

PL-2
**The Matrix Inverted:
A Primer in GLM Theory**

**2005 CAS Seminar on
Ratemaking**

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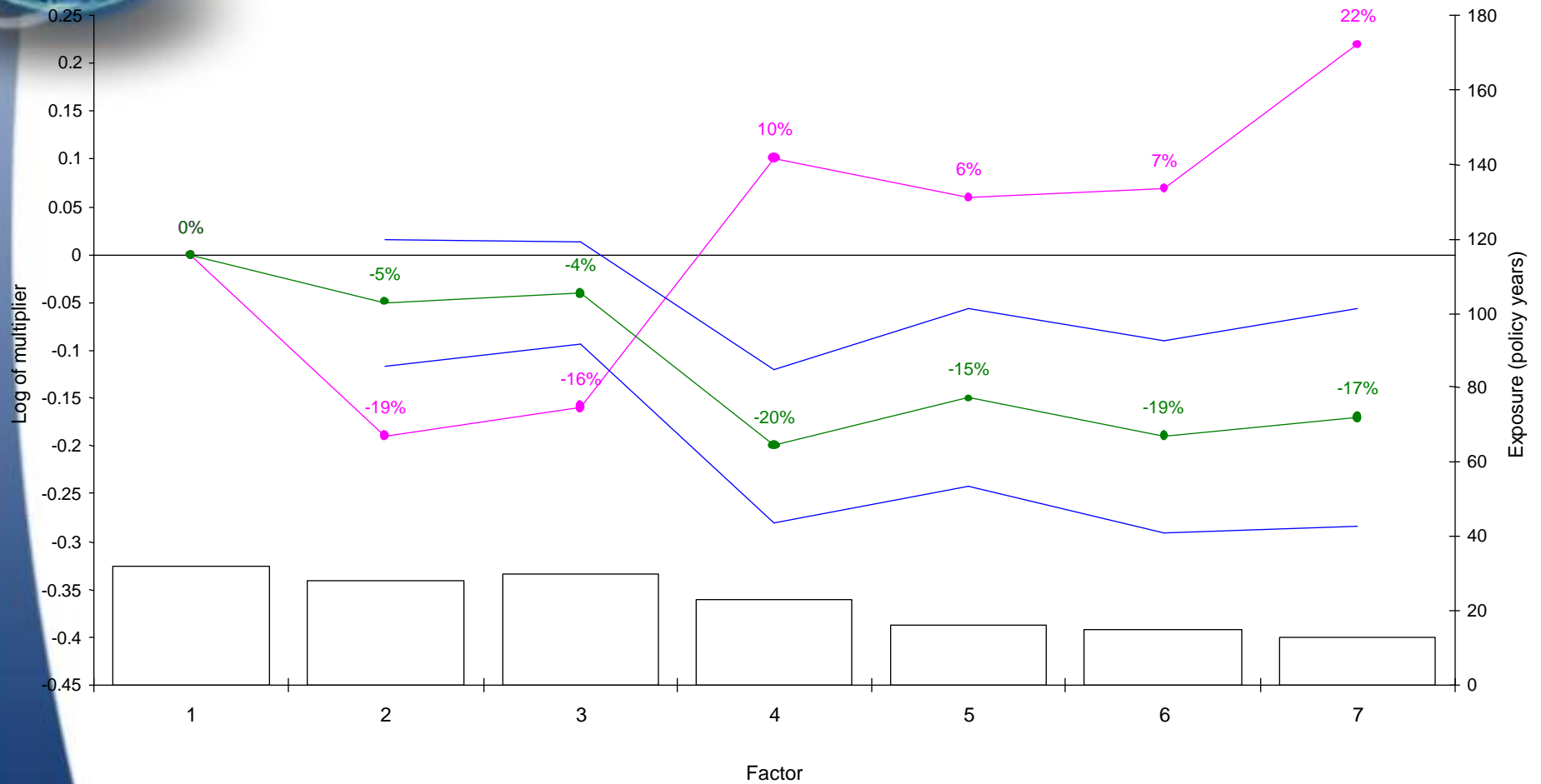


Generalized linear model benefits

- Consider all factors simultaneously
- Allow for nature of random process
- Provide diagnostics
- Robust and transparent



Example of GLM output (real UK data)



Exposure
 Oneway relativities
 Approx 2 SE from estimate
 GLM estimate





Agenda

- Formularization of GLMs
 - linear predictor, link function, offset
 - error term, scale parameter, prior weights
 - typical model forms
- Model testing
 - use only variables which are predictive
 - make sure model is reasonable
- Aliasing





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Linear models

- Linear model $Y_i = \mu_i + \text{error}$
- μ_i based on linear combination of measured factors
- Which factors, and how they are best combined is to be derived

$$\mu_i = \alpha + \beta \cdot \text{age}_i + \gamma \cdot \text{age}_i^2 + \delta \cdot \text{height}_i \cdot \text{age}_i$$



$$\mu_i = \alpha + \beta \cdot \text{age}_i + \gamma \cdot (\text{sex}_i = \text{female})$$



$$\mu_i = (\alpha + \beta \cdot \text{age}_i) * \exp(\delta \cdot \text{height}_i \cdot \text{age}_i)$$





Linear models - formularization

$$E[Y_i] = \mu_i = \sum X_{ij} \beta_j$$

$$\text{Var}[Y_i] = \sigma^2$$

$$Y_i \sim N(\mu_i, \sigma^2)$$





What is $\sum X_{ij}\beta_j$?

- **X** defines the explanatory variables to be included in the model
 - could be continuous variables - "variates"
 - could be categorical variables - "factors"
- $\underline{\beta}$ contains the parameter estimates which relate to the factors / variates defined by the structure of **X**
 - "the answer"





What is $\mathbf{X} \cdot \underline{\beta}$?

- Write $\sum X_{ij}\beta_j$ as $\mathbf{X} \cdot \underline{\beta}$
- Consider 3 rating factors
 - age of driver ("age")
 - sex of driver ("sex")
 - age of vehicle ("car")
- Represent $\underline{\beta}$ by $\alpha, \beta, \gamma, \delta, \dots$



What is $X\beta$?

- Suppose we wanted a model of the form:

$$\underline{\mu} = \alpha + \beta.\underline{age} + \gamma.\underline{age}^2 + \delta.\underline{car}^{27}.\underline{age}^{52\frac{1}{2}}$$

- $X\beta$ would need to be defined as:

$$\begin{pmatrix} 1 & age_1 & age_1^2 & car_1^{27}.age_1^{52\frac{1}{2}} \\ 1 & age_2 & age_2^2 & car_2^{27}.age_2^{52\frac{1}{2}} \\ 1 & age_3 & age_3^2 & car_3^{27}.age_3^{52\frac{1}{2}} \\ 1 & age_4 & age_4^2 & car_4^{27}.age_4^{52\frac{1}{2}} \\ 1 & age_5 & age_5^2 & car_5^{27}.age_5^{52\frac{1}{2}} \\ \dots & \dots & \dots & \dots \\ \dots & \dots & \dots & \dots \end{pmatrix} \cdot \begin{pmatrix} \alpha \\ \beta \\ \gamma \\ \delta \end{pmatrix}$$





What is $X.\beta$?

- Suppose we wanted a model of the form:

$$\underline{\mu} = \alpha + \beta_1 \text{ if } \underline{\text{age}} < 30$$

$$+ \beta_2 \text{ if } \underline{\text{age}} \text{ 30 - 40}$$

$$+ \beta_3 \text{ if } \underline{\text{age}} > 40$$

$$+ \gamma_1 \text{ if } \underline{\text{sex}} \text{ male}$$

$$+ \gamma_2 \text{ if } \underline{\text{sex}} \text{ female}$$





What is $X \cdot \beta$?

	Age			Sex	
	<30	30-40	>40	M	F
1	1	0	1	0	0
2	1	1	0	0	0
3	1	1	0	0	1
4	1	0	0	1	0
5	1	0	1	0	1
				
				

$$\begin{pmatrix} \alpha \\ \beta_1 \\ \beta_2 \\ \beta_3 \\ \gamma_1 \\ \gamma_2 \end{pmatrix} \cdot$$





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$$+ \gamma_1 \text{ if } \underline{\text{sex}} \text{ male}$$

$$+ \gamma_2 \text{ if } \underline{\text{sex}} \text{ female}$$





What is $X.\beta$?

- Suppose we wanted a model of the form:

$$\underline{\mu} = \alpha + \beta_1 \text{ if } \underline{\text{age}} < 30$$

$$+ \beta_2 \text{ if } \underline{\text{age}} 30 - 40$$

"Base levels" $+ \beta_3 \text{ if } \underline{\text{age}} > 40$

$$+ \gamma_1 \text{ if } \underline{\text{sex}} \text{ male}$$

$$+ \gamma_2 \text{ if } \underline{\text{sex}} \text{ female}$$





X.β having adjusted for base levels

		Age				Sex	
		<30	30-40	>40	M	F	
1	1	0	1	0	1	0	
2	1	1	0	0	1	0	
3	1	1	0	0	0	1	
4	1	0	0	1	1	0	
5	1	0	1	0	0	1	
						
						

β

α
β_1
β_2
β_3
γ_1
γ_2





Linear models - formularization

$$E[Y_i] = \mu_i = \sum X_{ij} \beta_j$$

$$\text{Var}[Y_i] = \sigma^2$$

$$Y_i \sim N(\mu_i, \sigma^2)$$





Generalized linear models

$$\mu_i = f(\alpha + \beta \cdot \text{age}_i + \gamma \cdot \text{age}_i^2 + \delta \cdot \text{height}_i \cdot \text{age}_i)$$

$$\mu_i = f(\alpha + \beta \cdot \text{age}_i + \gamma \cdot (\text{sex}_i = \text{female}))$$





Generalized linear models

$$\mu_i = g^{-1}(\alpha + \beta \cdot \text{age}_i + \gamma \cdot \text{age}_i^2 + \delta \cdot \text{height}_i \cdot \text{age}_i)$$

$$\mu_i = g^{-1}(\alpha + \beta \cdot \text{age}_i + \gamma \cdot (\text{sex}_i = \text{female}))$$





Generalized linear models

Linear Models

$$E[Y_i] = \mu_i = \sum X_{ij}\beta_j$$

$$\text{Var}[Y_i] = \sigma^2$$

Y from
Normal distribution

Generalized Linear Models

$$E[Y_i] = \mu_i = g^{-1}(\sum X_{ij}\beta_j + \xi_i)$$

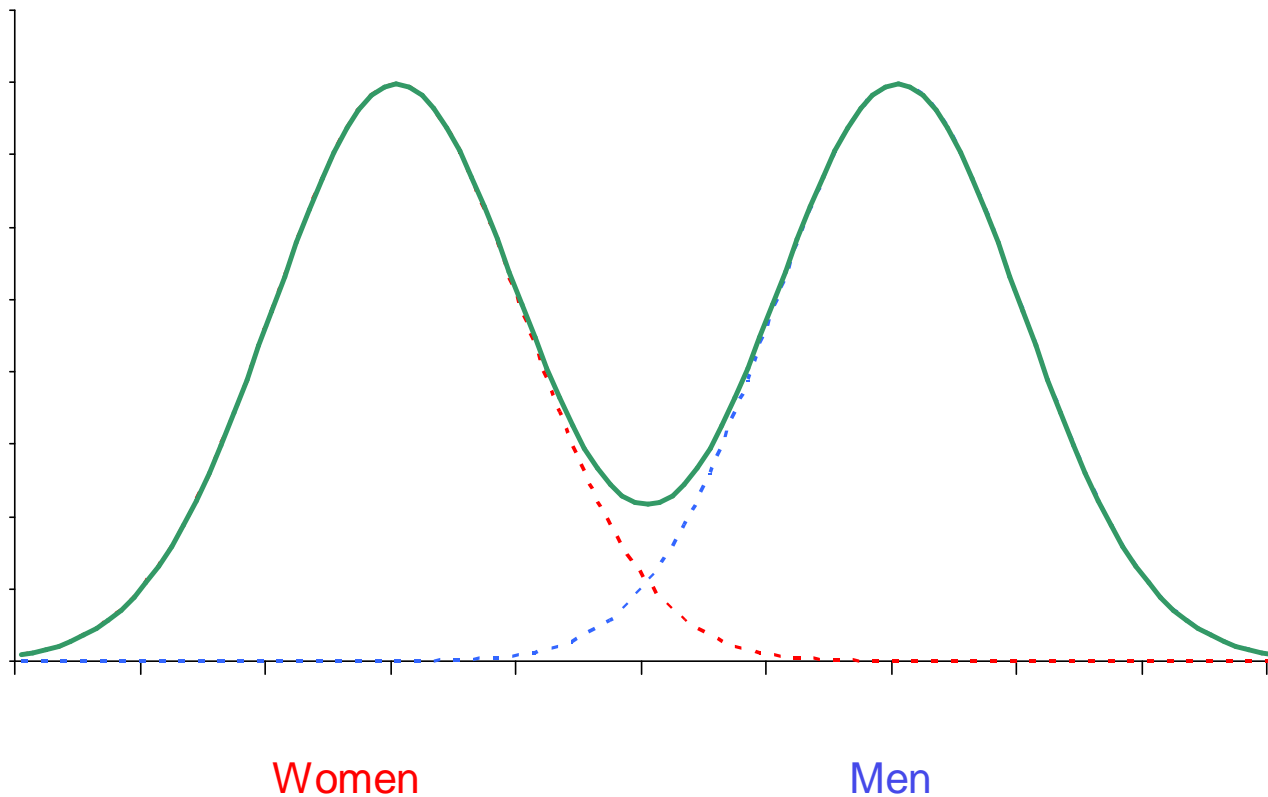
$$\text{Var}[Y_i] = \phi V(\mu_i)/\omega_i$$

Y from a distribution from the
exponential family



Generalized linear models

- Each observation i from distribution with mean μ_i





Generalized linear models

$$E[Y_i] = \mu_i = g^{-1}(\sum_j X_{ij}\beta_j + \xi_i)$$

$$\text{Var}[Y_i] = \phi \cdot V(\mu_i) / \omega_i$$





Generalized linear models

$$E[\underline{Y}] = \underline{\mu} = g^{-1}(\underline{X} \cdot \underline{\beta} + \underline{\xi})$$

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$





Generalized linear models

$$E[\underline{Y}] = \underline{\mu} = g^{-1}(\underline{X} \cdot \underline{\beta})$$

Observed thing
(data)

Some function
(user defined)

Some matrix based on data
(user defined)
as per linear models

Parameters to be
estimated
(the answer!)





What is $g^{-1}(\mathbf{X}.\underline{\beta})$?

$$\underline{Y} = g^{-1}(\mathbf{X}.\underline{\beta}) + \text{error}$$

Assuming a model with three categorical factors,
each observation can be expressed as:

$$Y_{ijk} = g^{-1}(\alpha + \beta_i + \gamma_j + \delta_k) + \text{error}$$

$$\beta_2 = \gamma_1 = \delta_3 = 0$$

age is in group i

sex is in group j

car is in group k





What is $g^{-1}(X, \underline{\beta})$?

- $g(x) = x \quad \Rightarrow \quad Y_{ijk} = \alpha + \beta_i + \gamma_j + \delta_k + \text{error}$
- $g(x) = \ln(x) \quad \Rightarrow \quad Y_{ijk} = e^{(\alpha + \beta_i + \gamma_j + \delta_k)} + \text{error}$
 $= A \cdot B_i \cdot C_j \cdot D_k + \text{error}$
where $B_i = e^{\beta_i}$ etc
- Multiplicative form common for frequency and amounts





Multiplicative model

\$207.10 x

Age	Factor
17	2.52
18	2.05
19	1.97
20	1.85
21-23	1.75
24-26	1.54
27-30	1.42
31-35	1.20
36-40	1.00
41-45	0.93
46-50	0.84
50-60	0.76
60+	0.78

Group	Factor
1	0.54
2	0.65
3	0.73
4	0.85
5	0.92
6	0.96
7	1.00
8	1.08
9	1.19
10	1.26
11	1.36
12	1.43
13	1.56

Sex	Factor
Male	1.00
Female	1.25

Area	Factor
A	0.95
B	1.00
C	1.09
D	1.15
E	1.18
F	1.27
G	1.36
H	1.44

$$E(\text{losses}) = \$ 207.10 \times 1.42 \times 0.92 \times 1.00 \times 1.15 = \$ 311.14$$





Generalized linear models

$$E[\underline{Y}] = \underline{\mu} = g^{-1}(\underline{X} \cdot \underline{\beta} + \xi)$$

"Offset"



Eg \underline{Y} = claim *numbers*

Smith: Male, 30, Ford, 1 years, 2 claims

Jones: Female, 40, VW, 1/2 year, 1 claim





What is ξ ?

- $g(x) = \ln(x)$
- $\xi_{ijk} = \ln(\text{exposure}_{ijk})$
- $E[Y_{ijk}] = e^{(\alpha + \beta_i + \gamma_j + \delta_k + \xi_{ijk})}$
 $= A \cdot B_i \cdot C_j \cdot D_k \cdot e^{(\ln(\text{exposure}_{ijk}))}$
 $= A \cdot B_i \cdot C_j \cdot D_k \cdot \text{exposure}_{ijk}$





Restricted models

$$E[Y] = \underline{\mu} = g^{-1} (\mathbf{X} \cdot \underline{\beta} + \xi)$$

Offset 

- Constrain model (eg territory, ABS discount)
- Other factors adjusted to compensate



Generalized linear models

$$E[\underline{Y}] = \underline{\mu} = g^{-1}(\underline{X} \cdot \underline{\beta} + \underline{\xi})$$

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$





Generalized linear models

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$

Normal: $\phi = \sigma^2$, $V(x) = 1 \Rightarrow \text{Var}[\underline{Y}] = \sigma^2 \cdot \underline{1}$

Poisson: $\phi = 1$, $V(x) = x \Rightarrow \text{Var}[\underline{Y}] = \underline{\mu}$

Gamma: $\phi = k$, $V(x) = x^2 \Rightarrow \text{Var}[\underline{Y}] = k \underline{\mu}^2$





The scale parameter

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$

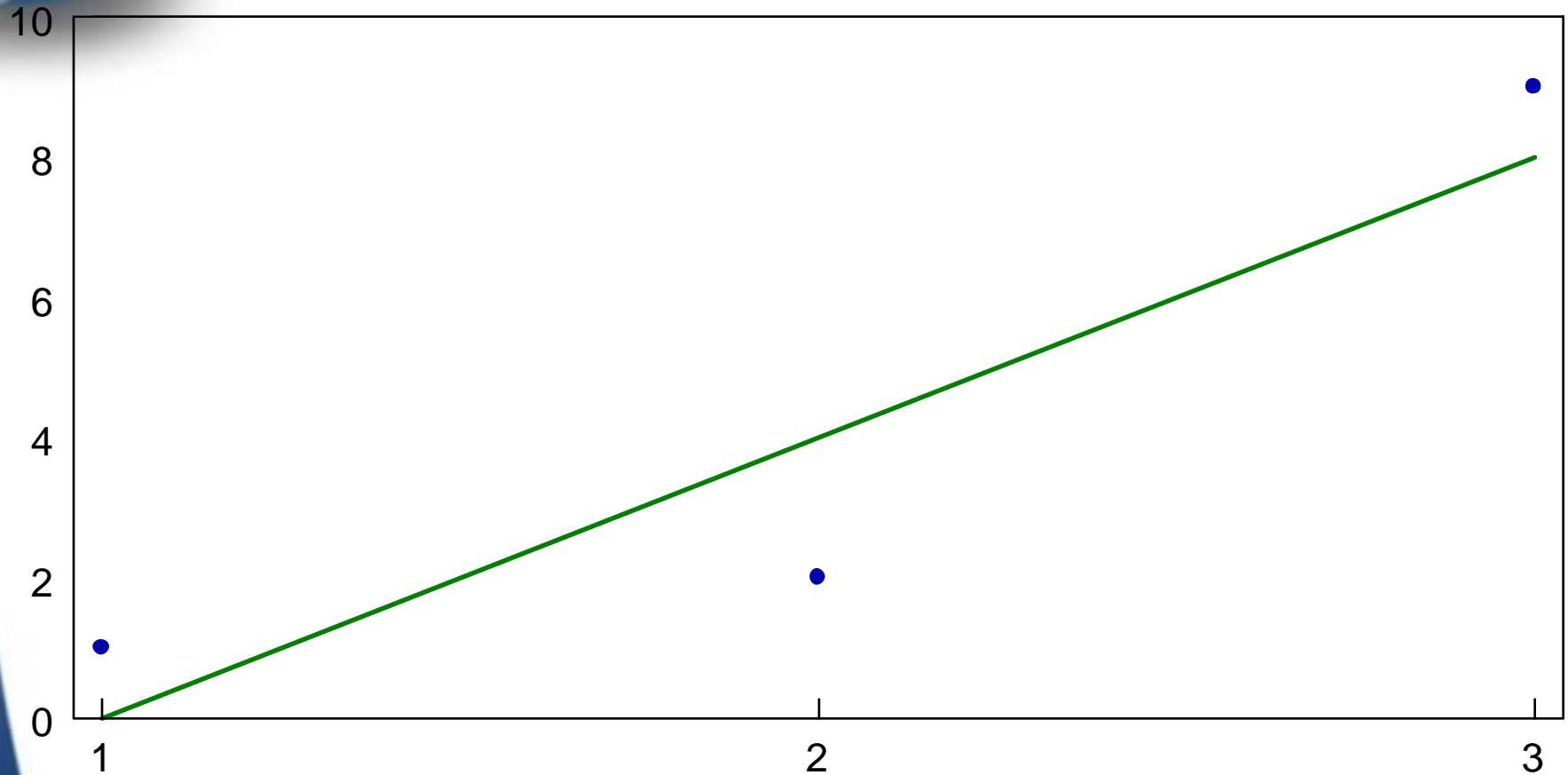
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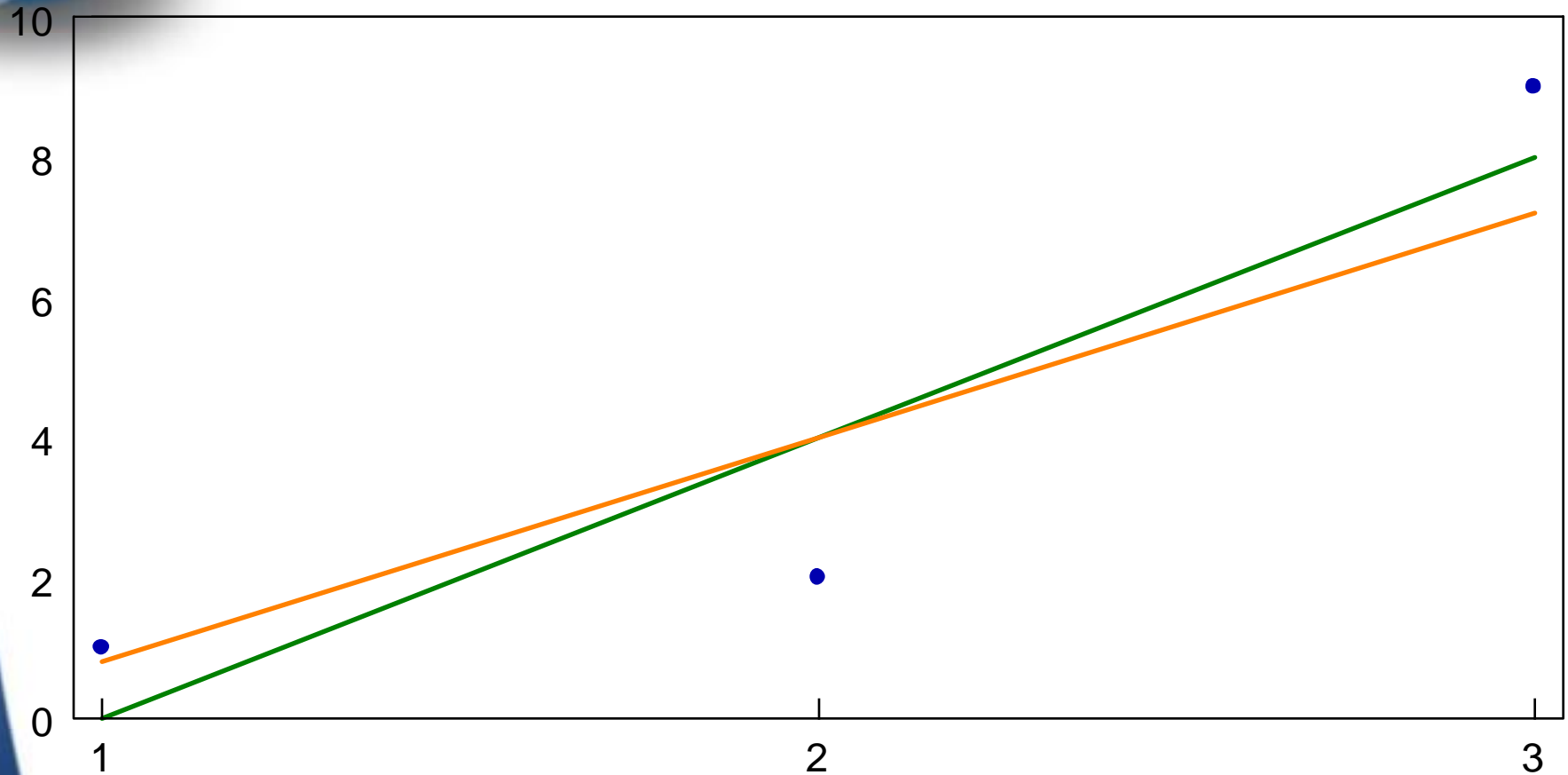
Example of effect of changing assumed error - 1



Data Normal
• —



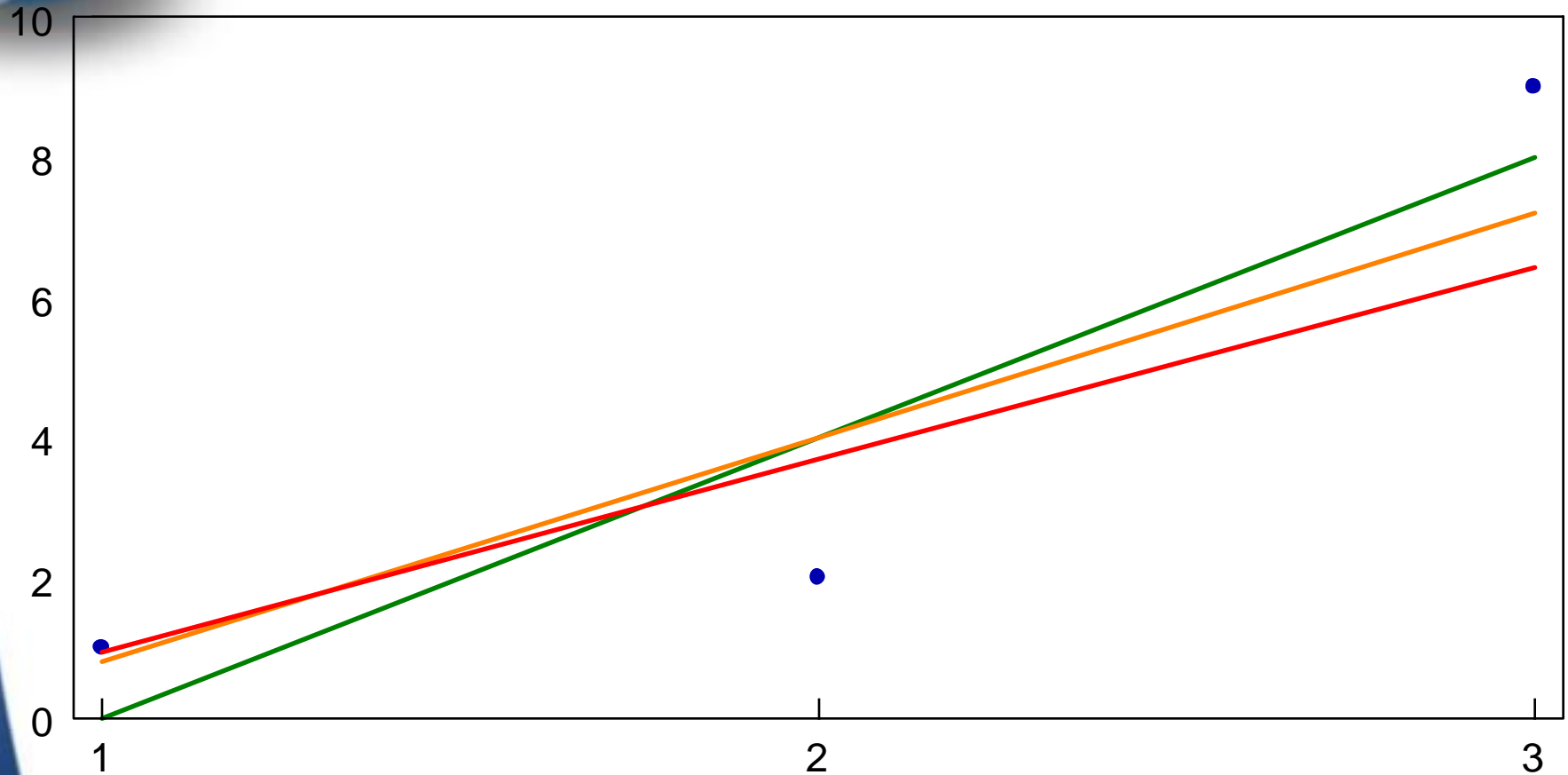
Example of effect of changing assumed error - 1



Data Normal Poisson
• — —



Example of effect of changing assumed error - 1



Data Normal Poisson Gamma
• — — —



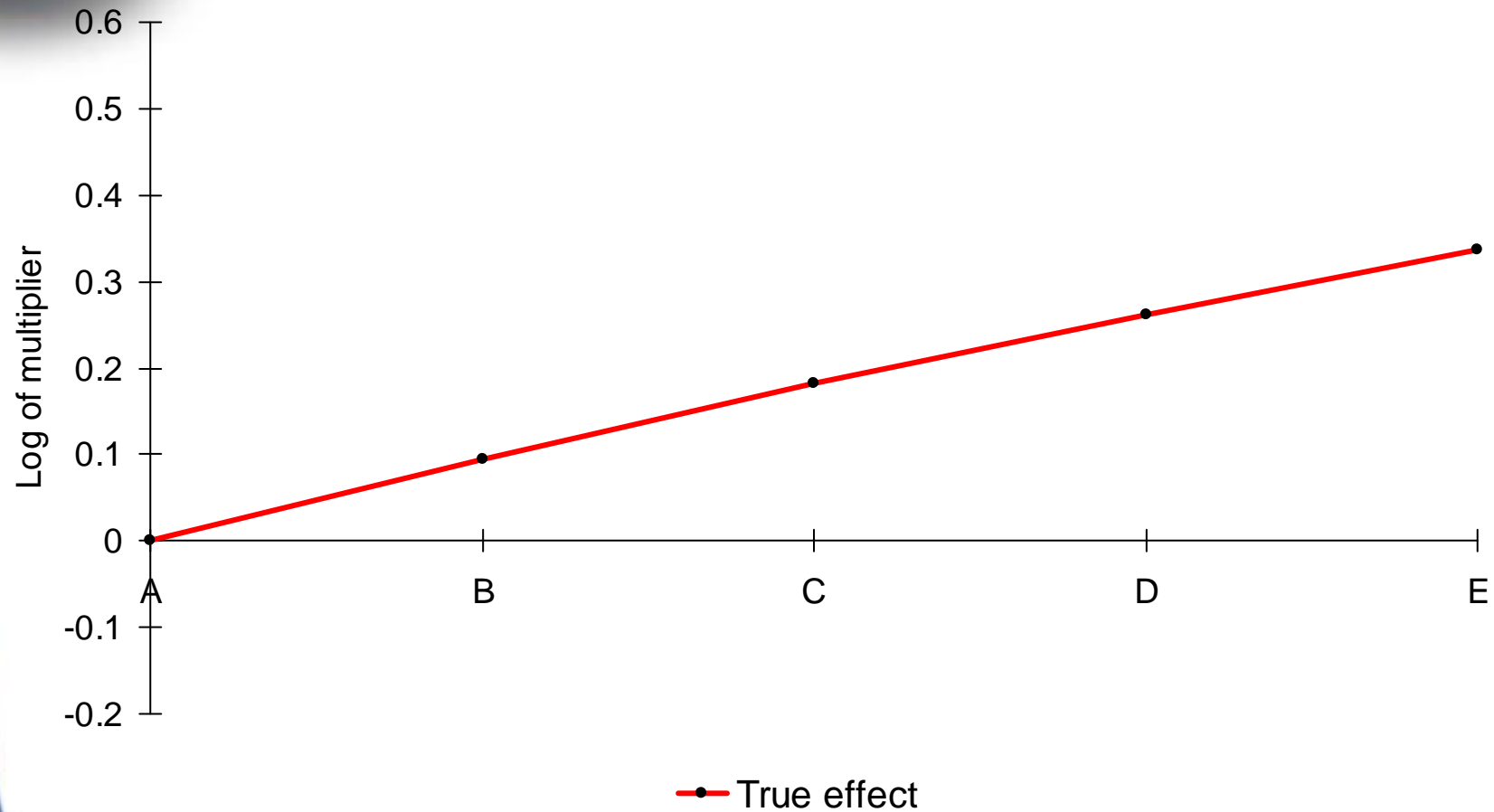


Example of effect of changing assumed error - 2

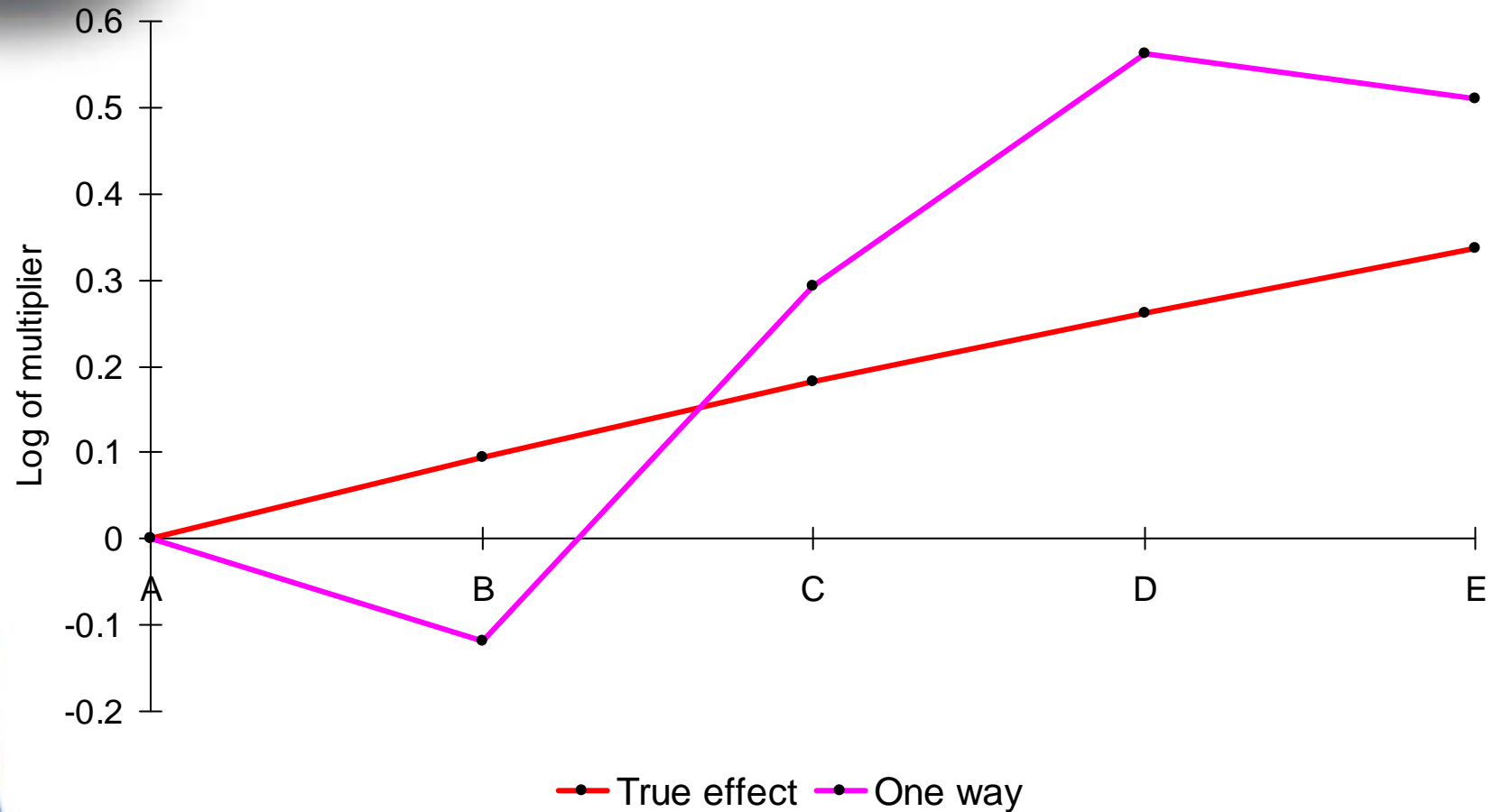
- Example portfolio with five rating factors, each with five levels A, B, C, D, E
- Typical correlations between those rating factors
- Assumed true effect of factors
- Claims randomly generated (with Gamma)
- Random experience analyzed by three models



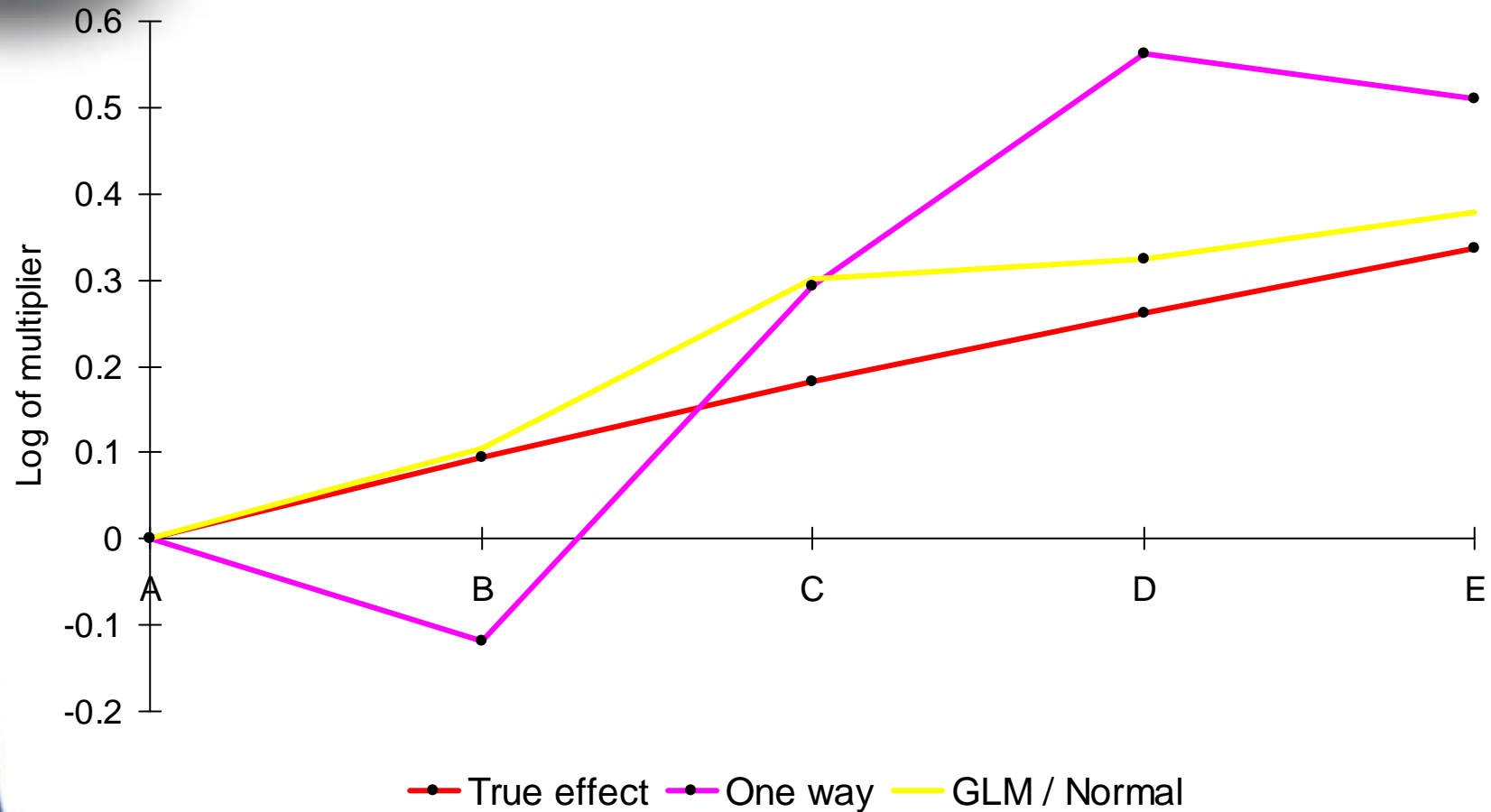
Example of effect of changing assumed error - 2



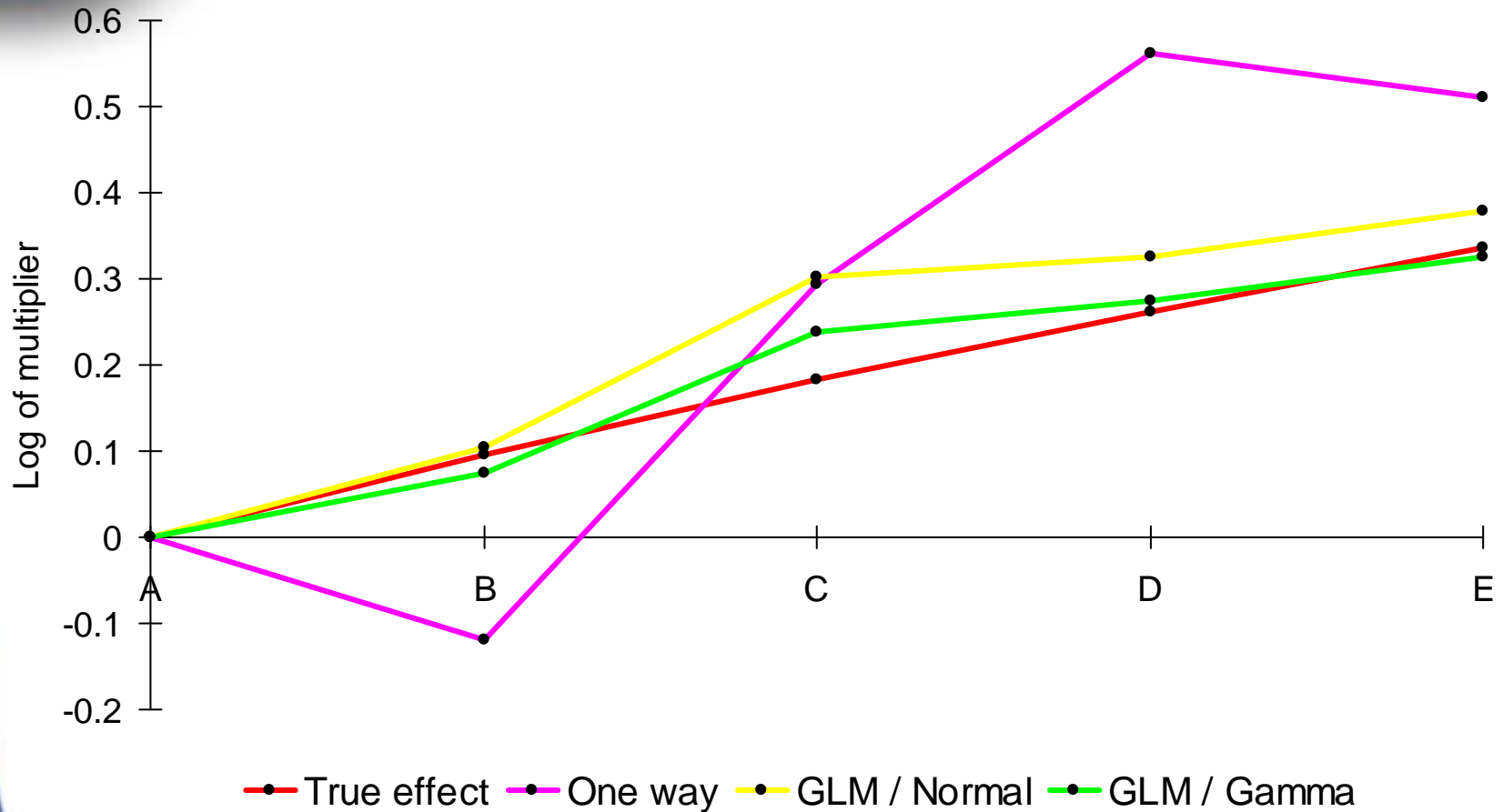
Example of effect of changing assumed error - 2



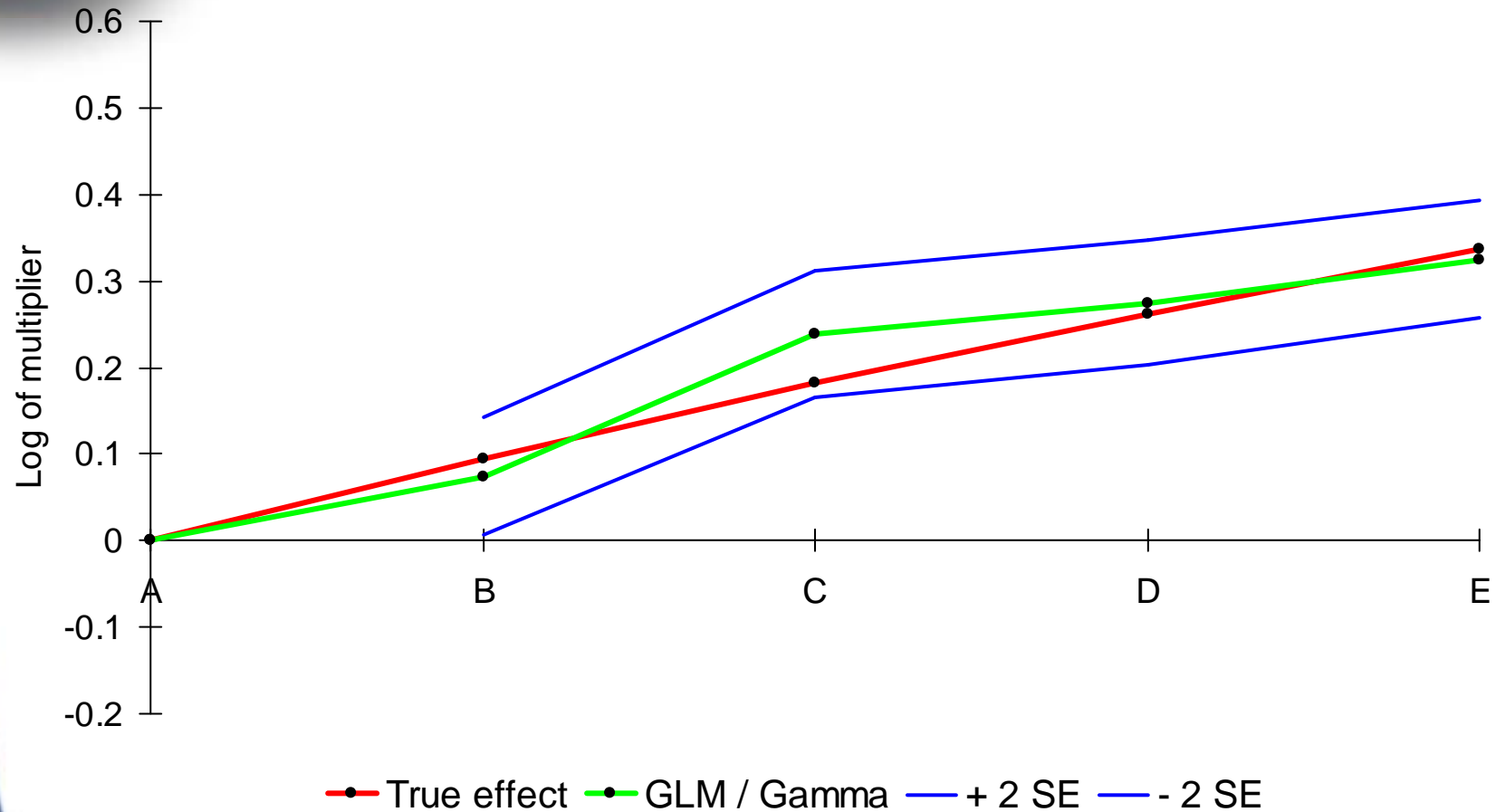
Example of effect of changing assumed error - 2



Example of effect of changing assumed error - 2



Example of effect of changing assumed error - 2





Prior weights

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$

- Exposure
- Other credibility

Eg \underline{Y} = claim *frequency*

Smith: Male, 30, Ford, 1 years, 2 claims, 100%

Jones: Female, 40, VW, 1/2 year, 1 claim, 100%



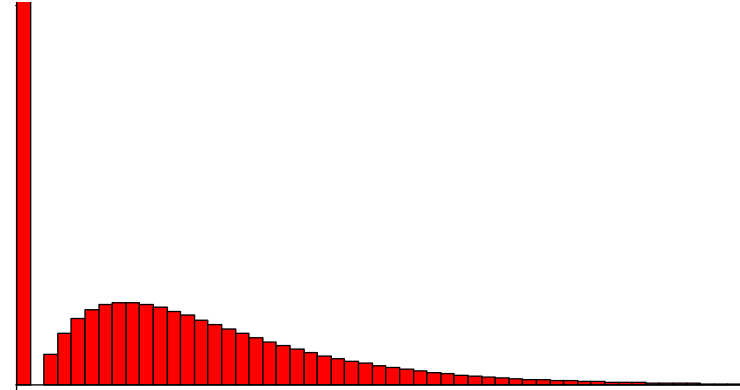
Typical model forms

\underline{Y}	Claim frequency	Claim number	Average claim amount	Probability (eg lapses)
$g(x)$	$\ln(x)$	$\ln(x)$	$\ln(x)$	$\ln(x/(1-x))$
Error	Poisson	Poisson	Gamma	Binomial
$\frac{\phi}{V(x)}$	$\frac{1}{x}$	$\frac{1}{x}$	estimate x^2	$\frac{1}{x(1-x)}$
$\underline{\omega}$	exposure	1	# claims	1
$\underline{\omega}$	0	$\ln(\text{exposure})$	0	0



Tweedie distributions

- Incurred losses have a point mass at zero and then a continuous distribution
- Poisson and gamma not suited to this
- Tweedie distribution has point mass and parameters which can alter the shape to be like Poisson and gamma above zero



$$f_Y(y; \theta, \lambda, \alpha) = \sum_{n=1}^{\infty} \frac{\{(\lambda \omega)^{1-\alpha} \kappa_{\alpha}(-1/y)\}^n}{\Gamma(-n\alpha)n!y} \cdot \exp\{\lambda \omega[\theta_0 y - \kappa_{\alpha}(\theta_0)]\} \quad \text{for } y > 0$$

$$p(Y = 0) = \exp\{-\lambda \omega \kappa_{\alpha}(\theta_0)\}$$





Generalized linear models

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Tweedie: $\phi = k$, $V(x) = x^p \Rightarrow \text{Var}[\underline{Y}] = k \underline{\mu}^p$





Tweedie distributions

Tweedie: $\phi = k$, $V(x) = x^p \Rightarrow \text{Var}[Y] = k\mu^p$

- Defines a valid distribution for $p < 0$, $1 < p < 2$, $p > 2$
- Can be considered as Poisson/gamma process for $1 < p < 2$
- Typical values of p for insurance incurred claims around, or just under, 1.5

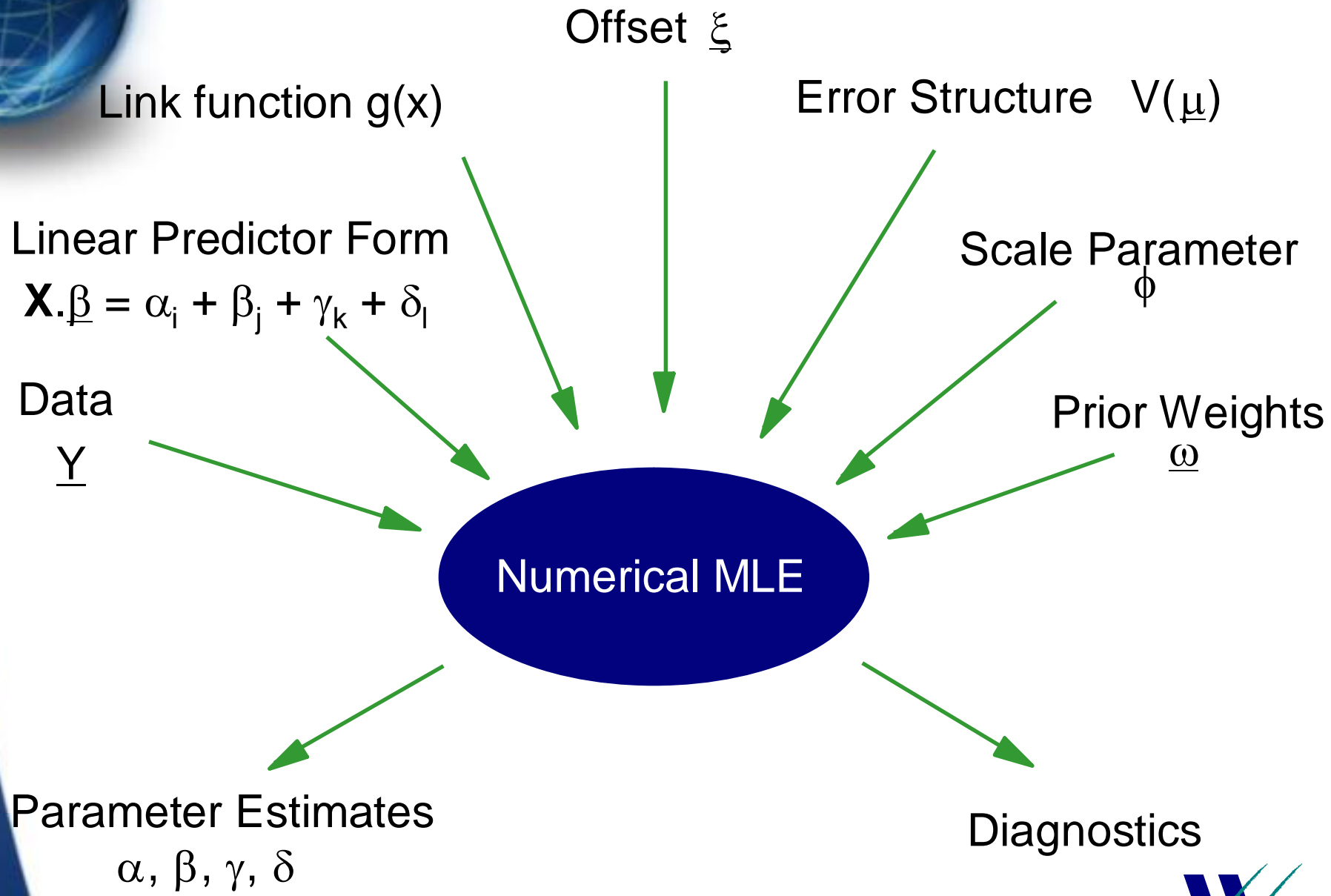




Tweedie conclusions

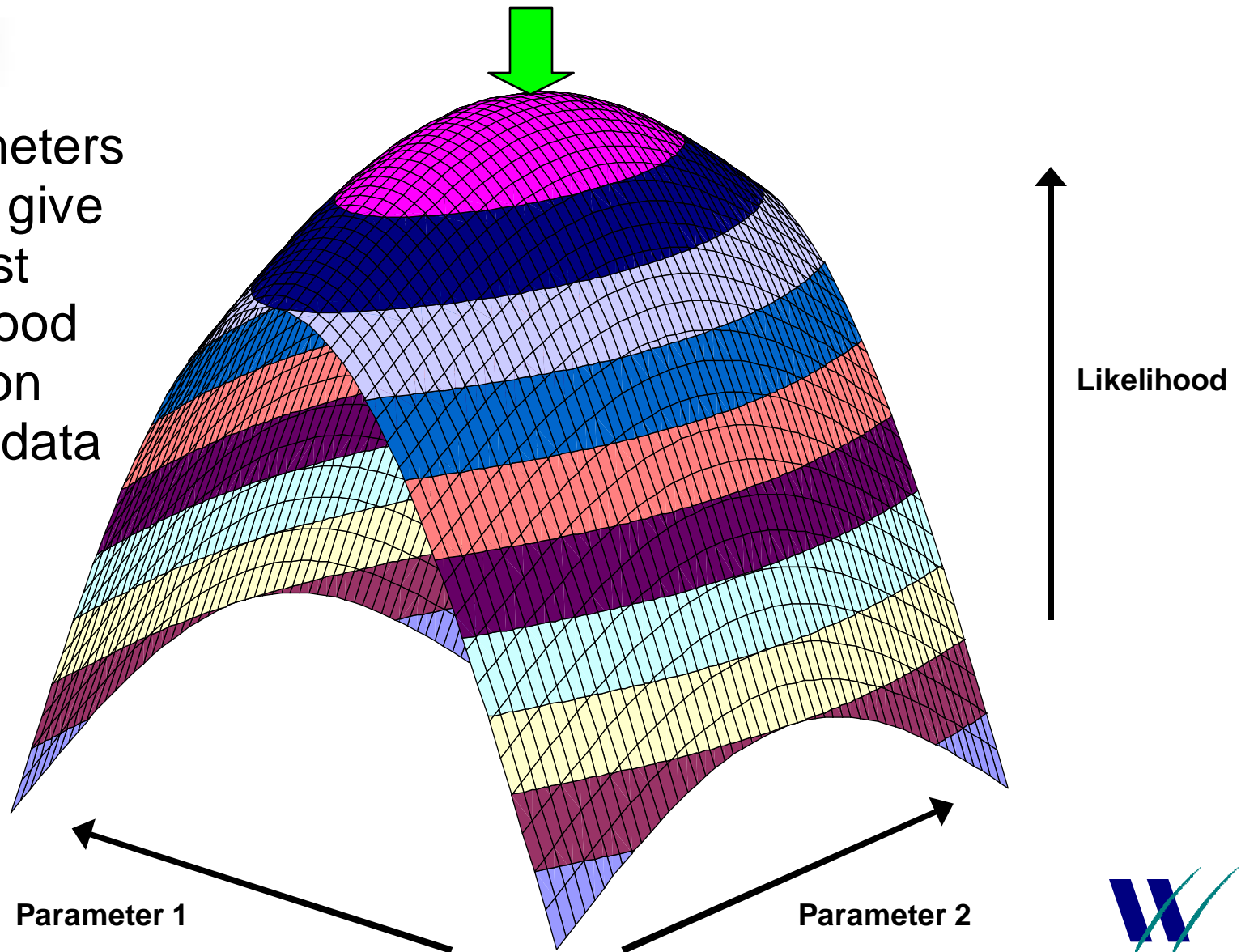
- Helpful when important to fit to pure premium
- Often similar results to traditional approach but differences may occur if numbers and amounts models have effects which are both large and insignificant
- No information about whether frequencies or amounts are driving result





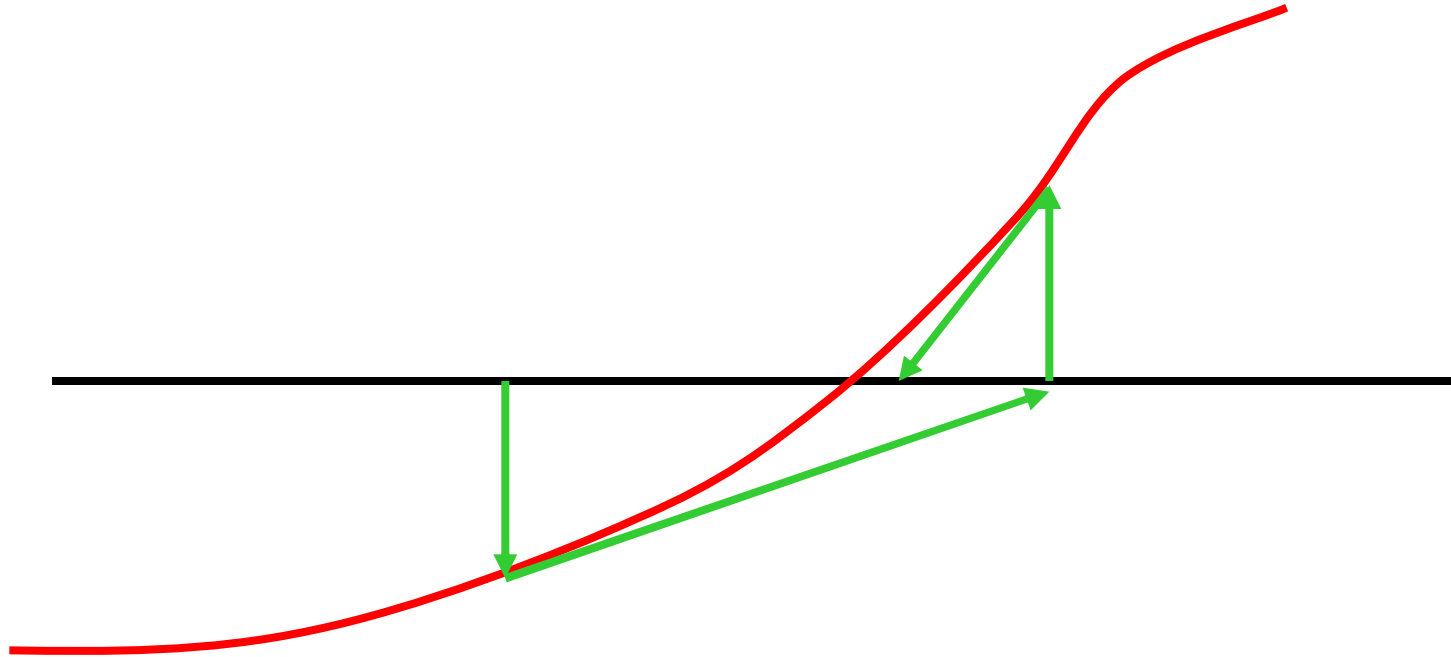
Maximum likelihood estimation

- Seek parameters which give highest likelihood function given data



Newton-Raphson

- In one dimension: $x_{n+1} = x_n - f'(x_n) / f''(x_n)$



- In n dimensions: $\underline{\beta}_{n+1} = \underline{\beta}_n - \mathbf{H}^{-1} \cdot \underline{s}$

where $\underline{\beta}$ is the vector of the parameter estimates (with p elements), \underline{s} is the vector of the first derivatives of the log-likelihood and \mathbf{H} is the $(p \times p)$ matrix containing the second derivatives of the log-likelihood





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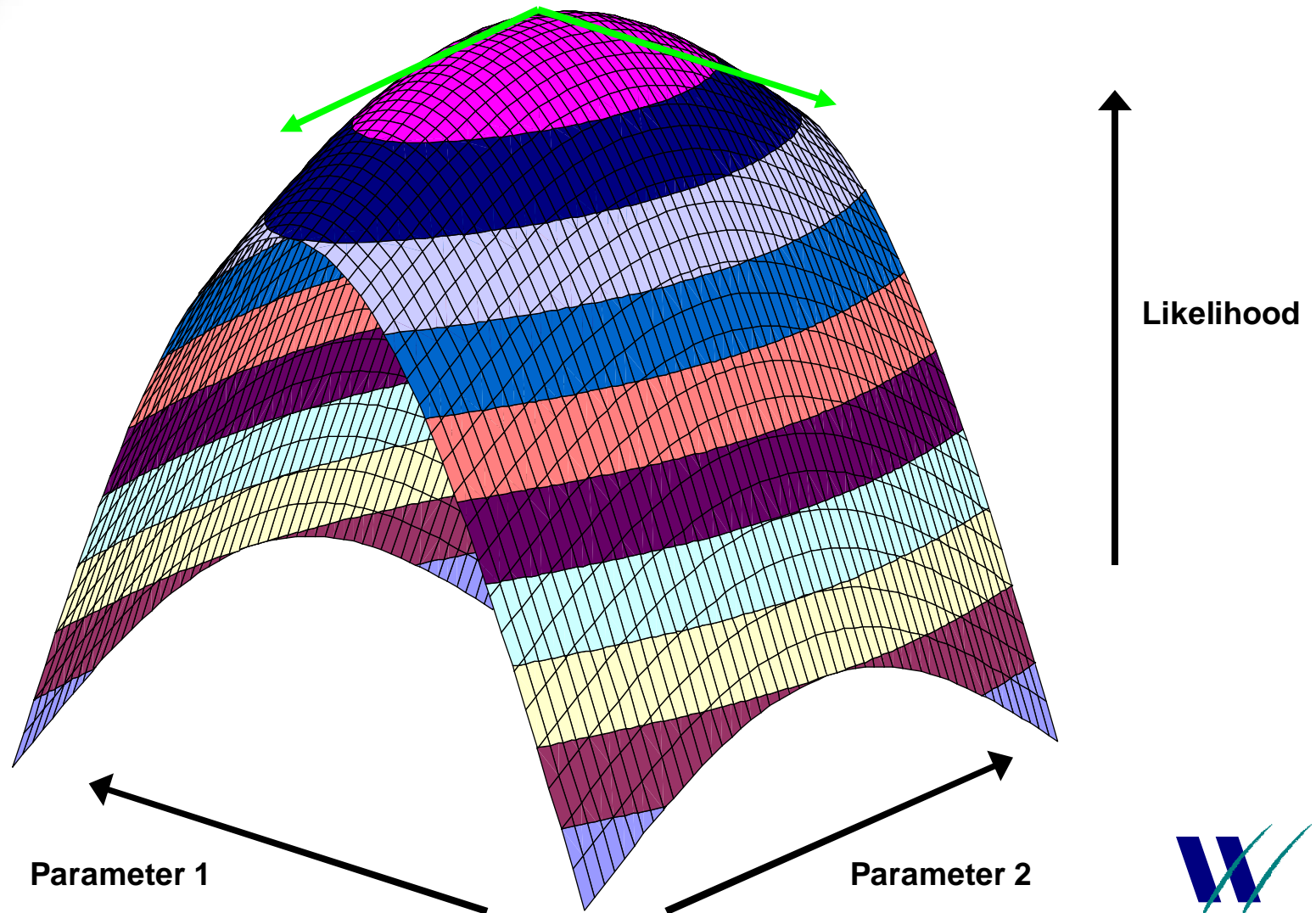
Model testing

- Use only those variables which are predictive
 - standard errors of parameter estimates
 - F tests / χ^2 tests on deviances
- Make sure the model is reasonable
 - histogram of deviance residuals
 - residual vs fitted value
 - Box Cox link function investigation

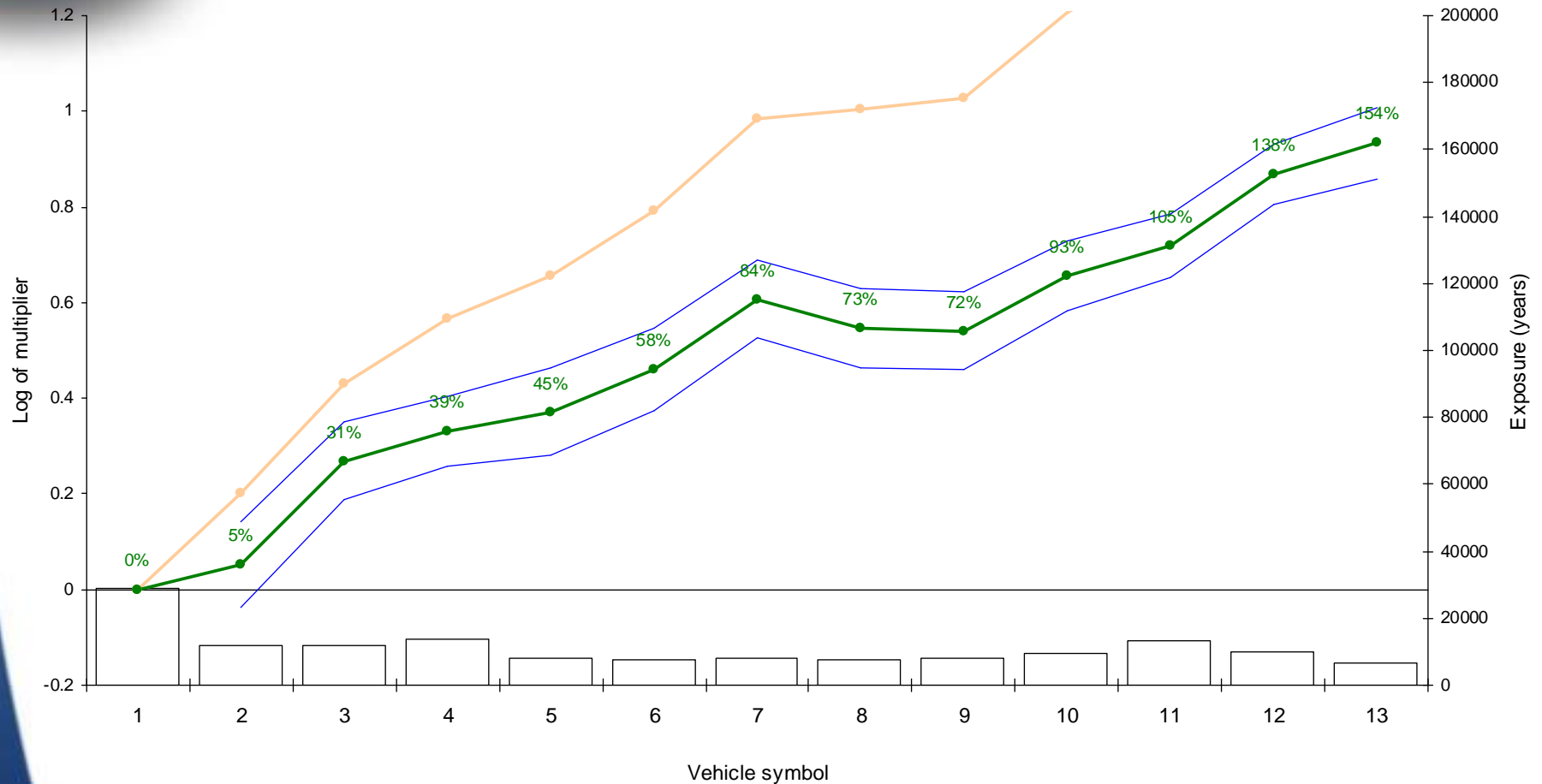


Standard errors

- Roughly speaking, for a parameter p : $SE = -1 / (\partial^2 / \partial p^2 \text{ Likelihood})$



GLM output (significant factor)

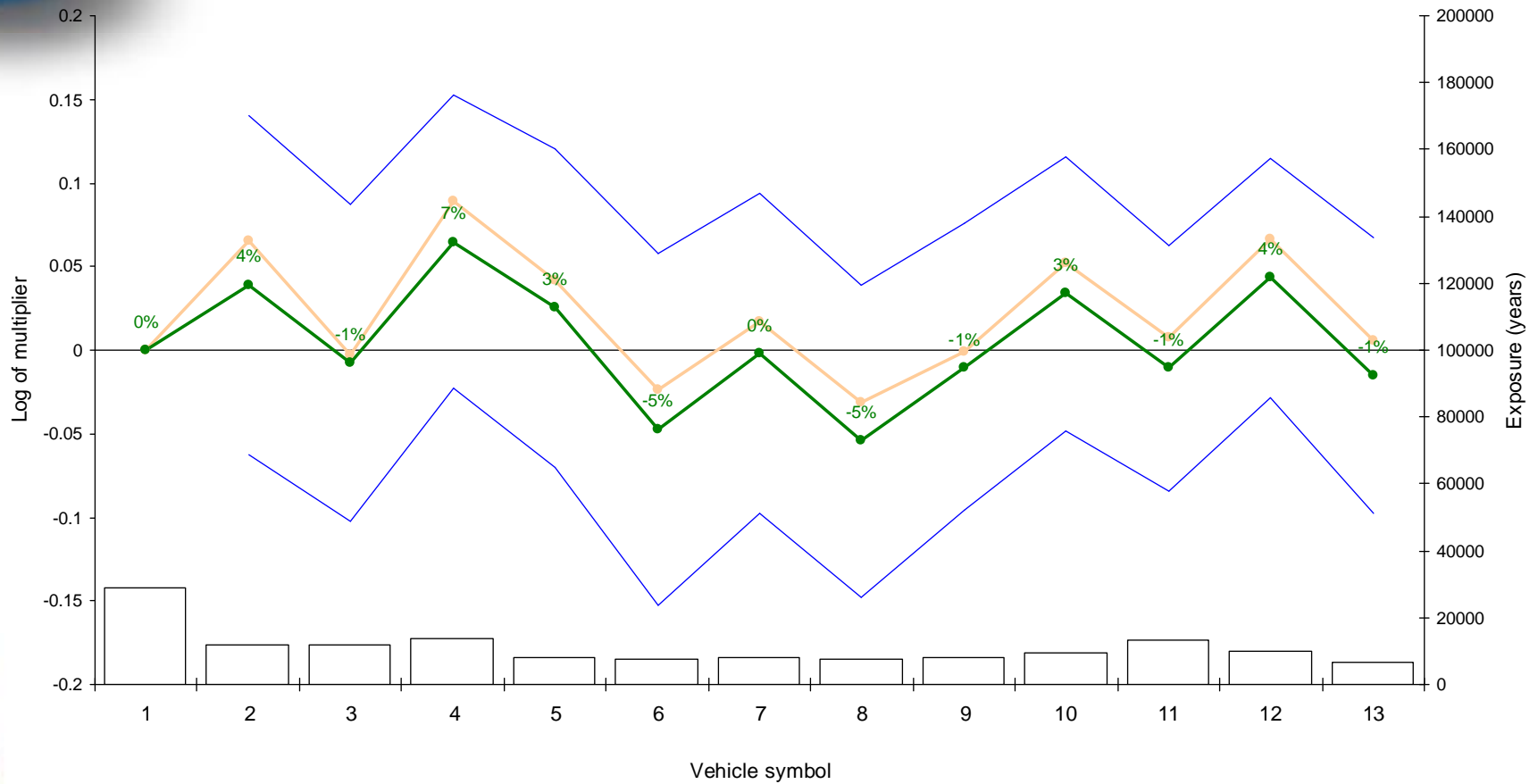


—●— Onew ay relativities —●— Approx 95% confidence interval —●— Parameter estimate

P value = 0.0%



GLM output (insignificant factor)

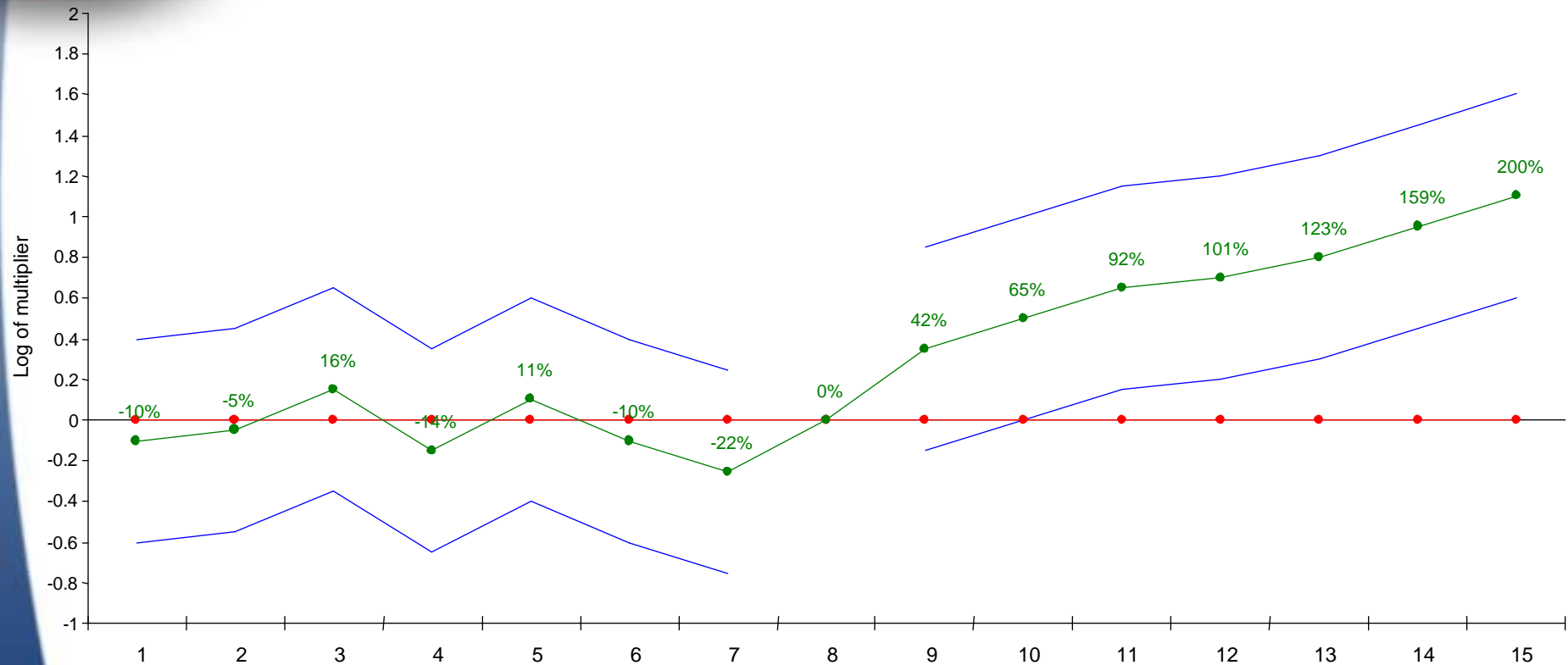


P value = 52.5%

—○— Oneway relativities ——— Approx 95% confidence interval —●— Parameter estimate



Awkward cases



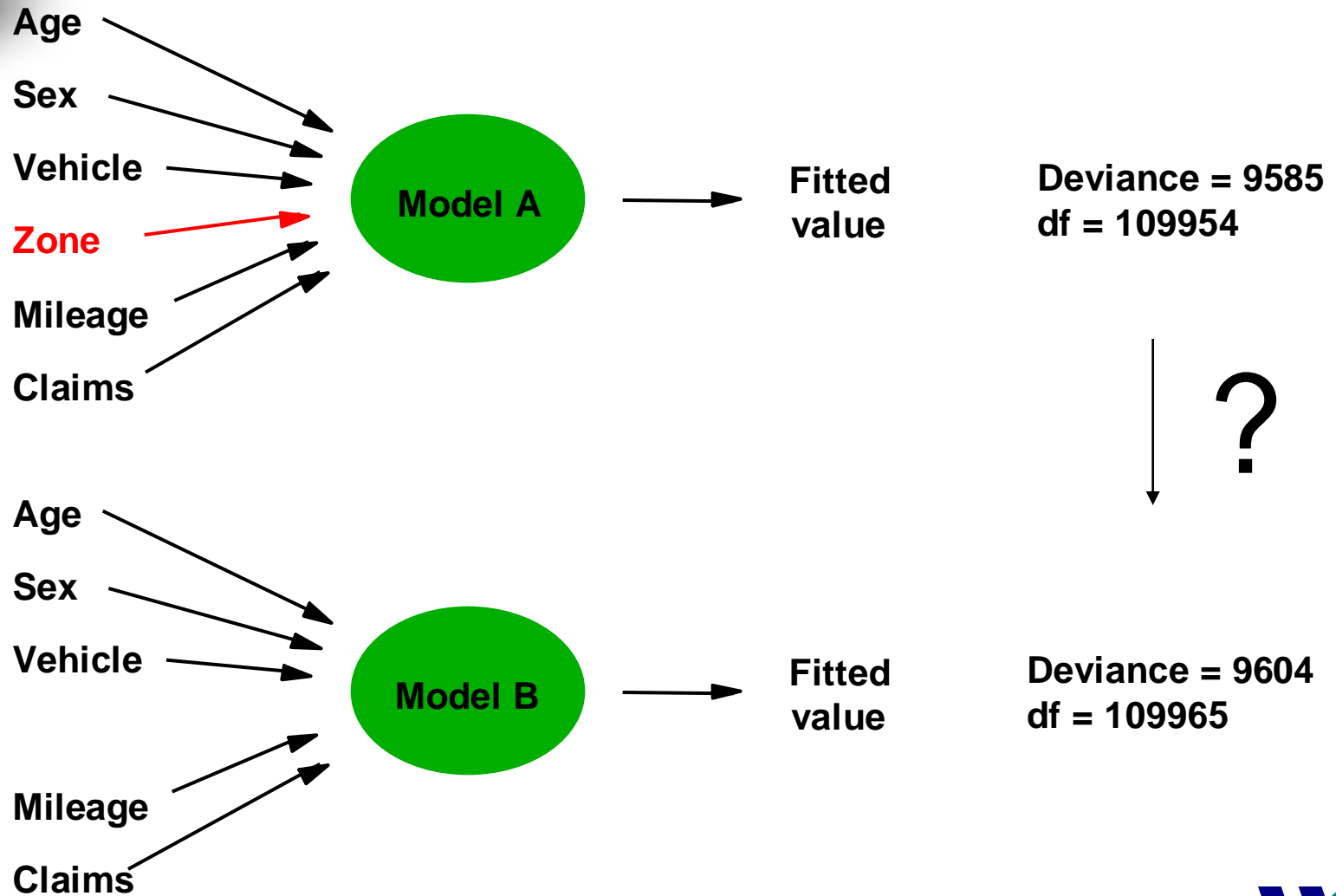


Deviances

- Single figure measure of goodness of fit
- Try model with & without a factor
- Statistical tests show the theoretical significance given the extra parameters



Deviances





Deviances

- If ϕ known, scaled deviance S output

$$S = \sum_{u=1}^n 2 \omega_u / \phi \int_{\mu_u}^{Y_u} (Y_u - \zeta) / V(\zeta) d\zeta$$

$$S_1 - S_2 \sim \chi^2_{d_1-d_2}$$

- If ϕ unknown, unscaled deviance $D = \phi.S$ output

$$\frac{(D_1 - D_2)}{(d_1 - d_2) D_3 / d_3} \sim F_{d_1-d_2, d_3}$$





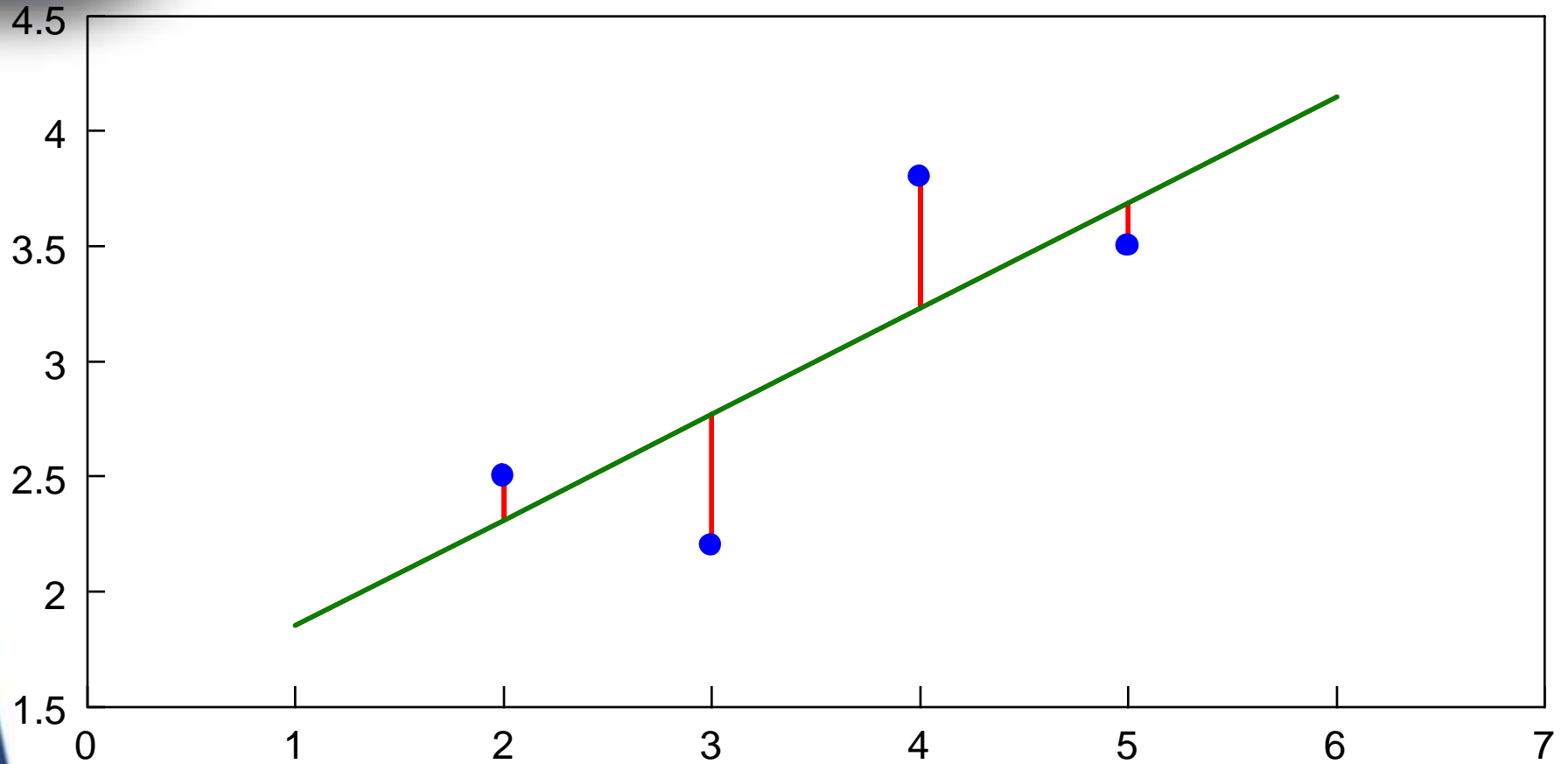
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Residuals



Residuals Fitted values Data



Residuals

- Several forms, eg

- standardized deviance

$$\text{sign} (Y_u - \mu_u) / (\phi (1-h_u))^{1/2}$$

$$\sqrt{2 \omega_u \int_{\mu_u}^{Y_u} (Y_u - \zeta) / V(\zeta) d\zeta}$$

- standardized Pearson

$$\frac{Y_u - \mu_u}{(\phi \cdot V(\mu_u) \cdot (1-h_u) / \omega_u)^{1/2}}$$

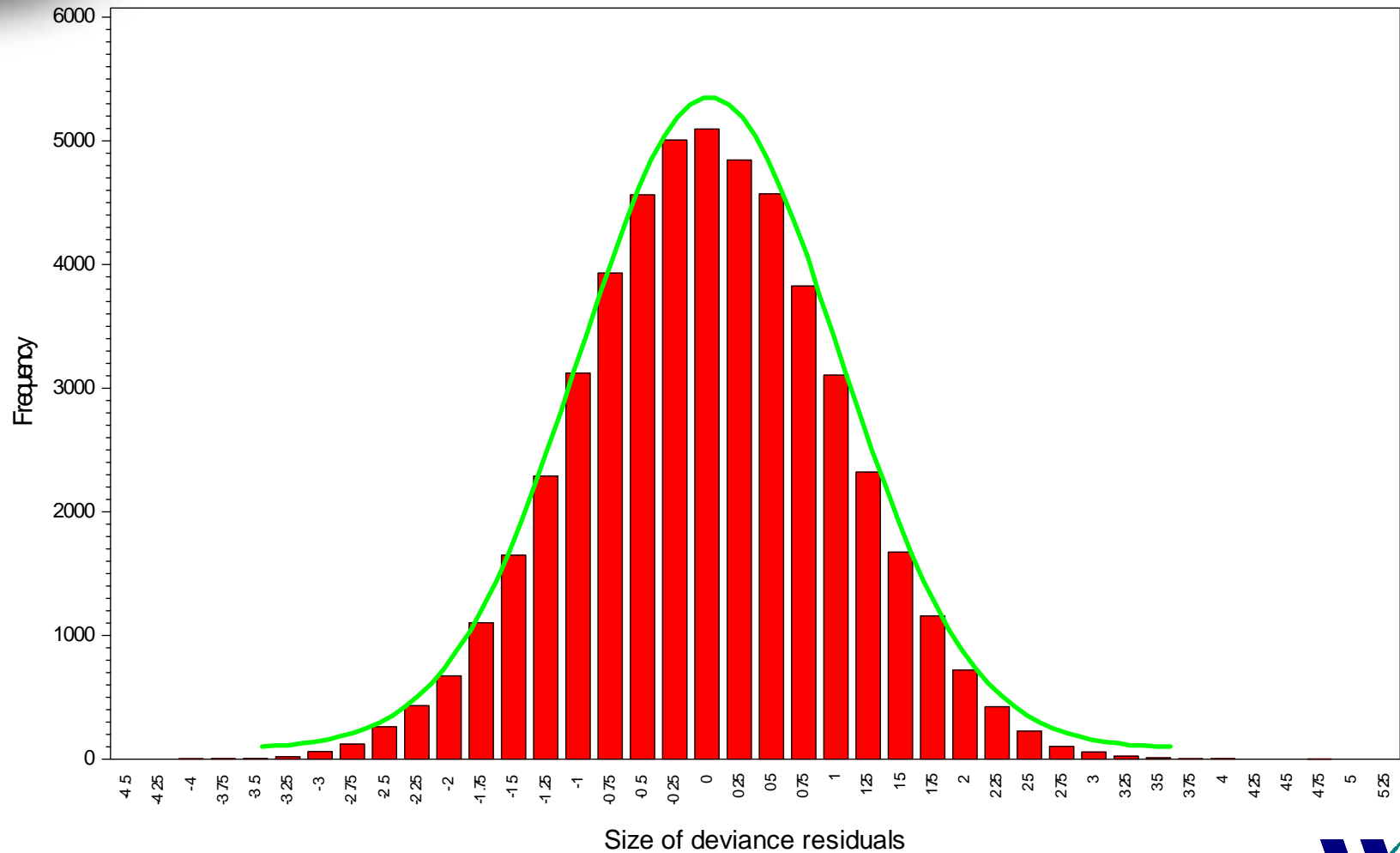
- Standardized deviance - Normal (0,1)
- Numbers/frequency residuals problematical



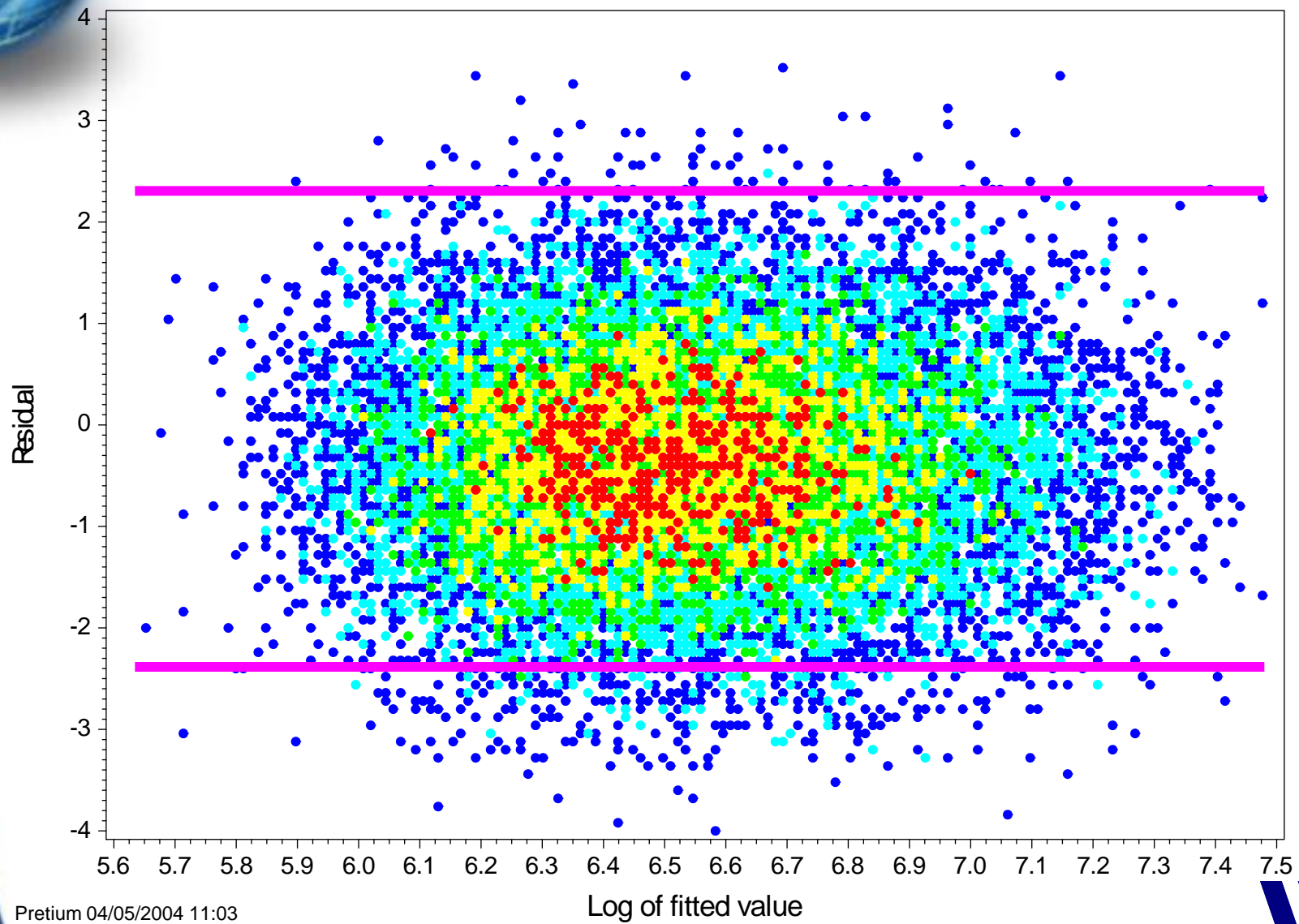


Residuals

Histogram of Deviance Residuals
Run 12 (Final models with analysis) Model 8 (AD amounts)



Residuals

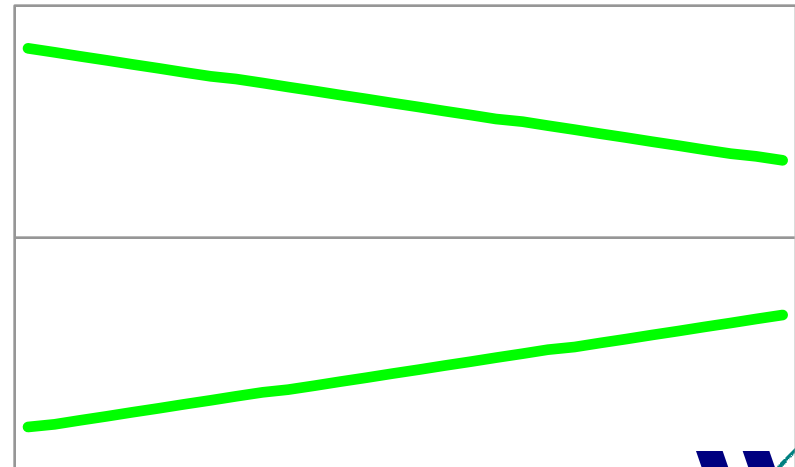
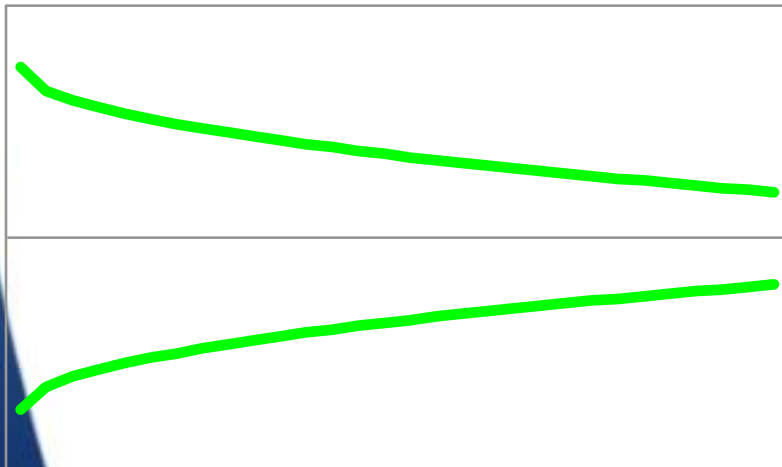
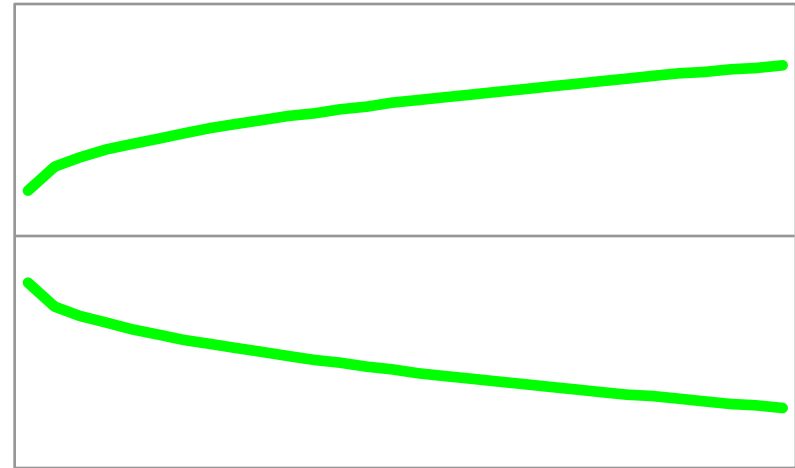
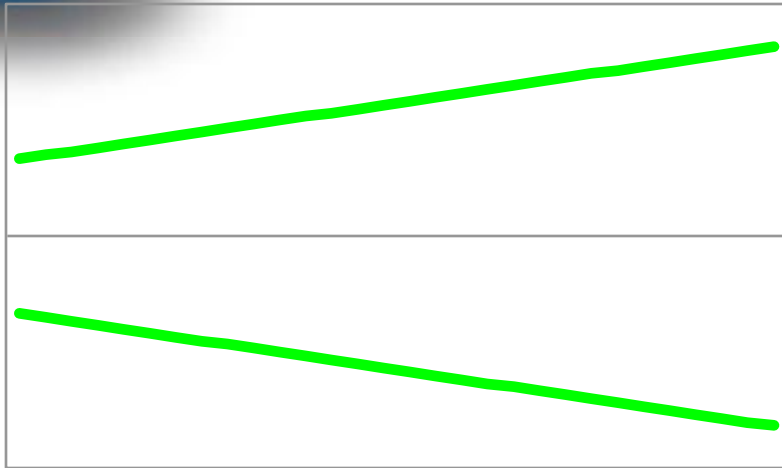


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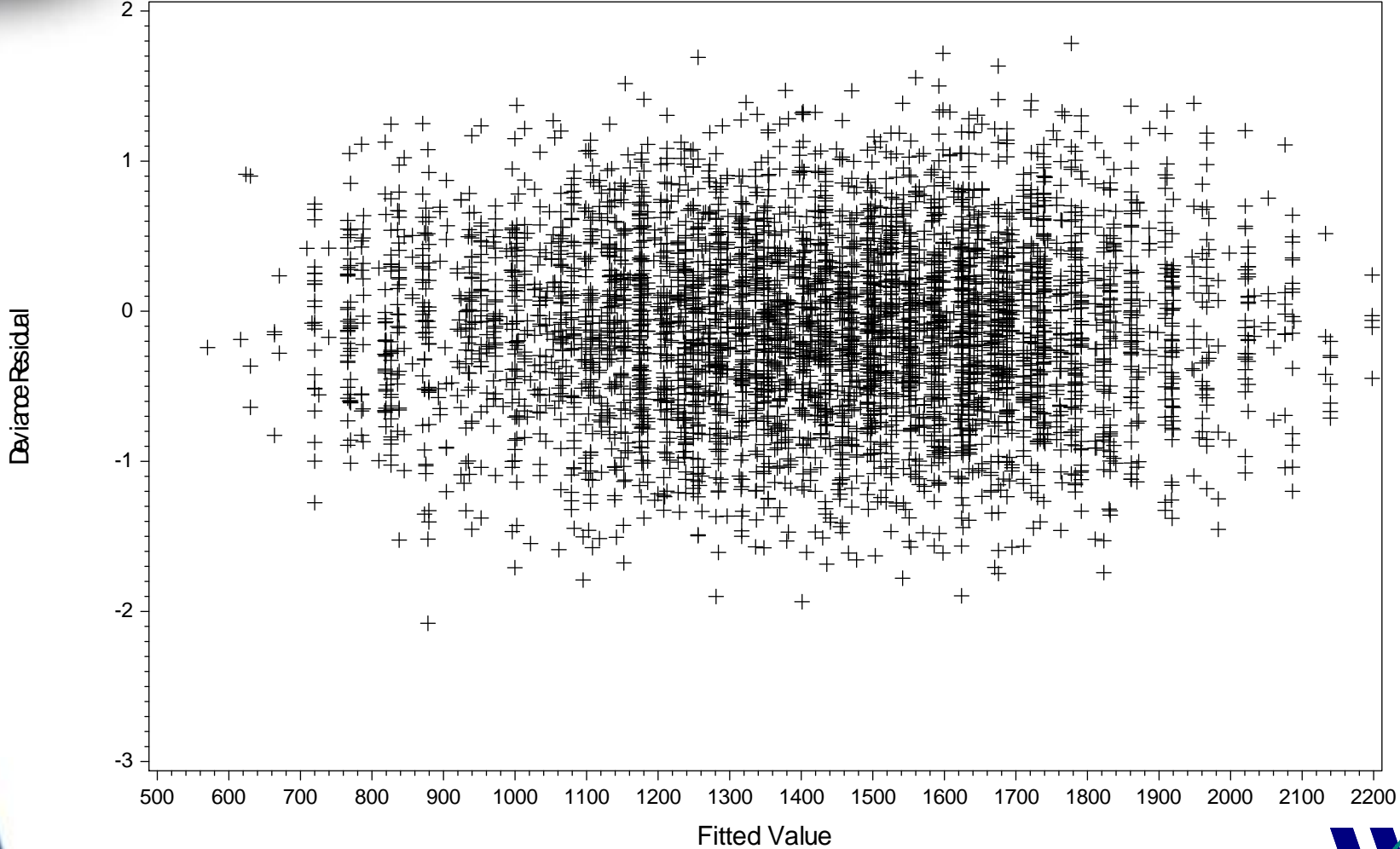
Residuals





Gamma data, Gamma error

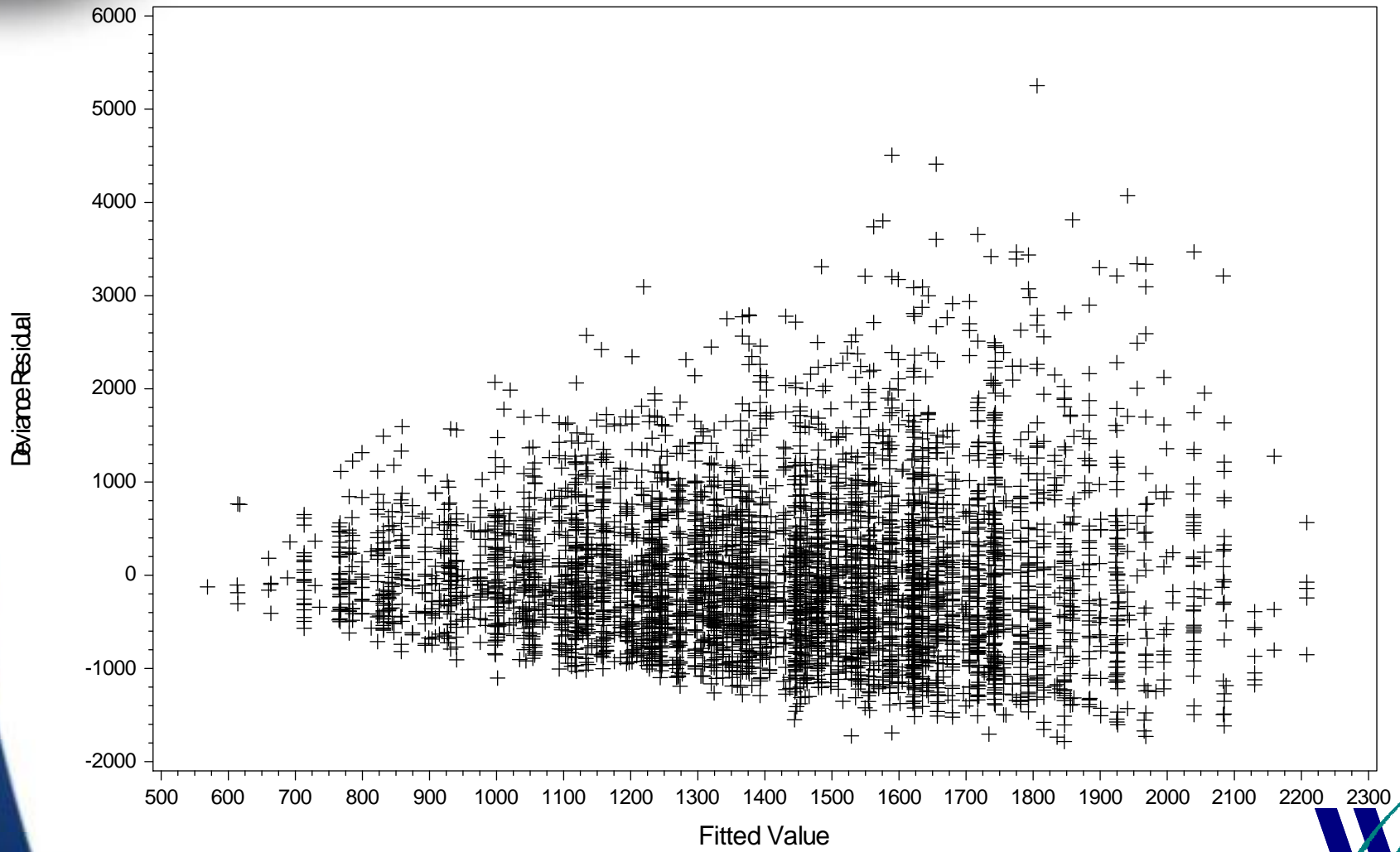
Plot of deviance residual against fitted value
Run 12 (All claim types, final models, N&A) Model 6 (Own damage, Amounts)





Gamma data, Normal error

Plot of deviance residual against fitted value
Run 12 (All claim types, final models, N&A) Model 7 (Own damage, Amounts)





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- Model testing
 - use only variables which are predictive
 - make sure model is reasonable
- Aliasing





Aliasing and "near aliasing"

- Aliasing
 - the removal of unwanted redundant parameters
- Intrinsic aliasing
 - occurs by the design of the model
- Extrinsic aliasing
 - occurs "accidentally" as a result of the data





Intrinsic aliasing

$X.\beta = \alpha + \beta_1$ if age 20 - 29

~~$+ \beta_2$ if age 30 - 39~~

$+ \beta_3$ if age 40 +

~~$+ \gamma_1$ if sex male~~

$+ \gamma_2$ if sex female

■ "Base levels"

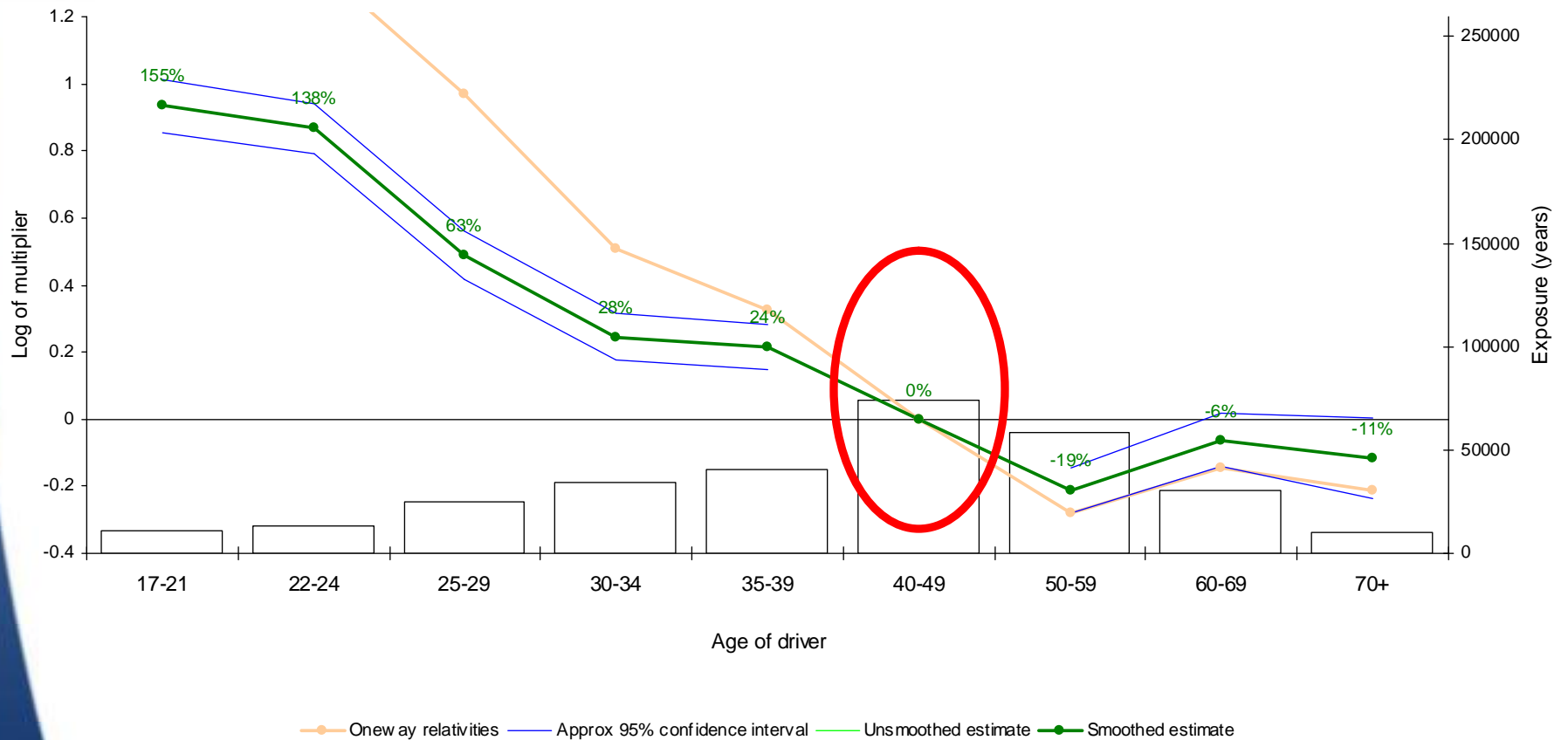




Intrinsic aliasing

Example job

Run 16 Model 3 - Small interaction - Third party material damage, Numbers



Extrinsic aliasing

- If a perfect correlation exists, one factor can alias levels of another
- Eg if doors declared first:

Exposure:	# Doors →	2	3	4 Selected base	5	Unknown
Color ↓						
Red Selected base		13,234	12,343	13,432	13,432	0
Green		4,543	4,543	13,243	2,345	0
Blue		6,544	5,443	15,654	4,565	0
Black		4,643	1,235	14,565	4,545	0
Unknown Further aliasing		0	0	0	0	3,242

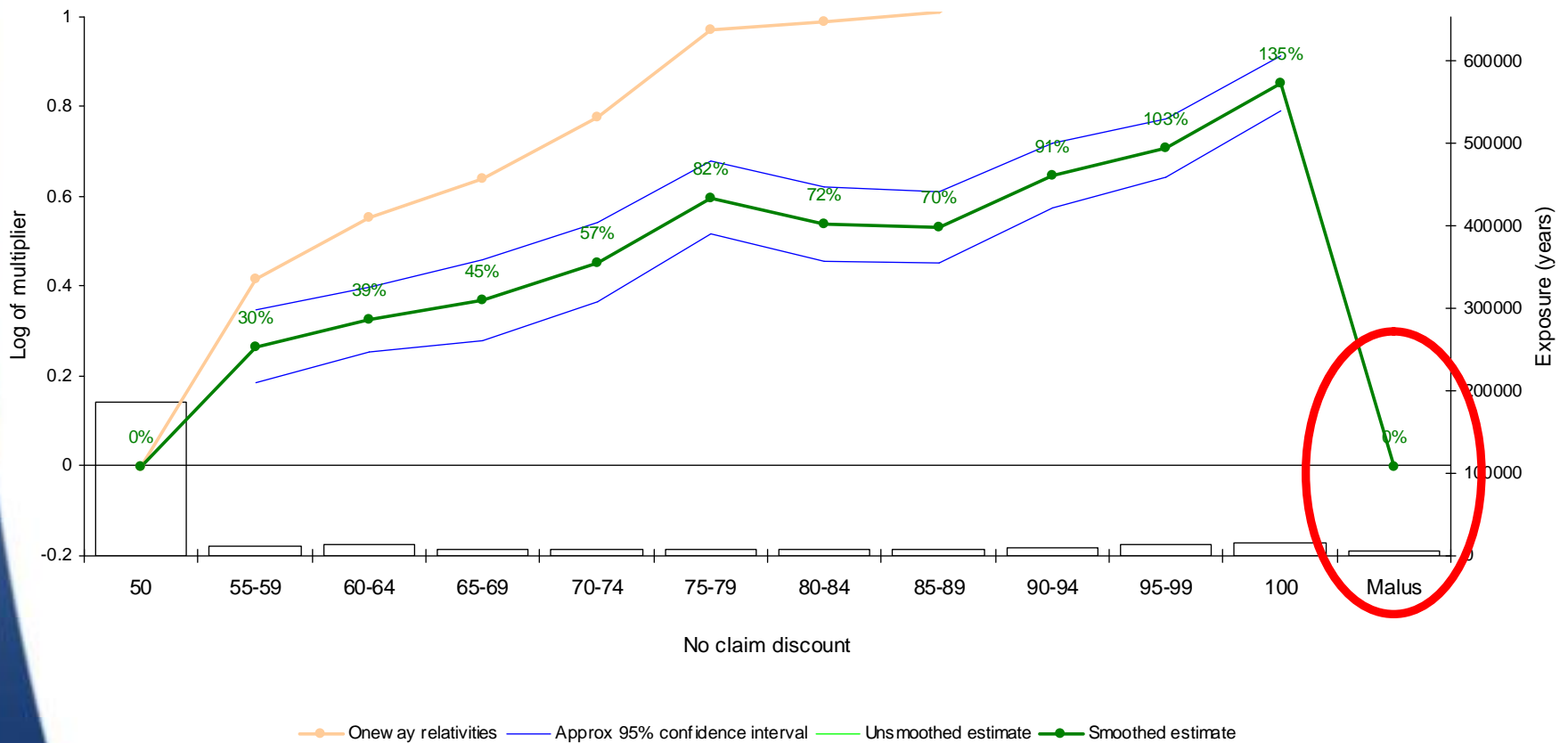
- This is the only reason the order of declaration can matter (fitted values are unaffected)



Extrinsic aliasing

Example job

Run 16 Model 3 - Small interaction - Third party material damage, Numbers



"Near aliasing"


- If two factors are almost perfectly, but not quite aliased, convergence problems can result as a result of low exposures (even though one-ways look fine), and/or results can become hard to interpret

Exposure:	# Doors →	2	3	4 Selected base	5	Unknown
Color ↓						
Red Selected base		13,234	12,343	13,432	13,432	0
Green		4,543	4,543	13,243	2,345	0
Blue		6,544	5,443	15,654	4,565	0
Black		4,643	1,235	14,565	4,545	2
Unknown		0	0	0	0	3,242

- Eg if the 2 black, unknown doors policies had no claims, GLM would try to estimate a very large negative number for unknown doors, and a very large positive number for unknown color



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