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PM-2

An Introduction to GLM Theory

**CAS Seminar on Ratemaking
Boston, March 17, 2008**

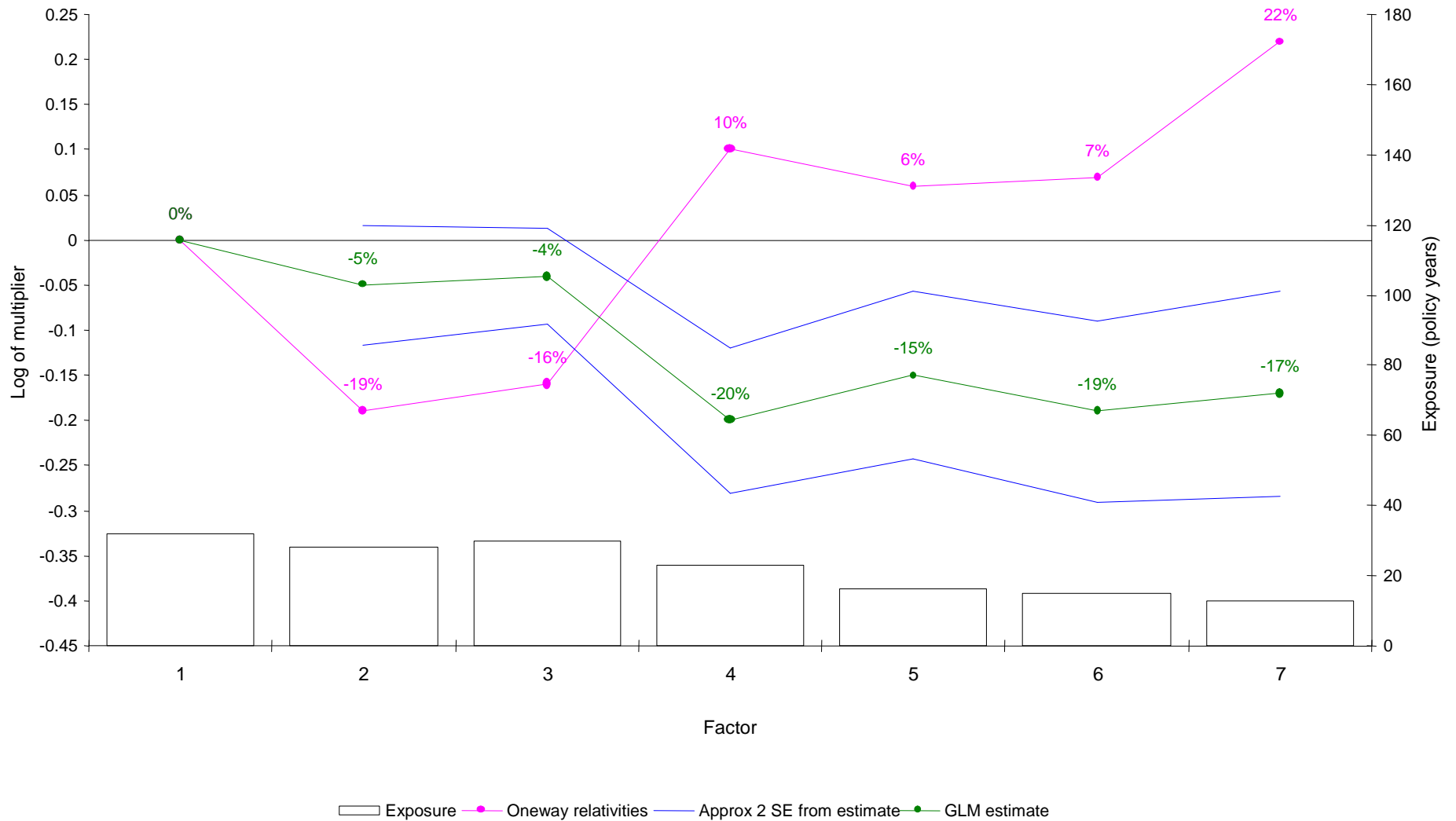
Claudine Modlin, FCAS, MAAA



Generalized linear model benefits

- Consider all factors simultaneously
- Allow for nature of random process
- Provide diagnostics
- Robust and transparent

Example of GLM output



Agenda

- GLM formulae
- Model testing
 - use only variables that are predictive
 - make sure model is reasonable
- Aliasing
- Refinements

Linear models

- Linear model $Y_i = \mu_i + \text{error}$
- μ_i based on linear combination of measured factors
- Which factors, and how they are best combined is to be derived

$$\mu_i = \alpha + \beta \cdot \text{age}_i + \gamma \cdot \text{age}_i^2 + \delta \cdot \text{height}_i \cdot \text{age}_i$$



$$\mu_i = \alpha + \beta \cdot \text{age}_i + \gamma \cdot (\text{sex}_i = \text{female})$$



$$\mu_i = (\alpha + \beta \cdot \text{age}_i) * \exp(\delta \cdot \text{height}_i \cdot \text{age}_i)$$



Linear models - formularization

$$E[Y_i] = \mu_i = \sum X_{ij} \beta_j$$

$$\text{Var}[Y_i] = \sigma^2$$

$$Y_i \sim N(\mu_i, \sigma^2)$$

What is $\sum X_{ij}\beta_j$?

- **X** defines the explanatory variables to be included in the model
 - could be continuous variables - "variates"
 - could be categorical variables - "factors"
- $\underline{\beta}$ contains the parameter estimates which relate to the factors / variates defined by the structure of **X**
 - "the answer"

What is $\mathbf{X}\cdot\beta$?

- Write $\sum X_{ij}\beta_j$ as $\mathbf{X}\cdot\beta$
- Consider 3 rating factors
 - age of driver ("age")
 - sex of driver ("sex")
 - age of vehicle ("car")
- Represent β by $\alpha, \beta, \gamma, \delta, \dots$

What is $X.\beta$?

- Suppose we wanted a model of the form:

$$\mu = \alpha + \beta.\underline{\text{age}} + \gamma.\underline{\text{age}}^2 + \delta.\underline{\text{car}}^{27}.\underline{\text{age}}^{52\frac{1}{2}}$$

- $X.\beta$ would need to be defined as:

$$\begin{pmatrix} 1 & \text{age}_1 & \text{age}_1^2 & \text{car}_1^{27}.\text{age}_1^{52\frac{1}{2}} \\ 1 & \text{age}_2 & \text{age}_2^2 & \text{car}_2^{27}.\text{age}_2^{52\frac{1}{2}} \\ 1 & \text{age}_3 & \text{age}_3^2 & \text{car}_3^{27}.\text{age}_3^{52\frac{1}{2}} \\ 1 & \text{age}_4 & \text{age}_4^2 & \text{car}_4^{27}.\text{age}_4^{52\frac{1}{2}} \\ 1 & \text{age}_5 & \text{age}_5^2 & \text{car}_5^{27}.\text{age}_5^{52\frac{1}{2}} \\ \dots & \dots & \dots & \dots \\ \dots & \dots & \dots & \dots \end{pmatrix} \cdot \begin{pmatrix} \alpha \\ \beta \\ \gamma \\ \delta \end{pmatrix}$$

What is $X.\beta$?

- Suppose we wanted a model of the form:

$$\begin{aligned}\underline{\mu} &= \alpha + \beta_1 \text{ if } \underline{\text{age}} < 30 \\ &+ \beta_2 \text{ if } \underline{\text{age}} \text{ 30 - 40} \\ &+ \beta_3 \text{ if } \underline{\text{age}} > 40 \\ &+ \gamma_1 \text{ if } \underline{\text{sex}} \text{ male} \\ &+ \gamma_2 \text{ if } \underline{\text{sex}} \text{ female}\end{aligned}$$

What is $X \cdot \beta$?

	Age			Sex	
	<30	30-40	>40	M	F
1	1	0	1	0	1
2	1	1	0	0	1
3	1	1	0	0	1
4	1	0	0	1	0
5	1	0	1	0	1
				
				

\cdot

α
β_1
β_2
β_3
γ_1
γ_2

What is $X.\beta$?

- Suppose we wanted a model of the form:

$$\begin{aligned}\underline{\mu} &= \alpha + \beta_1 \text{ if } \underline{\text{age}} < 30 \\ &+ \beta_2 \text{ if } \underline{\text{age}} 30 - 40 \\ &+ \beta_3 \text{ if } \underline{\text{age}} > 40 \\ &+ \gamma_1 \text{ if } \underline{\text{sex}} \text{ male} \\ &+ \gamma_2 \text{ if } \underline{\text{sex}} \text{ female}\end{aligned}$$

What is $X.\beta$?

- Suppose we wanted a model of the form:

$$\underline{\mu} = \alpha + \beta_1 \text{ if } \underline{\text{age}} < 30$$

$$~~+ \beta_2 \text{ if } \underline{\text{age}} 30 - 40~~$$

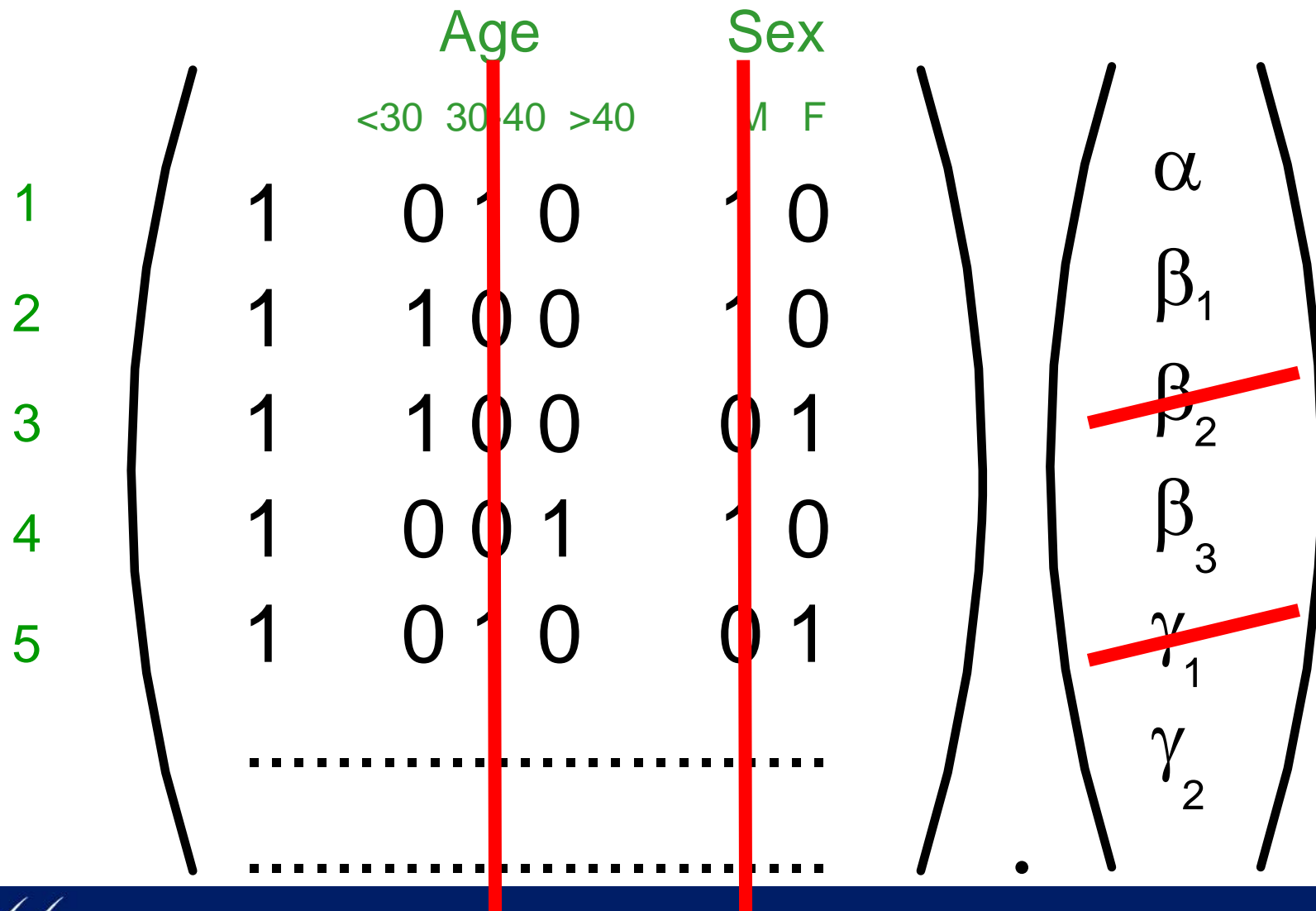
"Base levels"

$$+ \beta_3 \text{ if } \underline{\text{age}} > 40$$

$$~~+ \gamma_1 \text{ if } \underline{\text{sex}} \text{ male}~~$$

$$+ \gamma_2 \text{ if } \underline{\text{sex}} \text{ female}$$

$X \cdot \beta$ having adjusted for base levels



Linear models - formularization

$$E[Y_i] = \mu_i = \sum X_{ij} \beta_j$$

$$\text{Var}[Y_i] = \sigma^2$$

$$Y_i \sim N(\mu_i, \sigma^2)$$

Generalized linear models

$$\mu_i = f(\alpha + \beta \cdot \text{age}_i + \gamma \cdot \text{age}_i^2 + \delta \cdot \text{height}_i \cdot \text{age}_i)$$

$$\mu_i = f(\alpha + \beta \cdot \text{age}_i + \gamma \cdot (\text{sex}_i = \text{female}))$$

Generalized linear models

$$\mu_i = g^{-1}(\alpha + \beta \cdot \text{age}_i + \gamma \cdot \text{age}_i^2 + \delta \cdot \text{height}_i \cdot \text{age}_i)$$

$$\mu_i = g^{-1}(\alpha + \beta \cdot \text{age}_i + \gamma \cdot (\text{sex}_i = \text{female}))$$

Generalized linear models

Linear Models

$$E[Y_i] = \mu_i = \sum X_{ij}\beta_j$$

$$\text{Var}[Y_i] = \sigma^2$$

Y from
Normal distribution

Generalized Linear Models

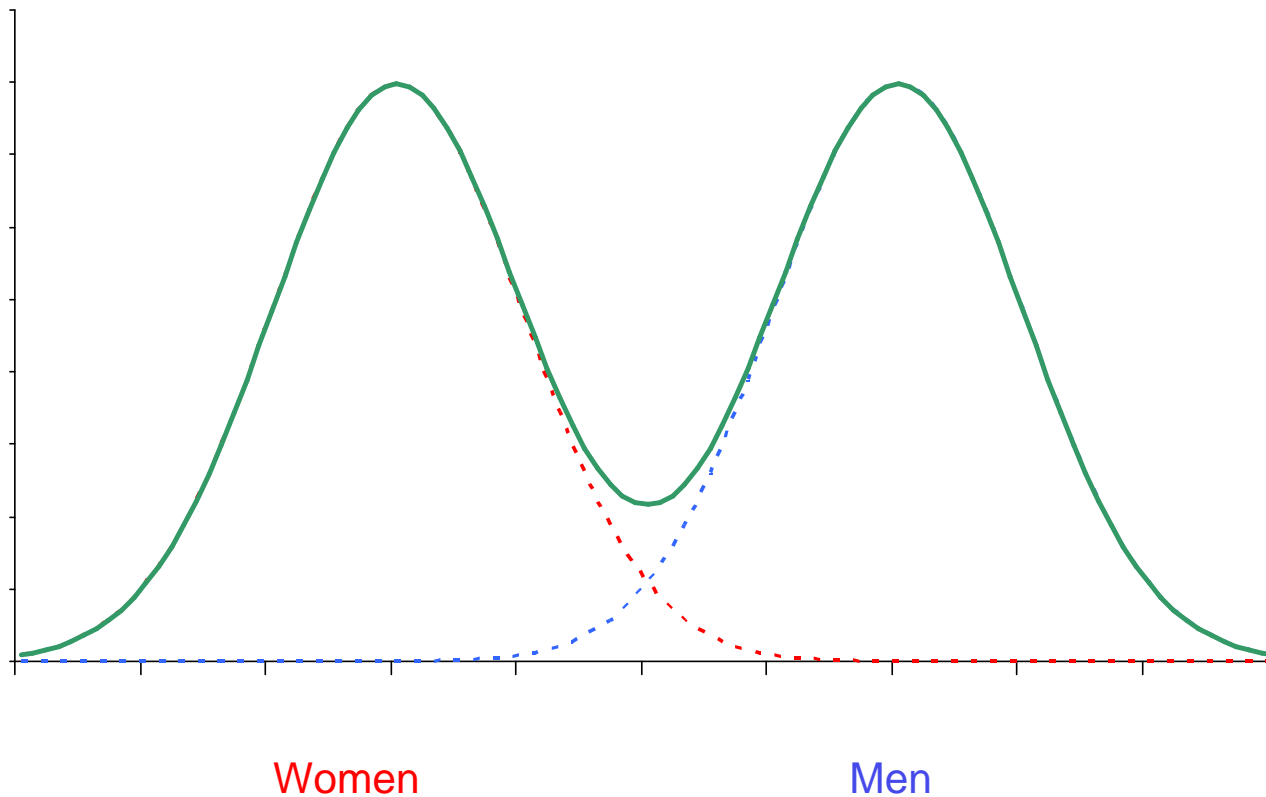
$$E[Y_i] = \mu_i = g^{-1}(\sum X_{ij}\beta_j + \xi_i)$$

$$\text{Var}[Y_i] = \phi V(\mu_i)/\omega_i$$

Y from a distribution from the
exponential family

Generalized linear models

- Each observation i from distribution with mean μ_i



Generalized linear models

$$E[\underline{Y}] = \underline{\mu} = g^{-1}(\underline{X} \cdot \underline{\beta} + \underline{\xi})$$

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$

Generalized linear models

$$E[\underline{Y}] = \underline{\mu} = g^{-1}(\underline{X} \cdot \underline{\beta})$$

Observed thing
(data)

Some function
(user defined)

Some matrix based on data
(user defined)

Parameters to be
estimated
(the answer!)

What is $g^{-1}(\mathbf{X}.\underline{\beta})$?

$$\underline{Y} = g^{-1}(\mathbf{X}.\underline{\beta}) + \text{error}$$

Assuming a model with three categorical factors,
each observation can be expressed as:

$$Y_{ijk} = g^{-1}(\alpha + \beta_i + \gamma_j + \delta_k) + \text{error}$$

$$\beta_2 = \gamma_1 = \delta_3 = 0$$

age is in group i

sex is in group j

car is in group k

What is $g^{-1}(\mathbf{X}.\underline{\beta})$?

- $g(x) = x \quad \Rightarrow \quad Y_{ijk} = \alpha + \beta_i + \gamma_j + \delta_k + \text{error}$
- $g(x) = \ln(x) \quad \Rightarrow \quad Y_{ijk} = e^{(\alpha + \beta_i + \gamma_j + \delta_k)} + \text{error}$
 $= A \cdot B_i \cdot C_j \cdot D_k + \text{error}$
where $B_i = e^{\beta_i}$ etc
- Multiplicative form common for frequency and amounts

Multiplicative model

\$ 207.10 x

Age	Factor
17	2.52
18	2.05
19	1.97
20	1.85
21-23	1.75
24-26	1.54
27-30	1.42
31-35	1.20
36-40	1.00
41-45	0.93
46-50	0.84
50-60	0.76
60+	0.78

Group	Factor
1	0.54
2	0.65
3	0.73
4	0.85
5	0.92
6	0.96
7	1.00
8	1.08
9	1.19
10	1.26
11	1.36
12	1.43
13	1.56

Sex	Factor
Male	1.00
Female	1.25

Area	Factor
A	0.95
B	1.00
C	1.09
D	1.15
E	1.18
F	1.27
G	1.36
H	1.44

$$E(\text{losses}) = \$ 207.10 \times 1.42 \times 0.92 \times 1.00 \times 1.15 = \$ 311.14$$

Generalized linear models

$$E[\underline{Y}] = \underline{\mu} = g^{-1}(\mathbf{X} \cdot \underline{\beta} + \xi)$$

"Offset"



Eg \underline{Y} = claim *numbers*

Smith: Male, 30, Ford, 1 years, 2 claims

Jones: Female, 40, VW, 1/2 year, 1 claim

What is ξ ?

- $g(x) = \ln(x)$
- $\xi_{ijk} = \ln(\text{exposure}_{ijk})$
- $E[Y_{ijk}] = e^{(\alpha + \beta_i + \gamma_j + \delta_k + \xi_{ijk})}$
 $= A \cdot B_i \cdot C_j \cdot D_k \cdot e^{(\ln(\text{exposure}_{ijk}))}$
 $= A \cdot B_i \cdot C_j \cdot D_k \cdot \text{exposure}_{ijk}$

Restricted models

$$E[Y] = \underline{\mu} = g^{-1} (\mathbf{X} \cdot \underline{\beta} + \xi)$$

Offset 

- Constrain model (eg increased limits, territory, amount of insurance, discounts)
- Other factors adjusted to compensate

Generalized linear models

$$E[\underline{Y}] = \underline{\mu} = g^{-1}(\underline{X} \cdot \underline{\beta} + \underline{\xi})$$

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$

Generalized linear models

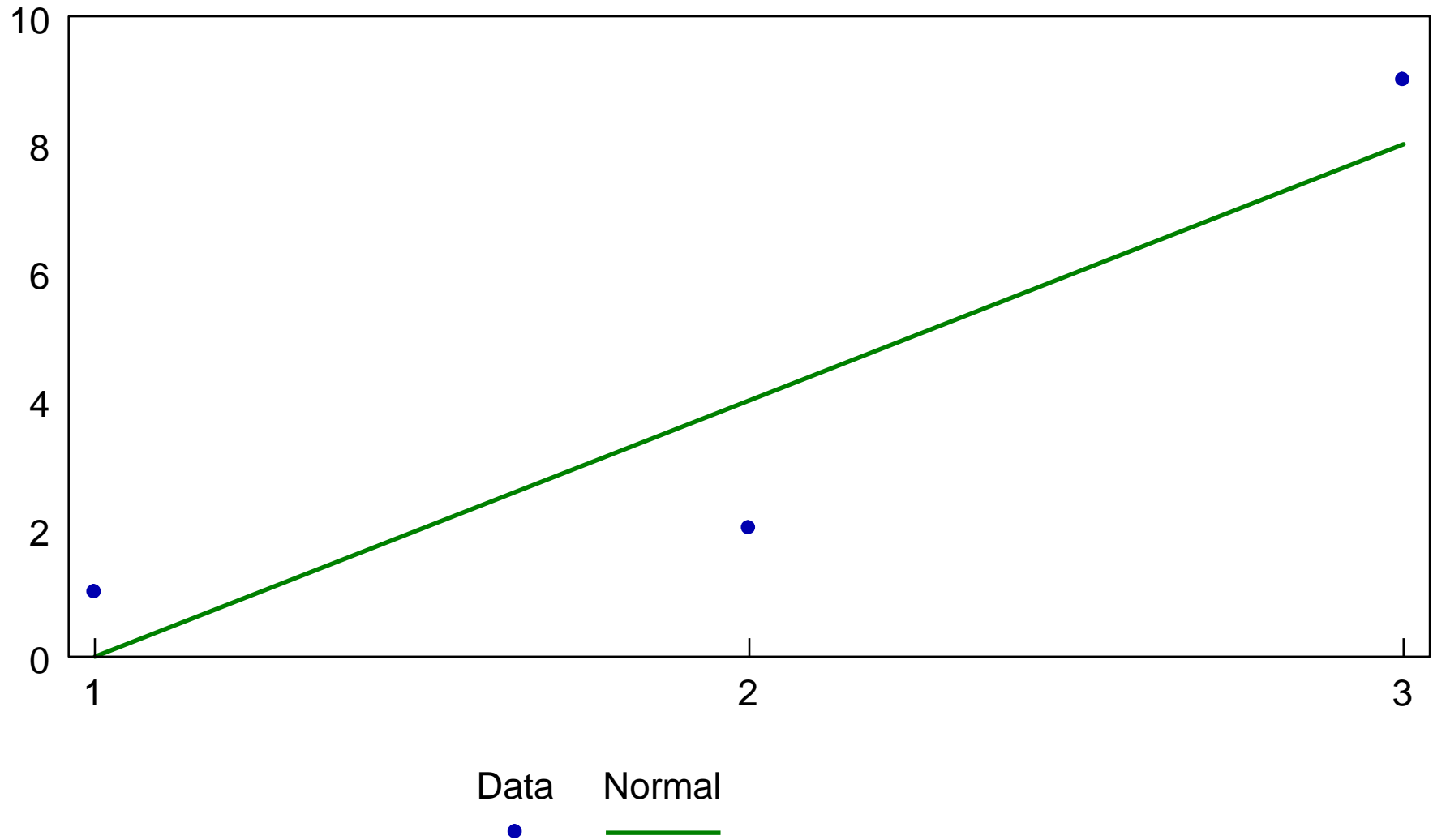
$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$

Normal: $\phi = \sigma^2$, $V(x) = 1 \Rightarrow \text{Var}[\underline{Y}] = \sigma^2 \cdot \underline{1}$

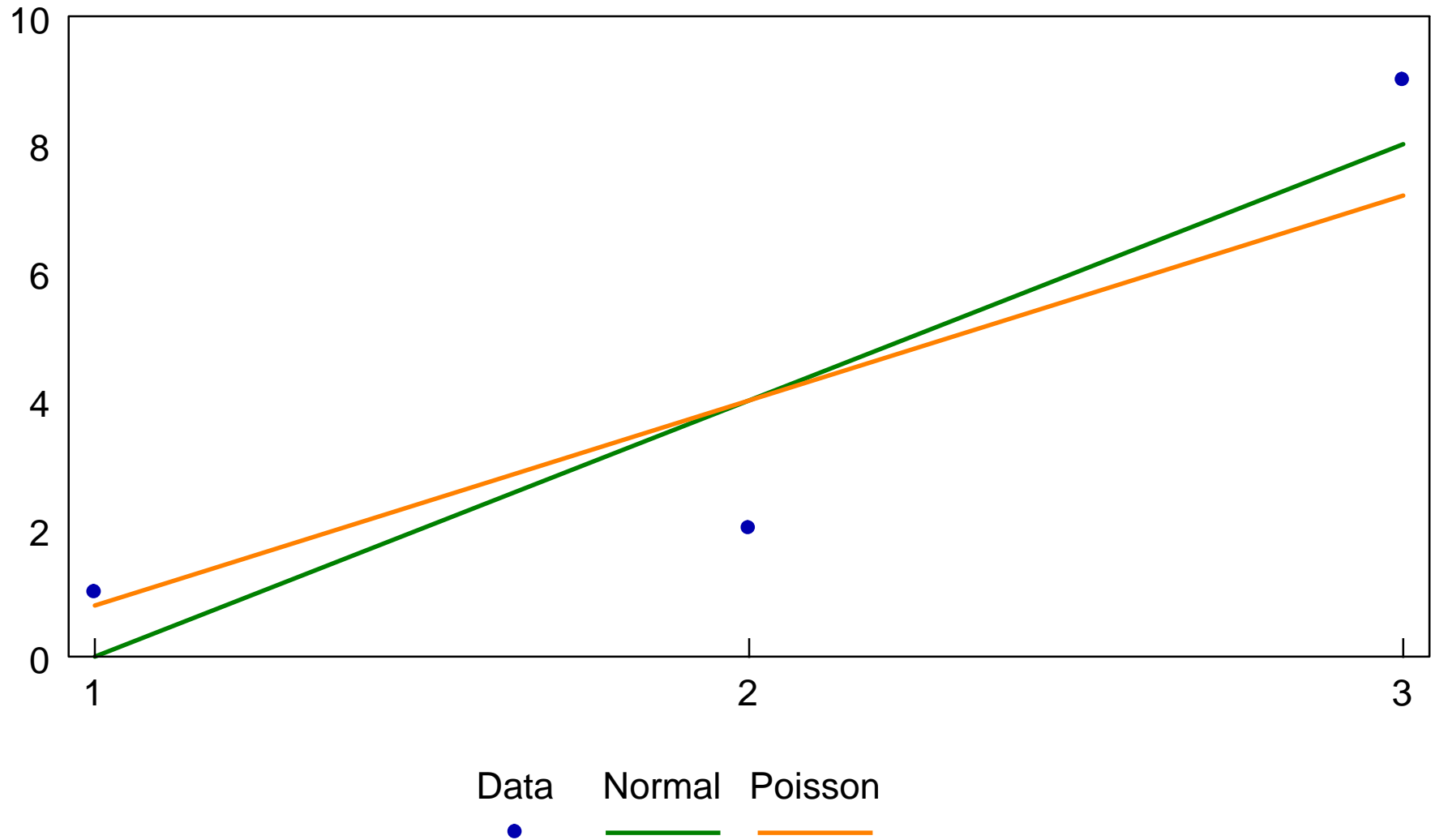
Poisson: $\phi = 1$, $V(x) = x \Rightarrow \text{Var}[\underline{Y}] = \underline{\mu}$

Gamma: $\phi = k$, $V(x) = x^2 \Rightarrow \text{Var}[\underline{Y}] = k \underline{\mu}^2$

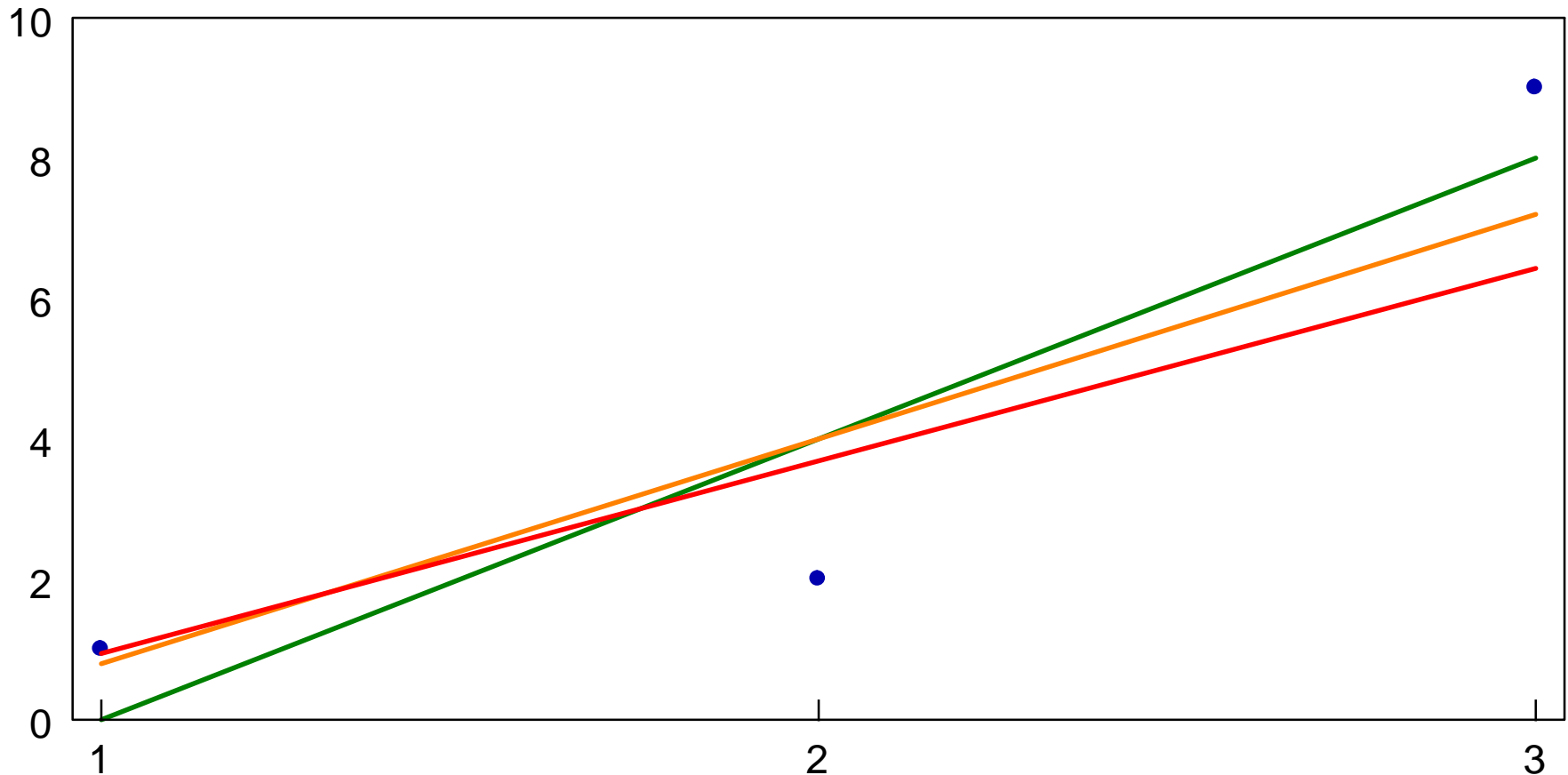
Example of effect of changing assumed error - 1



Example of effect of changing assumed error - 1



Example of effect of changing assumed error - 1

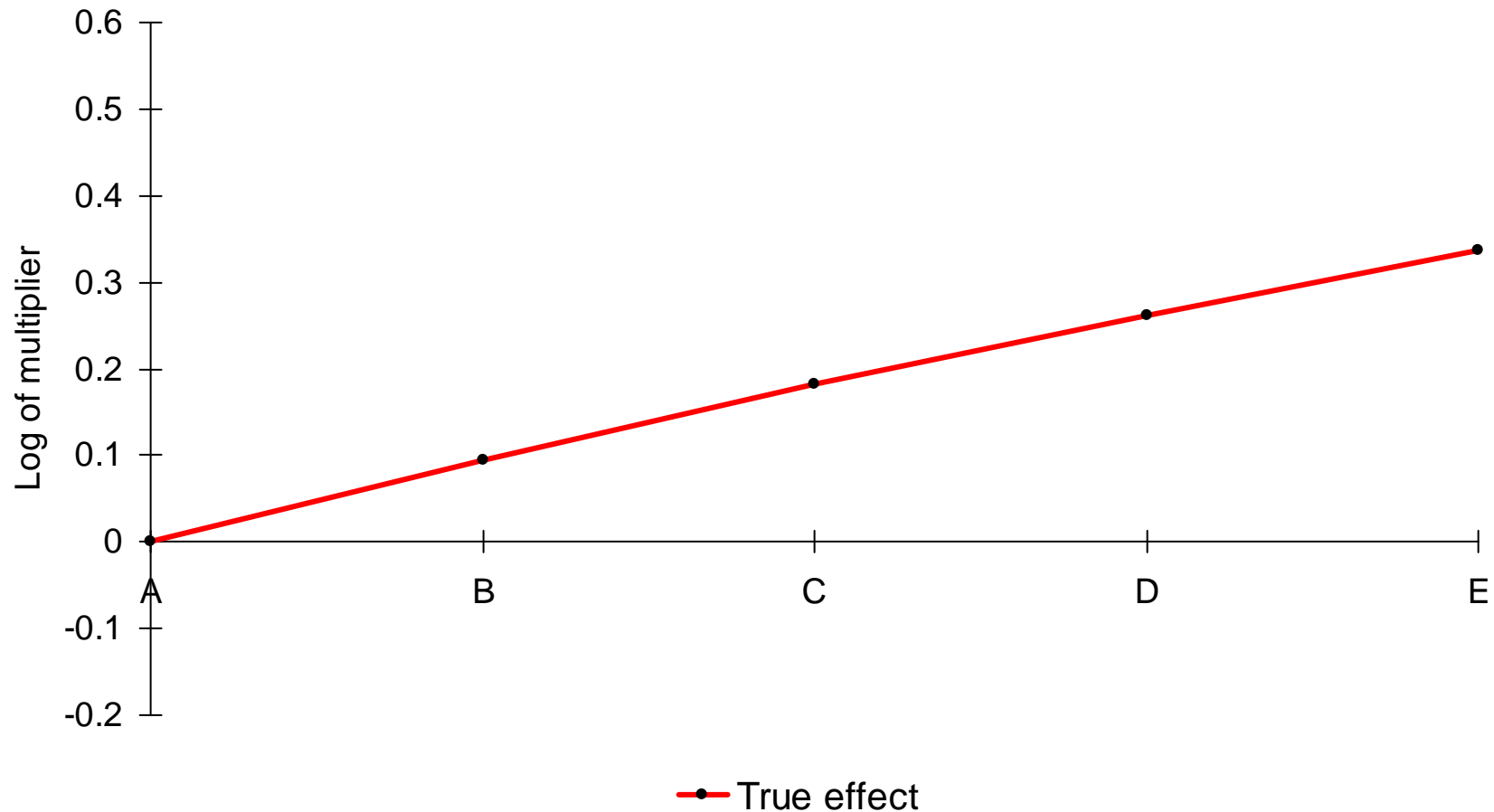


Data Normal Poisson Gamma
● — — —

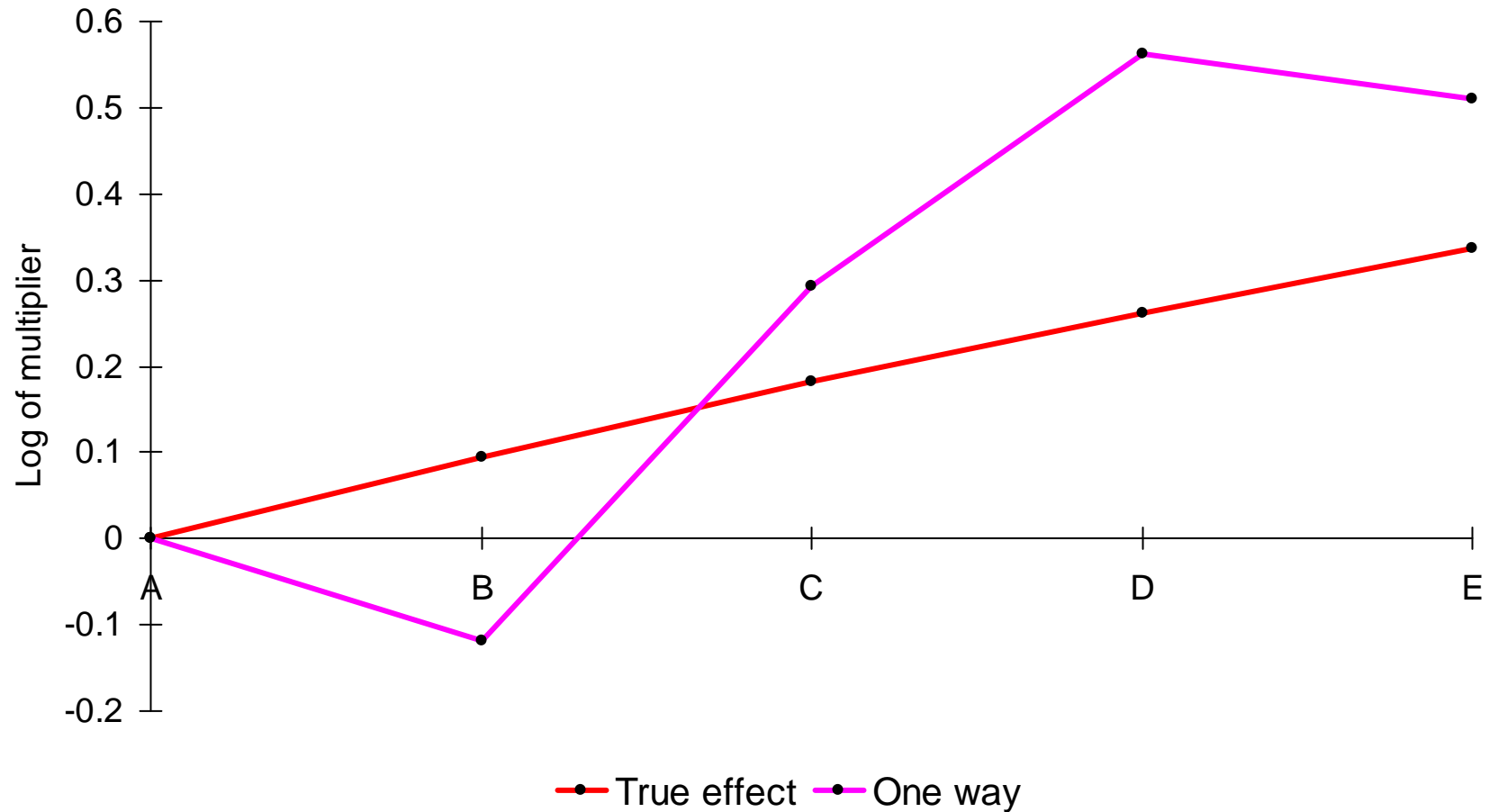
Example of effect of changing assumed error - 2

- Example portfolio with five rating factors, each with five levels A, B, C, D, E
- Typical correlations between those rating factors
- Assumed true effect of factors
- Claims randomly generated (with Gamma)
- Random experience analyzed by three models

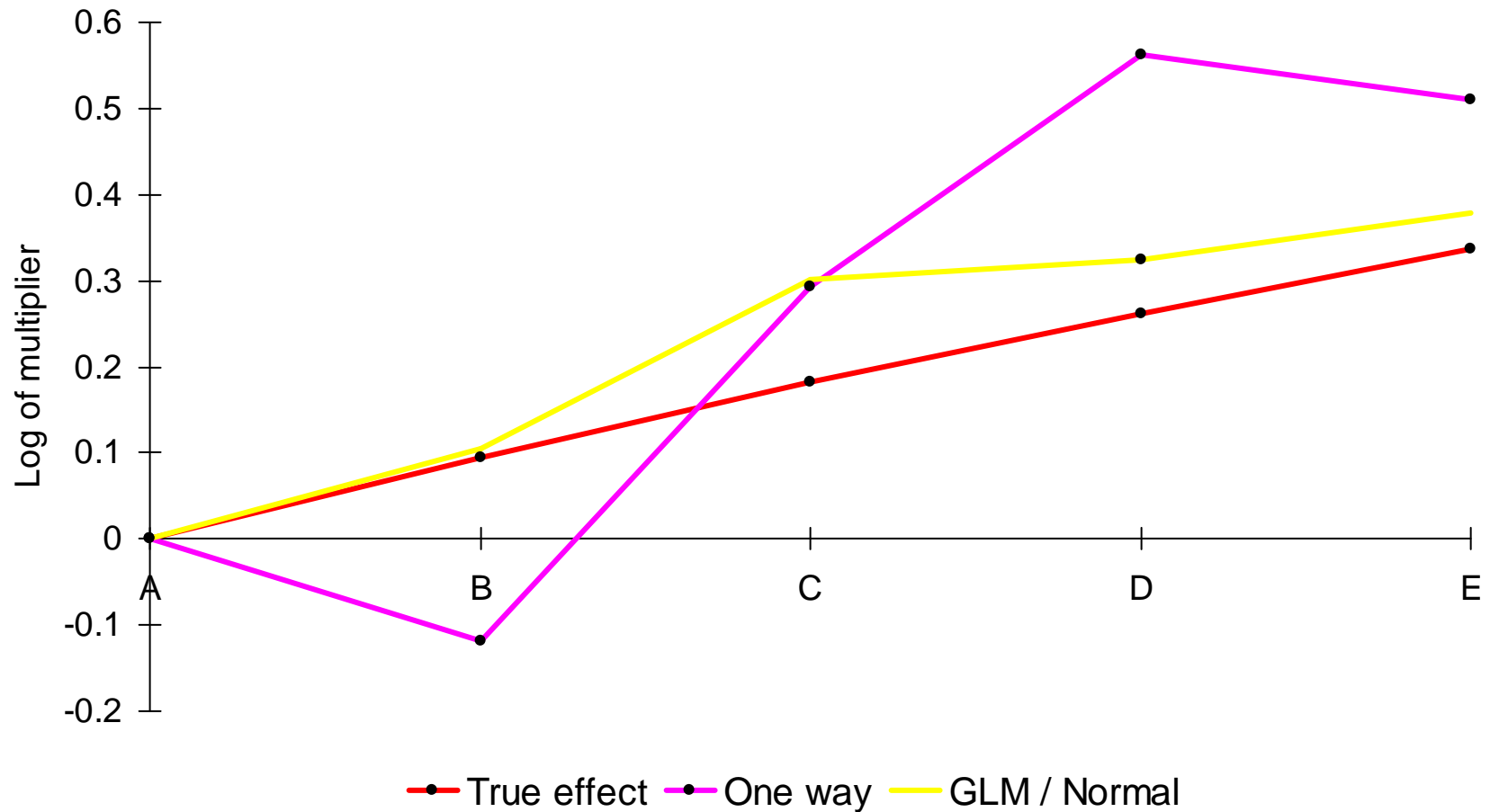
Example of effect of changing assumed error - 2



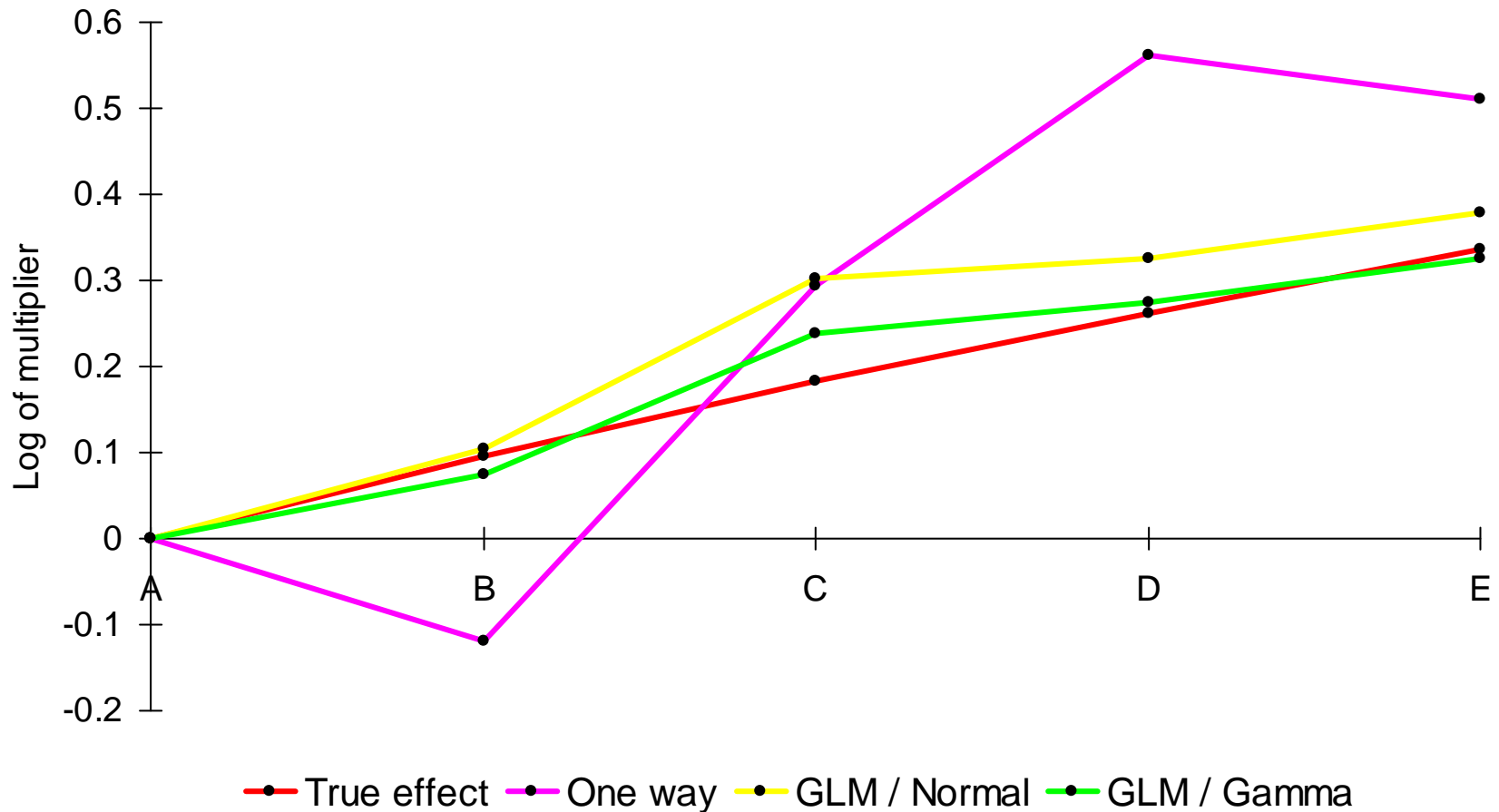
Example of effect of changing assumed error - 2



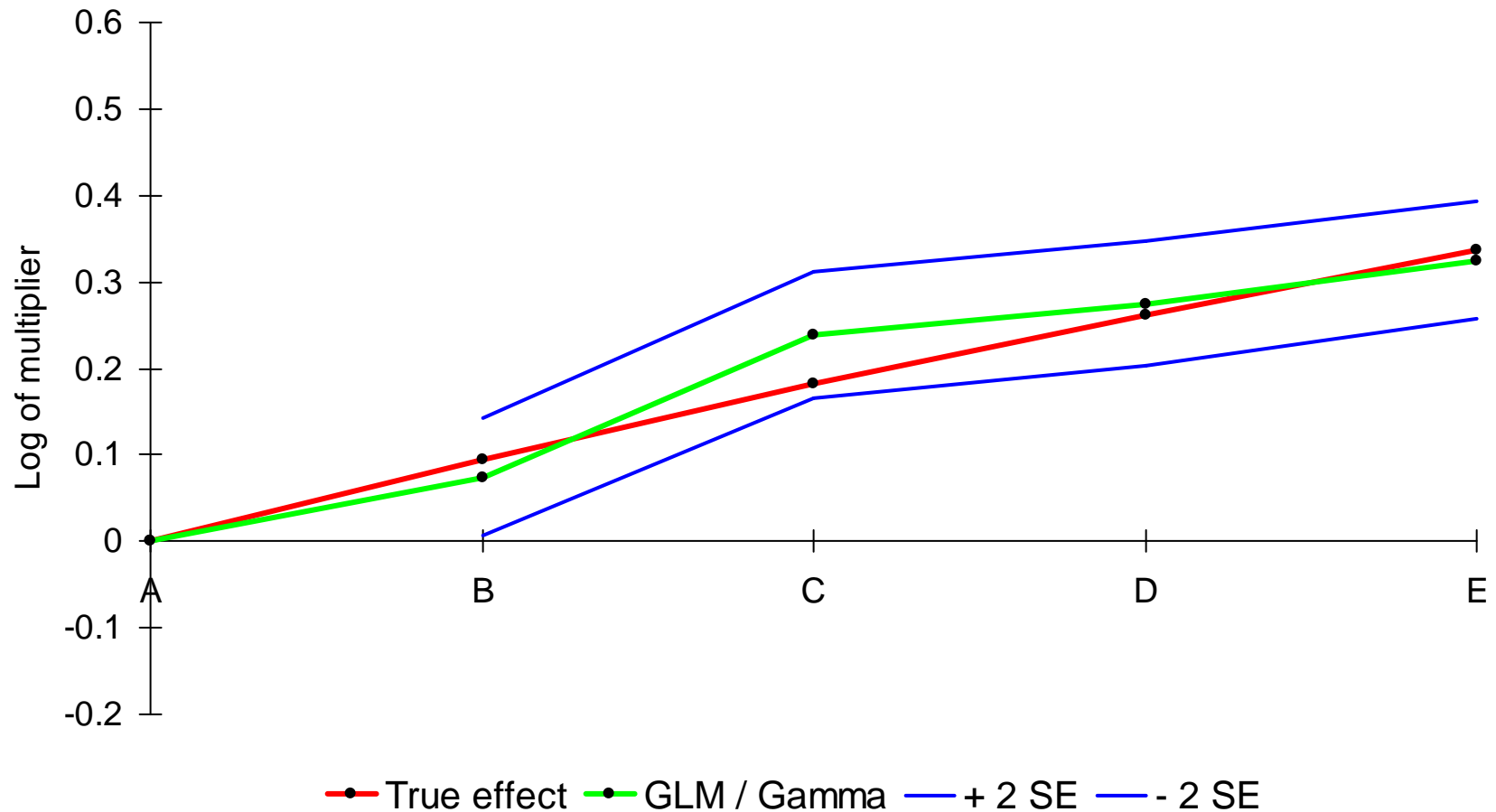
Example of effect of changing assumed error - 2



Example of effect of changing assumed error - 2



Example of effect of changing assumed error - 2



Prior weights

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$

- Exposure
- Other credibility

Eg \underline{Y} = claim *frequency*

Smith: Male, 30, Ford, 1 years, 2 claims, 100%

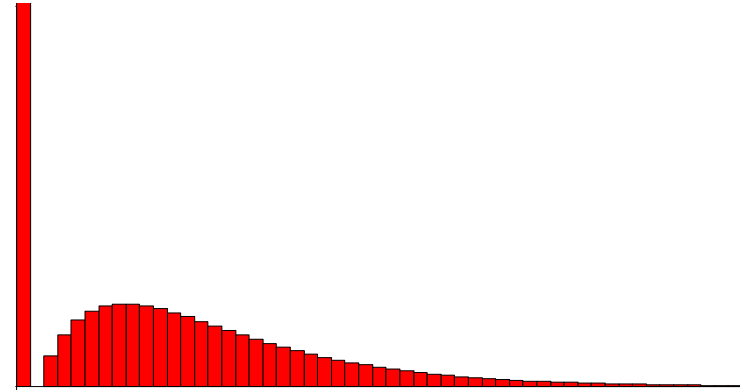
Jones: Female, 40, VW, 1/2 year, 1 claim, 100%

Typical model forms

\underline{Y}	Claim frequency	Claim number	Average claim amount	Probability (eg lapses)
$g(x)$	$\ln(x)$	$\ln(x)$	$\ln(x)$	$\ln(x/(1-x))$
Error	Poisson	Poisson	Gamma	Binomial
ϕ $V(x)$	$\frac{1}{x}$	$\frac{1}{x}$	estimate x^2	$\frac{1}{x(1-x)}$
$\underline{\omega}$	exposure	1	# claims	1
$\underline{\omega}$	0	$\ln(\text{exposure})$	0	0

Tweedie distributions

- Incurred losses have a point mass at zero and then a continuous distribution
- Poisson and gamma not suited to this
- Tweedie distribution has point mass and parameters which can alter the shape to be like Poisson and gamma above zero



$$f_Y(y; \theta, \lambda, \alpha) = \sum_{n=1}^{\infty} \frac{\{(\lambda\omega)^{1-\alpha} \kappa_{\alpha}(-1/y)\}^n}{\Gamma(-n\alpha)n!y} \cdot \exp\{\lambda\omega[\theta_0 y - \kappa_{\alpha}(\theta_0)]\} \quad \text{for } y > 0$$

$$p(Y = 0) = \exp\{-\lambda\omega\kappa_{\alpha}(\theta_0)\}$$

Generalized linear models

$$\text{Var}[\underline{Y}] = \phi \cdot V(\underline{\mu}) / \underline{\omega}$$

Normal: $\phi = \sigma^2$, $V(x) = 1 \Rightarrow \text{Var}[\underline{Y}] = \sigma^2 \cdot \underline{1}$

Poisson: $\phi = 1$, $V(x) = x \Rightarrow \text{Var}[\underline{Y}] = \underline{\mu}$

Gamma: $\phi = k$, $V(x) = x^2 \Rightarrow \text{Var}[\underline{Y}] = k \underline{\mu}^2$

Tweedie: $\phi = k$, $V(x) = x^p \Rightarrow \text{Var}[\underline{Y}] = k \underline{\mu}^p$

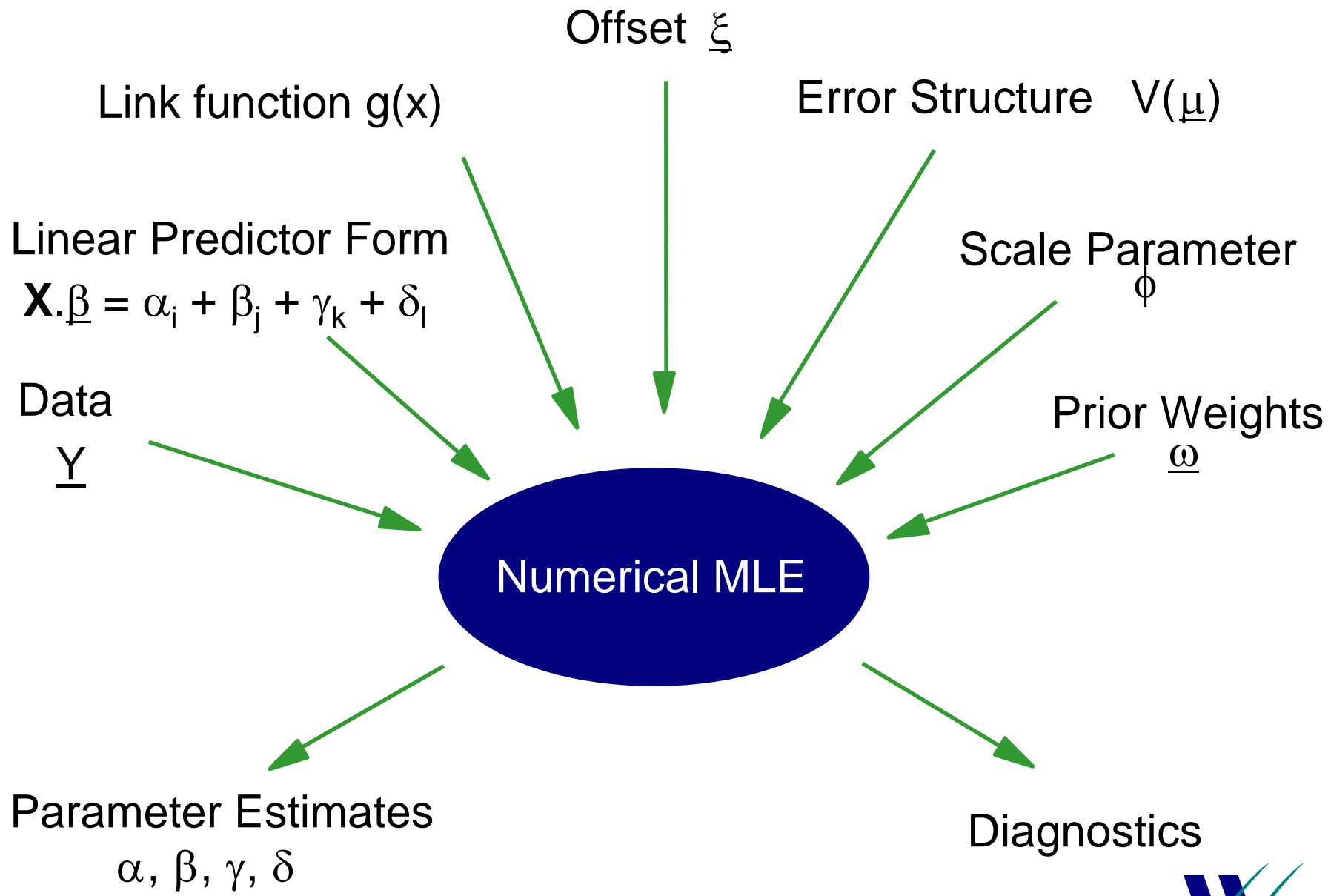
Tweedie distributions

Tweedie: $\phi = k$, $V(x) = x^p \Rightarrow \text{Var}[Y] = k\mu^p$

- Defines a valid distribution for $p < 0$, $1 < p < 2$, $p > 2$
- Can be considered as Poisson/gamma process for $1 < p < 2$
- Typical values of p for insurance incurred claims around, or just under, 1.5

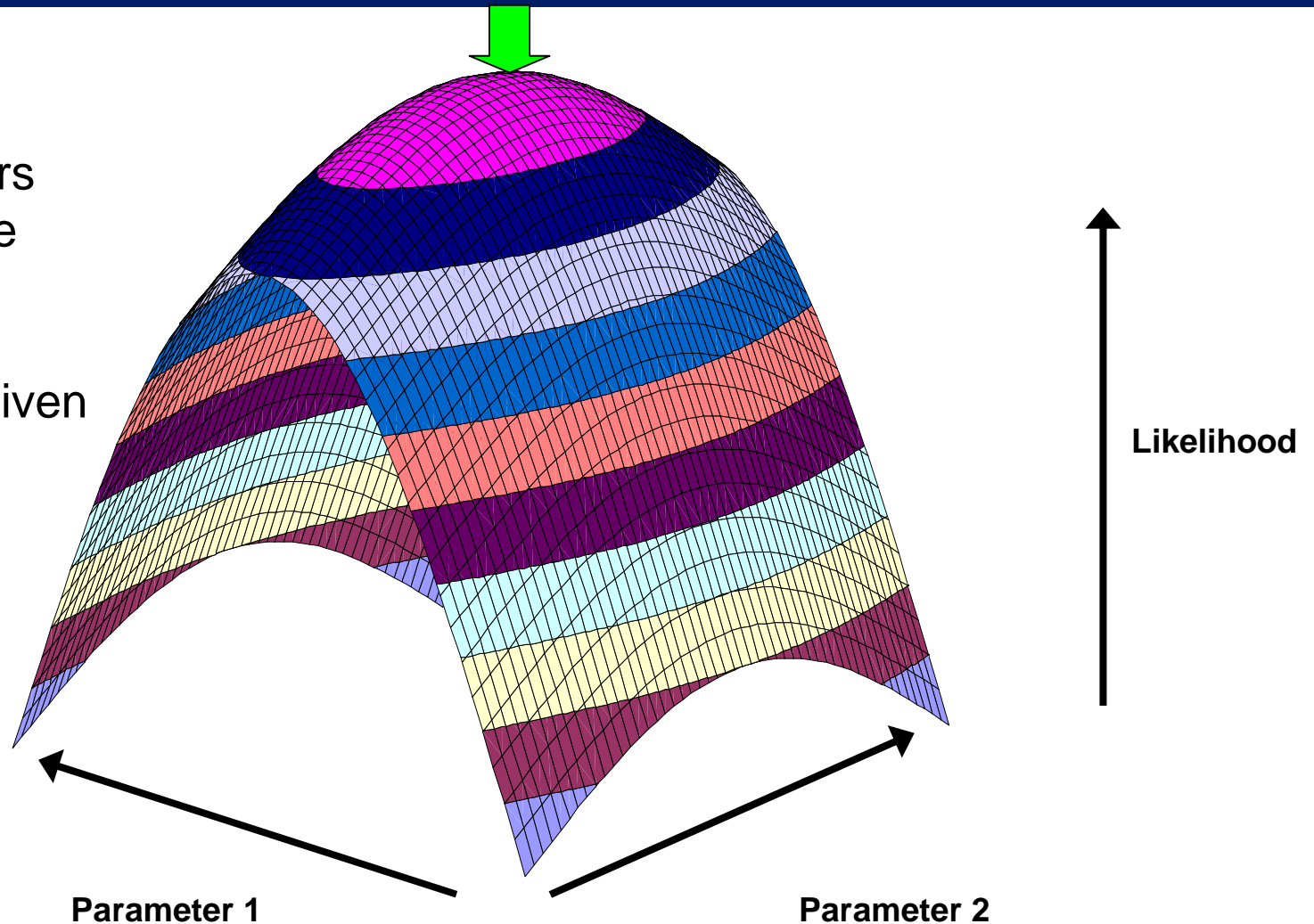
Tweedie distributions

- Helpful when important to fit to pure premium
- Often similar results to traditional approach but differences may occur if numbers and amounts models have effects which are both large and insignificant
- No information about whether frequencies or amounts are driving result



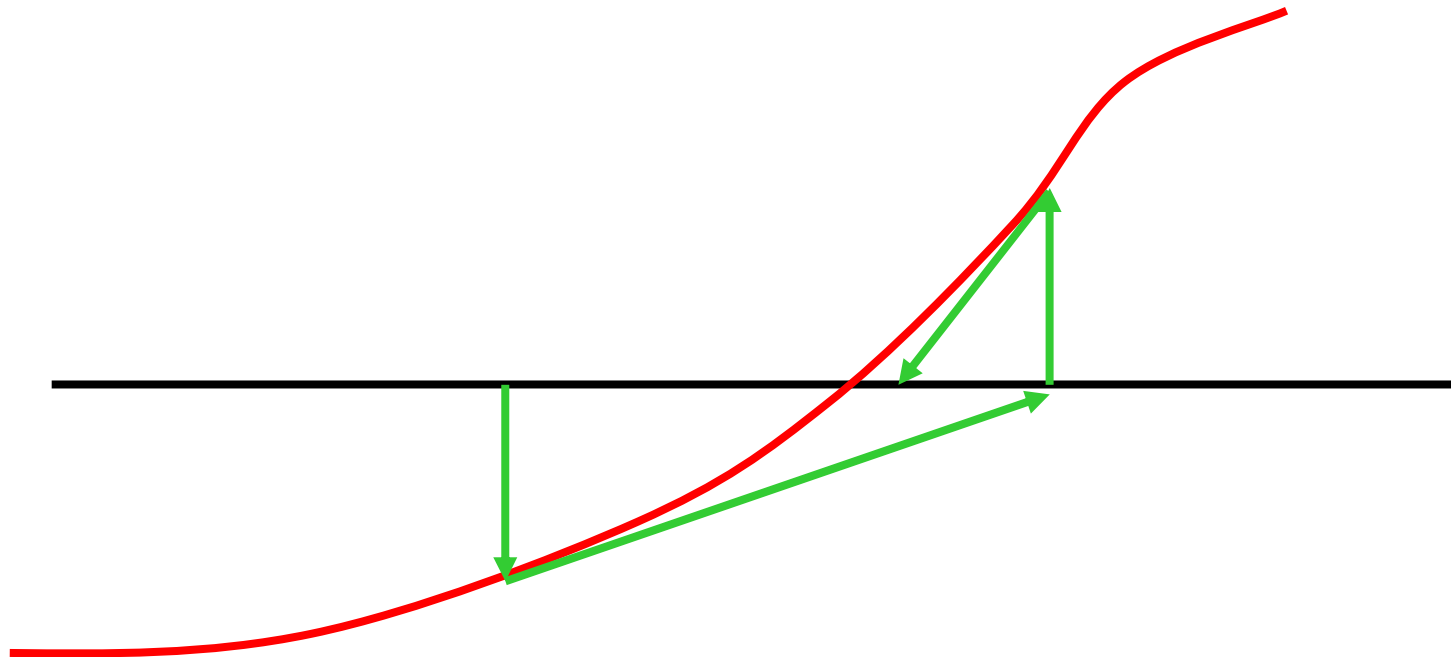
Maximum likelihood estimation

- Seek parameters which give highest likelihood function given data



Newton-Raphson

- In one dimension: $x_{n+1} = x_n - f'(x_n) / f''(x_n)$



- In n dimensions: $\underline{\beta}_{n+1} = \underline{\beta}_n - \mathbf{H}^{-1} \cdot \underline{s}$

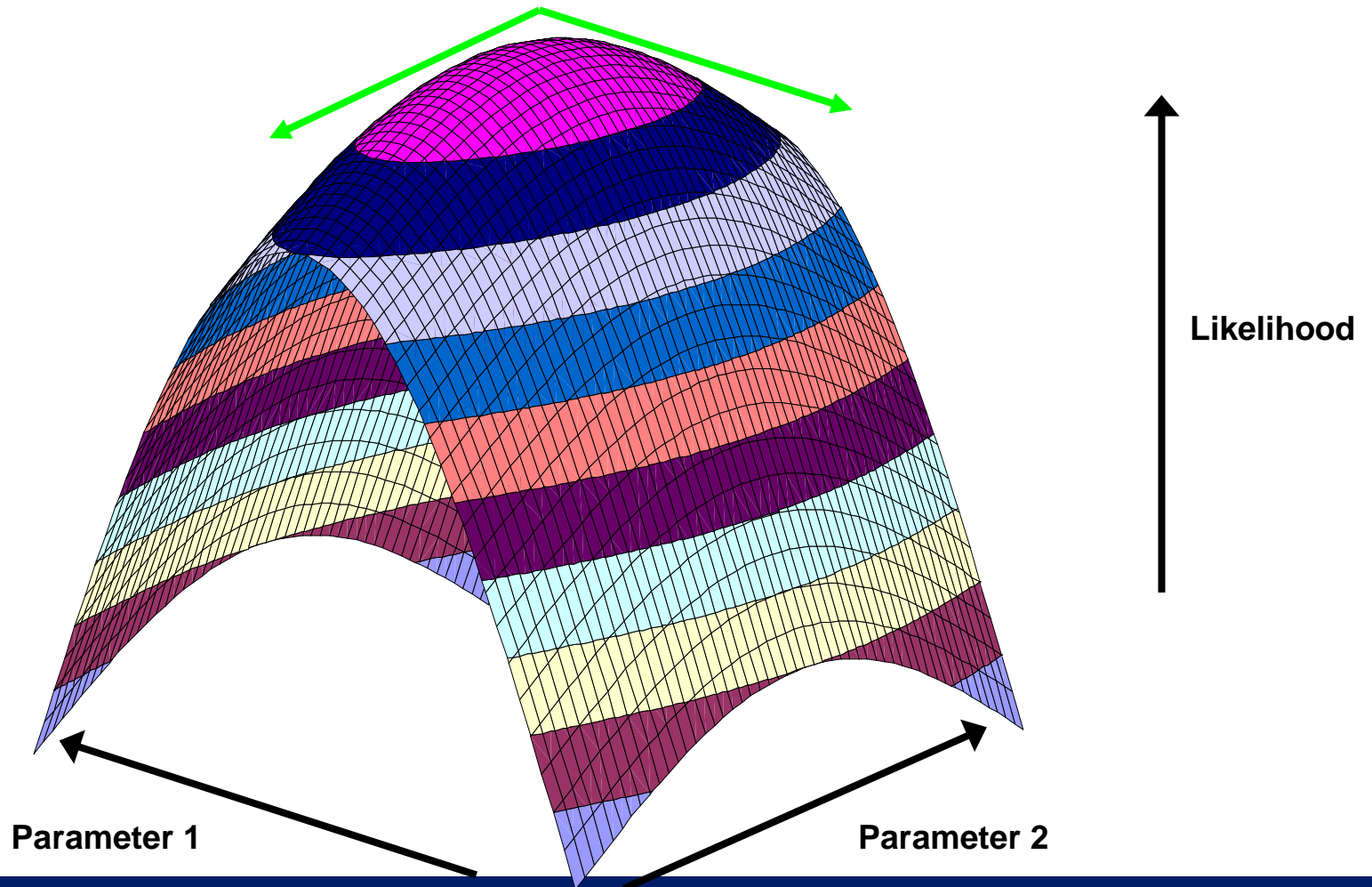
where $\underline{\beta}$ is the vector of the parameter estimates (with p elements), \underline{s} is the vector of the first derivatives of the log-likelihood and \mathbf{H} is the $(p \times p)$ matrix containing the second derivatives of the log-likelihood

Agenda

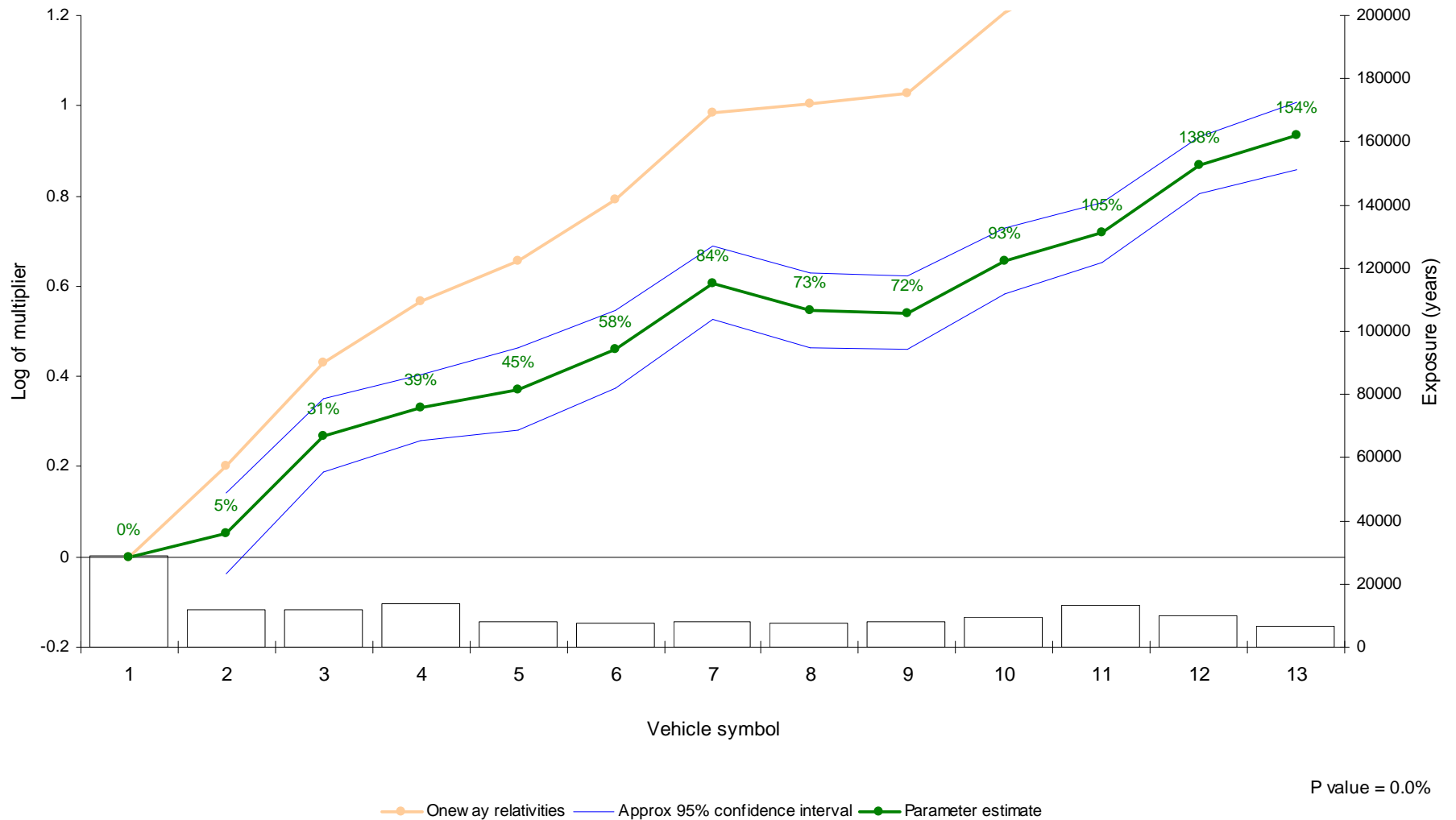
- GLM formulae
- Model testing
 - use only variables that are predictive
 - make sure model is reasonable
- Aliasing
- Model refinements

Standard errors

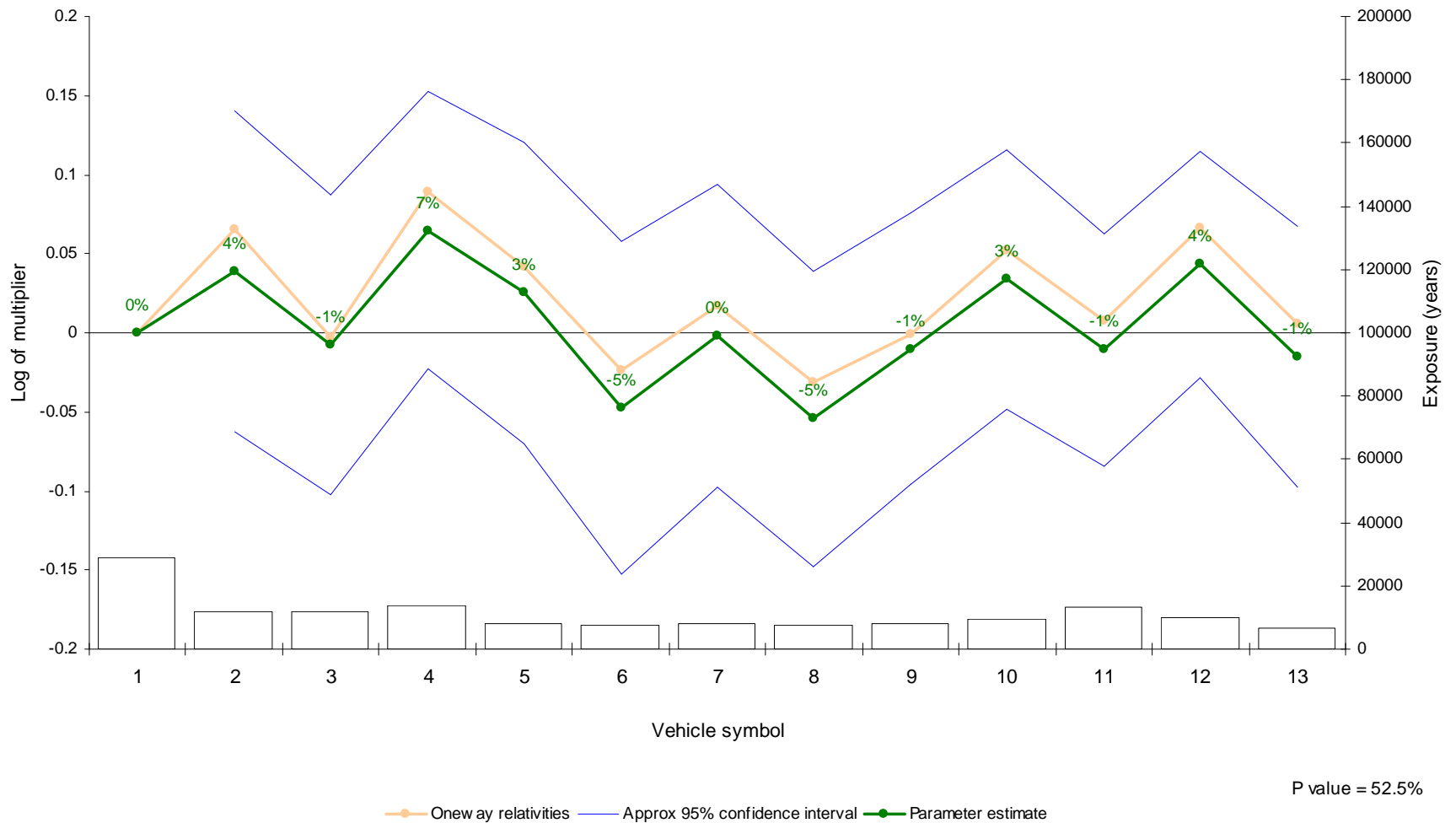
- Roughly speaking, for a parameter p : $SE = -1 / (\partial^2 / \partial p^2 \text{ Likelihood})$



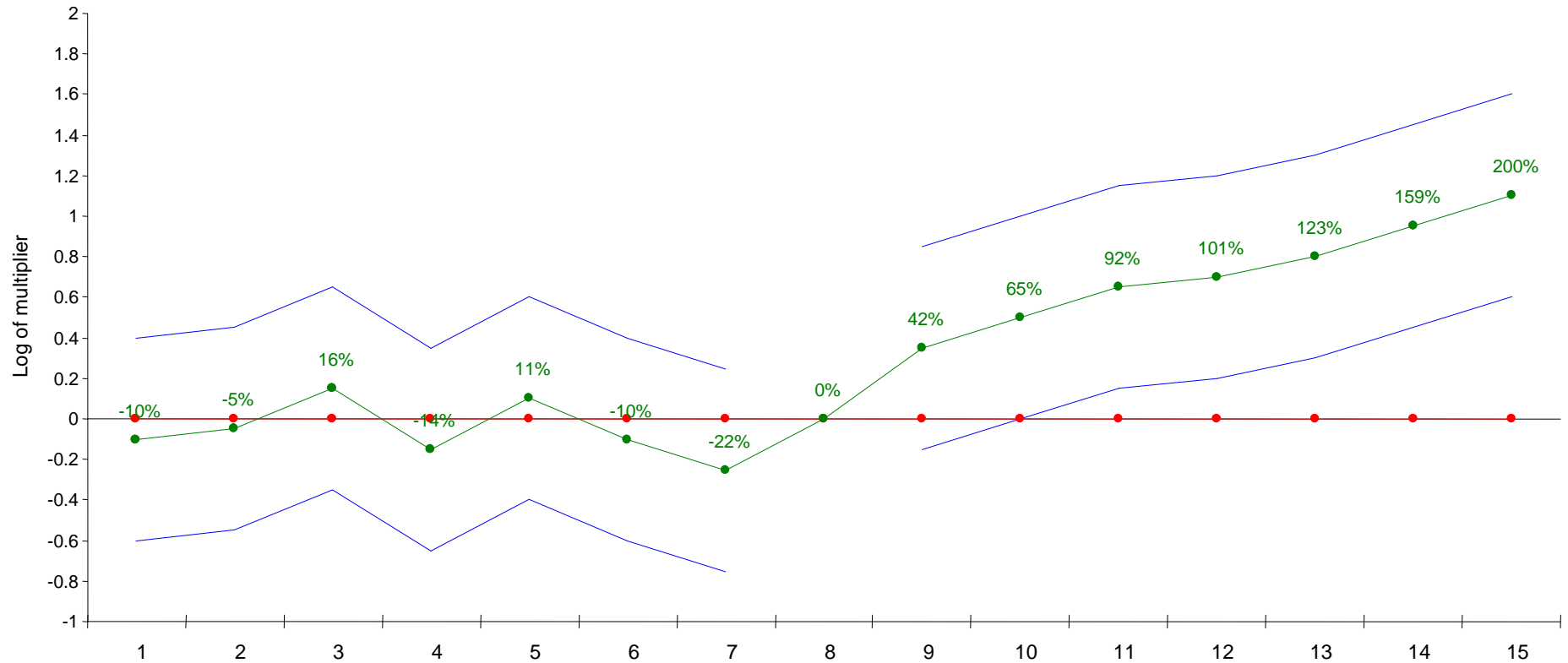
GLM output (significant factor)



GLM output (insignificant factor)



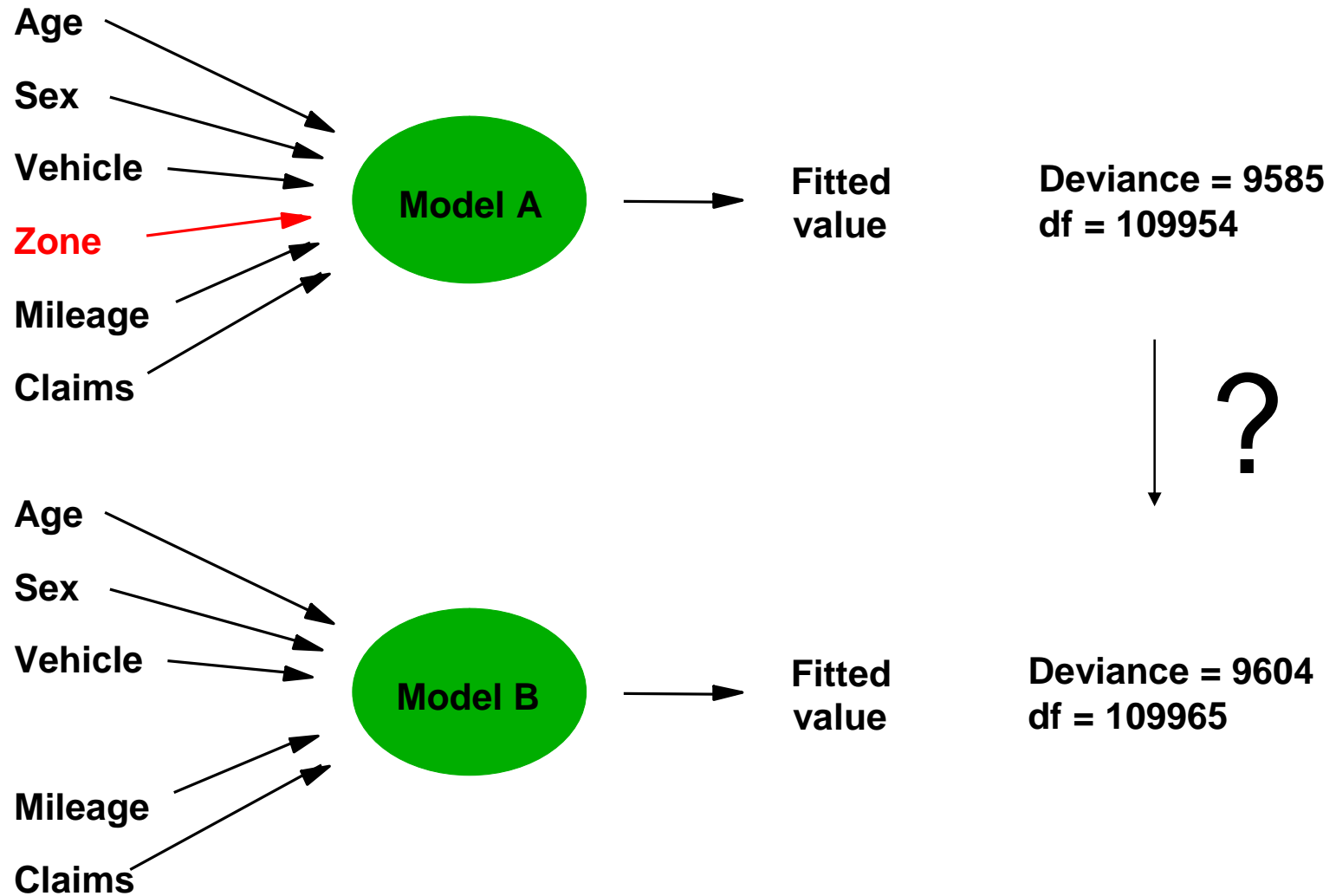
Awkward cases



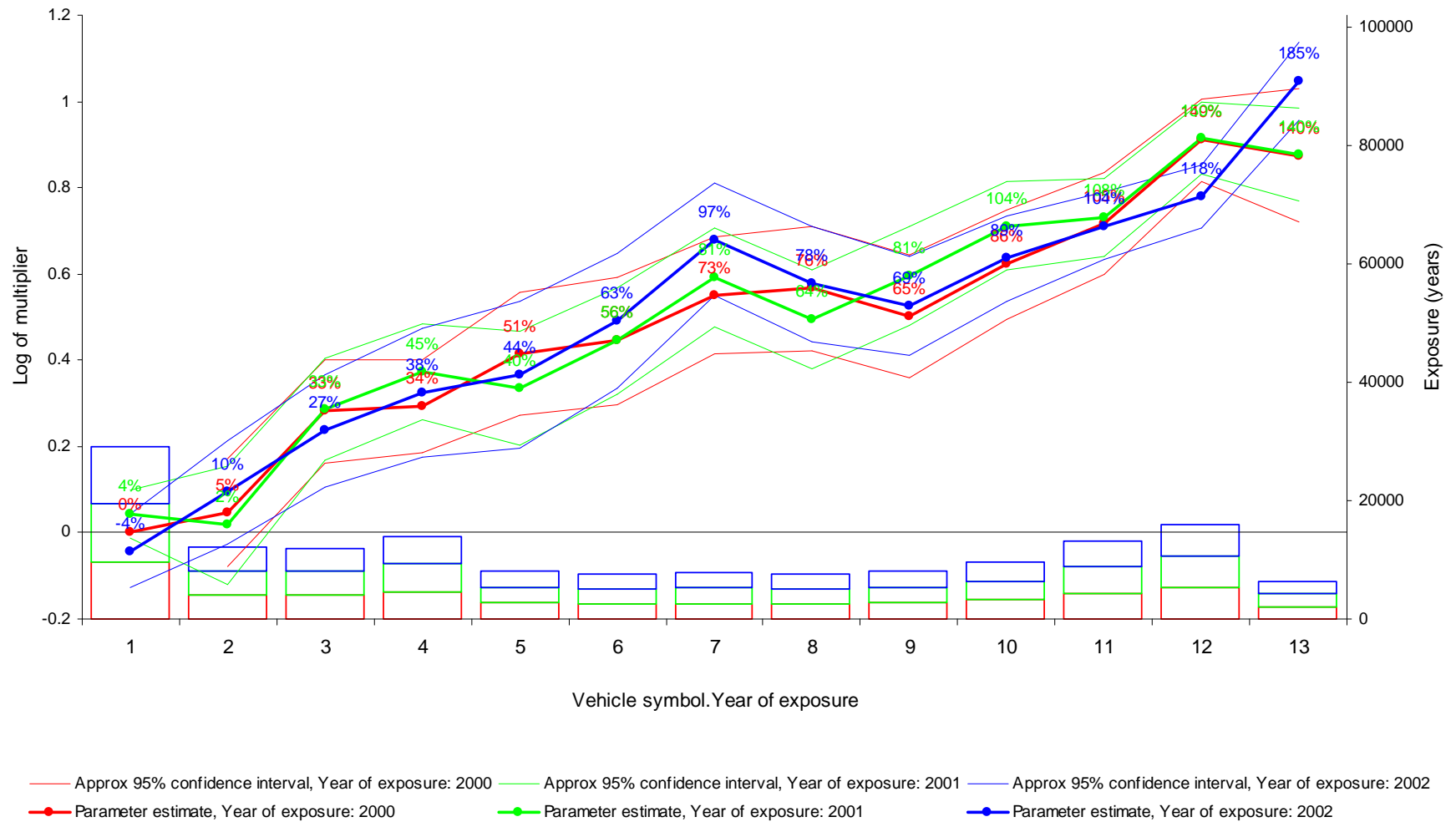
Deviances

- Single figure measure of goodness of fit
- Try model with & without a factor
- Statistical tests show the theoretical significance given the extra parameters

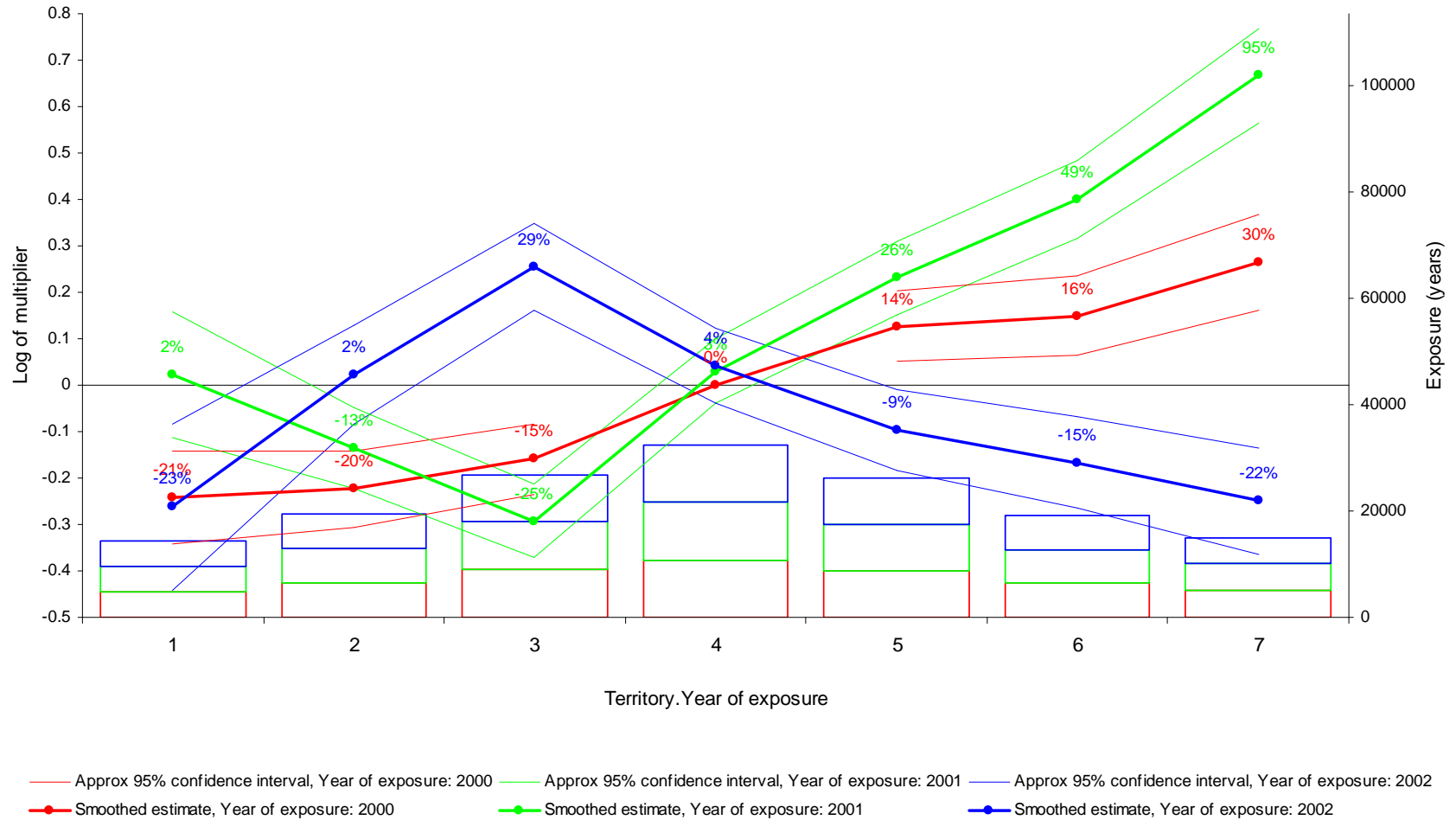
Deviances



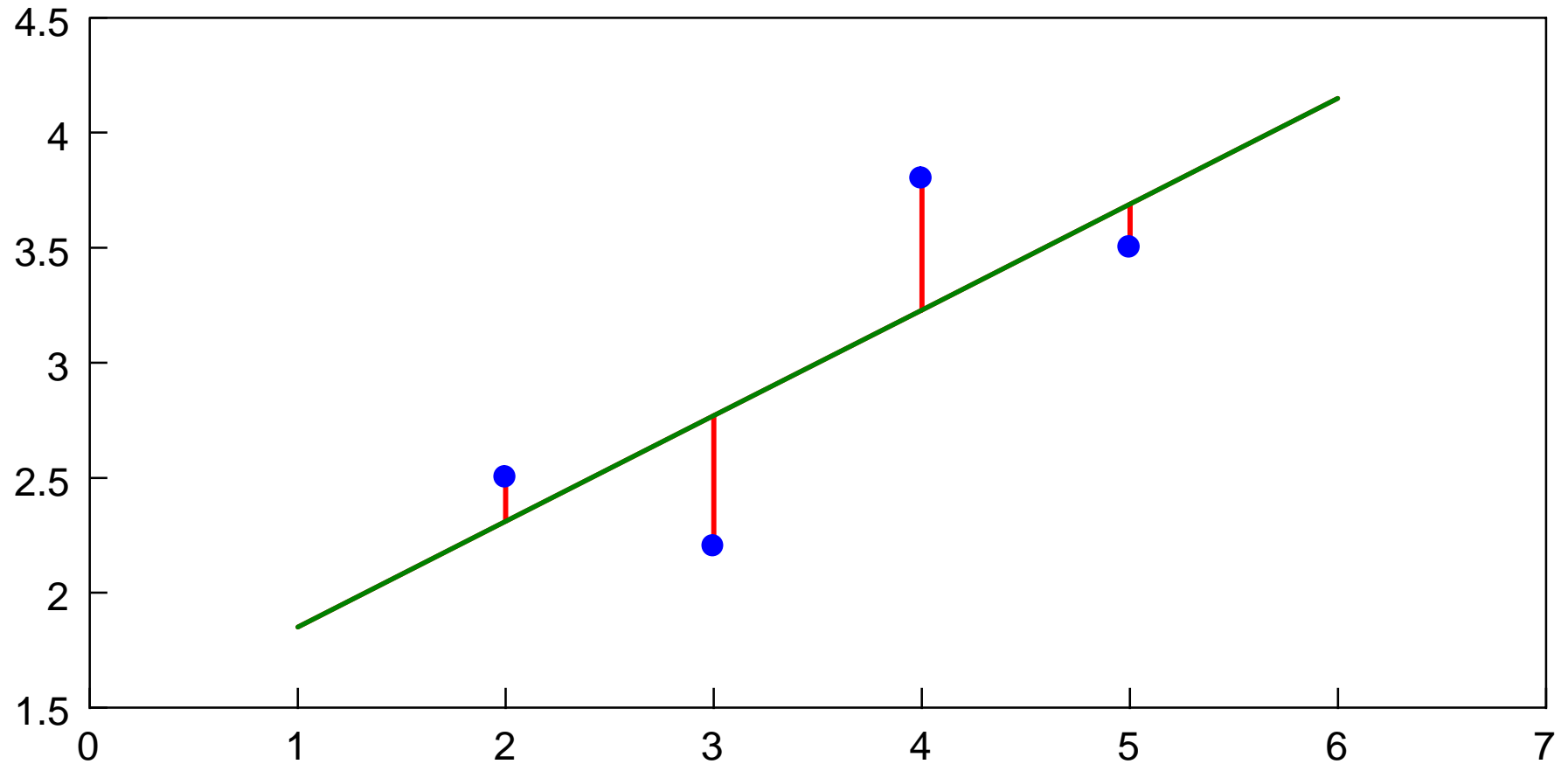
Consistency over time



Consistency over time



Residuals



Residuals Fitted values Data



Residuals

- Several forms, eg
 - standardized deviance

$$\text{sign} (Y_u - \mu_u) / (\phi (1-h_u))^{1/2} \sqrt{2 \omega_u \int_{\mu_u}^{Y_u} (Y_u - \zeta) / V(\zeta) d\zeta}$$

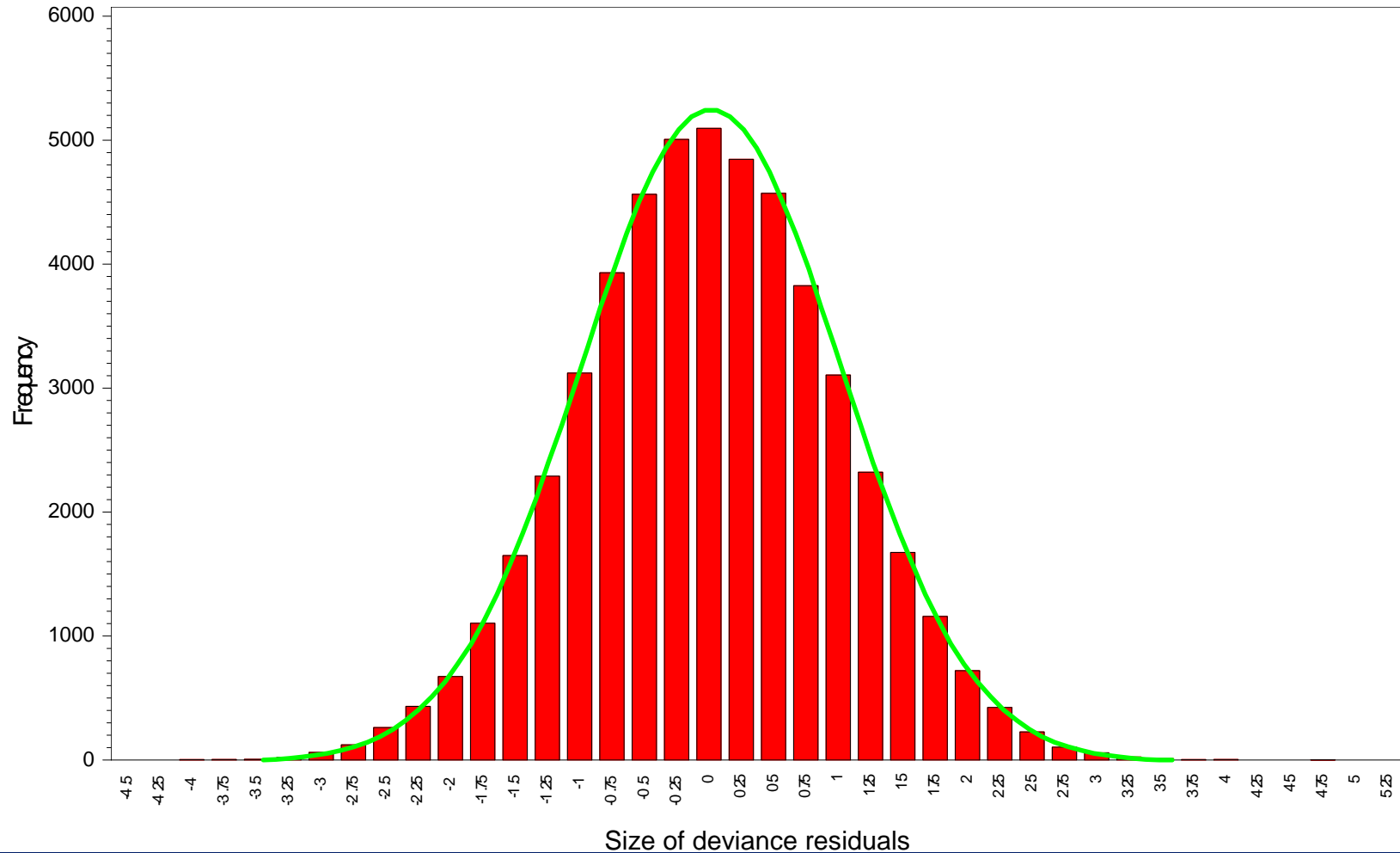
- standardized Pearson

$$\frac{Y_u - \mu_u}{(\phi \cdot V(\mu_u) \cdot (1-h_u) / \omega_u)^{1/2}}$$

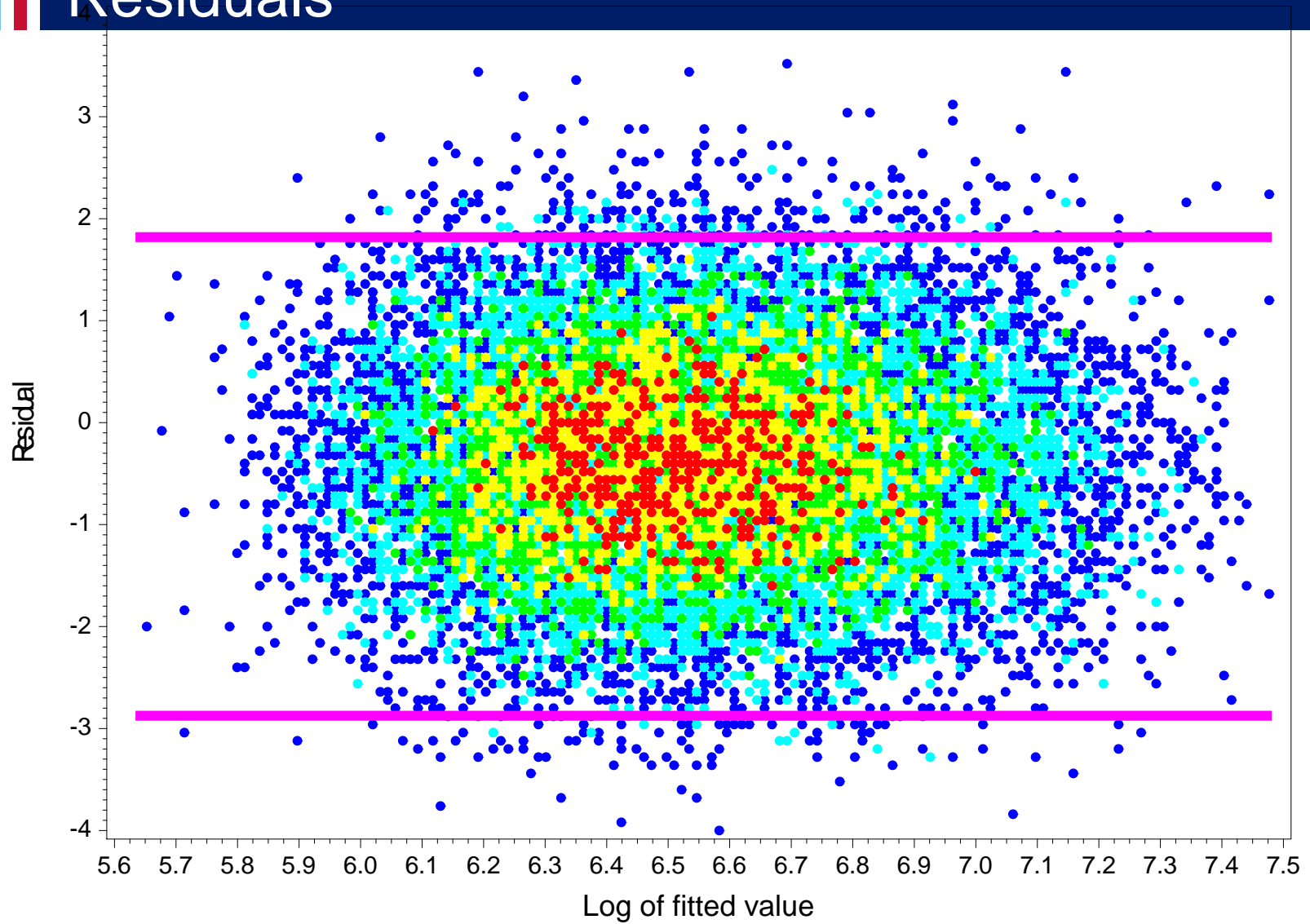
- Standardized deviance - Normal (0,1)
- Numbers/frequency residuals problematical

Residuals

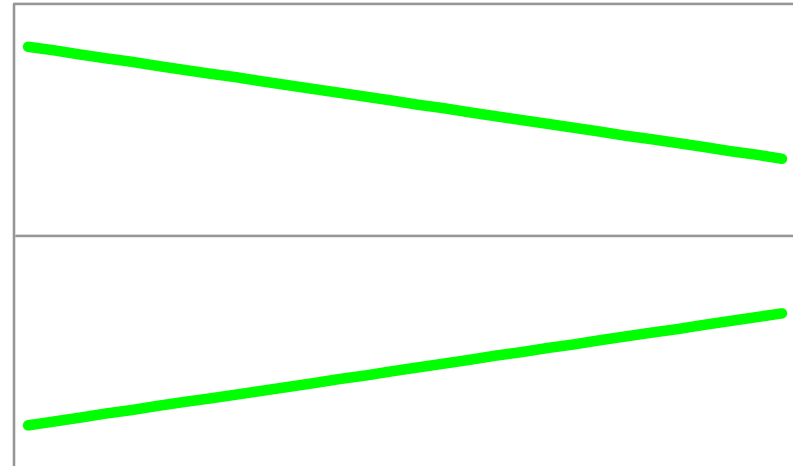
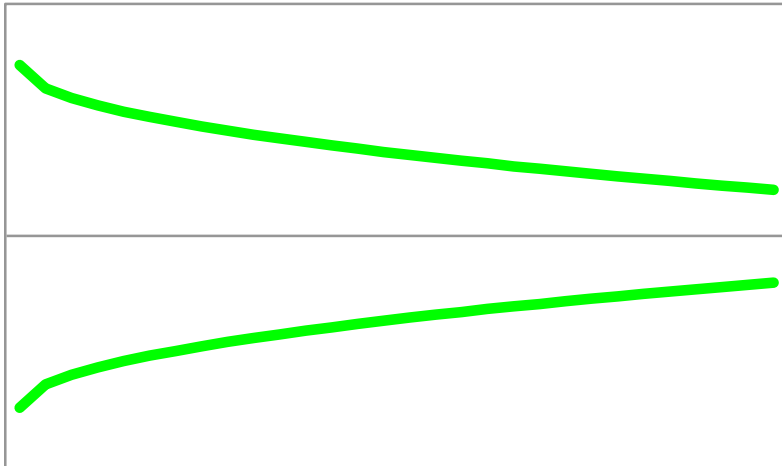
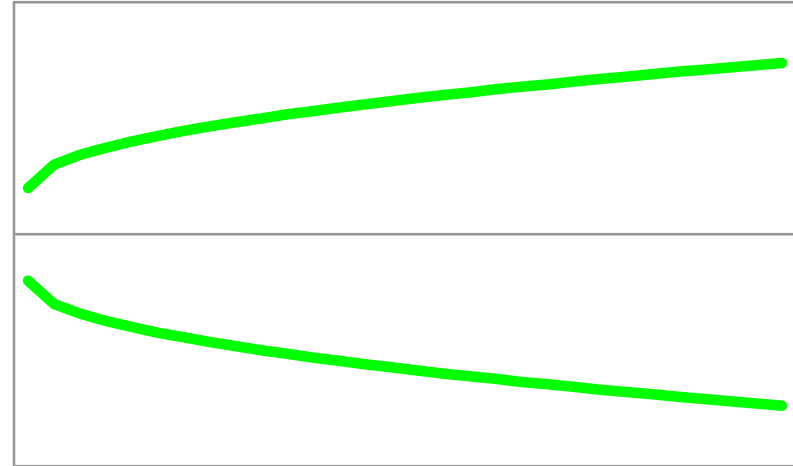
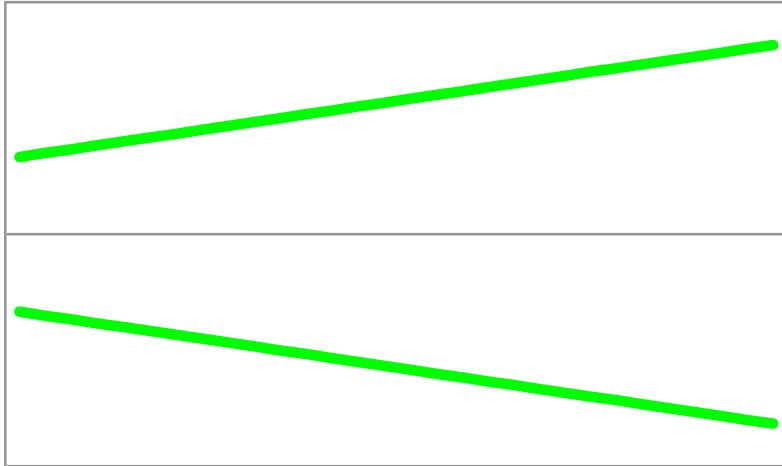
Histogram of Deviance Residuals
Run 12 (Final models with analysis) Model 8 (AD amounts)



Residuals

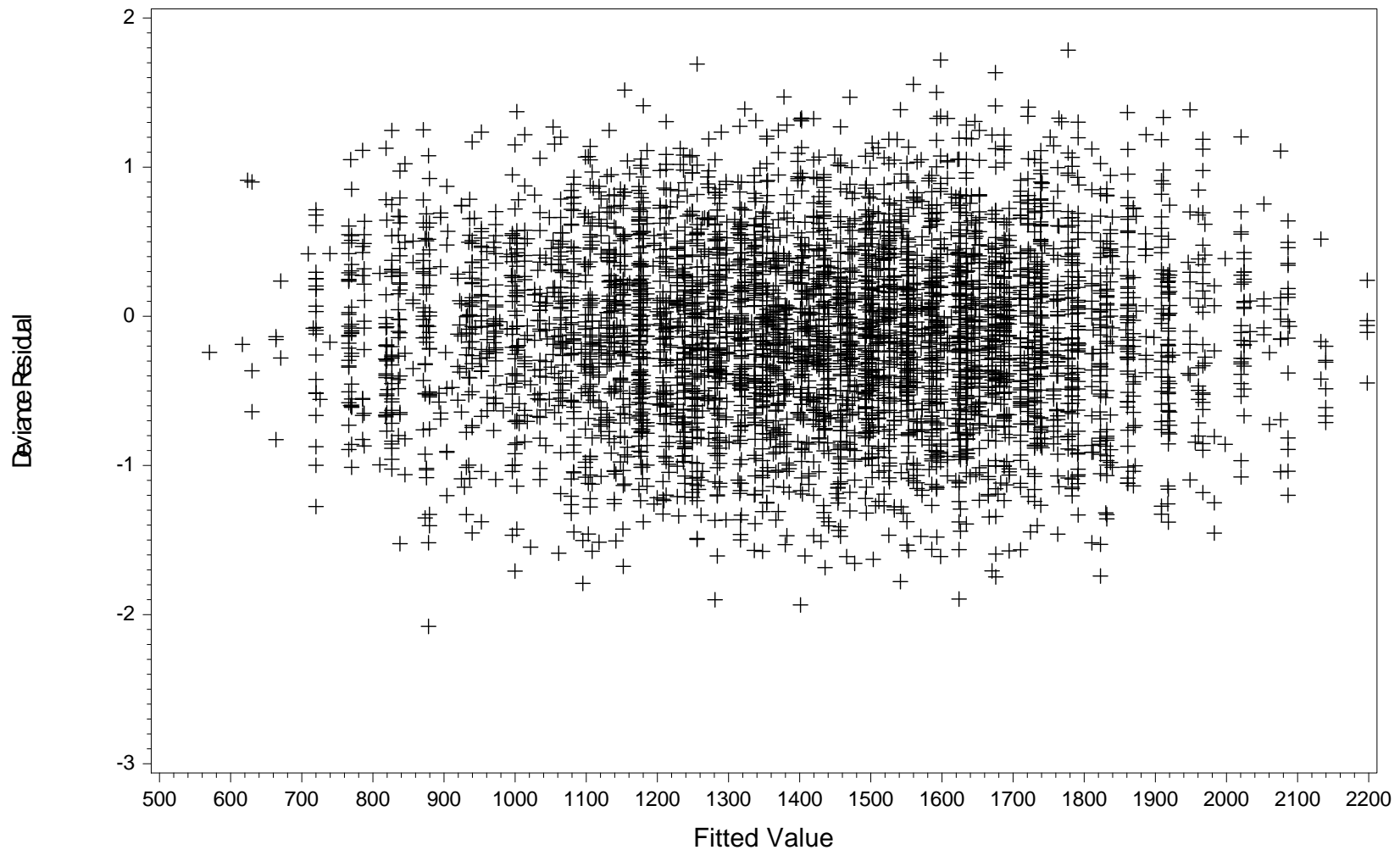


Residuals



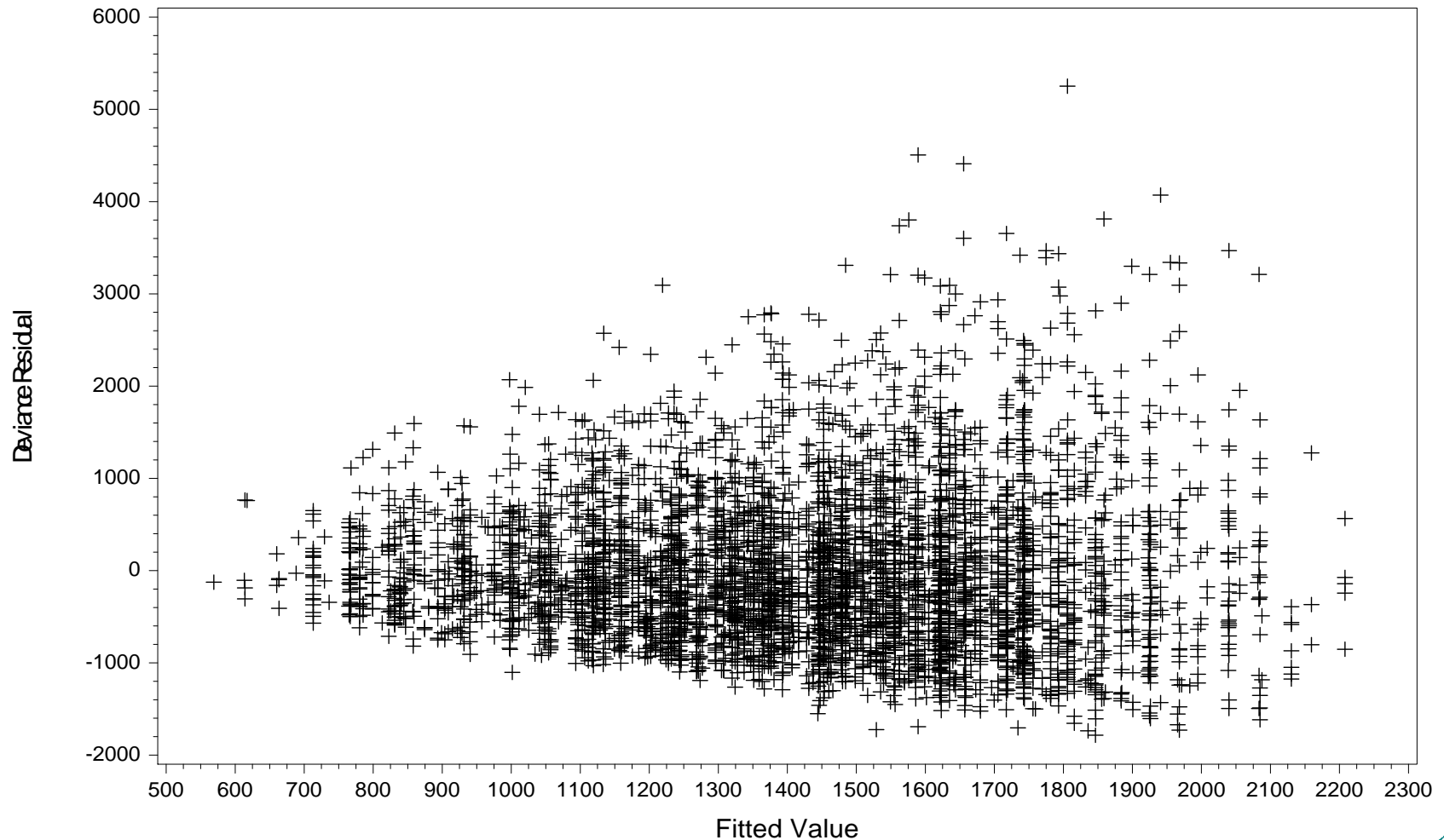
Gamma data, Gamma error

Plot of deviance residual against fitted value
Run 12 (All claim types, final models, N&A) Model 6 (Own damage, Amounts)



Gamma data, Normal error

Plot of deviance residual against fitted value
Run 12 (All claim types, final models, N&A) Model 7 (Own damage, Amounts)



Agenda

- GLM formulae
- Model testing
 - use only variables that are predictive
 - make sure model is reasonable
- Aliasing
- Model refinements

Aliasing and "near aliasing"

- Aliasing
 - the removal of unwanted redundant parameters
- Intrinsic aliasing
 - occurs by the design of the model
- Extrinsic aliasing
 - occurs "accidentally" as a result of the data

Intrinsic aliasing

$X.\beta = \alpha + \beta_1$ if age 20 - 29

~~$+ \beta_2$ if age 30 - 39~~

$+ \beta_3$ if age 40 +

~~$+ \gamma_1$ if sex male~~

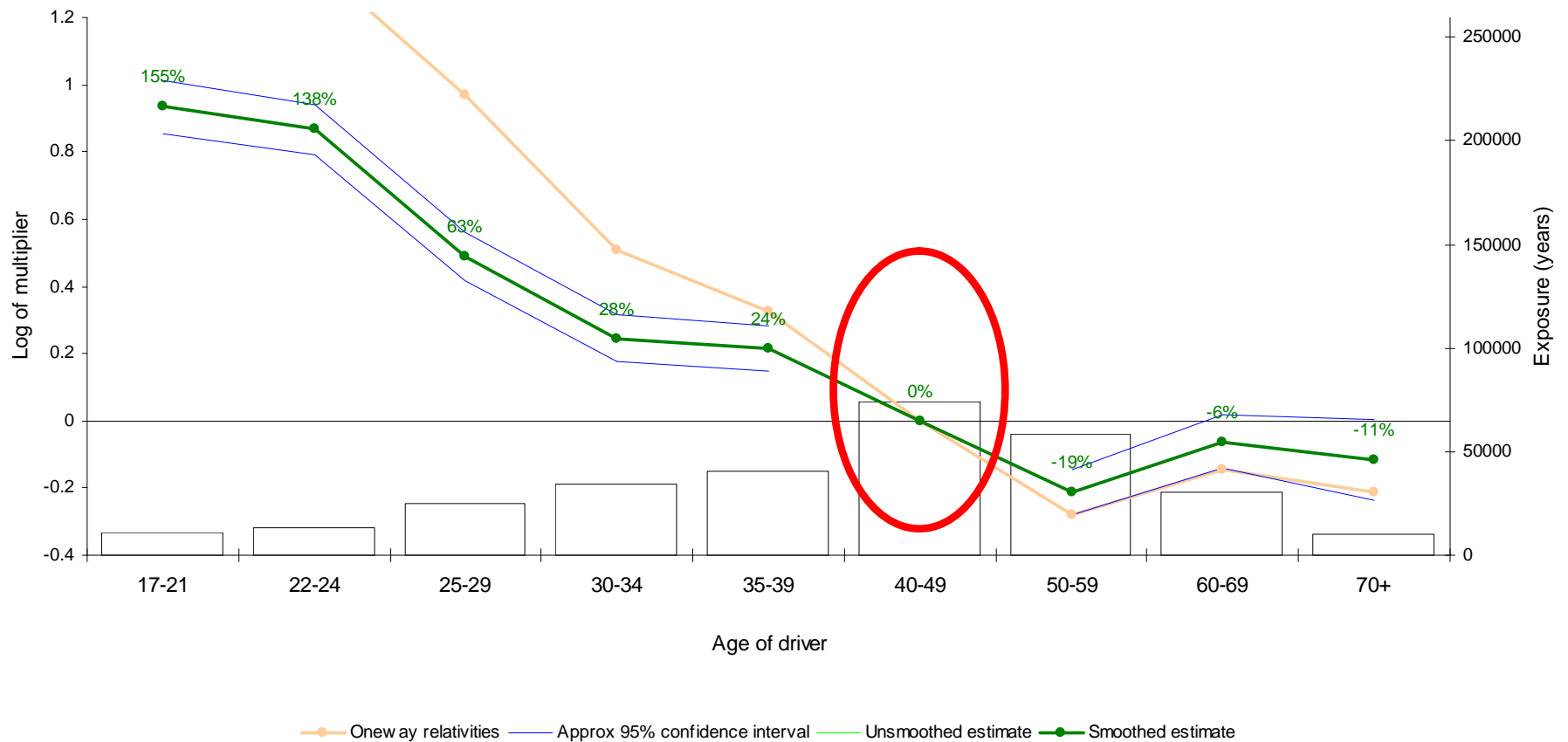
$+ \gamma_2$ if sex female

■ "Base levels"

Intrinsic aliasing

Example job

Run 16 Model 3 - Small interaction - Third party material damage, Numbers



Extrinsic aliasing

- If a perfect correlation exists, one factor can alias levels of another
- Eg if doors declared first:

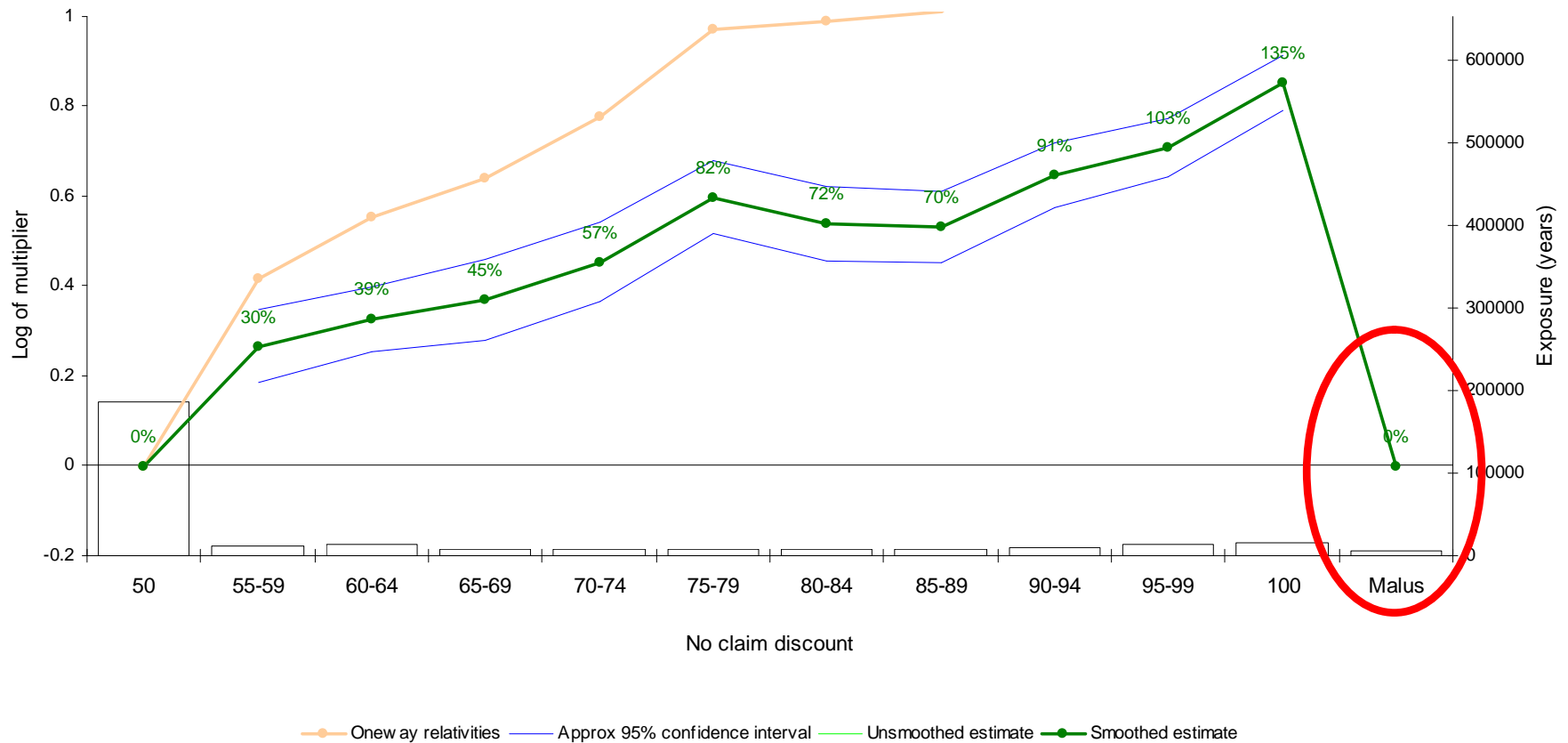
Exposure:	# Doors →	2	3	4 Selected base	5	Unknown
Color ↓						
Selected base Red		13,234	12,343	13,432	13,432	0
Green		4,543	4,543	13,243	2,345	0
Blue		6,544	5,443	15,654	4,565	0
Black		4,643	1,235	14,565	4,545	0
Further aliasing Unknown		0	0	0	0	3,242

- This is the only reason the order of declaration can matter (fitted values are unaffected)

Extrinsic aliasing

Example job

Run 16 Model 3 - Small interaction - Third party material damage, Numbers



"Near aliasing"

- If two factors are almost perfectly aliased, convergence problems can result as a result of low exposures and/or results can become hard to interpret

Exposure:	# Doors →	2	3	4 Selected base	5	Unknown
Color ↓						
Selected base Red		13,234	12,343	13,432	13,432	0
Green		4,543	4,543	13,243	2,345	0
Blue		6,544	5,443	15,654	4,565	0
Black		4,643	1,235	14,565	4,545	2
Unknown		0	0	0	0	3,242

- Eg if the 2 black, unknown doors policies had no claims, GLM would try to estimate a very large negative number for unknown doors, and a very large positive number for unknown color

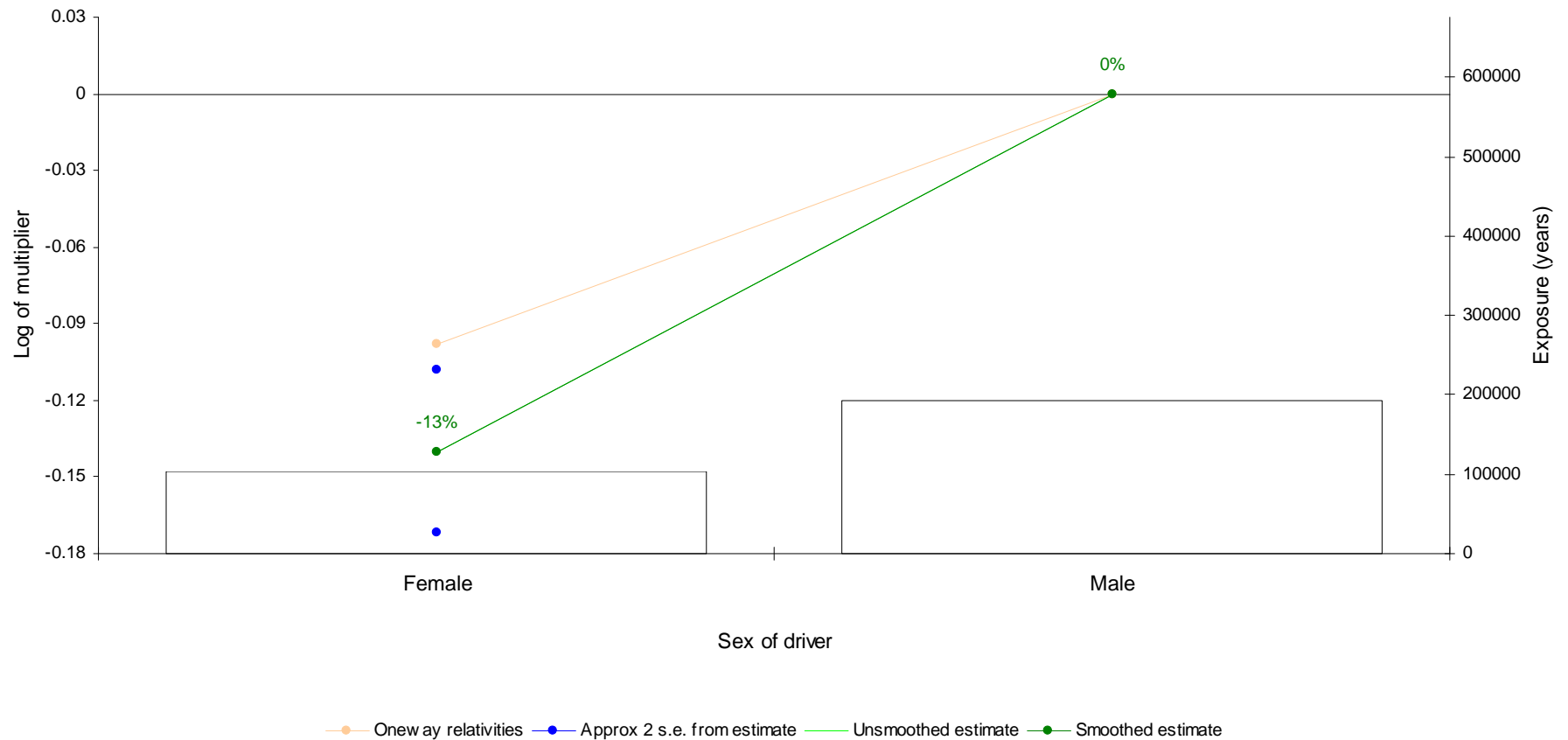
Agenda

- GLM formulae
- Model testing
 - use only variables that are predictive
 - make sure model is reasonable
- Aliasing
- Model refinements
 - Interactions
 - Splines
 - Restrictions

Interactions

Sample job

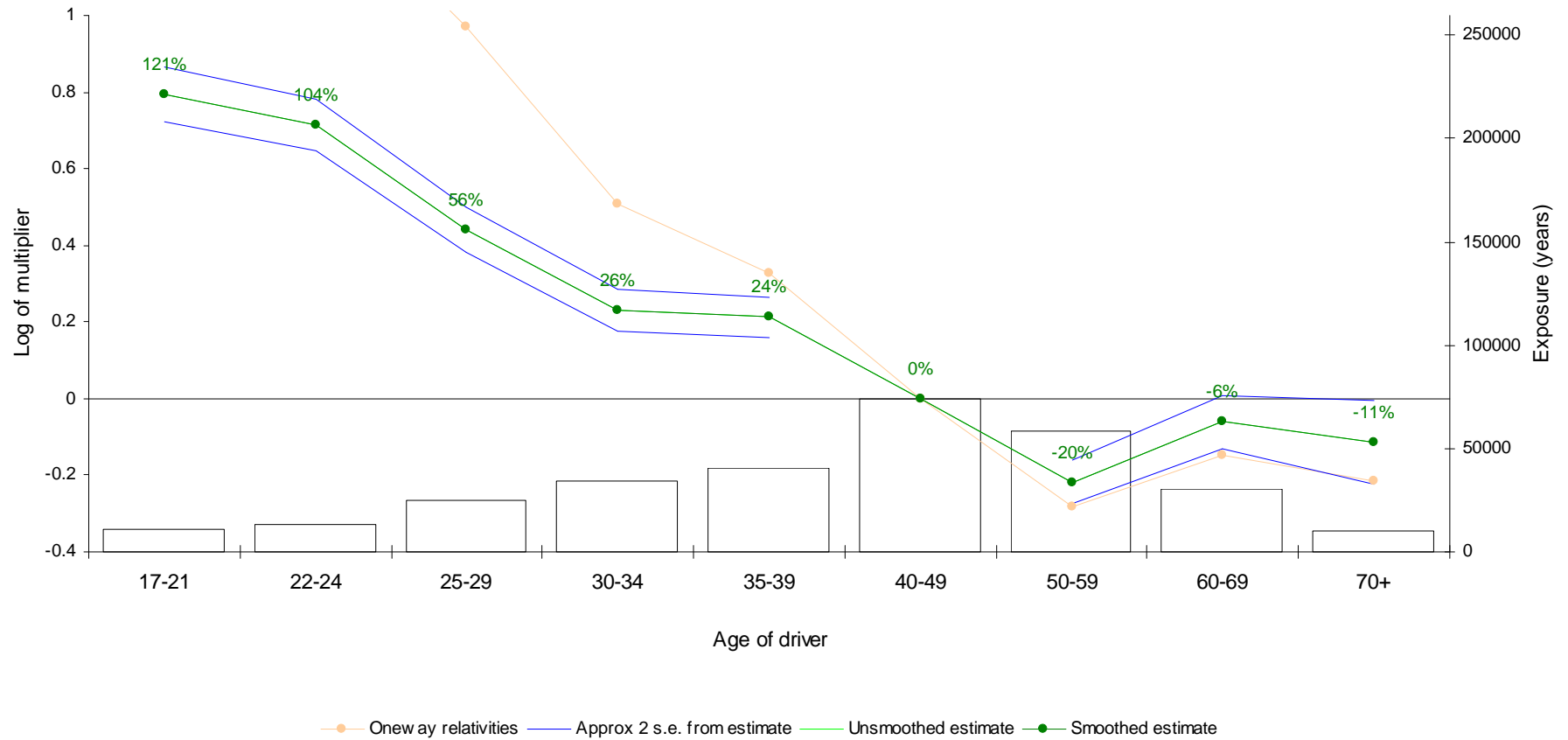
Run 23 Model 3 - Small interaction - Blah blah



Interactions

Sample job

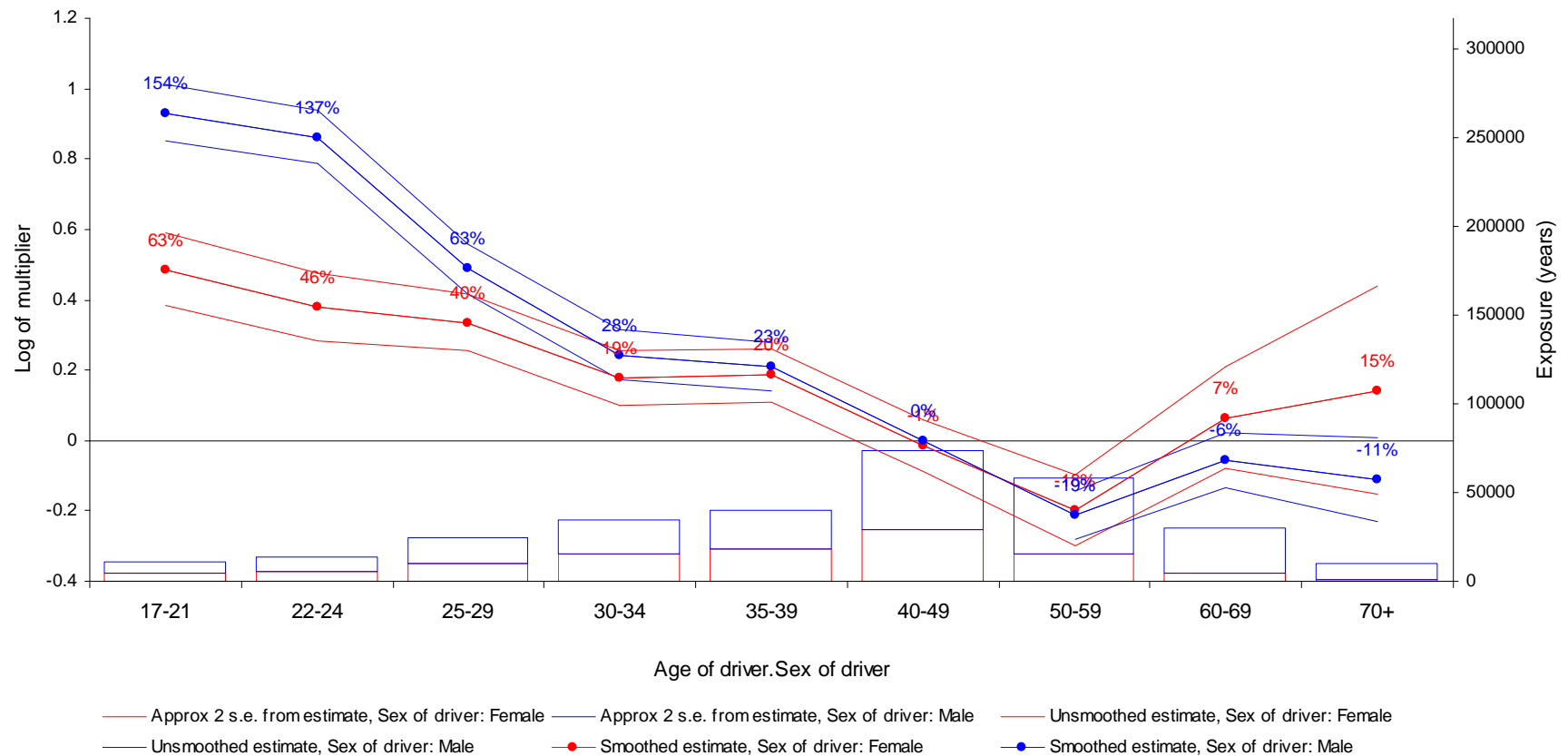
Run 23 Model 3 - No interaction



Interactions

Sample job

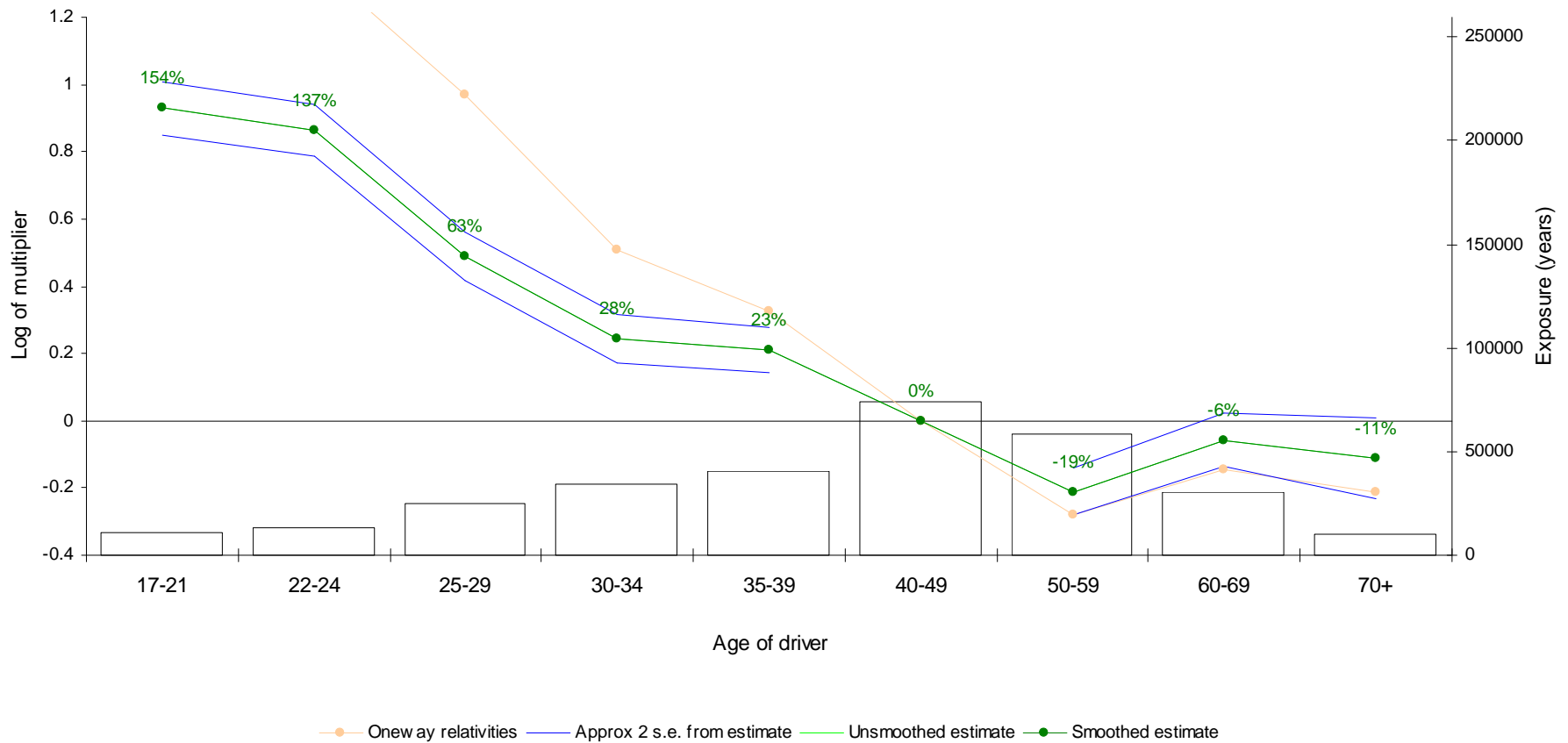
Run 19 Model 3 - Small interaction - Blah blah



Marginal interaction: Age effect

Sample job

Run 19 Model 3 - Small interaction

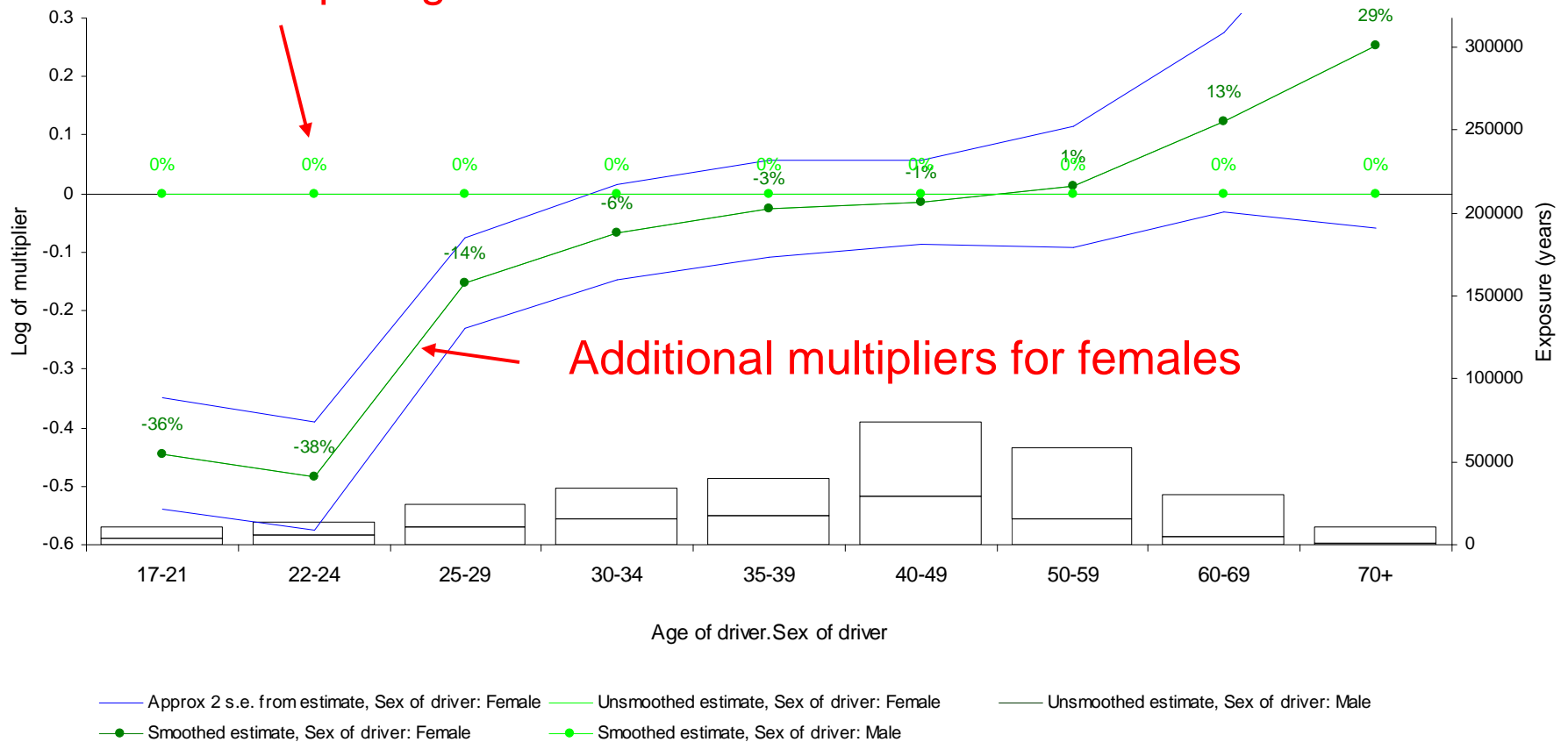


Marginal interaction: Age.Sex (ie additional female multipliers)

No additional loadings
required for males - already
made via simple age factor

Sample job

Run 19 Model 3 - Small interaction



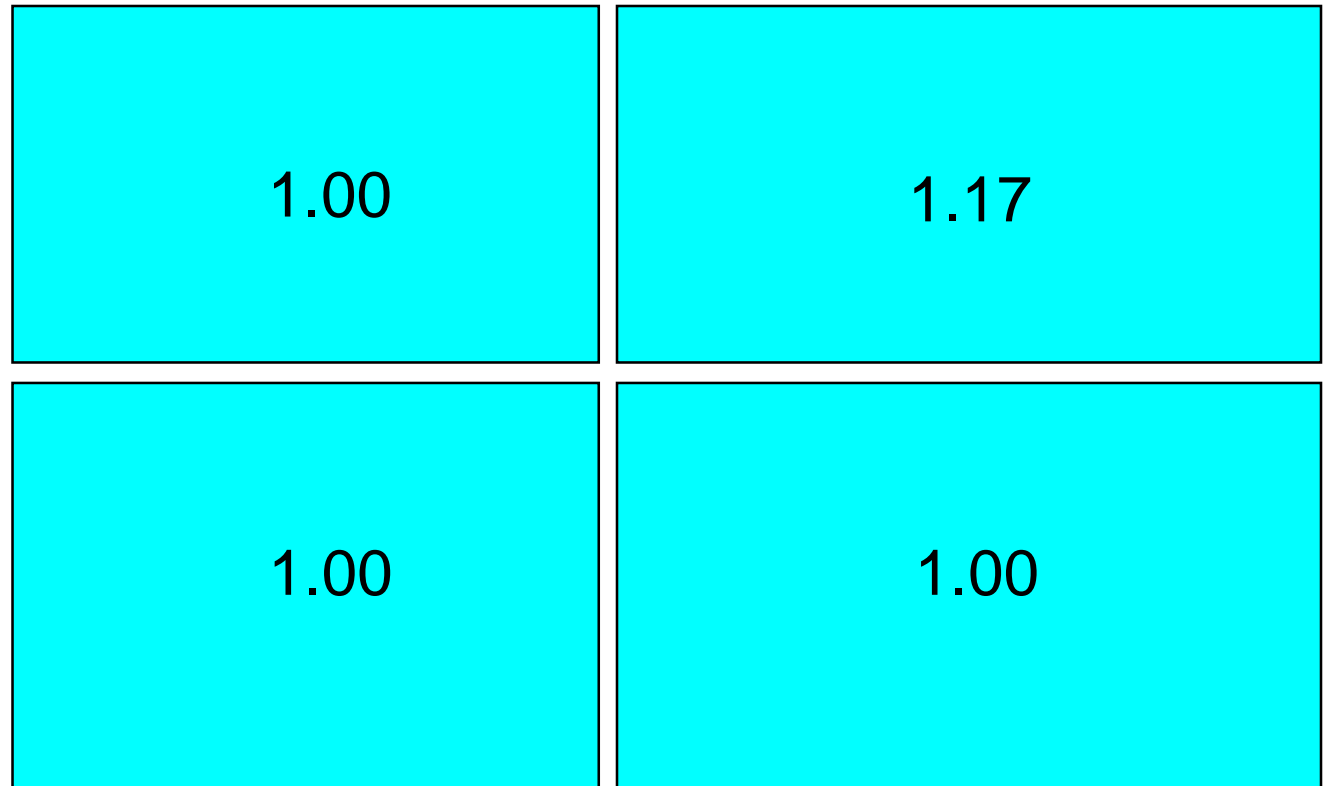
Interactions

Group >	1	2	3	4	5	6	7	8	9	10	11	12	13
Age v													
17	1.36	1.64	1.79	2.09	2.27	2.42	2.56	2.65	3.27	3.71	4.08	4.36	4.84
18	1.12	1.31	1.47	1.76	1.84	2.00	2.11	2.19	2.43	2.97	3.29	3.55	3.90
19	1.08	1.30	1.46	1.63	1.82	1.91	2.02	2.11	2.53	2.88	3.30	3.35	3.63
20	0.98	1.18	1.36	1.54	1.68	1.79	1.83	1.97	2.19	2.66	3.02	3.20	3.38
21-23	0.96	1.13	1.24	1.51	1.65	1.64	1.80	1.85	2.04	2.26	2.55	2.53	2.89
24-26	0.82	0.99	1.10	1.31	1.43	1.52	1.51	1.64	1.81	1.93	2.13	2.22	2.47
27-30	0.78	0.90	1.07	1.19	1.32	1.39	1.41	1.51	1.65	1.77	1.91	2.01	2.24
31-35	0.63	0.78	0.86	0.99	1.09	1.17	1.22	1.32	1.42	1.54	1.66	1.71	1.88
36-40	0.55	0.64	0.71	0.85	0.91	0.93	0.99	1.07	1.18	1.29	1.40	1.41	1.53
41-45	0.51	0.61	0.66	0.79	0.88	0.88	0.94	0.99	1.09	1.15	1.29	1.31	1.42
46-50	0.46	0.55	0.61	0.70	0.76	0.81	0.84	0.92	1.02	1.07	1.12	1.18	1.31
51-60	0.40	0.49	0.56	0.64	0.68	0.71	0.78	0.82	0.90	0.99	1.02	1.12	1.20
60+	0.43	0.52	0.55	0.67	0.72	0.73	0.78	0.83	0.93	0.98	1.04	1.11	1.25

Interactions

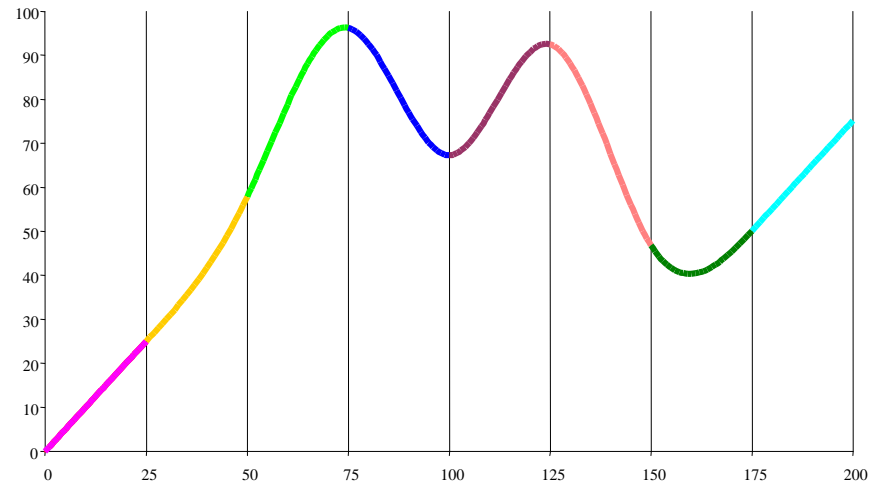
Group	1	2	3	4	5	6	7	8	9	10	11	12	13
Factor	0.54	0.65	0.73	0.85	0.92	0.96	1.00	1.08	1.19	1.26	1.36	1.43	1.56

Age	Factor
17	2.52
18	2.05
19	1.97
20	1.85
21-23	1.75
24-26	1.54
27-30	1.42
31-35	1.20
36-40	1.00
41-45	0.93
46-50	0.84
51-60	0.76
60+	0.78



Spline definition

- A series of polynomial functions, with each function defined over a short interval

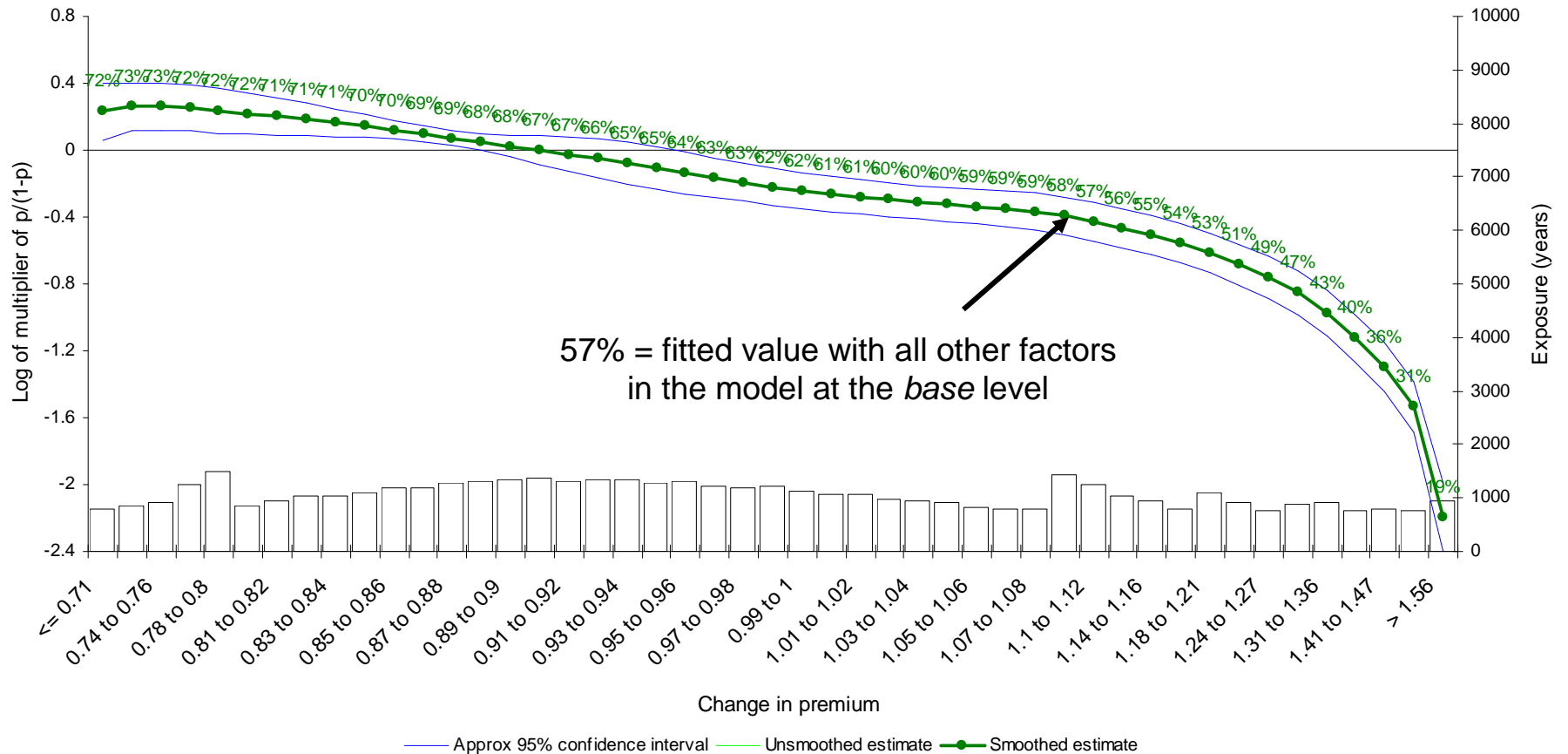


- Intervals are defined by $k+2$ knots
 - two exterior knots at extremes of data
 - variable number (k) of interior knots
- At each interior knot the two functions must join "smoothly"
- Regression splines are a form of generalized additive models

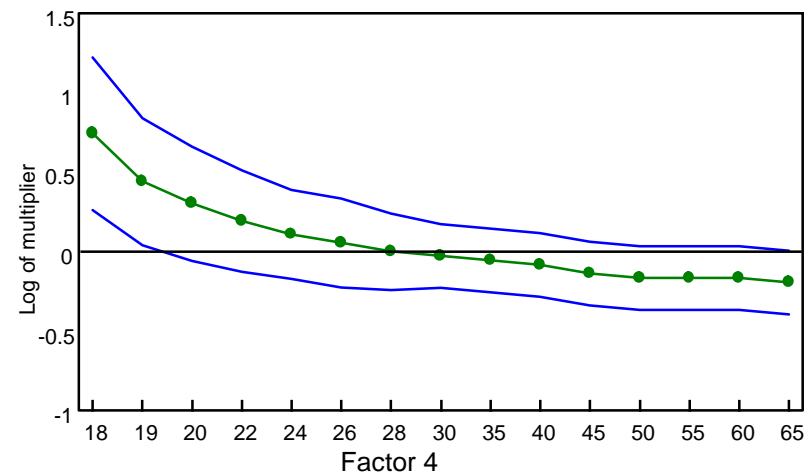
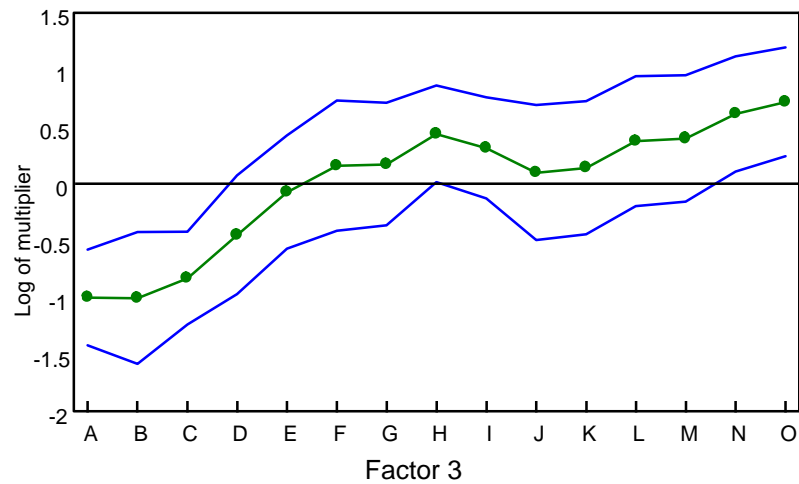
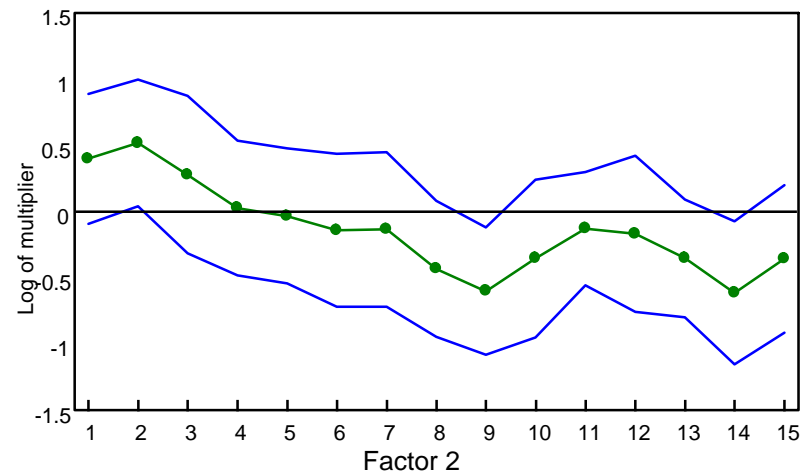
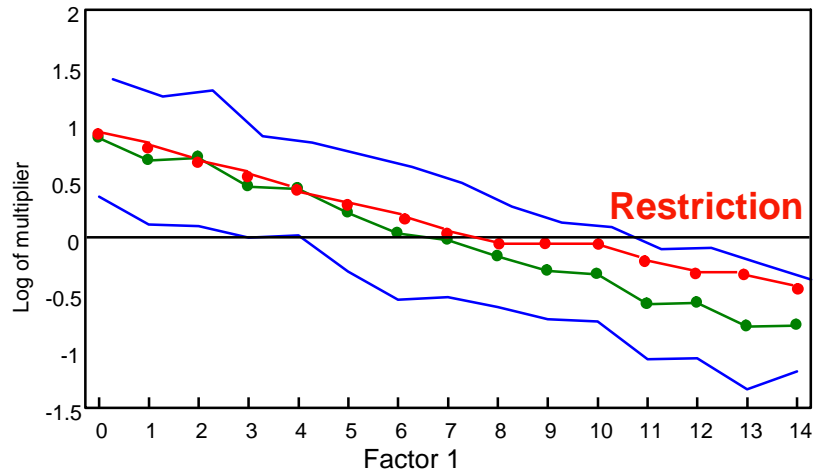
Example retention elasticity curve

Example retention analysis

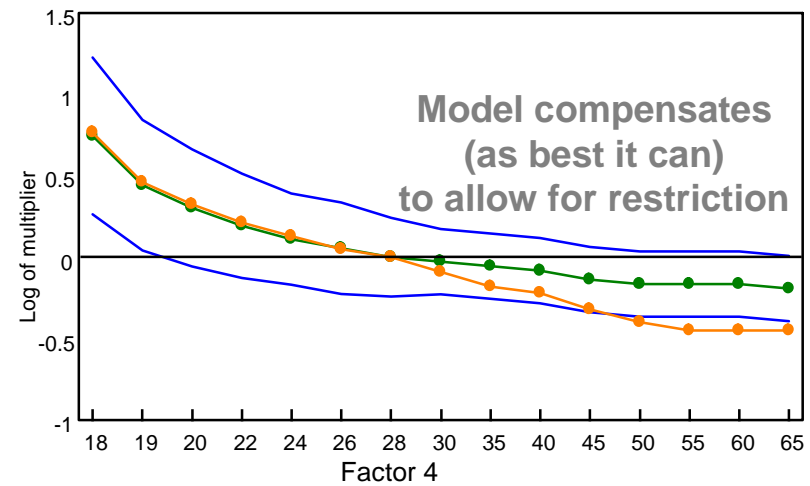
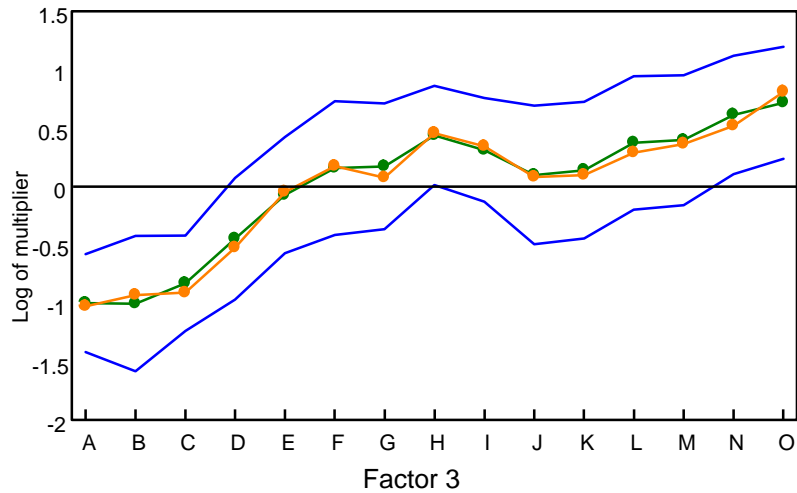
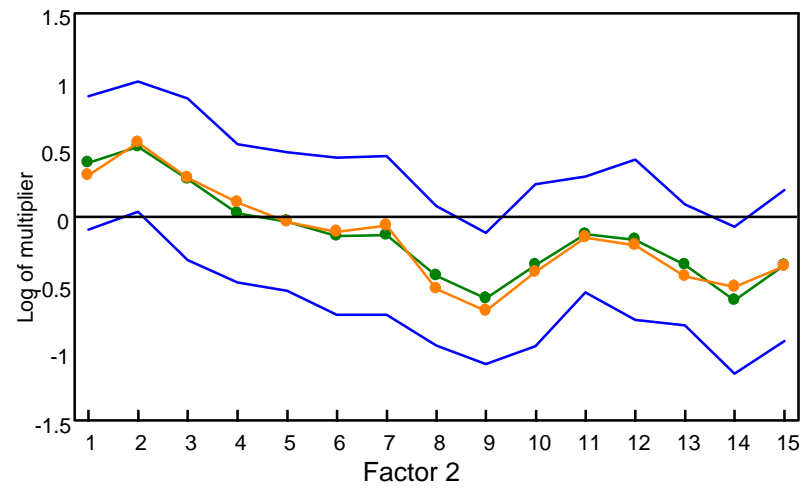
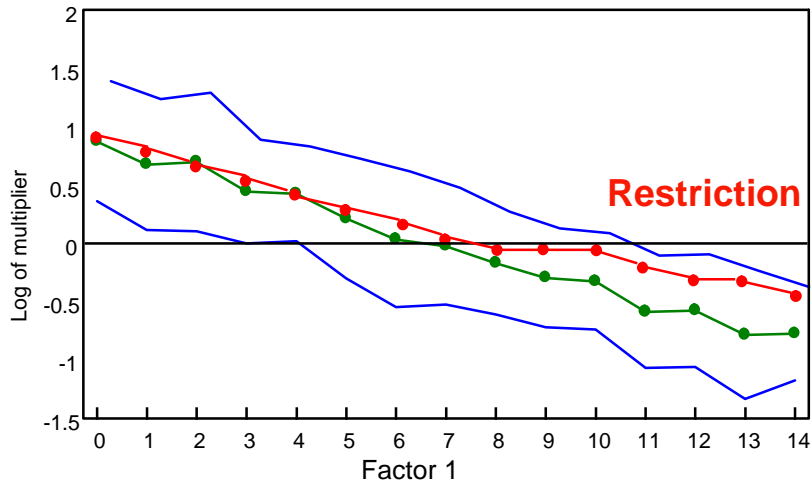
Run 2 Model 1 - Final model - Retention model



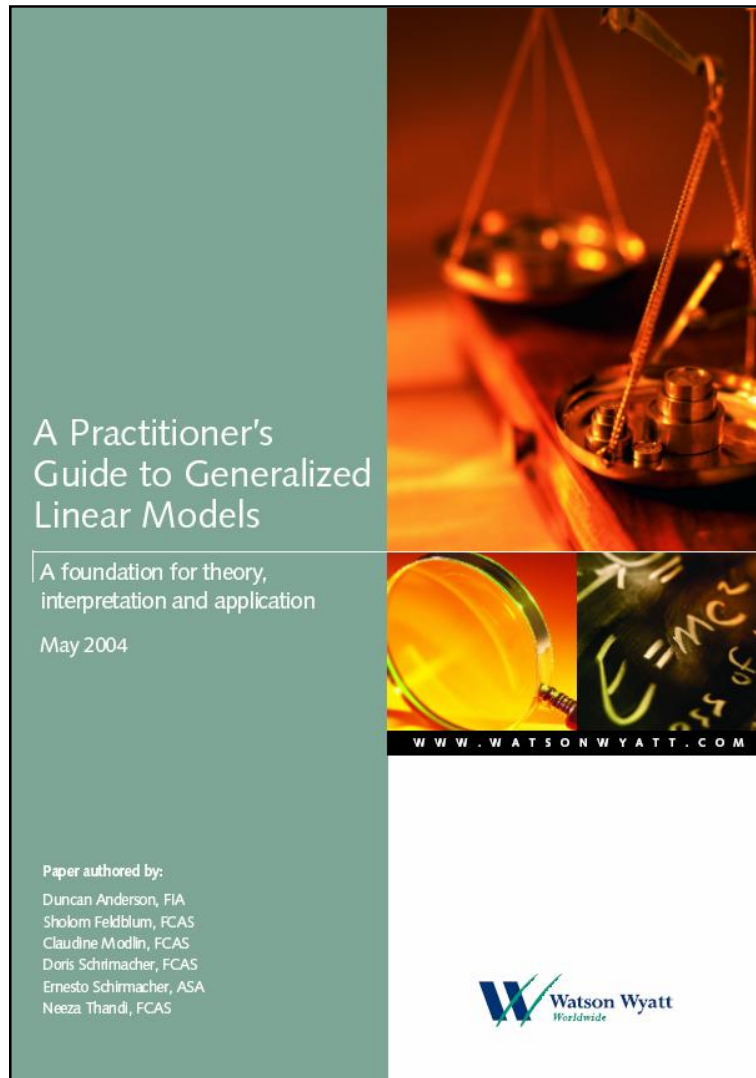
Restricted models



Restricted models



"A Practitioner's Guide to Generalized Linear Models"



- CAS 2004 Discussion Paper Program
- CAS Exam 9 syllabus as of 2006
- Copies available at www.watsonwyatt.com/glm

watsonwyatt.com



PM-2

An Introduction to GLM Theory

**CAS Seminar on Ratemaking
Boston, March 17, 2008**

Claudine Modlin, FCAS, MAAA

