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GLM II: Basic Modeling Strategy

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Casualty Actuarial Society
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Overview

Project Cycle

Modeling Cycle

Quick Review of GLMs

- Common GLM Model Forms

Personal Auto Claims Example

- Available Data

- Potential Models

Logistic Regression

- Summary Statistics

- Adjust for Exposure

- Adding Predictors

- Piecewise Linear Fits

- Residual Diagnostics

- Validation

Personal Injury Example

- Summary Statistics

- Exploratory Plots

- Model Fits

- Diagnostics

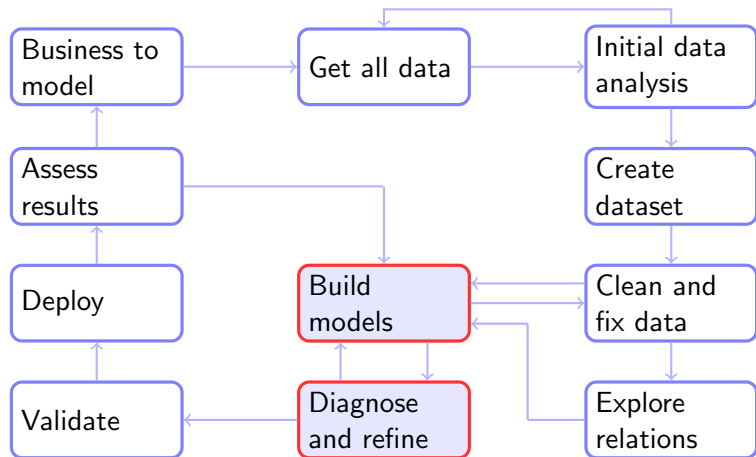
- Parameter Grouping

- Link Function

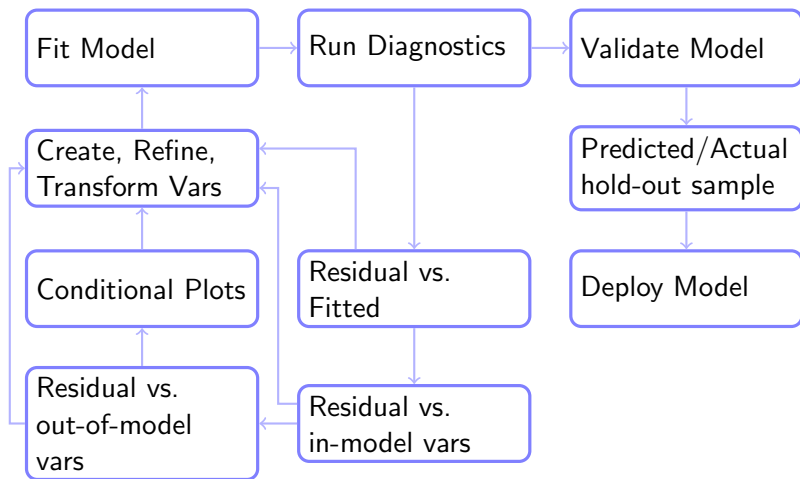
- Interactions

- Validation

Overall Project Cycle



Model Building Cycle



Basic GLM Specification

$$g(\mathbb{E}[y]) = \beta_0 + x_1\beta_1 + \cdots + x_k\beta_k + \text{offset}$$

1. The link function is g
2. The distribution of y is a member of the exponential family
3. The explanatory variables x_i may be continuous or indicator
4. The offset term can be used to adjust for exposure or to introduce known restrictions

Common Model Forms

	Target Variable			
	Claim Frequencies	Claim Counts	Average Claim Amount	Probability
Link $g(\mu)$	$\log(\mu)$	$\log(\mu)$	$\log(\mu)$	logit(μ)
Error	Poisson	Poisson	Gamma	Binomial
Variance $V(\mu)$	μ	μ	μ^2	$\mu(1 - \mu)$
Weights	Exposure	1	# claims	1
Offset	0	$\log(\text{Exposure})$	0	0

Personal Auto Claims

The dataset contains 59,876 policies taken out in 2004 or 2005. This dataset is a subset of the `car.csv` dataset featured in the book by de Jong & Heller [3]. I removed all records where vehicle body was not for personal use.

The available variables are:

1. Driver age
2. Gender
3. Garage location
4. Vehicle body
5. Vehicle age
6. Vehicle value
7. Exposure
8. Claim?
9. Number of claims
10. Total claim cost

Possible Models?

1. Binary model: will a claim be made?
2. Count model: how many claims?
3. Severity model: how costly will a claim be?
4. Conditional severity model: how costly will a claim be given that at least one claim is filed?

Build, Test, Validate

1. *Build*: used to create many models
2. *Test*: used to check intermediate models
3. *Validate*: used only once to check your final model

One rule of thumb: (50%, 25%, 25%).

Let us take three independent samples of 3 000 records each.

Logistic Model

Target variable: Did a policy file a claim?
claim = 0; no claim reported
claim = 1; one or more claims reported

We want to model the expected value of filing a claim as a linear combination of predictors.

$$\mathbb{E}[\text{claim}] = \text{linear combination of predictors}$$

Logistic Model

Target variable: Did a policy file a claim?
claim = 0; no claim reported
claim = 1; one or more claims reported

We want to model the expected value of filing a claim as a linear combination of predictors.

$$\mathbb{E}[\text{claim}] = \text{linear combination of predictors}$$

$$\log \left(\frac{\mathbb{E}[\text{claim}]}{1 - \mathbb{E}[\text{claim}]} \right) = \text{linear combination of predictors}$$

$$\text{logit}(\mu) = \log \left(\frac{\mu}{1 - \mu} \right) \quad \text{logit}^{-1}(\eta) = \frac{e^{\eta}}{1 + e^{\eta}} = \frac{1}{1 + e^{-\eta}}$$

Summary Statistics for Build dataset

Continuous Variables

	claim	exposure	veh.value
Min.	:0.000	0.003	0.220
1st Qu.	:0.000	0.224	0.980
Median	:0.000	0.441	1.470
Mean	:0.075	0.467	1.753
3rd Qu.	:0.000	0.706	2.070
Max.	:1.000	0.999	12.090

Vehicle value is in units of \$10,000.

Summary Statistics for Build dataset

Categorical Variables (record counts)

veh.body	veh.age	area	age.cat	gender
SEDAN: 1137	3: 892	C: 902	4: 727	F: 1808
HBACK: 940	4: 806	A: 757	3: 679	M: 1192
STNWG: 805	2: 744	B: 621	2: 566	
HDTOP: 87	1: 558	D: 350	5: 509	claim
COUPE: 28		E: 217	6: 273	0: 2775
CONVT: 2		F: 153	1: 246	1: 225
RDSTR: 1				

We know that

$$\mathbb{E}[\text{claim}] = \frac{225}{3000} = 7.5\%$$

Summary Statistics for Build dataset

Categorical Variables (record counts)

veh.body	veh.age	area	age.cat	gender
SEDAN:1137	3:892	C:902	4:727	F:1808
HBACK: 940	4:806	A:757	3:679	M:1192
STNWG: 805	2:744	B:621	2:566	
HDTOP: 87	1:558	D:350	5:509	claim
COUPE: 28		E:217	6:273	0:2775
CONVT: 2		F:153	1:246	1: 225
RDSTR: 1				

We know that

$$\mathbb{E}[\text{claim}] = \frac{225}{3000} = 7.5\%$$

Is this a good estimate?

What is the correct value for $\mathbb{E}[\text{claim}]$?

Let π be the probability of filing a claim over a one year period and let t_i be the exposure amount over the year.

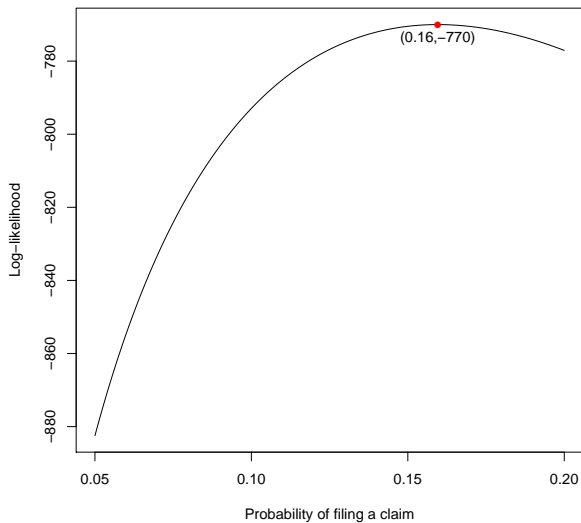
Let's estimate π via maximum likelihood.

$$\max_{\pi \in (0,1)} \left\{ \prod_{i=1}^{3000} (t_i \pi)^{\text{claim}_i} (1 - t_i \pi)^{1 - \text{claim}_i} \right\}$$

The log-likelihood is

$$\max_{\pi \in (0,1)} \left\{ \sum_{i=1}^{3000} \text{claim}_i \log(t_i \pi) + (1 - \text{claim}_i) \log(1 - t_i \pi) \right\}$$

Log-likelihood as a function of π



Logistic Regression—Null Model

Call:

```
glm(formula = claim ~ 1,  
     family = binomial(link = "logit"), ...)
```

Deviance Residuals: [...omited...]

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-2.51231	0.06932	-36.24	<2e-16 ***

Null deviance:	1598.3	on 2999	degrees of freedom
Residual deviance:	1598.3	on 2999	degrees of freedom

How to adjust for Exposure?

How do we do it when modeling claim counts (Poisson, log-link)?

$$\log \left(\frac{\mathbb{E}[\text{counts}]}{\text{exposure}} \right) = \text{linear predictor}$$

$$\log (\mathbb{E}[\text{counts}]) = \text{linear predictor} + \underbrace{\log (\text{exposure})}_{\text{offset term}}$$

How should we do it in logistic regression?

How to adjust for Exposure?

How do we do it when modeling claim counts (Poisson, log-link)?

$$\log \left(\frac{\mathbb{E}[\text{counts}]}{\text{exposure}} \right) = \text{linear predictor}$$

$$\log (\mathbb{E}[\text{counts}]) = \text{linear predictor} + \underbrace{\log (\text{exposure})}_{\text{offset term}}$$

How should we do it in logistic regression?

$$\text{logit} \left(\frac{\mathbb{E}[\text{claim}]}{\text{exposure}} \right) = \text{linear predictor}$$

Logistic Regression—Null Model *Exposure Adjusted*

Call:

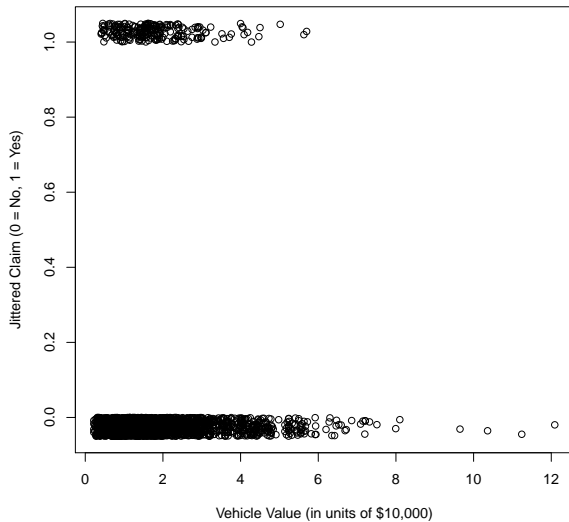
```
glm(formula = claim ~ 1,  
     family = binomial(link = logitexp(expo)), ...)
```

Coefficients:

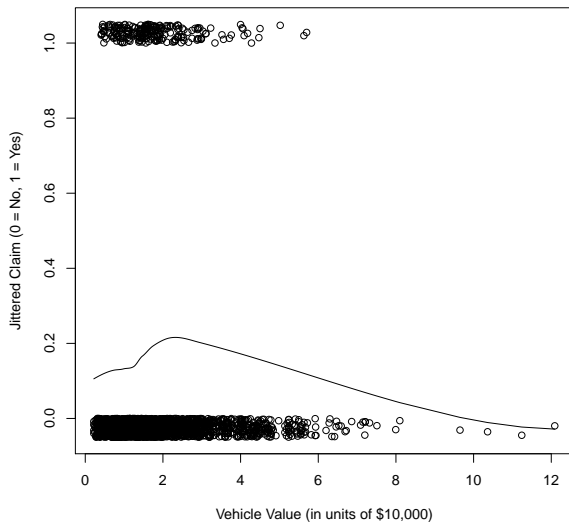
	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-1.66145	0.07532	-22.06	<2e-16 ***

Null deviance:	1598.3	on 2999	degrees of freedom
Residual deviance:	1540.0	on 2999	degrees of freedom

Occurrence of a claim vs. vehicle value



Occurrence of a claim vs. vehicle value with smoother



Logistic model with quadratic vehicle value

Call:

```
glm(formula = claim ~ veh.value + I(veh.value^2),  
     family = binomial(link = logitexp(expo)), ...)
```

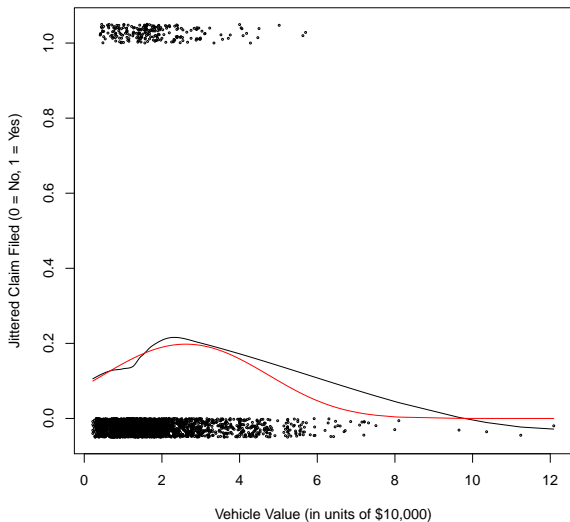
Deviance Residuals: [...omited...]

Coefficients:

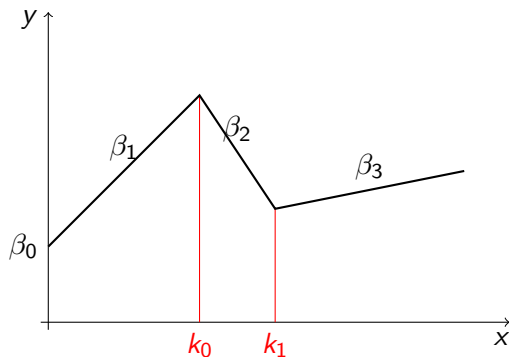
	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-2.35514	0.26906	-8.753	< 2e-16	***
veh.value	0.73131	0.25412	2.878	0.00400	**
I(veh.value^2)	-0.13991	0.05037	-2.778	0.00547	**

Null deviance: 1598.3 on 2999 degrees of freedom
Residual deviance: 1529.0 on 2997 degrees of freedom

Claim vs vehicle value with quadratic fit



How to create piecewise linear fits?



x	x^*	x^{**}
1.2	0	0
1.8	0	0
2.3	0.3	0
2.7	0.7	0
2.9	0.9	0
3.5	1.5	0.5
3.7	1.7	0.7
4.6	2.6	1.6

$$\mathbb{E}[y] = \beta_0 + \beta_1 x + \beta_2 \underbrace{\max(0, x - k_0)}_{\text{new variable } x^*} + \beta_3 \underbrace{\max(0, x - k_1)}_{\text{another variable } x^{**}}$$

Logistic model with piecewise linear fit

Call:

```
glm(formula = claim ~ vv + vv2 + vv3,  
     family = binomial(link = logitexp(expo)), ...)
```

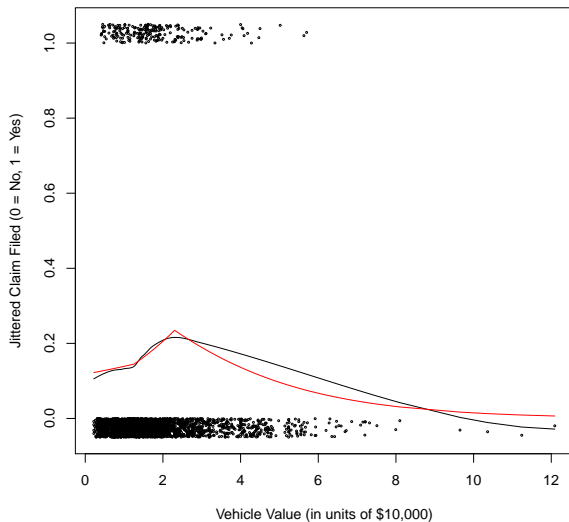
Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-2.0121	0.3762	-5.348	8.9e-08	***
veh.value	0.1874	0.3730	0.503	0.61530	
veh.value*	0.3812	0.5528	0.690	0.49045	
veh.value**	-0.9606	0.3556	-2.701	0.00691	**

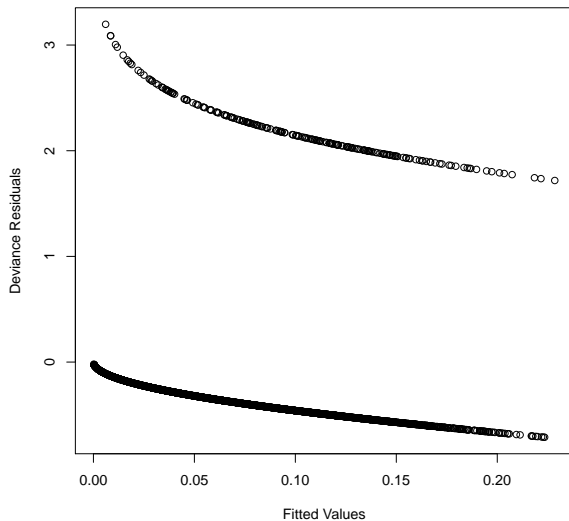
Null deviance: 1598.3 on 2999 degrees of freedom

Residual deviance: 1528.1 on 2996 degrees of freedom

Claim vs Vehicle Value with Piecewise Linear Fits



Residual Plot: deviance residuals vs. fitted values



Many flavors of residuals

Response $y - \hat{\mu}$

Working $(y - \hat{\mu})(\partial\eta/\partial\hat{\mu})$

Partial $(y - \hat{\mu})(\partial\eta/\partial\hat{\mu}) + x_k\hat{\beta}_k$

Pearson $(y - \hat{\mu})/\sqrt{V(\hat{\mu})}$ the variance.

Anscombe Transformed response residuals towards normality.

Deviance Signed square root contribution to the Deviance from each observation.

Quantile For each response variable find the equivalent standard normal deviate. Use randomization for discrete distributions.

Toppings can vary too!

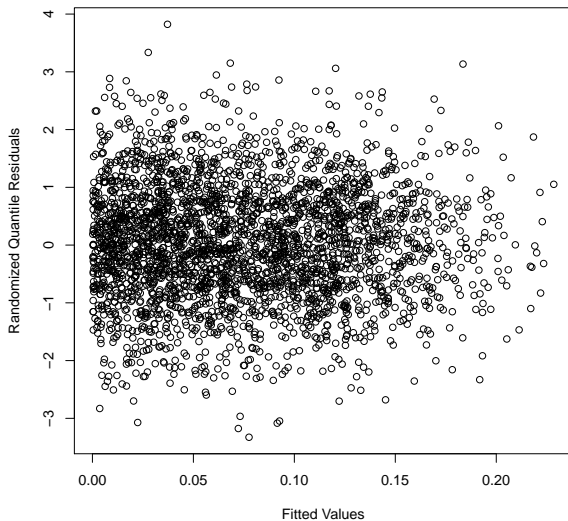
Modified The denominator of the residual has been modified to be a reasonable estimate of the response variance.

Standardized The variance of the residual has been standardized to take into account the correlation between the response and the fitted value.

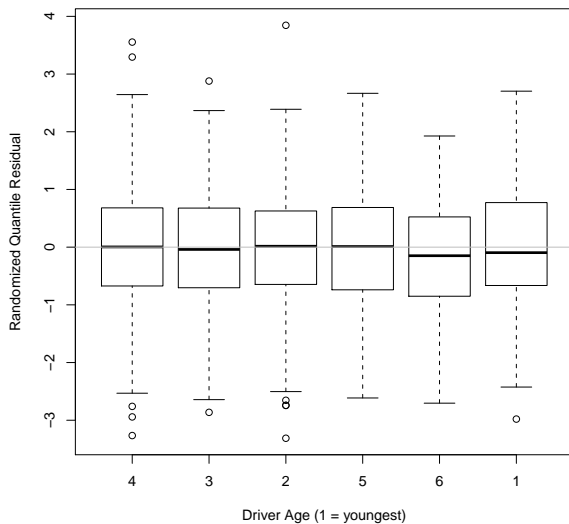
Studentized Residuals have been scaled by an estimate of the unknown scale parameter.

Adjusted The residual has been adjusted from its original definition to bring higher moments in line with the normal distribution.

Randomized Quantile Residual Plot



Quantile Residuals vs. Driver Age



Parameter Estimates Two Variable Model

Call:

```
glm(formula = claim ~ vv + vv2 + vv3 + age.cat,  
     family = binomial(link = logitexp(expo)), ...)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-1.6205	0.3981	-4.070	4.7e-05	***
vv	0.1149	0.3797	0.303	0.762232	
vv2	0.4440	0.5617	0.790	0.429285	
vv3	-0.9598	0.3608	-2.660	0.007808	**
age.cat 3	-0.4237	0.2177	-1.946	0.051602	.
age.cat 2	-0.2388	0.2264	-1.055	0.291546	
age.cat 5	-0.4734	0.2370	-1.998	0.045758	*
age.cat 6	-1.6637	0.4467	-3.724	0.000196	***
age.cat 1	0.1398	0.2791	0.501	0.616583	

Null deviance: 1598.3 on 2999 degrees of freedom
Residual deviance: 1503.4 on 2991 degrees of freedom

New Estimates with Merged Levels

Call:

```
glm(formula = claim ~ vv + vv2 + vv3 + age.cat2,  
     family = binomial(link = logitexp(expo)), ...)
```

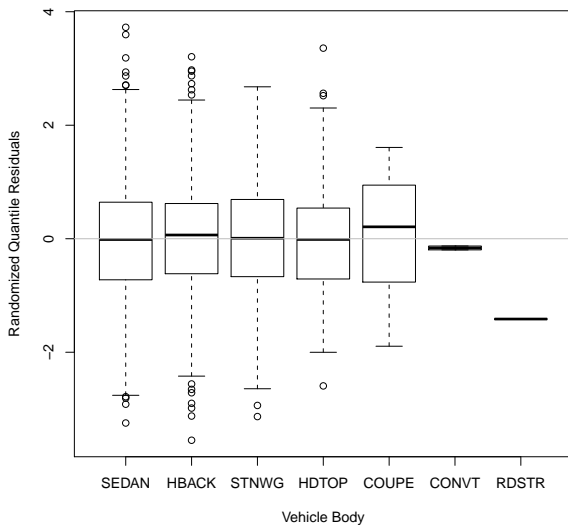
Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-1.86257	0.38295	-4.864	1.15e-06	***
vv	0.09459	0.37896	0.250	0.80289	
vv2	0.47624	0.56028	0.850	0.39532	
vv3	-0.98212	0.35880	-2.737	0.00620	**
age.cat2 5	-1.40331	0.43102	-3.256	0.00113	**
age.cat2 6	0.40216	0.25279	1.591	0.11163	

Null deviance: 1598.3 on 2999 degrees of freedom

Residual deviance: 1509.0 on 2994 degrees of freedom

Quantile Residuals vs. Vehicle Body



How well does our model predict?

Let the threshold probability be 15%. Then, on our build dataset, we have

		Predicted Claim		
		No	Yes	Total
Actual	No	2 568	207	2 775
	Claim	Yes	192	33
Total		2 760	240	3 000

$$\text{Error rate} = (207 + 192)/3000 = 0.133$$

$$\text{Sensitivity} = 33/225 = 0.147$$

$$\text{Specificity} = 2\,568/2\,775 = 0.925$$

Predictions against test and validate datasets?

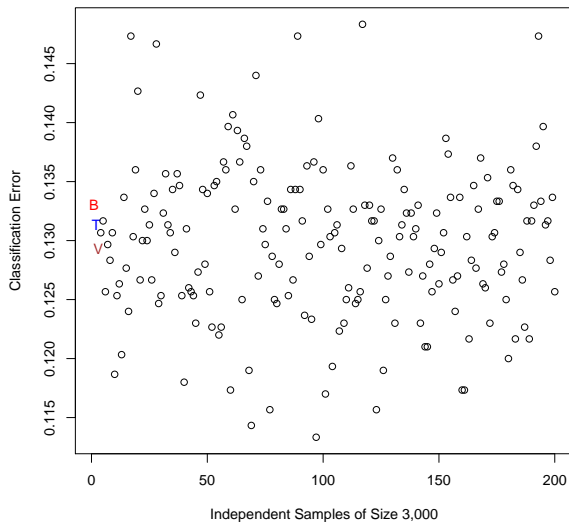
		Predicted Claim		
		No	Yes	Total
Actual	No	2 573	220	2 793
	Claim	Yes	174	33
Total		2 747	253	3 000

$$\text{Error rate} = (220 + 174)/3000 = 0.131$$

		Predicted Claim		
		No	Yes	Total
Actual	No	2 584	231	2 815
	Claim	Yes	157	28
Total		2 741	259	3 000

$$\text{Error rate} = (231 + 157)/3000 = 0.129$$

Classification error across many samples



Personal Injury Claims

The dataset (see [3]) contains 22,036 claims arising from accidents between July 1989 and January 1999. Claims settled with zero payment are not included. The variables in the dataset are:

1. Settlement amount (range: \$10 to \$4.5M)
2. Injury type (codes: 1, 2, 3, 4, 5, 6, 9)
3. Legal representation (codes: 1–Yes, 0–No)
4. Accident, reporting, and settlement month
5. Operational time

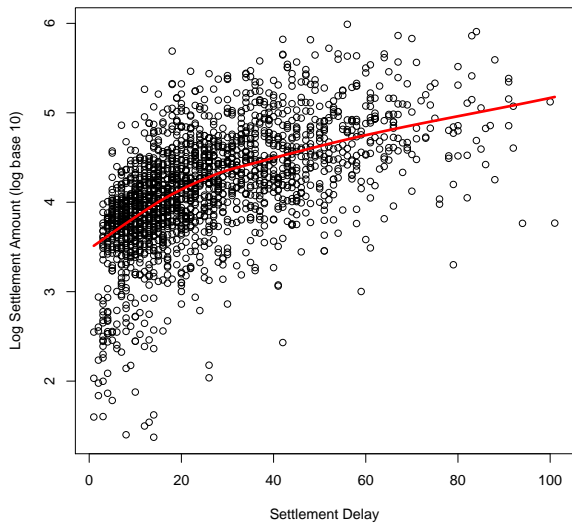
We will work with a random sample of 2,000 claims.

Summary Statistics (for random sample)

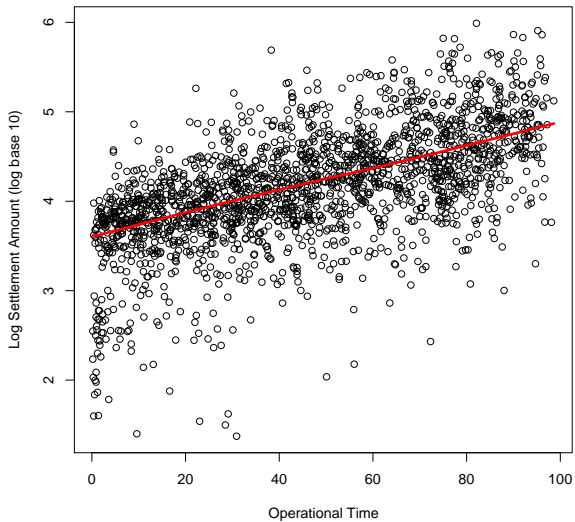
	Claim Amount
Minimum	24
1st Quartile	6,144
Median	14,222
Mean	37,525
3rd Quartile	35,435
Maximum	976,379

There are 172 records ($\approx 8.5\%$) with claim amounts greater than 100,000.

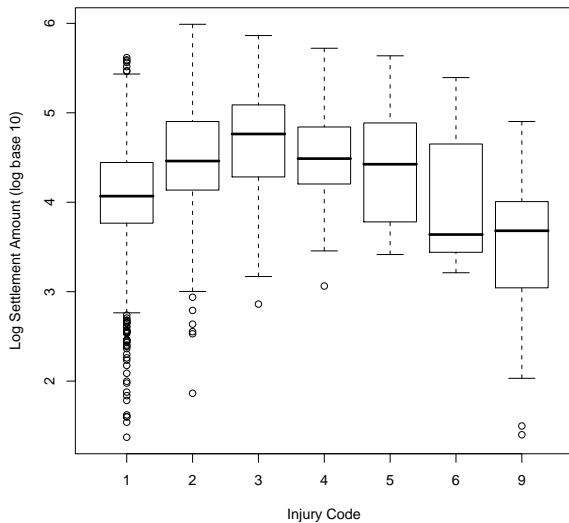
Exploratory Plots I



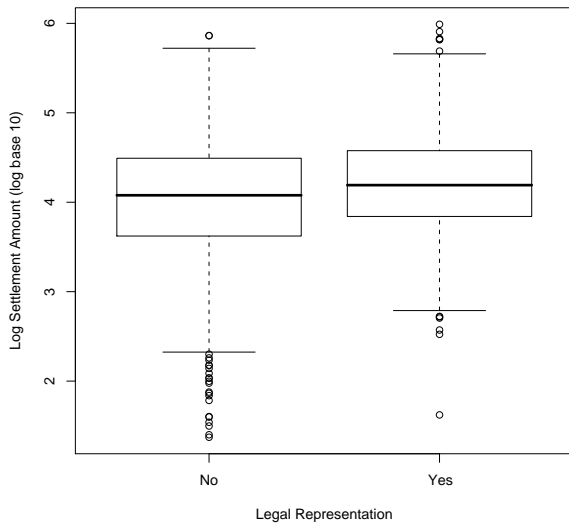
Exploratory Plots II



Exploratory Plots III



Exploratory Plots IV



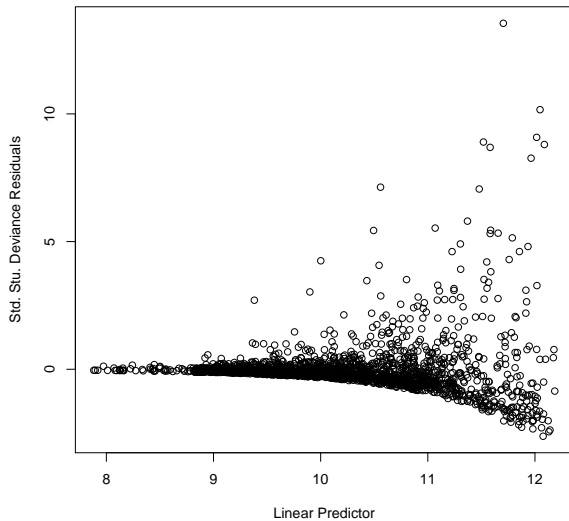
Normal log-link model

$$\log(\text{Settlement Amount}) = \text{Op.Time} + \text{Injury} + \text{Attorney}$$

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	8.817	0.138	63.99	< 2e-16
Op.Time	0.026	0.002	15.82	< 2e-16
injury 2	0.757	0.067	11.31	< 2e-16
injury 3	0.844	0.079	10.75	< 2e-16
injury 4	0.607	0.182	3.33	0.0009
injury 5	0.505	0.199	2.54	0.0113
injury 6	0.645	0.245	2.63	0.0086
injury 9	-0.942	0.554	-1.70	0.0892
attorney Yes	-0.017	0.057	-0.29	0.7705

Residual deviance: 7.9e+12 on 1991 degrees of freedom

Residual Check: Normal error



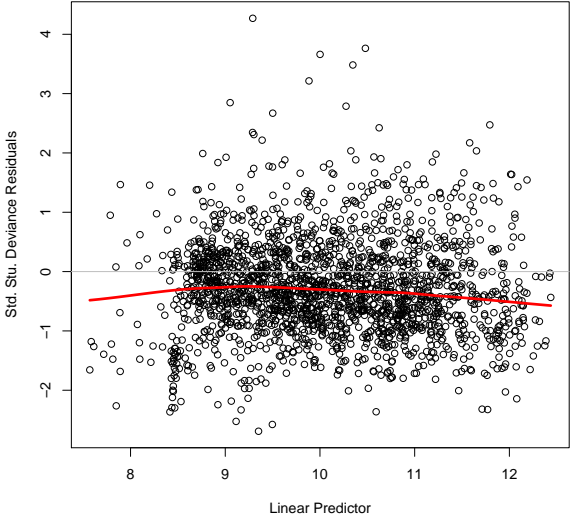
Gamma log-link model

$$\log(\text{Settlement Amount}) = \text{Op.Time} + \text{Injury} + \text{Attorney}$$

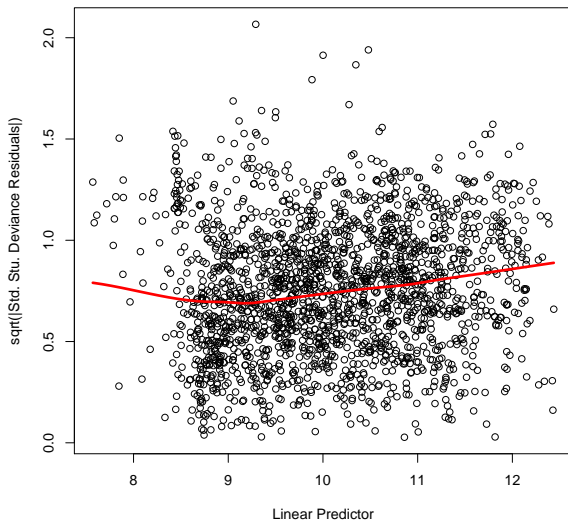
	Estimate	Std. Error	t value	Pr(> t)
Intercept)	8.425	0.064	130.69	< 2e-16
Op.Time	0.030	0.001	29.67	< 2e-16
injury 2	0.707	0.074	9.49	< 2e-16
injury 3	0.900	0.116	7.75	1.46e-14
injury 4	1.045	0.271	3.85	0.0001
injury 5	0.279	0.323	0.86	0.39
injury 6	0.199	0.247	0.80	0.42
injury 9	-0.864	0.129	-6.68	3.00e-11
attorney Yes	0.200	0.057	3.52	0.0004

Residual deviance: 2072.0 on 1991 degrees of freedom

Residual Check: Gamma error



Location-Spread Plot for Gamma Model



Analysis of Deviance Table

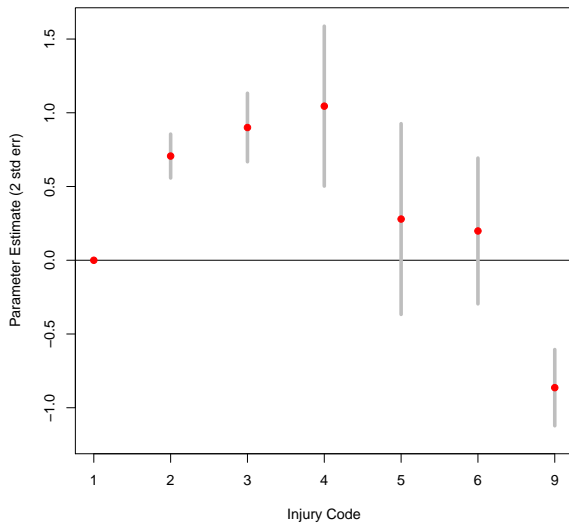
Model: Gamma, link: log

Response: settlement amount

Terms added sequentially (first to last)

	Df	Change in Deviance	Resid. Deviance	Resid. Df
(Intercept)			3894	1999
Op.Time	1	1502	2392	1998
injury	6	303	2089	1992
attorney	1	17	2072	1991

Injury Parameter Estimates



Grouping Injury Levels

Model	Injury levels	Deviance	Diff	q	Crit.Val.
1	1 2 3 4 5 6 9	2072			
2	1 <u>2 3 4</u> 5 6 9	2077	5	2	5.9
3	1 <u>2 3 4</u> <u>5 6</u> 9	2077	5	3	7.8
4	<u>1 5 6</u> <u>2 3 4</u> 9	2079	7	4	9.5
5	1 <u>2 3 4 5 6</u> 9	2086	14	4	9.5

Diff is the difference between the current model and model 1.

q is the number of restrictions in the current model compared to model 1.

Crit.Val. is the 0.95 quantile of the chi-squared distribution with q degrees of freedom.

Checking the Link Function

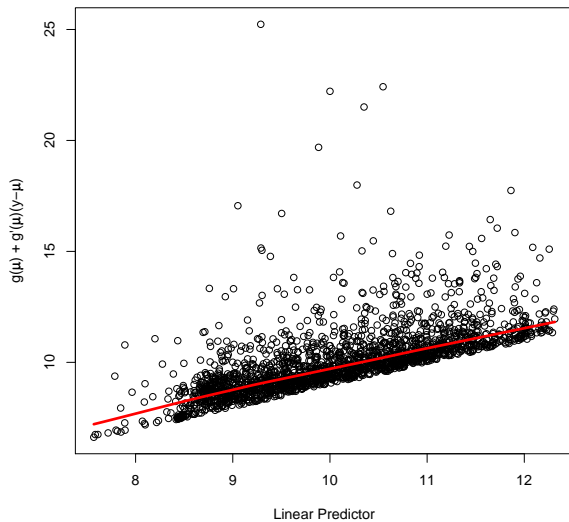
Two ways to assess the link function:

1. Embed the link function in a parametric family and compare model fit at various points.
2. We know that

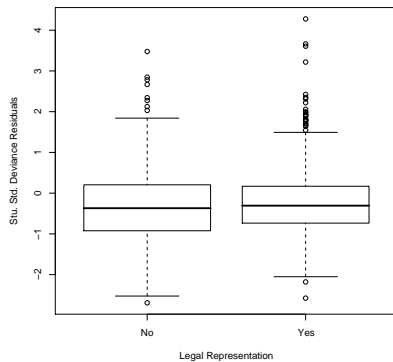
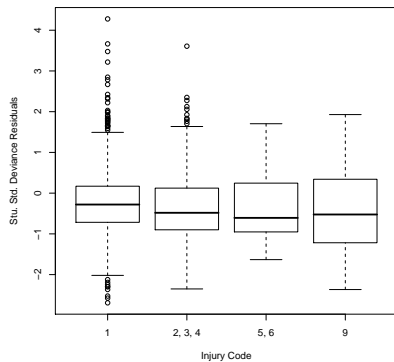
$$x_i\beta = g(y_i) \approx g(\mu_i) + g'(\mu_i)(y_i - \mu_i)$$

So plotting the linear predictor against the right-hand side of the above equation should give us a straight line.

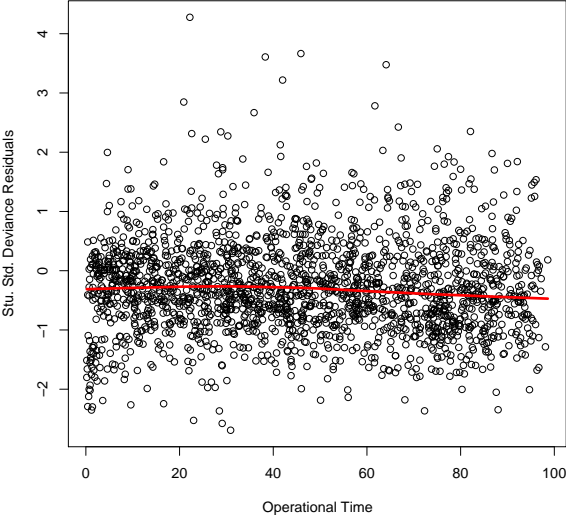
Checking the Link Function



Checking Explanatory Variables



Checking Explanatory Variables

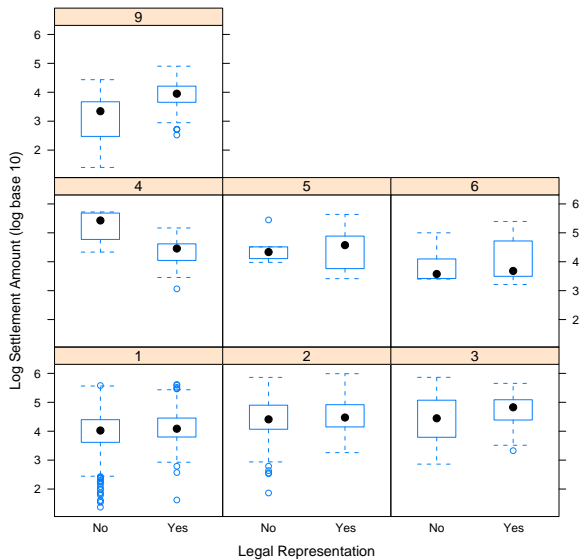


Interactions

We say that two explanatory variables x and z interact if the effect of x on the response variable depends on the values of z .

For our example, does the effect of attorney involvement depend on the type of injury?

Conditional Plot



Model Validation

Several model validation techniques:

1. Out-of-sample
2. Cross-validation
3. Bootstrap estimates of prediction errors

Out-of-Sample Validation

Predicted values compared against actual values for a new sample of 2,000 claims.

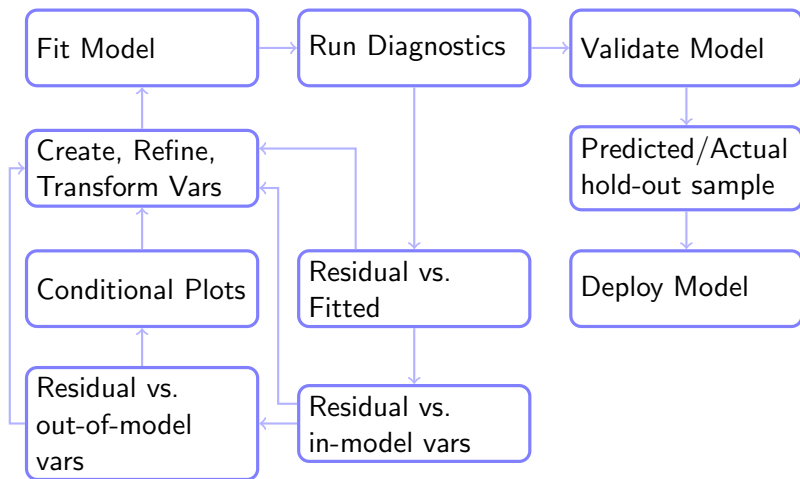
Predicted Range	Type	1st Qu.	Mean	Ratio A/P	3rd Qu.
(43800, 61500]	A	14770	45790		52880
	P	48150	52720	0.87	57460
(61500, 91600]	A	22800	77900		85350
	P	67180	74800	1.04	81860
(91600,232000]	A	42680	150700		171700
	P	106300	135000	1.12	156700

Only the last three groups of the table are shown.




The type column refers to actual (A) or predicted (P) values.

The column ratio A/P is the ratio of the actual mean divided by the predicted mean.





Model Building Cycle



References

-  John M. Chambers, William S. Cleveland, Beat Kleiner, and Paul A. Tukey.
Graphical Methods for Data Analysis.
The Wadsworth Statistics/Probability Series. Wadsworth International Group, Belmont, California, 1983.
-  W.S. Cleveland.
Visualizing Data.
Hobart Press, 1993.
-  Piet De Jong and Gillian Z. Heller.
Generalized Linear Models for Insurance Data.
Cambridge University Press, 2008.

References

-  Peter K. Dunn and Gordon K. Smyth.
Randomized quantile residuals.
Journal of Computational and Graphical Statistics,
5(3):236–244, 1996.
-  L. Fahrmeir and G. Tutz.
Multivariate Statistical Modelling Based on Generalized Linear Models.
Springer, 2001.
-  James Hardin and Joseph Hilbe.
Generalized Linear Models and Extensions.
Stata Press, College Station, Texas, 2001.
-  W.N. Venables and B.D. Ripley.
Modern Applied Statistics with S.
Springer New York, 2002.