

# Actionable predictive learning for insurance profit maximization

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- **Predictive Modeling** is a core strategic capability of many top insurers (widely applied in marketing, underwriting, pricing, claims management, fraud detection, etc.)
- **Common goal of models:** to predict a response variable using a collection of **observable attributes** (e.g., Age, Yrs. Licensed, Gender, Territory, Claims and Conviction History, etc.)

Tons of literature on the above, but less attention has been paid to:

- In many important settings, the values of certain attributes can be proactively chosen at the discretion of a decision maker – called **actionable attributes** or **“treatments”**. For instance, we can choose:
  - Which policyholders should be contacted to prevent them from switching to an alternative insurer?
  - Which Auto insurance clients should be offered a Life policy?
  - By how much should we change the rates at policy renewal?

- The values chosen for the actionable attributes have important **implications for the ultimate profitability of the insurance company**
- There is no “global” better action  $\Rightarrow$  Relevant in the context of **treatment heterogeneity effects**
- The objective is NOT to predict a response variable with high accuracy (as in predictive modeling), but to select the **optimal action** or treatment for each client
- **Optimal personalized treatment**  $\Rightarrow$  the one that maximizes the probability of a desirable outcome (e.g., Profits)
- Not addressed by traditional predictive modeling techniques (GLMs, CART, SVM, Neural Nets, etc.).

## A toy example: The red/blue envelope problem

- Consider a Client Retention Program aimed to increase the overall retention rate of an insurance portfolio
- Treatment consists in a promotion sent either in a red or blue envelope

Table: Treatment impact on the client's renewal outcome

Client Type	Red envelope	Blue envelope
A	NOT renew	NOT renew
B	Renew	Renew
C	NOT renew	Renew
D	Renew	NOT Renew

- Clients 'A' and 'B' are indifferent to the color of the envelope
- The optimal personalized treatment is to send a Blue envelope to 'C' clients and a Red envelope to 'D' clients

- Literature is relatively scarce and mostly published recently
- **Personalized Medicine:** (Qian and Murphy, 2011; Zhao et al., 2012; Su et al., 2009)
- **Marketing:** (Jaskowski and Jaroszewicz, 2012; Radcliffe and Surry, 2011; Lo, 2002)
- **Economics:** Imai and Ratkovic (2013)
- **Insurance:** Personalized treatments in the context of Pricing, Client Retention and Cross-Selling

Guelman, L. and Guillén, M. (2014). “A causal inference approach to measure price elasticity in automobile insurance”.

*Expert Systems with Applications* 41: 387–396.

Guelman, L., et al. (2013). “Uplift random forests”. *Cybernetics & Systems, forthcoming*.

Guelman, L., et al. (2013). “Optimal personalized treatment rules for marketing interventions: A review of methods, a new proposal, and an insurance case study.” *Submitted*.

The problem of selecting the optimal treatment is non-trivial...

- The **outcome of interest** – i.e., the optimal treatment – is **unknown** on a given training data set
- Each client can only be exposed to one treatment condition  $\Rightarrow$  we can only observe the response under the exposed condition.  
The counterfactual response is never observed  $\Rightarrow$  the “true” optimal treatment is not observed ([Holland, 1986](#))
- A key distinction for building personalized treatment learning models is between **randomized experiments** and **observational data**.

## Let's formalize the problem

- For now assume a controlled randomized experiment – i.e., clients are randomly assigned to two treatments, denoted by  $A \in \{0, 1\}$
- Let  $Y(a) \in \{0, 1\}$  denote a **binary potential response** of a client if assigned to treatment  $A = a$ ,  $a \in \{0, 1\}$
- The **observed response** is  $Y = AY(1) + (1 - A)Y(0)$
- Clients are characterized by a  $p$ -dimensional vector of baseline **predictors**  $\mathbf{X} = (X_1, \dots, X_p)^\top$
- Data consists of  $L$  i.i.d. realizations of  $(Y, A, \mathbf{X})$ ,  $\{(Y_\ell, A_\ell, \mathbf{X}_\ell), \ell = 1, \dots, L\}$ .



## Let's formalize the problem

- At the most granular level, the personalized treatment effect is a comparison between  $Y(1)$  and  $Y(0)$  on the same client. Usually,

$$Y_\ell(1) - Y_\ell(0) \quad \forall \ell = \{1, \dots, L\}$$

- But as discussed above, this is an unobserved quantity
- In practice, the best we can do is to estimate the personalized treatment effect by conditioning on clients with profile  $\mathbf{X} = \mathbf{x}$
- Thus, we define the *personalized treatment effect* (PTE) by

$$\begin{aligned} \tau(\mathbf{x}) &= E[Y_\ell(1) - Y_\ell(0) | \mathbf{X}_\ell = \mathbf{x}] \\ &= E[Y_\ell | \mathbf{X}_\ell = \mathbf{x}, A_\ell = 1] - E[Y_\ell | \mathbf{X}_\ell = \mathbf{x}, A_\ell = 0]. \end{aligned}$$

# The two-model approach to PTE estimation

- 1 Estimate  $E[Y|\mathbf{X}, A = 1]$  using the treated clients only
- 2 Estimate  $E[Y|\mathbf{X}, A = 0]$  using the control clients only
- 3 An estimate of the PTE for a client with predictors  $\mathbf{X}_\ell = \mathbf{x}$  is

$$\hat{\tau}(\mathbf{x}) = (\hat{Y}_\ell | \mathbf{X} = \mathbf{x}_\ell, A_\ell = 1) - (\hat{Y}_\ell | \mathbf{X} = \mathbf{x}_\ell, A_\ell = 0).$$

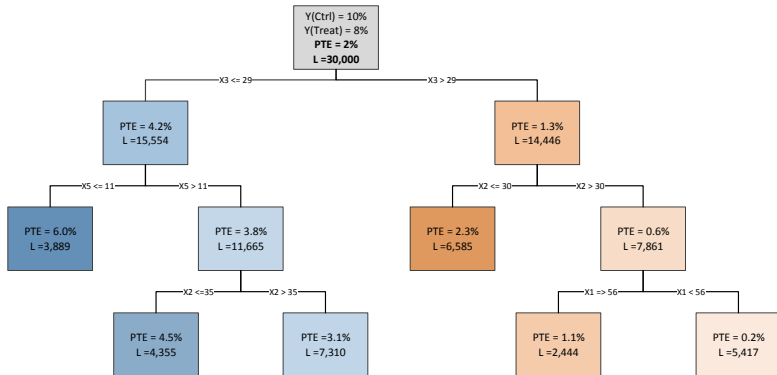
## Pros:

- Any conventional statistical or algorithmic binary classification method may serve to fit the models.

## Cons:

- Models developed to predict the wrong target!
  - The method emphasize the prediction accuracy on the response, not the accuracy in estimating the **change in the response caused by the treatment**
  - Relevant predictors for  $Y$  are usually different from relevant PTE predictors

# Causal Conditional Inference Tree



$Y(\text{Treat})$  = Attrition rate on treated clients

$Y(\text{Ctrl})$  = Attrition rate on control clients

$\text{PTE}$  = Personalized treatment effect:  $Y(\text{Ctrl}) - Y(\text{Treat})$

$L$  = Number of clients

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**Algorithm 1** Causal conditional inference tree

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- 1: **for** each terminal node **do**
  - 2:   Test the global null hypothesis  $H_0$  of no interaction effect between the treatment A and any of the  $p$  predictors at a level of significance  $\alpha$  based on a permutation test (Strasser and Weber, 1999)
  - 3:   **if** the null hypothesis  $H_0$  cannot be rejected **then**
  - 4:     **Stop**
  - 5:   **else**
  - 6:     Select the  $j^*$ -th predictor  $X_{j^*}$  with the strongest interaction effect (i.e., the one with the smallest adjusted  $P$  value)
  - 7:     Choose a partition  $\Omega^*$  of the covariate  $X_{j^*}$  in two disjoint sets  $\mathcal{M} \subset X_{j^*}$  and  $X_{j^*} \setminus \mathcal{M}$  based on the  $G^2(\Omega)$  split criterion
  - 8:   **end if**
  - 9: **end for**
- 

$$G^2(\Omega) = \frac{(L-4) \left\{ \overbrace{(\bar{Y}_{n_L}(1) - \bar{Y}_{n_L}(0))}^{\text{Left Node}} - \overbrace{(\bar{Y}_{n_R}(1) - \bar{Y}_{n_R}(0))}^{\text{Right Node}} \right\}^2}{\hat{\sigma}^2 \{1/L_{n_L}(1) + 1/L_{n_L}(0) + 1/L_{n_R}(1) + 1/L_{n_R}(0)\}}$$

# R implementation: The **uplift** package in CRAN

## The highlights:

- Implements various functions for training personalized treatment learning models (a.k.a., uplift)
- Currently 5 estimation methods are implemented
  - Causal conditional inference forests (`ccif`)
  - Uplift random forests (`upliftRF`)
  - Modified covariate method (`tian_transf`)
  - Modified outcome method (`rvtu`)
  - Uplift k-nearest neighbor (`upliftKNN`)
- Exploratory Data Analysis (EDA) tools designed for PTE models
- Functions for evaluating performance of PTE models
- Profiling results of PTE models
- PTE Monte Carlo simulations
- Package in continuous development

# A cross-sell example: Auto $\Rightarrow$ Property Insurance

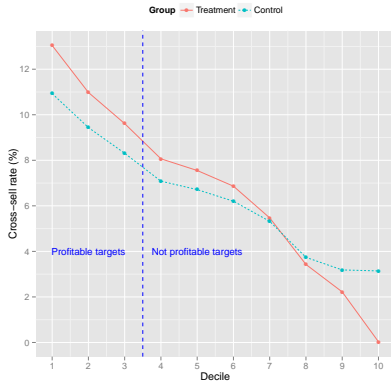
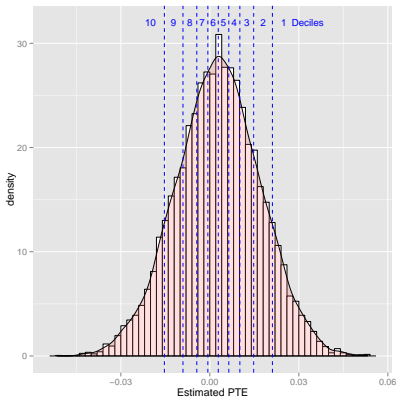
- A randomized experiment with a cross-sell binary “treatment”

Table: Cross-sell rates by group

	Treatment	Control
Purchased Home policy = N	30,184	3,322
Purchased Home policy = Y	789	75
Cross-sell rate	2.55%	2.21%

- The average treatment effect is **0.34%** (2.55% - 2.21%), which is not statistically significant ( $P$  value = 0.23)
- Can we identify a subgroup of clients for which the treatment was effective? If so, target those clients in the future.

# A cross-sell example: Auto $\Rightarrow$ Property



- Consider the existing portfolio of an insurer where the premium  $P_{\ell t}$  charged to policyholder  $\ell = \{1, \dots, L\}$  in year  $t$  is given by

$$P_{\ell t} = \hat{L}C_{\ell t} + E_{\ell t} + A_{\ell t}$$

where

$\hat{L}C_{\ell t}$  = Expected loss cost

$E_{\ell t}$  = Expenses

$A_{\ell t}$  = Profit loading

- **Loss cost** estimation has seen an enormous advance with predictive modeling
- **Profits** have remained obscure and rather forgotten.



# Ratemaking and personalized treatment learning

- We can think of  $A_\ell$  as an actionable attribute or “treatment” which can take values on a continuous scale
- The problem is to select the **optimal personalized treatment**: *the one that maximizes the overall profitability of the insurance portfolio* ( $\sum_{\ell=1}^L P_{\ell t} - L\hat{C}_{\ell t} - E_{\ell t}$ )
- Assuming  $L\hat{C}_\ell$  and  $E_\ell$  are exogenous, **then selecting the optimal  $A_\ell \Rightarrow$  selecting the optimal  $P_\ell$**
- The impact of a change in  $P_\ell$  on the overall profitability of the portfolio is a-priori uncertain as a big enough  $P_\ell$  will make a policyholder more likely to switch to an alternative insurer
- This requires understanding the precise impact of a change in  $P_\ell$  on the probability of renewal for each policyholder  $\ell$  – i.e., the **price elasticity**

# Price Elasticity as a missing data problem

- *Price elasticity involves a comparison of the potential renewal outcomes for alternative rate changes (the “treatments”) defined on the same policyholder*
- Due to the **fundamental problem of personalized treatment learning models**  $\Rightarrow$  each policyholder can only be exposed to one rate change value, so only one of the potential renewal outcomes is an observed outcome. *The counterfactual outcomes are never observed.*
- One way to think about the **counterfactual outcomes** is that their values are **“missing”** and therefore they should be multiply imputed to represent their uncertainty.

# Price Elasticity as a missing data problem

- To simplify, let's bin the rate change into five ordered values  $A = \{1 < \dots < 5\}$  and assume a 1-year horizon
- The entries  $r_{\ell a}$  below denote the observed renewal outcome  $\in \{0, 1\}$  of policyholder  $\ell = \{1, \dots, L\}$  when exposed to rate change level  $A = a$ ;  $a = \{1 < \dots < 5\}$
- Dots indicate counterfactual outcomes, which are missing
- The price elasticity estimation problem  $\equiv$  the problem of filling in the missing values in the client-by-rate change table with reliable estimates.

**Table: Client-by-Rate change table**

Client	Rate Change Level				
	Level 1	Level 2	Level 3	Level 4	Level 5
1	.	$r_{12}$	.	.	.
2	.	.	$r_{23}$	.	.
3	$r_{31}$	.	.	.	.
4	.	.	.	$r_{44}$	.
5	.	$r_{52}$	.	.	.
6	.	.	.	.	$r_{65}$
...	...	...	...	...	...
L	.	.	.	.	$r_{L5}$

## A key additional complexity

- **Reliable estimates of effects attributable to treatments require experimental data** (i.e., coming from randomized experiments)
- This means that for reliable price elasticity estimation, data must come from a randomized assignment of policyholders to rate change levels
- **This condition rarely holds in practice:** rate changes are mostly derived from a pricing modeling exercise  $\Rightarrow$  rate change is a deterministic function of the policyholder's observed risk characteristics
- Thus, we end up with **observational data** – i.e., not derived from experimentation
- **Policyholders exposed to different rate change levels are not directly comparable.**

## But...what is the problem?

- **The standard approach:** model the policyholder's lapse outcome as a function of the rate change and the policyholder's covariates
- The **key assumption:** the inclusion of those covariates adjust for the exposure correlations between price elasticity and other explanatory variables
- **Problem:** non-overlapping supports of  $X$  between policyholders exposed to different rate change levels
- **As an extreme example:** Assume policyholder's Age is associated with the lapse outcome

Age	Rate Change	
	5%	10%
< 25 yrs.	✓	✓
≥ 25 yrs.	✓	NA

- "✓" indicates whether historical data is available
- Clients  $\geq 25$  yrs. exposed to a 5% rate change don't have a good comparison in the 10% rate change group

## But...what is the problem?

- Regression analysis masks this fact and assumes that the estimated price elasticity model is good for all policyholders (even for those never observed under a specific rate change)
- In real data sets, extreme examples such as the above are rare, but **non-overlap situations are common**
- **Non-overlap** refers to the extent to which the distribution of the key renewal/lapse predictors differ across policyholders historically exposed to different rate change levels
- The problem is even worse with a **large number of predictors**, as groups may differ in a multivariate direction and so non-overlap problems are more difficult to detect

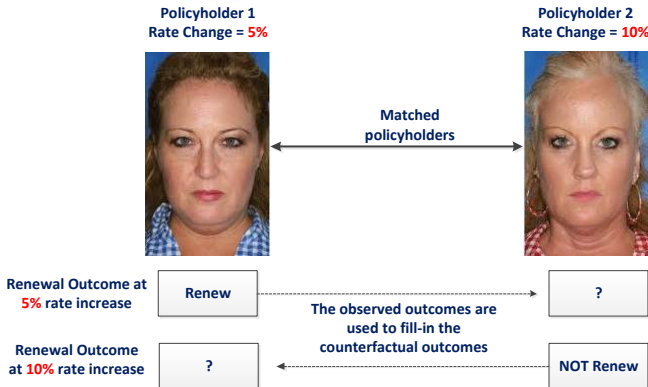
# Propensity scores and Matching algorithms

Some good news...

- Under certain data conditions ([Rosenbaum and Rubin, 1983](#)):
  - We can construct a randomized-type of experiment from observational data  $\Rightarrow$  helpful for determining **price elasticity at the portfolio level**
  - It's possible to infer the “missing” counterfactual renewal outcomes (and thus fill-in the missing values in the client-by-rate change table)  $\Rightarrow$  helpful for determining **price elasticity at the individual policyholder level**
- The key concepts are **propensity scores** ([Rosenbaum and Rubin, 1983](#)) and **matching algorithms** ([Gu and Rosenbaum, 1993](#))

# Matching: conceptual framework

- Let's say in the training data we have 2 policyholders which are very similar in terms of their relevant lapse predictors  $\mathbf{X}$  – i.e., about the same age, driving record, living in the same neighbourhood, etc.
- But, they have been exposed to different rate change levels – e.g., 5% and 10% (enough historical data may allow us to find such pair)





Matching algorithms have many variants. There are 3 key choices:

- 1 The **definition of distance** between two policyholders in terms of their characteristics
- 2 The choice of the **algorithm** used to form the matched pairs and make the distance small (greedy vs. optimal matching)
- 3 The **structure of the match** (i.e., the number of treated and control subjects that should be included in each match set)

In [Guelman and Guillén \(2014\)](#), we used **optimal pair matching**  
⇒ equivalent to finding a flow of minimum cost in a certain network (a standard combinatorial optimization problem)

# Propensity scores

- Even with a moderate number of predictors, exact matches on  $\mathbf{X}$  are not feasible  $\Rightarrow$  propensity scores come into play
- Given a binary treatment  $A \in \{0, 1\}$ , the **propensity score** is the conditional probability of assignment to treatment 1 given  $\mathbf{X}$ ,

$$\pi(\mathbf{X}_\ell) = P(A_\ell = 1 | \mathbf{X}_\ell)$$

- In a randomized experiment,  $\pi(\mathbf{X}_\ell) = 1/2 \quad \forall \mathbf{X}_\ell$
- In an observational study, the propensity score can be estimated (e.g., logistic regression)
- With more than two treatments, we could (i) consider all possible treatment dichotomies or (ii) build a multinomial response model.

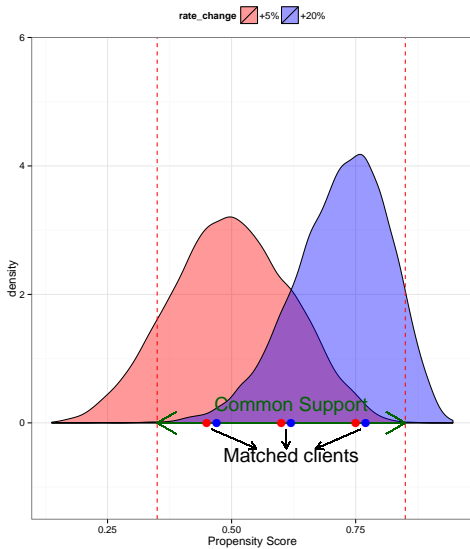
# Propensity scores - Balancing Property

- An important property of the propensity score allows us to **match only on the propensity score**
- The **Balancing Property**: Treatment  $A$  and the observed covariates  $\mathbf{X}$  are conditionally independent given the propensity score  $\pi(\mathbf{X})$ ,

$$A \perp \mathbf{X} | \pi(\mathbf{X})$$

i.e., conditional on the propensity score  $\pi(\mathbf{X})$ , the distribution of  $\mathbf{X}$  is similar for  $A=1$  and  $A=0$ .

# Propensity score for 20% vs. 5% rate change dichotomy



# Filling the Client-by-Rate change table

- 1 *Replace the actual renewal outcomes with probability estimates*

Client	Rate Change Level				
	Level 1	Level 2	Level 3	Level 4	Level 5
1	.	$\hat{r}_{12}$	.	.	.
2	.	.	$\hat{r}_{23}$	.	.
3	$\hat{r}_{31}$	.	.	.	.
...	...	...	...	...	...
L	.	.	.	.	$\hat{r}_{L5}$

- 2 *Infer the counterfactual renewal outcomes from the matched pairs (as far as the overlap situation permits)*

Client	Rate Change Level				
	Level 1	Level 2	Level 3	Level 4	Level 5
1	$\hat{r}_{11}$	$\hat{r}_{12}$	$\hat{r}_{13}$	$\hat{r}_{14}$	$\hat{r}_{15}$
2	$\hat{r}_{21}$	$\hat{r}_{22}$	$\hat{r}_{23}$	$\hat{r}_{24}$	.
3	$\hat{r}_{31}$	.	.	.	.
...	...	...	...	...	...
L	$\hat{r}_{L1}$	$\hat{r}_{L2}$	$\hat{r}_{L3}$	$\hat{r}_{L4}$	$\hat{r}_{L5}$

# Filling the Client-by-Rate change table

## 3 Develop a “global model” of the response.

- Develop a global model  $\hat{r}_{\ell t}(\mathbf{x}_\ell)$ , obtained by fitting the estimates  $\hat{r}_{\ell t}$  of the observed responses, plus the estimates of a subset of the counterfactual responses on the vector of observed characteristics  $\mathbf{x}_\ell$  and rate change level  $a = \{1 < \dots < 5\}$
- This model allows us to predict the renewal outcome for each rate change  $A = a$  and value of  $\mathbf{X}$ .

**Table: Client-by-Rate change table filled with “global” renewal probability estimates**

Client	Rate Change Level				
	Level 1	Level 2	Level 3	Level 4	Level 5
1	$\hat{r}_{11}$	$\hat{r}_{12}$	$\hat{r}_{13}$	$\hat{r}_{14}$	$\hat{r}_{15}$
2	$\hat{r}_{21}$	$\hat{r}_{22}$	$\hat{r}_{23}$	$\hat{r}_{24}$	$\hat{r}_{25}$
3	$\hat{r}_{31}$	$\hat{r}_{32}$	$\hat{r}_{33}$	$\hat{r}_{34}$	$\hat{r}_{35}$
...	...	...	...	...	...
L	$\hat{r}_{L1}$	$\hat{r}_{L2}$	$\hat{r}_{L3}$	$\hat{r}_{L4}$	$\hat{r}_{L5}$

- The proposed framework to fill-in the counterfactual renewal outcomes with probability estimates allows us to **more efficiently solve the Economic Price Optimization problem**
- **The problem:** which rate change should we expose each policyholder to maximize the overall expected profit of the portfolio subject to a fixed overall retention rate?
- Recall that: An **Optimal personalized treatment** is the one that maximizes the probability of a desirable outcome (treatment  $\equiv$  rate change and the outcome  $\equiv$  profits)

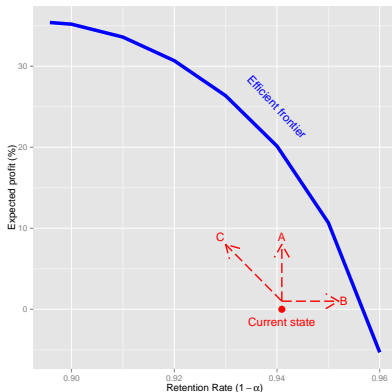
# The optimization problem: An integer program

Maximize an expected profit function

$$Z_{\ell a} \text{Max}_{\forall \ell \forall a} \sum_{\forall \ell} \sum_{\forall a} Z_{\ell a} \left[ P_{\ell} (1 + RC_a) (1 - L \hat{R}_{\ell a}) (1 - \hat{r}_{\ell a}) \right]$$

subject to a retention constraint

$$\begin{aligned} \sum_{\forall a} Z_{\ell a} &= 1 \quad \forall \ell \\ Z_{\ell a} &\in \{0, 1\} \\ \sum_{\forall \ell} \sum_a Z_{\ell a} \hat{r}_{\ell a} / L &\leq \alpha. \end{aligned}$$





# Wrapping up

- We introduced the concept of predictive learning with actionable attributes (in the context of marketing and pricing intervention activities)
- The values chosen for these attributes have important implications for the ultimate profitability of the insurer
- Off-the-shelf predictive modeling algorithms can generally not be used to tackle learning with actionable attributes
- The nature of the data is key: experimental vs. observational (experimental data is more common in marketing than in pricing interventions)
- Discussed methods and tools useful for each data context.

Your turn...

