



# Parameterizing Collective Risk Models

CAS Seminar on Reinsurance 2013

*R*<sup>2</sup>

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6/7/2013

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## Overview

- Collective Risk Model (CRM) for multiple lines of business with correlation.
- Well-Trodden Ground:
  - Wang
  - Meyers and Collaborators
  - Mildenhall
  - Homer-Rosengarten
  - Many Others
- Correlation: By common shock method as found in several of the references above – with a twist.
- Along the way point out some underappreciated aspects of CRM.
- Actually parameterizing simulation *method* consistent with the *model*.

# Overview

- Requirements:
  - **Efficient** as to runtime.
  - **Efficient** as to parameterization – relatively low number of parameters,
  - Simulate **small and large** losses – and reflect the appropriate dependency. Generate individual large losses and small losses in the aggregate.
  - Reflect **correlation** between lines/years.
  - **Consistent** with underlying CRM.

## CRM - Setup

- CV: For any random variable  $Y$ , the **coefficient of variation**, or CV is
  - $v(Y) = \sqrt{\text{Var}(Y)}/E(Y)$
- CV is unit-less, makes for nice formulas.
- **Collective Risk Model**,
  - $Z = X_1 + \dots + X_N$ ,  $X_i$  iid,  $X, N$  independent.
- Where  $Z$  = aggregate losses,  $X$  = severity, and the random variable  $N$  is the claim count, or “frequency”
- Independence of  $X, N$  could be violated by inhomogeneous data.
- Large/Small Losses – Threshold  $T$  such that (severity) losses  $\geq T$  are “large”, losses  $< T$  are “small”.

## CRM – Contagion Factor, Moments

- Induced CRMs
  - $Z_L = X_{1,L} + \dots + X_{N,L}, Z_S = X_{1,S} + \dots + X_{N,S}$
- Contagion Parameter – Set  $c = \mathbf{v}^2(N) - 1/E(N)$ . Then  $c$  is invariant in the sense  $c = c_L = c_S$  (follows from independence if  $X, N$ )
- Assume  $c > 0$  (positive contagion).
- Moments of CRM:
  - $E(Z) = E(N)E(X)$
  - $\mathbf{v}(Z) = \sqrt{(\mathbf{v}^2(X) + 1)/E(N) + c}$
- It follows that  $\mathbf{v}(Z) \rightarrow \sqrt{c}$  as  $E(N) \rightarrow \infty$

## CRM – Large, Small, Total Losses

- Correlation:

$$\rho(Z_S, Z_L) = c / (\mathbf{v}(Z_S)\mathbf{v}(Z_L))$$

(common shock based on identical mixing distributions)

- Total Variation:

$$E^2(Z)(\mathbf{v}^2(Z) - c) = E^2(Z_L)(\mathbf{v}^2(Z_L) - c) + E^2(Z_S)(\mathbf{v}^2(Z_S) - c).$$

## CV Interval

- Interval for  $v(Z)$  :

$$\sqrt{c + \frac{E^2(Z_L)}{E^2(Z)}(v^2(Z_L) - c)} \leq v(Z) \leq \sqrt{c + \frac{E^2(Z_L)}{E^2(Z)}(v^2(Z_L) - c) + \frac{T}{E(Z)}\left(1 - \frac{E(Z_L)}{E(Z)}\right)} \quad (*)$$

$$\sqrt{c} \leq v(Z_S) \leq \sqrt{c + \frac{T}{E(Z_S)}}$$

- Inequality is sharp.
- Proof** : Dividing the total variation equation by  $E^2(Z)$  immediately gives the left-hand inequality in (\*).

To prove the right-hand side, must show that

$$\frac{E^2(Z_S)}{E^2(Z)} (v^2(Z_S) - c) \leq \frac{T}{E(Z)} \left(1 - \frac{E(Z_L)}{E(Z)}\right), \text{ which reduces to}$$

$$(v^2(Z_S) - c) \leq \frac{T}{E(Z_S)}$$

## CV Interval

We use the following:

**Fact:** If  $Y$  is a non-negative random variable with support on  $[0, T]$ , then  $Var(Y) \leq E(Y)(T - E(Y))$ .

**Proof of Fact:**

$\frac{T^2}{4} \geq E\left(\left(\frac{T}{2} - Y\right)^2\right) = \frac{T^2}{4} - TE(Y) + E(Y^2) = \frac{T^2}{4} - TE(Y) + E^2(Y) + Var(Y)$ , which gives the result.

Using fact:

$$\begin{aligned} (v^2(Z_S) - c) &= \frac{E(X_S^2)}{E(N_S)E^2(X_S)} = \frac{1}{E(N_S)} \left(1 + \frac{Var(X_S)}{E^2(X_S)}\right) \leq \frac{1}{E(N_S)} \left(1 + \frac{E(X_S)(T - E(X_S))}{E^2(X_S)}\right) \\ &= \frac{T}{E(N_S)E(X_S)} = \frac{T}{E(Z_S)}, \text{ as required.} \end{aligned}$$



## CV Interval

For sharpness note that if we hold  $E(Z_S)$  fixed while letting  $E(N_S) \rightarrow \infty$ , then  $v^2(Z_S) \rightarrow c$ , so that  $v(Z) \rightarrow \sqrt{c + \frac{E^2(Z_L)}{E^2(Z)}(v^2(Z_L) - c)}$ , which is the left-hand side of (\*). Furthermore if we take  $X_S$  to be a 2-point distribution with masses at  $X_S = 0$  and  $X_S = T$  (with probability  $p = \frac{E(Z_S)}{E(N_S)T}$ ), then equality holds for the right-hand side of (\*).

## Mixed Poisson CRM

- We now assume that the claim count r.v  $N$  is of *mixed Poisson* type, meaning  $N \sim \text{Poisson}[E(N)G]$ , where  $G$  is a r.v with mean 1.
- To draw from  $N$ :
  - 1. Draw  $g$  from  $G$ .
  - 2. Draw from  $\text{Poisson}[E(N)g]$ .
- $\text{Var}(G) = c$ . Will use the notation  $G[c]$
- $N_L, N_S$  are also mixed Poisson with the same *mixing distribution*  $G$ .
- Example:  $G \sim \text{gamma}$ . Then  $N \sim \text{Negative Binomial}$ .
- Fact (“Severity is Irrelevant”):  $Z/E(Z) \xrightarrow{D} G$  as  $E(N) \rightarrow \infty$

## Simulation Method - CAD Algorithm with Frequency, “Severity” and Serial Common Shock

- Ref:Homer-Rosengarten (2011), Meyers-Klinker-LaLonde (2003)
- **Full Info CAD** (Have  $N, X$ )
  - Draw from  $N$  (i.e. draw from  $G$  and then from  $Poisson[E(N)G]$ )
  - Draw  $N_L$  from  $Bin(N, q)$ , where  $q = 1 - CDF_X(T)$ .  $N_S = N - N_L$ .
  - Draw  $X_{1,L}, \dots, X_{N,L}$  large losses.  $Z_L = X_{1,L} + \dots + X_{N,L}$
  - Draw  $\widetilde{Z}_S$  from Conditional Aggregate Distribution (eg, lognormal) matching  $k \geq 2$  moments of  $Z_S|N_S$ .
  - $\widetilde{Z} = \widetilde{Z}_S + Z_L$
- H-R Paper:  $\widetilde{Z}/E(Z), \widetilde{Z}_S/E(Z_S) (Z_L/E(Z_L)) \xrightarrow{D} G$ . This generalizes the “severity is irrelevant” result. Also, the method generates the correct dependence between large and small losses

## Simulation Method

- **Limited Info CAD** (Don't have  $N, X$ )
  - Draw from  $G$  only.
  - Draw  $N_L$  from  $Poisson[E(N_L)G]$
  - Draw large losses as previously.
  - Draw  $\widetilde{Z}_S$  from CAD matching first **two** moments of  $Z_S|G$
- **Minimum Parameterization:**  $G[c], E(N_L), X_L, E(Z), \mathbf{v}(Z)$
- Can then eliminate severity,  $N_S$  from equations for first two moments of  $Z_S|G$ .
- To wit,  $E(Z_S|G) = GE(Z_S), \mathbf{v}(Z_S|G) = \sqrt{(\mathbf{v}^2(Z_S) - c)/G}$
- **But**, it is not automatic that this minimal parameterization is consistent with CRM

## Simulation Method

- To address, suppose we have all the minimal parameters except  $\nu(Z)$ . We can then evaluate the lhs and rhs of inequality (\*)

$$\sqrt{c + \frac{E^2(Z_L)}{E^2(Z)}(\nu^2(Z_L) - c)} \leq \nu(Z) \leq \sqrt{c + \frac{E^2(Z_L)}{E^2(Z)}(\nu^2(Z_L) - c) + \frac{T}{E(Z)}\left(1 - \frac{E(Z_L)}{E(Z)}\right)}$$

- Any choice for  $\nu(Z)$  within this interval is a) possible and b) consistent with MP CRM.

# Beginning of Example – R<sup>2</sup> Ins Co.

Loss Parameters												
Non-Cat												
LoB	Premium	E(Z)	Loss Ratio	v(Z)	T	c	E(N <sub>T</sub> )	E(Z <sub>T</sub> )	X <sub>T</sub>	v(Z <sub>T</sub> )	E(Z <sub>Q</sub> )	v(Z <sub>Q</sub> )
GL	110,000,000	65,000,000	59.1%	0.2000	1,000,000	0.03	3.500	5,457,138	Empirical	0.7349	59,542,862	0.1940
WC	90,000,000	45,000,000	50.0%	0.2200	1,000,000	0.02	3.000	6,568,231	Empirical	0.7604	38,431,769	0.2065
CAL	40,000,000	22,000,000	55.0%	0.2750	1,000,000	0.04	0.250	512,500	Empirical	3.2929	21,487,500	0.2668
Umb	9,000,000	6,500,000	72.2%	0.5200	1,000,000	0.02	3.000	4,248,825	Empirical	0.7444	2,251,175	0.4525
PropNon-Cat	300,000,000	175,000,000	58.3%	0.1600	1,000,000	0.02	14.000	30,534,169	Empirical	0.3734	144,465,831	0.1513
<b>Total Non-Cat</b>	<b>549,000,000</b>	<b>313,500,000</b>	<b>57.10%</b>	<b>0.1139</b>			<b>23.750</b>	<b>47,320,864</b>		<b>0.2877</b>	<b>266,179,136</b>	<b>0.1096</b>
Cat												
SmallCat	549,000,000	40,000,000	7.3%	0.4300	2,000,000	0.16	10.000	4,000,000	Lognormal	0.4300	-	-
MajorCat (Net)	549,000,000	25,000,000	4.6%	1.9000		Inf 1.00	-	-	N\A	-	25,000,000	1.9000
<b>Total Inc Cat</b>	<b>549,000,000</b>	<b>378,500,000</b>	<b>68.94%</b>	<b>0.1685</b>			<b>33.750</b>	<b>51,320,864</b>		<b>0.2847</b>	<b>291,179,136</b>	<b>0.195</b>

# R<sup>2</sup> Ins Co. – Mean, CV Parameters

Section C: If Model Choice a ) = "Full Info" - ignore; b) = "Limited Info 1" - mean and cv of aggregate total losses; c) = "Limited Info 2" - mean and cv of aggregate small losses.

For case b) and c), first parameterize large loss CRM in Section D below. Selection is then guided by Mu Min, CV Min, and CV Max.

Moments of:	All Losses	All Losses	All Losses	All Losses	All Losses	Small Losses	All Losses
Mean Mu Min[Yr]	5,457,138		512,500				0
Mean Mu[Yr]	65,000,000	45,000,000	22,000,000	6,500,000	175,000,000		0 25,000,000
CV Nu Min[Yr]	0.18331	0.17862	0.21423	0.49824	0.15375	0.40006	1.00007
CV Nu Max[Yr]	0.21841	0.22552	0.30043	0.54914	0.16845	99999.00006	6.40317
CV Nu[Yr]	0.200000	0.220000	0.275000	0.520000	0.160000	0.430000	1.900000

Section D - "Full Info" - Total Loss CRM; otherwise Large Loss CRM

Parameters Of:	Large Loss CRM	Large Loss CRM	Large Loss CRM	Large Loss CRM	Large Loss CRM	Large Loss CRM	Large Loss CRM
E(Claim Count) Lambda [Yr]	3.5000	3.0000	0.2500	3.0000	14.0000	10.0000	0.0000
Severity Max[Yr]	999,999,999,999	999,999,999,999	999,999,999,999	#####	999,999,999,999	999,999,999,999	#####
Severity Distribution	Empirical	Empirical	Empirical	Empirical	Empirical	Lognormal	Empirical
See Below	0.0	0.0	0.0	0.0	0.0		0.0
shift						2,000,000.0	
mu						14.4	
sigma						0.5	
min						2,000,000.0	
max						#####	

Severity Values and Weights used for Empirical And Mixed Exponential severity distributions

Severity Values[Yr]							
	1,000,000	1,000,000	1,000,000	1,000,000	1,000,000	2,000,000	2,000,000
	1,250,000	1,250,000	2,000,000	1,250,000	1,250,000		
	1,500,000	1,500,000	10,000,000	1,500,000	1,500,000		

## Common Shock Correlation

- Correlate LoBs modeled with MP CRM/CAD method.
- LoBs are organized into covariance groups. Only Lobs within the same covariance group co-vary with one another.
- Frequency, “severity”, and serial common shock.



## Frequency Common Shock

- General Idea: Common draw from mixing distribution.
- Need to allow that LoBs might have different mixing distributions.
- Solution is draw common uniforms and use these to invert the mixing distributions ( $g = F_G^{-1}(u)$ ).
- Remaining problem is that this will tend to generate very high correlation.
- Usual solution is to assume that  $G$  is an independent product, ie
  - $G[c] = G_1[c_1]G_2[c_2]$
  - Then apply common shock only to  $G_1$ .
  - Note that  $c = c_1 + c_2 + c_1c_2$

## Frequency Common Shock

- Variant is the “twisted product”  $G[c] = G_1[c_1] \times G_2[c_2]$  defined by  $G = G_1 G_2 [c_2/G_1]$ .
- That is, to draw from  $G$ :
  - Draw  $g_1$  from  $G_1$ .
  - Draw  $g_2$  from  $G_2[c_2/g_1]$ .
  - $g = g_1 g_2$ .
- Nice thing about twisted product is  $c = c_1 + c_2$ .
- **Parameter:** FrCoVarWt =  $w$ ,  $0 \leq w \leq 1$ . Varies by LoB.
- In twisted product set  $c_1 = wc$ ,  $c_2 = (1 - w)c$  (where  $G_i[0] \equiv 1$ ).

## Serial Common Shock

- Bring in uniforms necessary to invert  $G_1$ 's for frequency c.s. These vary by covariance group and year.
- Also bring in uniforms for  $G_2$ 's – varying by LoB and year.
- Reason for  $G_2$ 's is generate sufficient correlation between years but within LoB.
- Flip a weighted coin.
- For year  $j, j \geq 2$ , if coin flip comes up “heads” use the uniforms from year  $j - 1$ . Otherwise use year  $j$ .
- **Parameter** – FrSerialCoVarWt – the weight for the coin flip. Can vary by covariance group or LoB. Usually by covariance group.

# Serial Common Shock

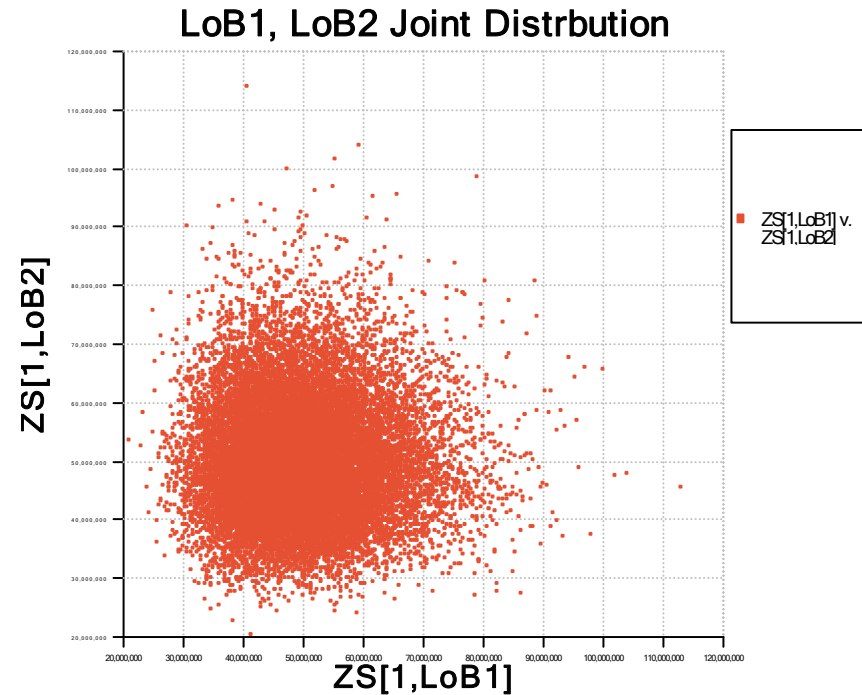
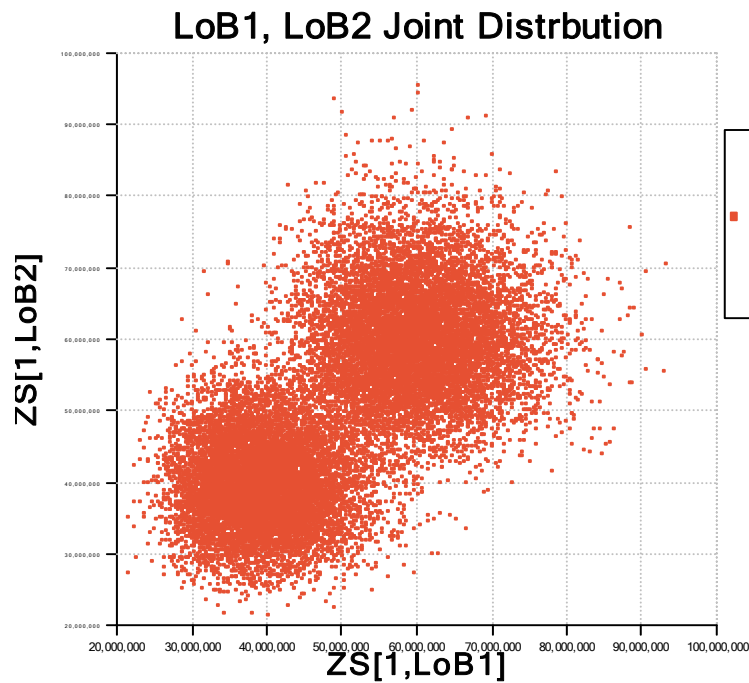
- Summary
  - $G_1$  correlates non-identical LoBs, both within-year and serially.
  - $G_2$  - serial correlation for identical LoBs.
  - Serial correlation decays by FrSerialCoVarWt.

## “Severity” Common Shock

- Really it's c.s. applied to the conditional aggregate distribution generating  $\widetilde{Z}_S$ .
- By H-R, the particular distribution family used doesn't matter.
- Assume lognormal, with  $Mu, Sigma$  the conditional parameters.
- **Parameters:** ZSCoVarWt, ZSSerialCoVarWt.
- Express CAD as a product of Lognormals
- $CAD = \text{logn}[\cdot 5Mu, Sigma\sqrt{ZSCovarWt}] \text{logn}[\cdot 5Mu, Sigma\sqrt{1 - ZSCovarWt}]$
- Play same game as previously.

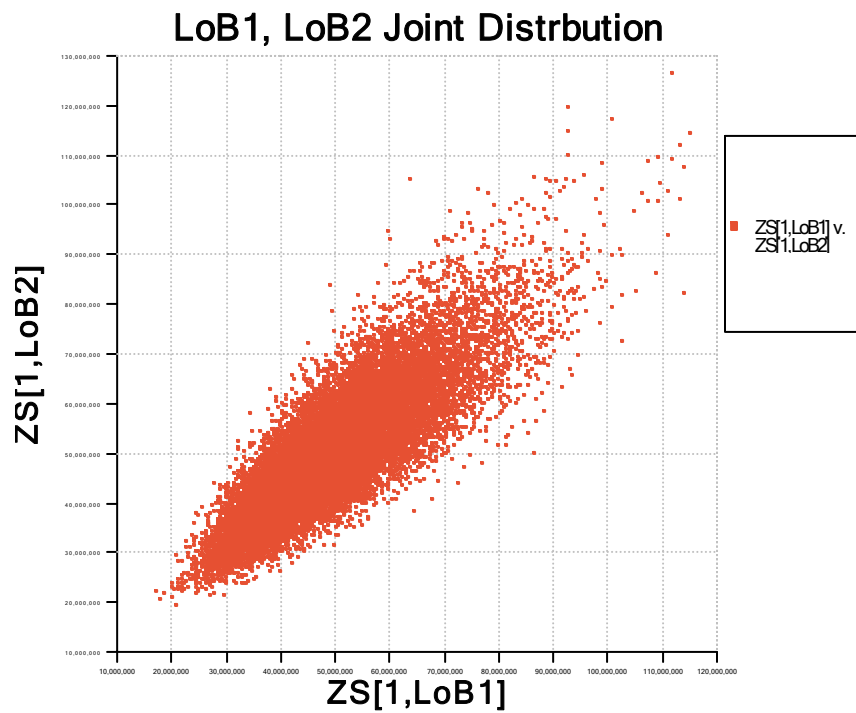
## Why do we need ZSCoVarWt?

- Example: Identical LoBs LoB1, LoB2
- $FrCoVarWt = .85$ ,  $ZSCoVarWt = 0$ ,  $G_1 = 1 \pm \sqrt{c}$ , with probability .5.
- $c = O(\mathbf{v}^2)$  - High Correlation  $c = 0 (\mathbf{v}^2 \gg c)$  - No Correlation



# Why ZSCoVarWt?

- $FrCoVarWt = 0$ ,  $ZSCoVarWt = .85$ ,  $c = 0$



## Why ZSCoVarWt?

- For Identical LoBs:

	FrCoVarWt=1 ZSCoVarWt=0	FrCoVarWt=0 ZSCoVarWt=1
$v^2 \rightarrow c$	$\rho \rightarrow 1$	$\rho \rightarrow 0$
$v^2 \gg c$	$\rho \rightarrow 0$	$\rho \rightarrow 1$

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## More Tricks

- Can increase skewness by adding shift to mixing distributions.
  - Shifted lognormal:  $G = s + \text{Logn} \left[ \ln \left( \frac{(1-s)^2}{\sqrt{c+(1-s)^2}} \right), \sqrt{\ln \left( 1 + \frac{c}{(1-s)^2} \right)} \right]$
  - Skewness =  $\frac{\sqrt{c}}{(1-s)} \left( 3 + \frac{c}{(1-s)^2} \right)$
- Can use discrete mixing distribution to create a mass at 0, for example.

# R<sup>2</sup> Ins Co. – Correlation Parameters, Mixing Distributions

CAD Large Loss/Small Loss Simulation - Elements that vary by year are indicated with [Yr]

Section A - Setup

Classes	GL	WC	CAL	Umb	PropNon-Cat	SmallCat	MajorCat
Large Loss Threshold[Yr]	1,000,000	1,000,000	1,000,000	1,000,000	1,000,000	2,000,000	999,999,999
ModelChoice	Limited Info1	Limited Info1	Limited Info1	Limited Info1	Limited Info1	Limited Info2	Limited Info1
CoVarGroup	1	1	1	1	2	3	4
FrCoVarWT	0.5000	0.3500	0.2500	1.0000	0.2500	0.5000	0.6108
ZSCoVarWT	1	0.075	0.075	1	0	0	0
FrSerialCoVarWT	0.3				0	0	0
ZSSerialCovarWeight	0.3				0	0	0

Section B - Parameterize Mixing Distribution. Mixing Dist of the form  $G = G1$  compound with  $G2$ .  $G$  has mean 1 and variance  $c$  ( $G1$  mean 1 and variance  $c1$ ).

Note that  $c1 = FrCoVarWt * c$  So, choice of  $FrCoVarWt = 1$  (and  $p = 1$  for Wtd Sum) means  $c1 = c$  and  $G = G1$ ; Choice of  $FrCoVarWt = 0$  (and  $p = 0$  for Wtd. Sum) means  $c2 = c$  and  $G = G2$ .

Note that the  $G1$ 's by class co-vary within the defined covariance groups.

Compounding Type	TwistedProduct	TwistedProduct	TwistedProduct	TwistedProduct	TwistedProduct	TwistedProduct	StraightProduct
	0.5	0.5	0.5	0.5	0.5	0.5	0.5
c	0.03000	0.02000	0.04000	0.02000	0.02000	0.16000	1.00000
G1Choice	Lognormal	Lognormal	Lognormal	Lognormal	Lognormal	Lognormal	DiscreteUniform
shift	0	0	0	0	0	0	
p							0.3892
m							1
G2Choice	Lognormal	Lognormal	Lognormal	Lognormal	Lognormal	Lognormal	Lognormal
shift	0	0	0	0	0	0	0.75

## R<sup>2</sup> Ins Co. – Correlation Parameters, Mixing Distributions

- Casualty lines co-vary. (Covariance group 1)
- Non-Cat Property, Cats are independent (CoVar groups 2-4).
- Mixing Distributions all of form  $G = \text{logn} \times \text{logn}$  except Major Cat
- Major Cat (Net of Cat XoL)
  - From Cat modelling know  $Prob(0), E(\text{Major Cat}), \nu, \gamma = \text{skewness}$ .
  - Modeling Solution:  $G = \text{Discrete Uniform} * (\text{shift logn}), c = 1$
  - Parameters of Discrete Uniform (including  $c_1 = FrCoVarWt$ ) set up to match probability mass at 0.
  - Shifted lognormal set up to match skewness.
- Umbrella: parameters set up to give higher correlation with GL than other casualty lines.

# Correlation Matrix

CorrZ		1							
		GL	WC	CAL	Umb	PropNon-Cat	SmallCat	MajorCat	
1	GL	1.0000	0.2981	0.2923	0.2755	-0.0110	-0.0017	0.0004	
	WC	0.2981	1.0000	0.1654	0.1453	-0.0033	-0.0075	-0.0079	
	CAL	0.2923	0.1654	1.0000	0.1345	-0.0057	0.0034	-0.0007	
	Umb	0.2755	0.1453	0.1345	1.0000	-0.0003	0.0056	0.0030	
	PropNon-Cat	-0.0110	-0.0033	-0.0057	-0.0003	1.0000	0.0015	-0.0028	
	SmallCat	-0.0017	-0.0075	0.0034	0.0056	0.0015	1.0000	-0.0072	
	MajorCat	0.0004	-0.0079	-0.0007	0.0030	-0.0028	-0.0072	1.0000	
2	GL	0.2672	0.0780	0.1010	0.0834	-0.0053	0.0006	0.0072	
	WC	0.0846	0.2368	0.0494	0.0395	-0.0069	-0.0132	0.0011	
	CAL	0.0814	0.0482	0.2754	0.0488	-0.0032	-0.0095	0.0130	
	Umb	0.0882	0.0287	0.0435	0.0459	0.0051	0.0147	0.0008	
	PropNon-Cat	-0.0005	0.0212	0.0058	0.0234	-0.0030	0.0012	-0.0033	
	SmallCat	-0.0004	-0.0045	0.0075	0.0067	-0.0131	-0.0134	0.0002	
	MajorCat	0.0084	0.0070	-0.0042	0.0087	0.0046	0.0143	-0.0020	
3	GL	0.0768	0.0318	0.0309	0.0234	0.0007	-0.0099	-0.0002	
	WC	0.0293	0.0655	0.0235	0.0082	-0.0045	-0.0057	0.0029	
	CAL	0.0343	0.0322	0.0967	0.0113	-0.0020	-0.0107	0.0092	
	Umb	0.0324	0.0167	0.0165	0.0167	-0.0039	-0.0065	-0.0004	
	PropNon-Cat	-0.0053	0.0029	0.0085	-0.0010	0.0080	-0.0027	-0.0002	
	SmallCat	0.0015	-0.0011	0.0074	0.0028	0.0067	0.0025	0.0021	
	MajorCat	0.0042	0.0108	0.0054	0.0026	0.0054	0.0008	-0.0037	

## Reinsurance Cover for R<sup>2</sup> Insurance Co.

- Aggregate Stop-Loss – Term: 2 years
- Subject Losses:
  - 100% of Non-Cat losses limited to \$1m per risk
  - 50% of \$1m *xs* \$1m per risk
  - 100% of Cat losses limited to 20% of Subject Premium (\$549m) per year.
- Subject Loss Ratio: 64.5%
- Coverage: 15% *xs* 75% of SP
- Premium: 5% of SP (33% RoL), 30% of which is margin, the remainder to an experience account.
- Profit Commission: 100% of residual EA

## Reinsurance Cover Results

- NPV basis (Have also developed payout patterns by LoB).
- Low parameter model allows for efficient sensitivity testing.
- Key Stats (Reinsurer PoV):

Key Stats w\Sensitivity Testing				
	Base	c's, CVs Up	CoVar Wts Up	Skewness Up
NPV(Profit/Loss)	12,851,332	9,773,105	12,281,670	12,776,752
Prob(Econ. Loss)	11.59%	17.51%	12.32%	11.73%
TVaR(95)	(35,406,559)	(50,849,025)	(41,391,711)	(35,944,163)
TVaR(97.5)	(47,754,569)	(66,138,266)	(56,187,593)	(48,446,883)
<b>RoRaC(95)</b>	9.46%	6.00%	8.08%	9.34%
<b>RoRaC(97.5)</b>	7.53%	5.08%	6.48%	7.45%
ERD	-4.40%	-8.00%	-5.23%	-4.52%