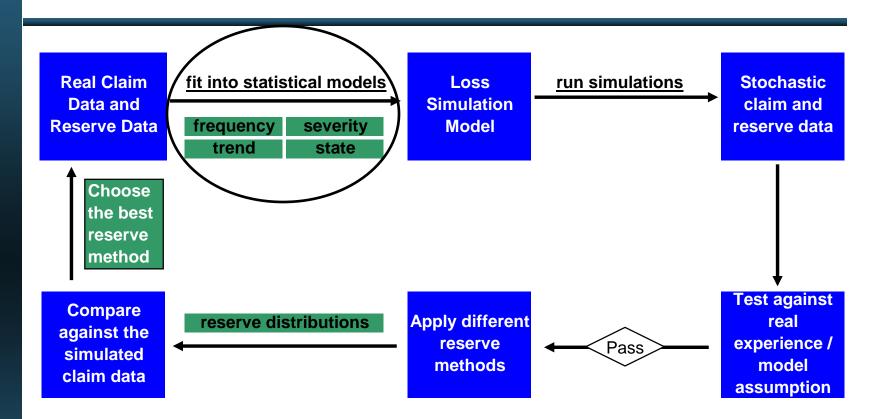
Model Testing



Test against model assumption

- ✓ Negative binomial frequency distribution
- ✓ Correlation
- ✓ Severity trend
- ✓ Case reserve adequacy distribution

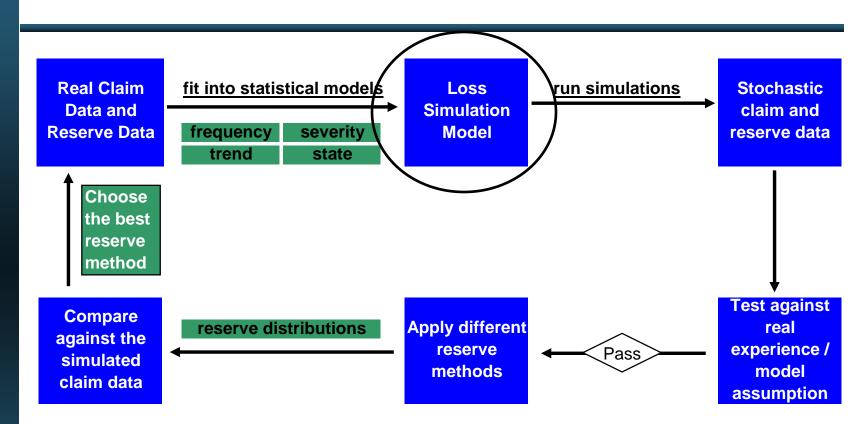
Real Data Model Fitting



Fit real claim data to statistical models

- ✓ frequency
- ✓ Severity
- ✓ Trend
- ✓ Correlation

Model Enhancement



Two-state regime-switching distribution

- Switch between states at specified probability
- ✓ Each state represents a distinct distribution

II. Model Testing

DAY ONE

9 AM



Negative Binomial Frequency Testing

Frequency simulation

 \checkmark One Line with annual frequency Negative Binomial (size=100, prob.=0.4)

- ✓Monthly exposure: 1
- ✓ Frequency Trend: 1
- ✓ Seasonality: 1
- ✓ Accident Year: 2000
- ✓ Random Seed: 16807
- ✓ No. of Simulations: 1000

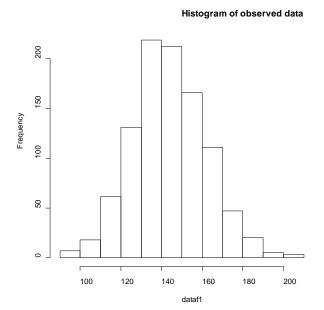
R code extract

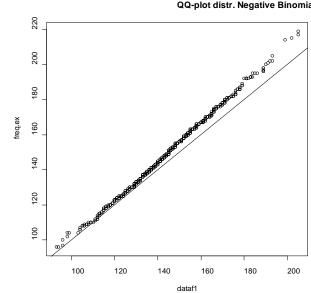
draw histogram hist(dataf1,main="Histogram of observed data")

OOPlot

freq.ex<-(rnbinom(n=1000,size=100,prob=0.4)) ggplot(dataf1,freg.ex,main="QQ-plot distr. Negative Binomial") abline(0,1) ## a 45-degree reference line is plotted

Histogram and QQ plot







Negative Binomial Frequency Testing

• Goodness of fit test - Pearson's χ^2

	χ^{2}	p value
Pearson	197.4	0.64

Maximum likelihood (ML) estimation

	size	μ
Estimation	117.2	144.2
S.D.	9.5	0.57

	Model Assumption	ML estimation
Size	100	117
Prob.	0.4	0.448
Mean (µ) 150	144.2
Variance	375	321.5

R code extract

Goodness of fit test library(vcd) #load package vcd gf<-goodfit(dataf1,type="nbinom",par=list(size=100,prob=0.4))

Maximum likelihood estimation gf<-goodfit(dataf1,type= "nbinom",method= "ML") fitdistr(dataf1, "Negative Binomial")

DAY ONE



Correlation

Correlation among frequencies of different lines

- Gaussian Copula
- Clayton Copula
- Frank Copula
- Gumbel Copula
- *t* Copula
- Correlation between claim size and report lag

- Gaussian Copula

- Clayton Copula
- Frank Copula
- Gumbel Copula
- t Copula

Use R package "copula"

Frequencies – Frank Copula

Gumbel Copula: $C_{\theta}^{n}(u) = -\frac{1}{\theta} \ln(1 + \frac{(e^{-\theta u_{1}} - 1)(e^{-\theta u_{2}} - 1)\cdots(e^{-\theta u_{n}} - 1)}{(e^{-\theta} - 1)^{n-1}} \qquad \theta > 0$

- Ui: marginal cumulative distribution function (CDF)
- C(u): joint CDF
- Frequencies simulation
 - Two Lines with annual frequency Poisson ($\lambda = 96$)
 - Monthly exposure: 1
 - Frequency Trend: 1
 - Seasonality: 1
 - Accident Year: 2000
 - Random Seed: 16807
 - Frequency correlation: $\Theta = 8$, n = 2
 - # of Simulations: 1000
- Test Method
 - Scatter plot
 - Goodness-of-fit test_

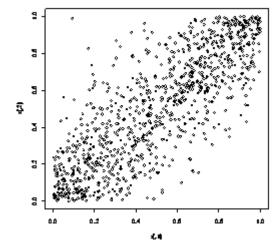
1. Parameter estimation based on maximum likelihood and inverse of Kendall's tau 2. Cramer-von Mises (CvM) statistic $S_n^{(k)} = \sum_{n=1}^{n} \{C_n^{(k)}(\widehat{U}_i^{(k)}) - C_{\theta_n}^{(k)}(\widehat{U}_i^{(k)})\}^2$

3. *p* value by parametric bootstrapping

Frequencies – Frank Copula

Scatter plot



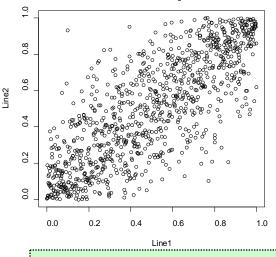


Goodness-of-fit test

Maximum Likelihood method

- Parameter estimate(s): 7.51
- Std. error: 0.28 CvM statistic: 0.016 with *p*-value 0.31
- Inversion of Kendall's tau method Parameter estimate(s): 7.54
 Std. error: 0.31
 CvM statistic: 0.017 with *p*-value 0.20

Simulated Frequencies



R code extract

construct a Gumbel copula object gumbel.cop <- gumbelCopula(3, dim=2)

parameter estimation

fit.gumbel<-fitCopula(gumbel.cop,x,method="ml") fit.gumbel<-fitCopula(gumbel.cop,x,method="itau")

#Copula Goodness-of-fit test

gofCopula(gumbel.cop, x, N=100, method = "mpl") gofCopula(gumbel.cop, x, N=100, method = "itau")

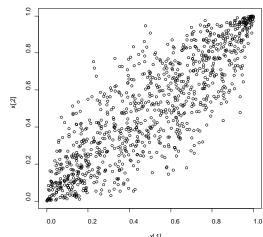
Claim Size and Report Lag – Normal Copula

Normal Copula a.k.a. Gaussian Copula: $C_{\Sigma}^{n}(u) = \Phi_{\Sigma}(\Phi^{-1}(u_{1}), \dots, \Phi^{-1}(u_{n}))$

- Σ : correlation matrix
- Φ : normal cumulative distribution function
- Claim simulation
 - One Line with annual frequency Poisson ($\lambda = 120$)
 - Monthly exposure: 1
 - Frequency Trend: 1.05
 - Seasonality: 1
 - Accident Year: 2000
 - Random Seed: 16807
 - Payment Lag: Exponential with rate = 0.00274, which implies a mean of 365 days.
 - Size of entire loss: Lognormal with $\mu = 11.17$ and $\sigma = 0.83$
 - Correlation between payment lag and size of loss: normal copula with correlation = 0.85, dimension 2
 - # of Simulations: 10

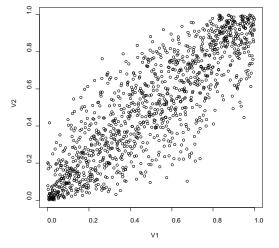
Claim Size and Report Lag – Normal Copula

Scatter plot



Normal Copula (0.85)

Simulated claim size vs. report lag



Goodness-of-fit test

Maximum Likelihood method
 Parameter estimate(s): 0.83
 Std. error: 0.01
 CvM statistic: 0.062 with *p*-value 0.05

Inversion of Kendall's tau method Parameter estimate(s): 0.85 Std. error: 0.01 CvM statistic: 0.029 with *p*-value 0.015

DAY THREE



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The LSM has two ways to model it

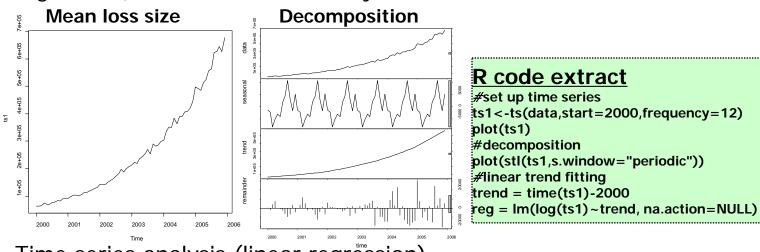
- Trend factor (cum) - α (Persistency of the force of the trend) $trend = (cum_{acc_date}) \left(\frac{cum_{pmt_date}}{cum_{acc_date}} \right)^{\alpha} = (cum_{acc_date})^{1-\alpha} (cum_{pmt_date})^{\alpha}$

• Trend factor Test Parameters

- One Line with annual frequency Poisson ($\lambda = 96$)
- Monthly exposure: 1
- Frequency Trend: 1
- Seasonality: 1
- Accident Year: 2000 to 2005
- Random Seed: 16807
- Size of entire loss: Lognormal with $\mu = 11.17$ and $\sigma = 0.83$
- Severity trend: 1.5
- # of Simulations: 300

Trend factor Test

 Decomposition of Time Series by Loess (Locally weighted regression) into trend, seasonality, and remainder



Time series analysis (linear regression)

Log(Mean Loss Size) = Intercept + trend * (time - 2000) + error term

Coefficients:

	Estimate	Std. Error	t value	Pr(> <i>t</i>)
(Intercept)	11.034162	0.007526	1466.1	<2 <i>e</i> -16
trend	0.405552	0.002196	184.7	<2 <i>e</i> -16
Residual standard error: 0.03226 on 70 degrees of freedom				
Multiple R-squared: 0.998, Adjusted R-squared: 0.9979				
<i>F</i> -statistic: 3.412e+04 on 1 and 70 DF, <i>p</i> -value: < 2.2 <i>e</i> -16				

exp(0.405552) = 1.50013 vs. model input 1.5

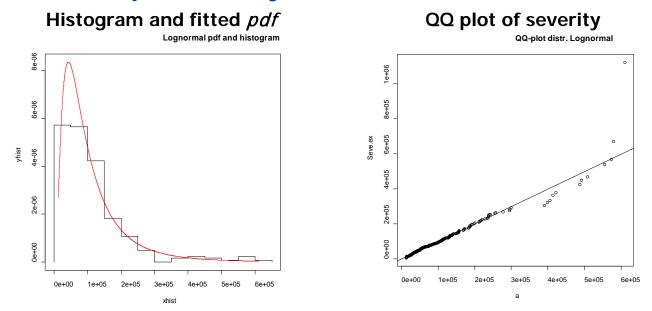
• Trend persistency α Test Parameters

- One Line with annual frequency Poisson ($\lambda = 96$)
- Monthly exposure: 1
- Frequency Trend: 1
- Seasonality: 1
- Accident Year: 2000 to 2001
- Random Seed: 16807
- Size of entire loss: Lognormal with m = 11.17 and s = 0.83
- Severity trend: 1.5
- Alpha = 0.4
- # of Simulations: 1000

But how do we test it?

Choose the loss payments with report date during the 1st month and payment date during the 7th month. The severity trend is $(1.5^{1/12})^{(1-0.4)} \cdot (1.5^{7/12})^{0.4} \approx 1.122$ The expected loss size is $1.122 \cdot e^{11.17+0.83^2/2} \approx 112,175$

• Trend persistency α Test



Maximum likelihood estimation (mean of severity=113,346)

	meanlog	sdlog
Estimation	11.32	0.80
Standard Deviation	0.052	0.037

Normality test of log (severity)
 Kolmogorov-Smirnov test: *p*-value = 0.82
 Anderson-Darling normality test: *p*-value = 0.34

R code extract

#Kolmogorov-Smirnov Tests
ks.test(a,"plnorm", meanlog=11.32,
 sdlog=0.8)
#Anderson-Darling Test
library(nortest) ## package loading
ad.test(datas1.norm)

DAY FOUR

9 AM



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Case Reserve Adequacy

In the LSM, the case reserve adequacy (CRA) distribution attempts to model the reserve process by generating case reserve adequacy ratio at each valuation date

- Case reserve = generated final claim amount \times case reserve adequacy ratio

Case Reserve Simulation

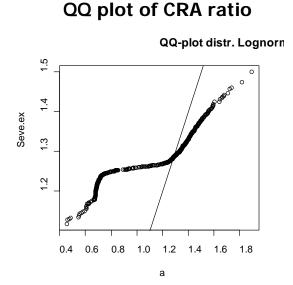
- One Line with annual frequency Poisson ($\lambda = 96$)
- Monthly exposure: 1
- Frequency Trend: 1
- Seasonality: 1
- Accident Year: 2000 to 2001
- Random Seed: 16807
- Size of entire loss: Lognormal with $\mu = 11.17$ and $\sigma = 0.83$
- Severity trend: 1
- P(0) = 0.4
- Est P(0) = 0.4
- # of Simulations: 8

Test 40% time point (60×report date + 40%×final payment date) case reserve adequacy ratio

```
Mean: e^{0.25+0.05^2/2} \approx 1.2856
```

Case Reserve Adequacy

Case Reserve Adequacy Test



- Maximum likelihood estimation meanlog sdlog Estimation 0.08 0.32 Standard Deviation 0.014 0.010
- Normality test of log (CRA ratio)
 Kolmogorov-Smirnov test: p-value = 0.00
 Anderson-Darling normality test: p-value = 0.00

Where went wrong?

case reserve is generated on the simulated valuation dates.

Linear interpolation method is used to get case reserve ratio at 40% time point.

On the report date, a case reserve of 2,000 is allocated for each claim.

If the second valuation date > 40% time point, linear interpolation method is not appropriate.

III. Real Data





Wait a minute Tom! I want you to think about how to use real claim data for model calibration during the weekend!



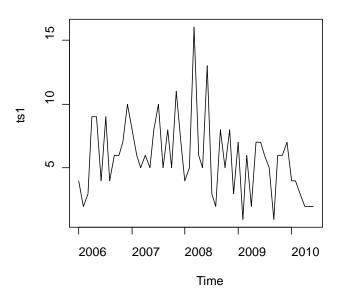
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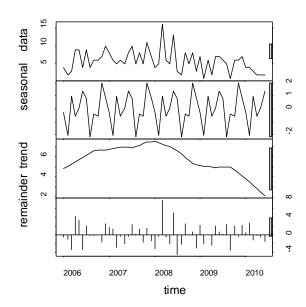
Marine claim data for distribution fitting, trend analysis, and correlation analysis

- two product lines: Property and Liability
- data period: 2006 2010
- accident date, payment date, and final payment amount
- Fit the frequency
 - Draw time series and decomposition

Historical Frequency

Decomposition



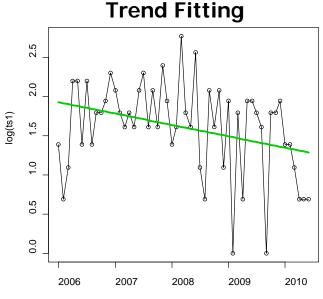


• Fit the frequency (continued)

- Linear regression for trend analysis

Log(Monthly Frequency) = Intercept + trend * (time – 2006) + error term <u>Coefficients:</u>

Estimate Std. Error t value Pr(>|t|)(Intercept) 1.93060 0.15164 12.732 <2e-16 trend -0.14570 0.05919 -2.462 0.0172 Residual standard error: 0.5649 on 52 degrees of freedom. Multiple R-squared: 0.1044, Adjusted R-squared: 0.08715. *F*-statistic: 6.06 on 1 and 52 DF, *p*-value: 0.01718.



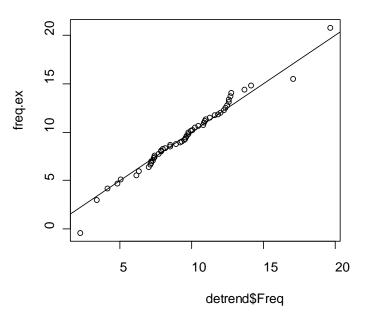
• Fit the frequency (continued)

 Detrend the frequency and fit to the lognormal distribution meanlog sdlog
 Estimation 9.5539259 3.1311762
 Standard Deviation 0.4260991 0.3012976

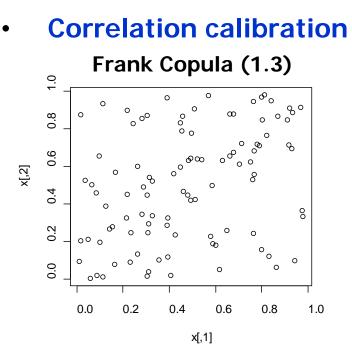
Normality test of log (detrended freq.)
 Kolmogorov-Smirnov test: p-value = 0.84

QQ plot of detrended freq.

QQ-plot distr. normal

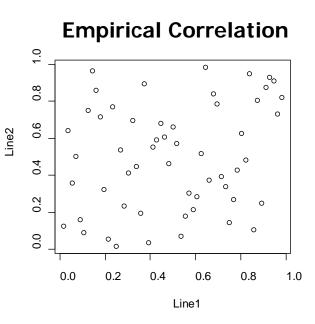


• Fit the Severity



Maximum Likelihood method
 Parameter estimate(s): 1.51
 CvM statistic: 0.027 with *p*-value 0.35

 Inversion of Kendall's tau method Parameter estimate(s): 1.34
 CvM statistic: 0.028 with *p*-value 0.40



What is missing? Historical reserve data which are essential for case reserve adequacy modeling.

IV. Model Enhancement

Sometimes the frequency and severity distribution are not stable over time

- Structural change
- Cyclical pattern
- Idiosyncratic character

•The model

- Two distinct distributions represent different states
- Transition rules from one state to another

 P_{11} : state 1 persistency, the probability that the state will be 1 next month given that it is 1 this month.

 P_{12} : the probability that the state will be 2 next month given that it is 1 this month.

 P_{21} : the probability that the state will be 1 next month given that it is 2 this month.

 P_{22} : state 2 persistency, the probability that the state will be 2 next month given that it is 2 this month.

 Π_1 : steady probability of state 1. Π_2 : steady probability of state 2.

$$\begin{pmatrix} \Pi_1 & \Pi_2 \end{pmatrix} \begin{pmatrix} P_{11} & P_{12} \\ P_{21} & P_{22} \end{pmatrix} = \begin{pmatrix} \Pi_1 & \Pi_2 \end{pmatrix}$$

 $P_{11} = 1 - P_{12}$

$$P_{21} = 1 - P_{22}$$
$$\Pi_1 + \Pi_2 = 1$$

The Simulation

- Steps
- 1. Generate uniform random number randf₀ on range [0,1].
- **2.** If rand $f_0 < \Pi_1$, state of first month state is 1, else, it is 2.
- 3. Generate uniform random number randf_i on range [0,1].
- 4. For previous month state I, if rand $f_i < P_{i1}$, then state is 1, else it is 2.
- 5. Repeat step 3 and 4 until the end of the simulation is reached.
- Test Parameters
- ✓ State 1: Poisson Distribution (λ = 120)
- \checkmark State 2: Negative Binomial Distribution (size = 36, prob = 0.5)
- ✓ Assume the trend, monthly exposure, and seasonality are all 1
- ✓ State 1 persistency: 0.5
- ✓ State 2 persistency: 0.7
- ✓ Seed: 16807

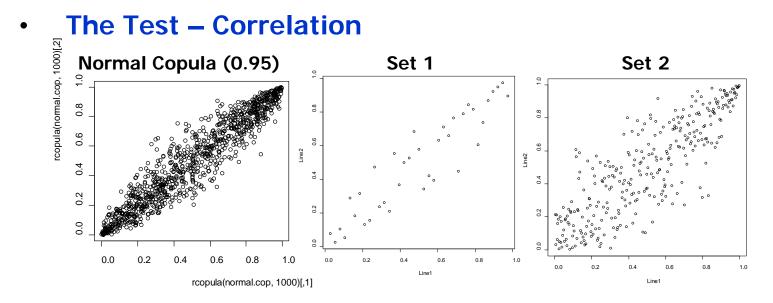
$$\Pi_{1} = \frac{1 - P_{22}}{2 - P_{11} - P_{22}} = \frac{1 - 0.7}{2 - 0.5 - 0.7} = 0.375$$
$$\Pi_{2} = \frac{1 - P_{11}}{2 - P_{11} - P_{22}} = \frac{1 - 0.7}{2 - 0.5 - 0.7} = 0.625$$

Random Number (RN)	State	Criteria
0.634633548790589	2	RN>0.375
0.801362191326916	1	RN>0.7
0.529508789768443	2	RN>0.5
0.0441845036111772	2	RN<0.7
0.994539848994464	1	RN>0.7
0.21886122901924	1	RN<0.5
0.0928565948270261	1	RN<0.5
0.797880138037726	2	RN>0.5
0.129500501556322	2	RN<0.7
0.24027365935035	2	RN<0.7
0.797712686471641	1	RN>0.7
0.0569291599094868	1	RN<0.5

• The Test – Transition Matrix

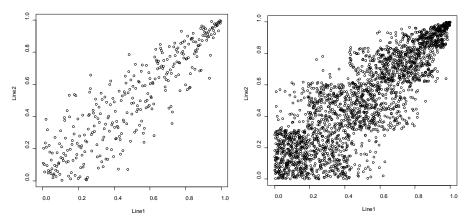
- Frequency State 1: Poisson ($\lambda = 120$); State 1 persistency: 0.2 State 2: Negative Binomial (size = 36, prob = 0.5); State 2 persistency: 0.9 Line 1 Frequency Line 2 Frequency $\begin{pmatrix} P_{11} & P_{12} \\ P_{21} & P_{22} \end{pmatrix} = \begin{pmatrix} 0.15 & 0.85 \\ 0.1 & 0.9 \end{pmatrix}$ $\begin{pmatrix} P_{11} & P_{12} \\ P_{21} & P_{22} \end{pmatrix} = \begin{pmatrix} 0.2 & 0.8 \\ 0.1 & 0.9 \end{pmatrix}$ $(\Pi_1 \quad \Pi_2) = (11.11\% \quad 88.89\%)$ $(\Pi_1 \quad \Pi_2) = (10.53\% \quad 89.47\%)$ Non Zero Cases: State 1: 391 State 1: 410 State 2: 2797 State 2: 2733 **Probability of Zero Cases:** State 1: 0.005% (*e*⁻¹⁰) State 1: 0.005% (*e*⁻¹⁰) State 2: 0.125 (prob³) State 2: 0.135 (e⁻²) Estimated all Cases: Non Zero Cases/ (1 – Probability of Zero Cases) State 1: 410 State 1: 391 State 2: 3188 (2797/(1-0.125)) State 2: 3161 (2733/(1-0.135)) Total Cases: # of simulations * 12 months = 3600

```
Steady-state probability (compared with P_1 \& P_2)State 1: 391/3600 = 10.86%State 1: 410/3600 = 11.4%State 2: 1-10.86% = 89.14%State 2: 1-11.4% = 88.6%
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Set 3





Set 1: State 1 for line 1 and state 1 for line 1 Set 2: State 1 for line 1 and state 2 for line 2

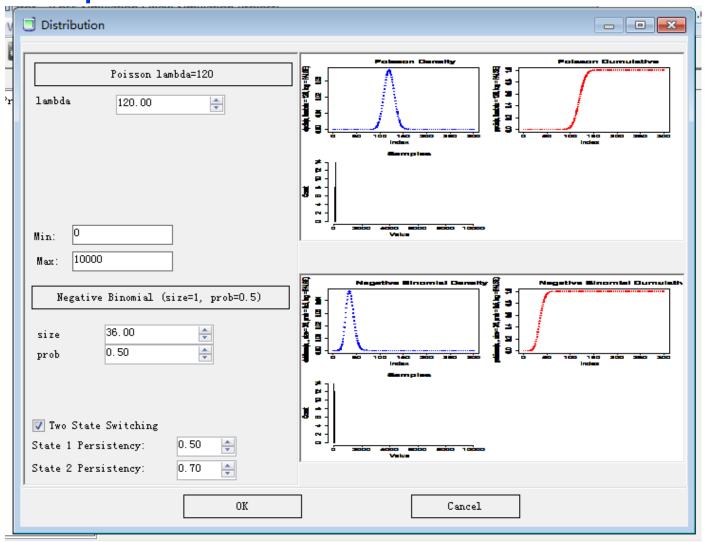
Set 3: State 2 for line 1 and state 1 for line 1

Set 4: State 2 for line 2 and state 2 for line 2

Goodness-of-fit test is also conducted.

Interface

• Input



Interface

Output

- Additional column in claim and transaction output files to record the state
- Showing state and random number while simulating

 Start Simulation 	↓ 0KB/S ↑	0KB/S	×
Sunmary Claims Loss Triangles			
Simulation Project New Simulation Project: Number of Iterations: 1 Start Date: 2000/1/1 0:00:00 Frequency Correlation Copula: normal Correlation=c() Dim=1 // / / Kendom Seed: 16807 line: 1 month: 1 Frequency: Negative Binomial (size=3, prob=0.5) state: 2 prestate: 1 pres: 2 rand: 0.634633548790589 line: 1 month: 2 Frequency: Negative Binomial (size=3, prob=0.5) state: 2 prestate: 1 pres: 2 rand: 0.529508789768443 line: 1 month: 3 Frequency: Negative Binomial (size=3, prob=0.5) state: 2 prestate: 1 pres: 2 rand: 0.529508789768443 line: 1 month: 4 Frequency: Negative Binomial (size=3, prob=0.5) state: 2 prestate: 2 pres: 2 rand: 0.529508789768443 line: 1 month: 5 Frequency: Poisson lambda=10 state: 1 prestate: 2 pres: 1 rand: 0.994539848994464 line: 1 month: 5 Frequency: Poisson lambda=10 state: 1 prestate: 1 pres: 1 rand: 0.994539848994464 line: 1 month: 6 Frequency: Poisson lambda=10 state: 1 prestate: 1 pres: 1 rand: 0.928559848270261 line: 1 month: 8 Frequency: Negative Binomial (size=3, prob=0.5) state: 2 prestate: 1 pres: 2 rand: 0.797880138037726 line: 1 month: 8 Frequency: Negative Binomial (size=3, prob=0.5) state: 2 prestate: 1 pres: 2 rand: 0.129500501556322 line: 1 month: 10 Frequency: Negative Binomial (size=3, prob=0.5) state: 2 prestate: 2 pres: 2 rand: 0.24027365935035 line: 1 month: 11 Frequency: Poisson lambda=10 state: 1 prestate: 2 pres: 1 rand: 0.797712686471641 line: 1 month: 11 Frequency: Poisson lambda=10 state: 1 prestate: 2 pres: 1 rand: 0.797712686471641 line: 1 month: 11 Frequency: Poisson lambda=10 state: 1 prestate: 2 pres: 1 rand: 0.0586291599094868 	sistency: 1 state	: 2 rand:	
Progress:			
Claim Output File D:\LS\RS\TSS\co.csv 🔎 Transaction Output File: D:\LS\RS\TSS\to.csv			<u> </u>
Number of Iterations: 1 Run	Stop	C109	5 e
Ready			

THREE MONTHS LATER



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V. Further Development

Further Development

Case reserve adequacy test shows that the assumption is not consistent with simulation data.

This may be caused by the linear interpolation method used to derive 40% time point case reserve.

It is suggested revising the way in which valuation date is determined in the LSM. In addition to the simulated valuation dates based on the waiting-period distribution assumption as in the LSM, some deterministic time points can be added as valuation dates.

In the LSM, 0%, 40%, 70%, and 90% time-points, case reserve adequacy distribution can be input into the model. Therefore, 0%, 40%, 70% and 90% time points may be added as deterministic valuation dates.

Thank you!