

VR-5 Assessing & Improving the Quality of Reserve Ranges

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Presentation

- Hindsight Testing
- Potential Problems with Ranges
- The Role of Judgment



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Hindsight Testing

- Percentiles of outcome distributions lend themselves most easily to testing.
- Consider the following example of a modeled reserve distribution:

Probability Percentile	Expected Outcome (\$,000s)	Difference from Mean
10%	7,608	(2,392)
20%	8,300	(1,700)
30%	8,839	(1,161)
40%	9,326	(674)
50%	9,806	(194)
60%	10,310	310
70%	10,879	879
80%	11,584	1,584
90%	12,639	2,639
Mean	10,000	-

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Hindsight Testing

- Suppose the following quarter, the hindsight opinion of all accident periods previously analyzed increases by \$2M
- Is this consistent with our range?

Probability Percentile	Expected Outcome (\$,000s)	Difference from Mean
10%	7,608	(2,392)
20%	8,300	(1,700)
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80%	11,584	1,584
90%	12,639	2,639
Mean	10,000	-

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Hindsight Testing

- If we assume that our analysis is efficient (i.e. all available information is included, bias removed, etc.) then changes in our opinion over time should be independent, i.e. a random walk.
- This suggests that reserve variability should be proportional to the square root of time elapsed.
- Therefore a reasonable adjustment is to modify the full payout outcome variability by (1/avg payout length)

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Hindsight Testing

- Returning to our example, let's assume that the projected payout stream averages 3 years (12 quarters) for the runoff of our reserves.
- We would then expect the variability to be $(1/12)^5 = 0.289$ of what the complete runoff variability would be.

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Hindsight Testing

- Adding this adjustment in this way is not perfect, because we would also expect the skewness to increase as we shorten the time period, but it is a good start.
- A \$2M increase is definitely at the high end of expectations, in this case in the top 10% of outcomes

Probability Percentile	Expected Outcome (\$,000s)	Difference from Mean	Implied Quarterly Variance*
10%	7,608	(2,392)	(891)
20%	8,300	(1,700)	(491)
30%	8,839	(1,161)	(335)
40%	9,326	(674)	(195)
50%	9,806	(194)	(54)
60%	10,310	310	90
70%	10,879	879	254
80%	11,584	1,584	457
90%	12,639	2,639	782
Mean	10,000	-	

*Assumes Average Payout of 3 years

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Hindsight Testing

- Qualifying adjustments in such a way provides a useful way of normalizing reserve adjustments across time and analysis
- Results should be uniform, and correlation across calendar quarters should be zero
- Both of these can be tested

	Analysis Line		
	A	B	C
Cal Qtr 1	60-70%	50-60%	0-10%
Cal Qtr 2	50-60%	30-40%	10-20%
Cal Qtr 3	70-80%	20-30%	40-50%
Cal Qtr 4	50-60%	90-100%	20-30%
Cal Qtr 5	70-80%	50-60%	70-80%
Cal Qtr 6	50-60%	0-10%	30-40%
Cal Qtr 7	90-100%	70-80%	50-60%
Cal Qtr 8	70-80%	70-80%	60-70%
Cal Qtr 9	30-40%	70-80%	40-50%
Cal Qtr 10	90-100%	40-50%	10-20%

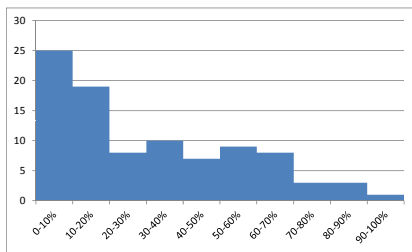
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Hindsight Testing

- If each percentile bin is of equal width, they each have the same expected number of observations. A Chi-Square test can be performed using:
 $\Sigma[(\text{Actual} - \text{Expected})^2 / \text{Expected}]$

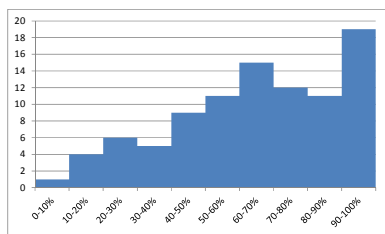
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Estimates Biased High



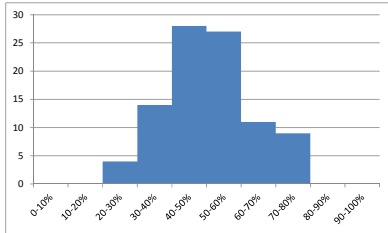
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Estimates Biased Low



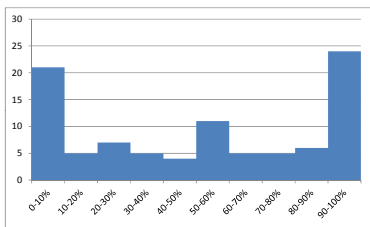
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Ranges too Wide



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Ranges too Narrow



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Potential Problems with Ranges

- Correlation
 - Between development periods
 - Between accident periods
 - Between methods
- Effect of skewed distributions
- Tail Factors

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Basic Data

- A typical triangle of losses will serve as an example

	Age1	Age2	Age3	Age4	Age5	Age6	Age7	Age8	Age9	Age10
2000	2,800,405	3,768,316	4,045,263	4,236,909	4,289,114	4,321,256	4,336,731	4,344,159	4,348,544	4,348,968
2001	2,608,018	3,573,193	3,591,428	4,194,788	4,256,829	4,286,767	4,313,103	4,317,923	4,318,432	
2002	2,551,421	3,596,466	3,973,491	4,158,497	4,212,343	4,240,059	4,254,540	4,260,569		
2003	2,307,563	3,293,323	3,260,122	3,368,538	3,430,394	3,446,014	3,454,098			
2004	2,220,241	3,076,331	3,314,683	3,437,104	3,481,601	3,498,406				
2005	2,518,198	3,482,141	3,732,331	3,867,741	3,910,920					
2006	2,624,262	3,596,594	3,877,256	3,993,855						
2007	2,823,468	3,872,470	4,126,395							
2008	2,443,839	3,176,659								
2009	1,954,178									
Wtd Average	1.966142	1.084493	1.041331	1.012783	1.006244	1.003914	1.001416	1.000565	1.000098	
Age to Ult	1.581721	1.157801	1.067596	1.025223	1.012283	1.006002	1.00208	1.000663	1.000098	

- For this illustration we will assume development is finished at ten years, and certain in the tenth year

Variability of Development Factors

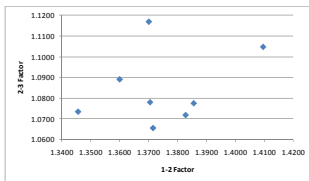
	1-2	2-3	3-4	4-5	5-6	6-7	7-8	8-9	9-10
2000	1.3456	1.0735	1.0474	1.0123	1.0075	1.0036	1.0017	1.0010	1.0001
2001	1.3701	1.1170	1.0509	1.0148	1.0070	1.0061	1.0011	1.0001	
2002	1.4096	1.1048	1.0466	1.0129	1.0066	1.0034	1.0014		
2003	1.3600	1.0891	1.0394	1.0124	1.0047	1.0022			
2004	1.3856	1.0775	1.0369	1.0129	1.0048				
2005	1.3828	1.0718	1.0363	1.0112					
2006	1.3705	1.0780	1.0301						
2007	1.3745	1.0656							
2008	1.2999								

SD of Age-Age Factors 0.030533 0.017823 0.007437 0.001191 0.001279 0.001667 0.000298 0.00063
 SD of Age-Ultimate Factors 0.045487 0.020948 0.008078 0.002561 0.002231 0.001812 0.000698 0.00063
 (assuming independence)

- Combination of variability factors without making a distributional assumption, but assuming independence

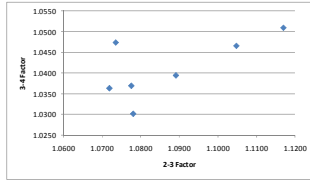
Independent?

- Scatter plots of paired historical development factors can illustrate lack of independence.
- Here the second development factor is shown against the first one.



Independent?

- ... and here the third development factor against the second



Correlation matrix

- A correlation matrix can be created from looking at such pairs:

Correlation Coefficient	1-2	2-3	3-4	4-5	5-6	6-7	7-8	8-9
1-2	1.000	0.305	(0.056)	0.038	(0.173)	0.047	(0.375)	(1.000)
2-3	0.305	1.000	0.668	0.836	0.277	0.655	(0.968)	(1.000)
3-4	(0.056)	0.668	1.000	0.705	0.924	0.902	(0.768)	(1.000)
4-5	0.038	0.836	0.705	1.000	0.319	0.931	(0.963)	(1.000)
5-6	(0.173)	0.277	0.924	0.319	1.000	0.642	0.501	1.000
6-7	0.047	0.655	0.902	0.931	0.642	1.000	(0.840)	(1.000)
7-8	(0.375)	(0.968)	(0.768)	(0.963)	0.501	(0.840)	1.000	1.000
8-9	(1.000)	(1.000)	(1.000)	(1.000)	1.000	(1.000)	1.000	1.000

- Unfortunately, statistical significance is often lacking due to limited data (ex. 8-9 factor)

Lognormal Distribution Model

- Not always usable, but helpful to illustrate effect of correlation on reserve variability

Ln(Age to Age Factor)	1-2	2-3	3-4	4-5	5-6	6-7	7-8	8-9	9-10	
2000	0.2969	0.0709	0.0463	0.0122	0.0075	0.0036	0.0017	0.0010	0.0001	
2001	0.3149	0.1307	0.0497	0.0147	0.0070	0.0061	0.0011	0.0001		
2002	0.3433	0.0987	0.0455	0.0129	0.0066	0.0034	0.0014			
2003	0.3075	0.0854	0.0386	0.0123	0.0047	0.0022				
2004	0.3261	0.0747	0.0362	0.0129	0.0048					
2005	0.3241	0.0694	0.0356	0.0111						
2006	0.3152	0.0751	0.0296							
2007	0.3159	0.0635								
2008	0.2623									
Average	0.3118	0.0812	0.0402	0.0127	0.0061	0.0038	0.0014	0.0006	0.0001	
Std Dev	0.0226	0.0163	0.0071	0.0032	0.0013	0.0017	0.0003	0.0006		
Sum of Ln(Factor)	0.4579 μ								$\text{Exp}(\mu + 5\sigma^2)$	1.5814
(Sum of SD(Ln(Factor))) ² * 5	0.0289 σ								$\text{Exp}(\mu - 5\sigma^2) * (\text{Exp}(\sigma^2) - 1) * 5$	0.0457
	Estimated 2008 AT SD (Assumes Independence)									89,277

Residual Terms

Difference of ln from Average	1-2	2-3	3-4	4-5	5-6	6-7	7-8	8-9
2000	(0.0149)	(0.0103)	0.0061	(0.0004)	0.0014	(0.0002)	0.0003	0.0004
2001	0.0031	0.0295	0.0095	0.0020	0.0009	0.0023	(0.0003)	(0.0004)
2002	0.0315	0.0185	0.0053	0.0002	0.0004	(0.0004)	0.0000	
2003	(0.0043)	0.0042	(0.0016)	(0.0004)	(0.0014)	(0.0017)		
2004	0.0143	(0.0065)	(0.0040)	0.0002	(0.0013)			
2005	0.0123	(0.0118)	(0.0046)	(0.0016)				
2006	0.0034	(0.0060)	(0.0106)					
2007	0.0041	(0.0177)						
2008	(0.0495)							

- These terms will be used to develop a covariance matrix

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Testing the hypothesis that $\rho=0$

- t statistic is useful:

$$r \cdot \sqrt{(n-2) / (1-r^2)} \sim t(n-2)$$

where r is the sample correlation coefficient and n is the number of observed pairs.

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Grouping by Lags

1 period apart

(0.0149)	(0.0103)	(0.0016)	(0.0004)	measured correlation
0.0031	0.0295	(0.0040)	0.0002	0.362136
0.0315	0.0185	(0.0046)	(0.0016)	number of observations
(0.0043)	0.0042	(0.0004)	0.0014	35
0.0143	(0.0065)	0.0020	0.0009	t-statistic
0.0123	(0.0118)	0.0002	0.0004	2.23
0.0034	(0.0060)	(0.0004)	(0.0014)	probability under null hypothesis
0.0041	(0.0177)	0.0002	(0.0013)	0.032535
(0.0103)	0.0061	0.0014	(0.0002)	
0.0295	0.0095	0.0009	0.0023	
0.0185	0.0053	0.0004	(0.0004)	
0.0042	(0.0016)	(0.0014)	(0.0017)	
(0.0065)	(0.0040)	(0.0002)	0.0003	
(0.0118)	(0.0046)	0.0023	(0.0003)	
(0.0060)	(0.0106)	(0.0004)	0.0000	
0.0061	(0.0004)	0.0003	0.0004	
0.0095	0.0020	(0.0003)	(0.0004)	
0.0053	0.0002			

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Correlation by Lag

- Repeating with other lags, results are summarized below:

lag	measured correlation	probability under null	selected correlation
1	0.362	0.0325	0.3621
2	0.076	0.7071	0.0000
3	0.195	0.3972	0.0000
4	0.202	0.4896	0.0000
5	-0.061	0.8767	0.0000
6	-0.698	0.1903	0.0000

- May wish to add reasonable constraints such as that the correlation for a given lag is no greater than a smaller lag.

Correlation Matrix

- Build selected correlation matrix

Correlation Matrix	1-2	2-3	3-4	4-5	5-6	6-7	7-8	8-9
1-2	1.0000	0.3621	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
2-3	0.3621	1.0000	0.3621	0.0000	0.0000	0.0000	0.0000	0.0000
3-4	0.0000	0.3621	1.0000	0.3621	0.0000	0.0000	0.0000	0.0000
4-5	0.0000	0.0000	0.3621	1.0000	0.3621	0.0000	0.0000	0.0000
5-6	0.0000	0.0000	0.0000	0.3621	1.0000	0.3621	0.0000	0.0000
6-7	0.0000	0.0000	0.0000	0.0000	0.3621	1.0000	0.3621	0.0000
7-8	0.0000	0.0000	0.0000	0.0000	0.0000	0.3621	1.0000	0.3621
8-9	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.3621	1.0000

Covariance Matrix

- Build the covariance matrix using $\sigma_{i,j} = \rho_{i,j}\sigma_i\sigma_j$

Covariance Matrix	0.02259	0.01633	0.00715	0.00118	0.00127	0.00166	0.00030	0.00063
0.02259	0.00051	0.000134	0	0	0	0	0	0
0.01633	0.000134	0.000267	4.23E-05	0	0	0	0	0
0.00715	0	4.23E-05	5.11E-05	3.04E-06	0	0	0	0
0.00118	0	0	3.04E-06	1.38E-06	5.41E-07	0	0	0
0.00127	0	0	0	5.41E-07	1.62E-06	7.64E-07	0	0
0.00166	0	0	0	0	7.64E-07	2.76E-06	1.79E-07	0
0.00030	0	0	0	0	0	1.79E-07	8.83E-08	6.782E-08
0.00063	0	0	0	0	0	0	6.78E-08	3.97E-07

- Summing elements of sections of this matrix provides σ^2 for the age-to-ultimate lognormal.
- Highlighted area represents variability from age three.

Impact of Correlation

Year	Age	μ	σ	mean = $\exp(\mu - 5\sigma^2)$	st dev = $\text{mean} * (\exp(\sigma^2) - 1)^{1/2}$	Current Loss	Ultimate	Estimated Standard Deviation of Reserve and Independence of Development	
2001	9	0.0001	1.0001	0.0000	0.0000	4,318,432			
2002	8	0.0007	3.97E-07	1.0007	0.0006	4,260,569	2,686	2,686	
2003	7	0.0021	6.21E-07	1.0021	0.0008	3,454,098	2,727	2,411	
2004	6	0.0059	3.73E-06	1.0059	0.0019	3,498,406	6,800	6,335	
2005	5	0.0120	6.88E-06	1.0121	0.0027	3,910,920	10,380	8,723	
2006	4	0.0247	9.34E-06	1.0250	0.0031	3,993,855	12,513	10,226	
2007	3	0.0649	6.65E-05	1.0671	0.0087	4,126,395	35,903	33,329	
2008	2	0.1461	0.000418	1.1375	0.0237	3,176,659	75,152	66,383	
2009	1	0.4579	0.001195	1.5817	0.0547	1,954,178	106,884	89,277	
Combination of Accident Years (Assuming Independence)							136,698	117,026	

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Accident Year Correlation

- Residual terms from applying age to age factors to history to project historical losses one period forward and subtract actual loss

57,435	41,451	(24,451)	1,954	(5,361)	1,438	(1,286)	(1,931)
(10,270)	(116,328)	(38,390)	(8,421)	(3,358)	(9,558)	1,289	1,931
(110,862)	(73,148)	(20,777)	(690)	(1,414)	2,114	(3)	
13,505	(13,883)	6,328	1,458	5,199	6,006		
(43,166)	21,377	14,787	(562)	4,934			
(41,525)	44,028	18,851	6,261				
(11,479)	23,226	43,652					
(15,211)	73,273						
161,972							

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Accident Year Correlation

- Correlation across accident years can be observed within a calendar period


57,435	41,451	(24,451)	1,954	(5,361)	1,438	(1,286)	(1,931)
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Accident Year Correlation

- Correlation across accident years can be observed within a calendar period

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(10,270)	(116,324)	(38,390)	(8,421)	(3,358)	(9,558)	1,289	1,931
(149,862)	(73,148)	(20,777)	(690)	(1,414)	2,114	(3)	
13,505	(13,883)	6,328	1,458	5,199	6,006		
(43,166)	21,377	14,787	(562)	4,934			
(41,925)	44,028	18,851	6,261				
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Accident Year Correlation

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57,435	41,451	(24,451)	1,954	(5,361)	1,438	(1,286)	(1,931)
(10,270)	(116,324)	(38,390)	(8,421)	(3,358)	(9,558)	1,289	1,931
(149,862)	(73,148)	(20,777)	(690)	(1,414)	2,114	(3)	
13,505	(13,883)	6,328	1,458	5,199	6,006		
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(15,211)	73,273						
161,972							




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Grouping by Lags

Lag = 1					measured correlation
(10,270)	41,451	14,787	1,458		0.65144
(110,862)	(116,324)	18,851	(562)		
13,505	(73,148)	43,652	6,261		
(43,166)	(13,883)	(8,421)	(5,361)		number of observations
(41,925)	21,377	(690)	(3,358)		35
(11,479)	44,028	1,458	(1,414)		
(15,211)	23,226	(562)	5,199		t-statistic
161,972	73,273	6,261	4,934		4.93
(116,324)	(24,451)	(3,358)	1,438		
(73,148)	(38,390)	(1,414)	(9,558)		probability under null hypothesis
(13,883)	(20,777)	5,199	2,114		2.25E-05
21,377	6,328	4,934	6,006		
44,028	14,787	(9,558)	(1,286)		
23,226	18,851	2,114	1,289		
73,273	43,652	6,006	(3)		
(38,390)	1,954	1,289	(1,931)		
(20,777)	(8,421)	(3)	1,931		
6,328	(690)				




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Comparison of Indications


- Simple method sometimes used -Standard Deviation of Indications
- Implicit assumptions that each method error has the same standard deviation σ , and that the method errors are uncorrelated
- Let L_i be the indicated ultimate loss for method i .

$$S = \frac{\sum_{i=1}^n (L_i - \bar{L})^2}{n-1}$$

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Independent Indications

- $E(S) = \sigma$
- If the selected loss is a straight average of the individual indications, the final estimate would actually have a standard deviation of $\sigma/n^{.5}$ with an estimate provided by $S/n^{.5}$


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Correlated Indications

- $E(S) = \sigma \sqrt{1 - \text{Avg}(\rho_{ij} | i \neq j)}$
- Example of indication correlation:

	Incurred			
	Paid LDF	LDF	Paid BF	Incurred BF
Paid LDF	1	0.5	0.75	0.4
Incurred LDF	0.5	1	0.5	0.85
Paid BF	0.75	0.5	1	0.65
Incurred BF	0.4	0.85	0.65	1

- $S = 0.63 * \sigma$

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Correlated Indications

- Additionally the standard deviation of a straight average of the indications is different:

$$\sigma \sqrt{\frac{\sum A_i^2}{n^2}}$$

- In our example this is 0.70σ

	True Amount	Estimate (Expected Value)
Individual Method	σ	0.63σ
Average	0.70σ	$0.63 * s / 2 = .37 \sigma$

Problems with Skewed Distributions

- Simple example – Small excess general liability triangle with limited or no losses to date
- You know the potential for significant losses is there, but it so far it hasn't happened.
- A priori loss ratios and development factors are being used to develop a B-F reserve estimate.
- **Don't** base your estimate of reserve variability on historical deviations from the **historical** mean
- Base your estimate of reserve variability on historical deviations from the **expected** mean

Tail Factors

- Large tail factors represent significant uncertainty.
- Even if these factors are based on historical observation of larger triangles, the data is old and the environment has changed.
- Estimating the standard deviation of tail development by comparison with similarly sized factors earlier in the development pattern is likely to understate the variability.
- Consider making adjustments, for example assuming that tail development is perfectly correlated with the development that comes before it.

The Role for Judgment

- Informed judgment plays an important role in most actuarial analyses.
- Most actuaries are comfortable with using judgment when selecting reserve estimates (i.e. means)
- Since estimating reserve variability can be challenging, judgment is important here as well.
- How have you seen loss development estimates change in the past? How comfortable do you feel about the analysis including, any inputs such as reference patterns, etc.
- Put yourself in the position of someone personally guaranteeing the results to remove the process from being purely technical.

Summary

- Our ranges are meaningless unless we hold ourselves accountable for them
- Hindsight testing of ranges can yield significant benefits, particularly over time
- The implicit assumptions that we make when develop our ranges (such as independence) can be problematic
- Judgment should play an important role in estimating ranges
